

A Benchmark Approach to Investing, Pricing and Hedging

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Pl. & Heath (2006, 2010), A Benchmark Approach to Quantitative Finance. Springer

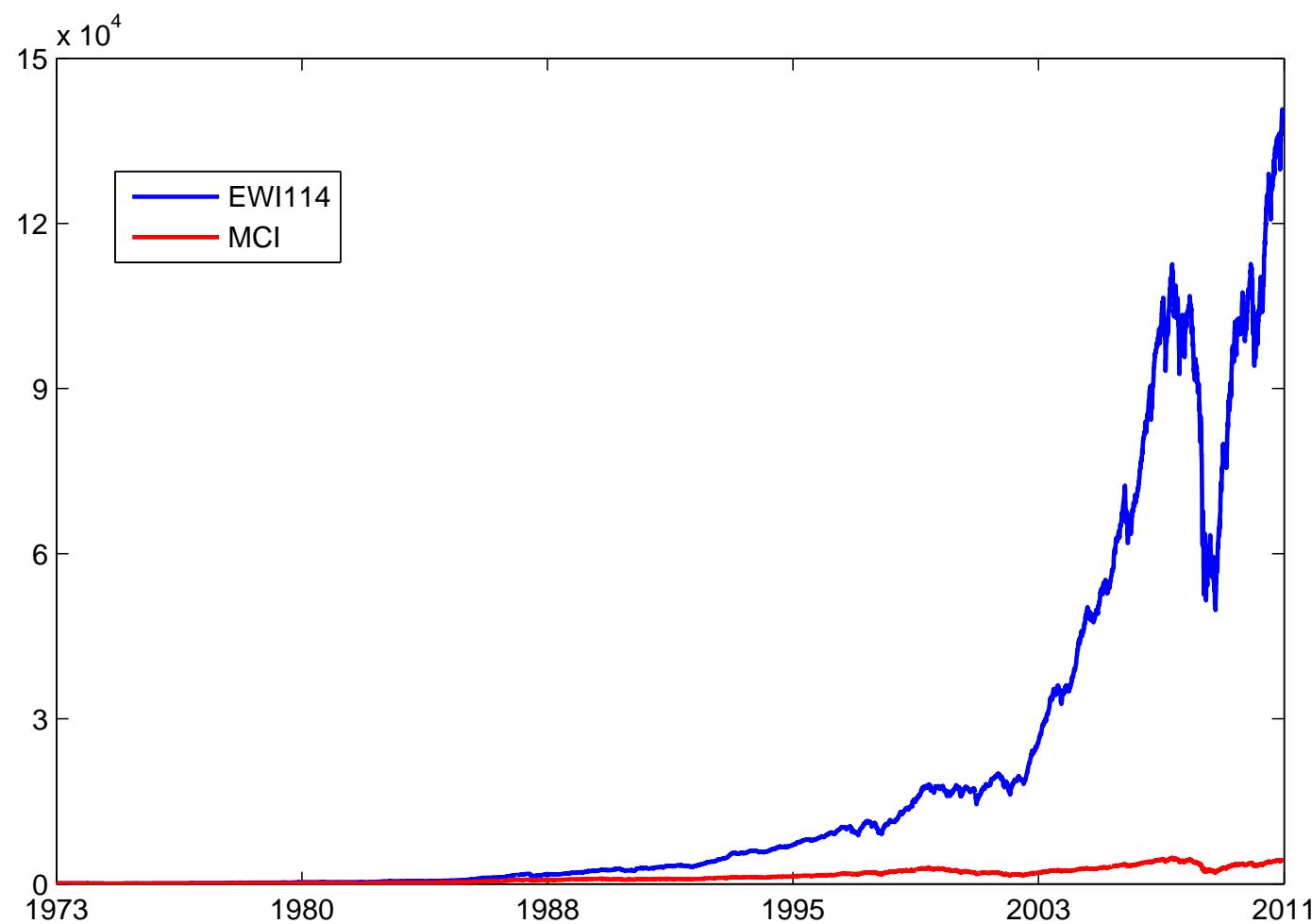
Pl. & Bruti-Liberati (2010), Numerical Solution of Stochastic Differential Equations with Jumps in Finance. Springer

Baldeaux & Pl. (2013). Functionals of Multidimensional Diffusions with Applications to Finance. Springer

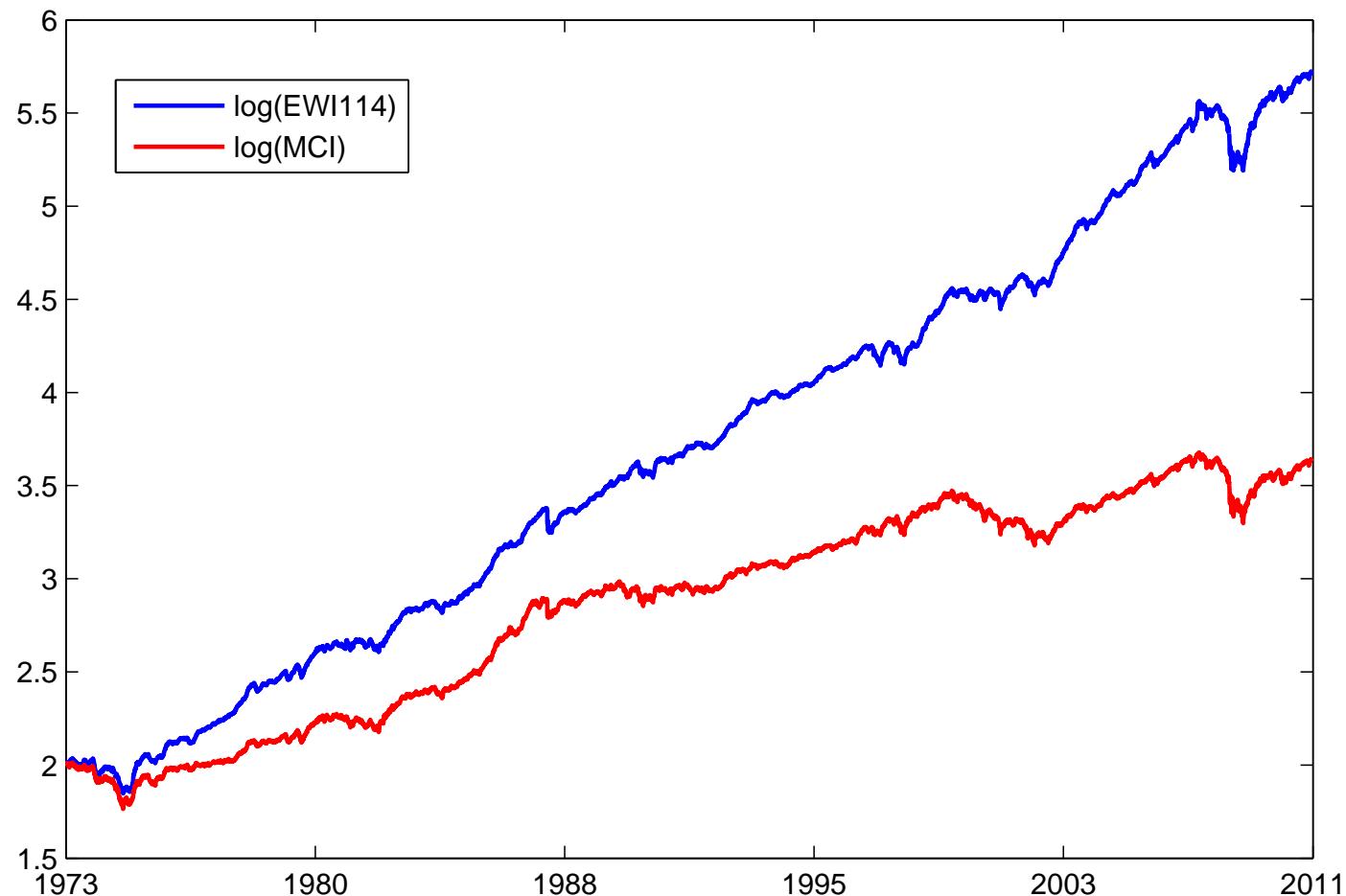
1 Best Performing Portfolio as Benchmark

- with focus on long term
- if possible pathwise
- should prefer more for less
- diversified portfolio
- robust, almost model independent construction
- sustainable (strictly positive long only)
- criterion independent from denomination
- criterion independent from sequence of drawdown events

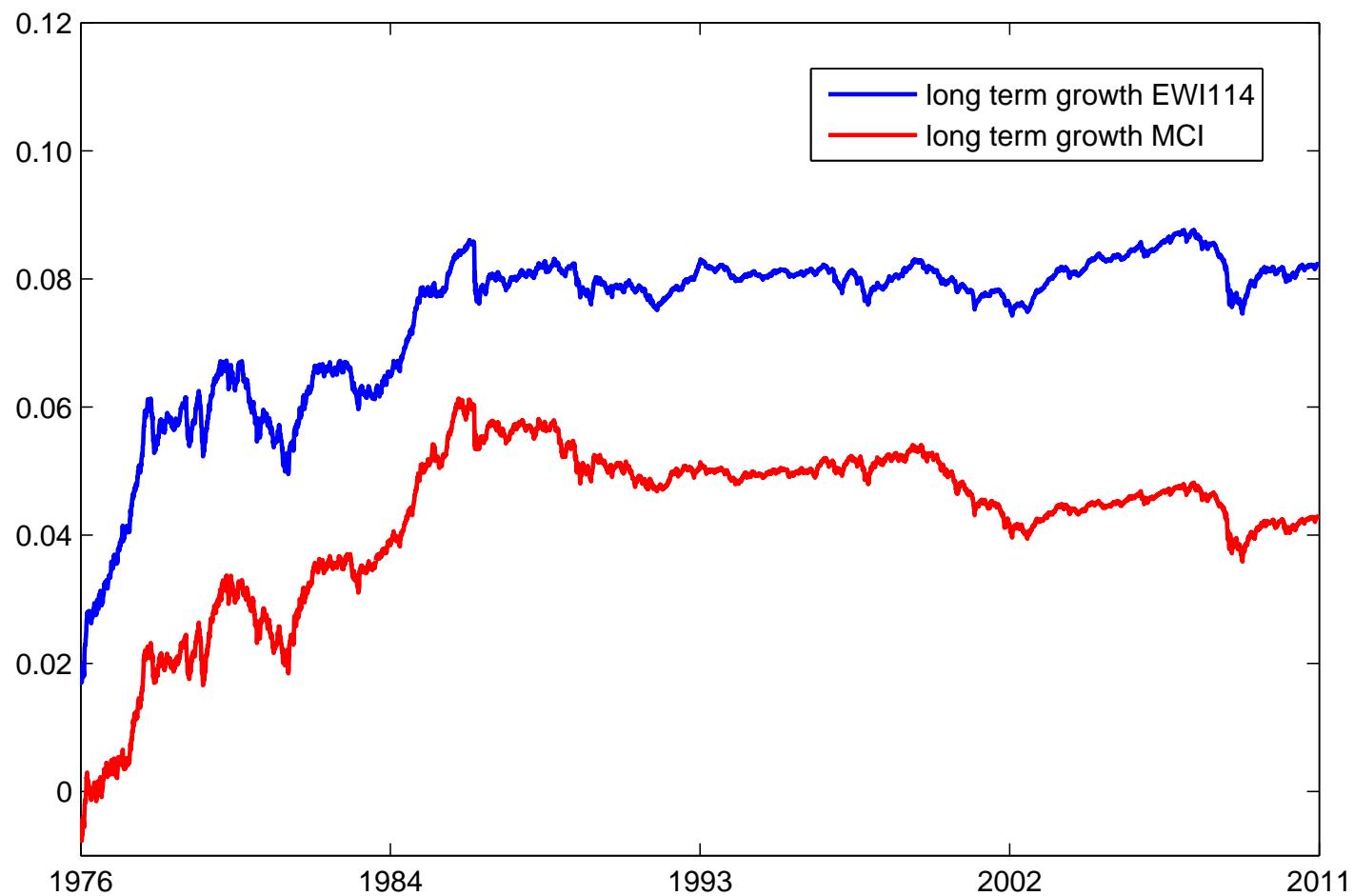
⇒ **maximize logarithm of portfolio**



EWI141 and MCI



logarithms of EWI114 and MCI



Long term growth of EWI114 and MCI

Key idea:

Make the “**best**” performing portfolio $S_t^{\delta*}$ numéraire or **benchmark**

Portfolio:

$$S_t^\delta = \sum_{j=0}^d \delta_t^j S_t^j$$

S_t^j - j th constituent (e.g. cum dividend stocks)

Long Term Growth:

$$\lim_{t \rightarrow \infty} \frac{1}{t} \log \left(\frac{S_t^\delta}{S_0^\delta} \right) \leq \lim_{t \rightarrow \infty} \frac{1}{t} \log \left(\frac{S_t^{\delta*}}{S_0^{\delta*}} \right)$$

a.s. pathwise

$S^{\delta*}$ is the **numéraire portfolio** (NP)

Long (1990)

Main Assumption:

There exists NP $S_t^{\delta*}$,

which means

$$\frac{S_t^\delta}{S_t^{\delta*}} \geq E_t \left(\frac{S_s^\delta}{S_s^{\delta*}} \right) \quad (*)$$

for all $0 \leq t \leq s < \infty$ and all nonnegative portfolios S_t^δ ,

$E_t(\cdot)$ - real world conditional expectation under information given at time t

$S_t^{\delta*}$ - numéraire portfolio (NP), see Long (1990)

Benchmarking:

make NP numéraire and benchmark

$$\hat{S}_t^\delta = \frac{S_t^\delta}{S_t^{\delta*}} \quad - \quad \text{benchmarked portfolio}$$

$$\hat{S}_t^j = \frac{S_t^j}{S_t^{\delta*}} \quad - \quad j\text{th benchmarked security}$$

Existence of NP \Leftrightarrow Supermartingale Property

$$\hat{S}_t^\delta \geq E_t(\hat{S}_s^\delta) \quad (*)$$

for all $0 \leq t \leq s < \infty$ and all nonnegative S_t^δ ,

- The key property of financial market!
- model independent
- makes also in the short term “best” performance of $S^{\delta*}$ precise
- existence of NP, extremely general, Karatzas and Kardaras (2007)

Outperforming Long Term Growth

- *long term growth rate*

$$g^\delta = \limsup_{t \rightarrow \infty} \frac{1}{t} \ln \left(\frac{S_t^\delta}{S_0^\delta} \right)$$

strictly positive portfolio S^δ

pathwise characteristic

Theorem *The NP S^{δ_*} achieves the maximum long term growth rate.*
For any strictly positive portfolio S^δ

$$g^\delta \leq g^{\delta_*}.$$

NP outperforms long term growth pathwise!

- Pathwise outperformance asymptotically over time!
- Ideal long term investment if no constraints!

Proof: Consider strictly positive portfolio S^δ , with $S_0^\delta = S_0^{\delta_*} = x > 0$. By supermartingale property $(*)$ of \hat{S}^δ and Doob (1953), for any $k \in \{1, 2, \dots\}$ and $\varepsilon \in (0, 1)$ one has

$$\exp\{\varepsilon k\} P \left(\sup_{k \leq t < \infty} \hat{S}_t^\delta > \exp\{\varepsilon k\} \right) \leq E_0 \left(\hat{S}_k^\delta \right) \leq \hat{S}_0^\delta = 1.$$

For fixed $\varepsilon \in (0, 1)$

$$\sum_{k=1}^{\infty} P \left(\sup_{k \leq t < \infty} \ln \left(\hat{S}_t^\delta \right) > \varepsilon k \right) \leq \sum_{k=1}^{\infty} \exp\{-\varepsilon k\} < \infty.$$

By Lemma of Borel and Cantelli there exists k_ε such that for all $k \geq k_\varepsilon$ and $t \geq k$

$$\ln \left(\hat{S}_t^\delta \right) \leq \varepsilon k \leq \varepsilon t.$$

Therefore, for all $k > k_\varepsilon$

$$\sup_{t \geq k} \frac{1}{t} \ln (\hat{S}_t^\delta) \leq \varepsilon,$$

which implies

$$\limsup_{t \rightarrow \infty} \frac{1}{t} \ln \left(\frac{S_t^\delta}{S_0^\delta} \right) \leq \limsup_{t \rightarrow \infty} \frac{1}{t} \ln \left(\frac{S_t^{\delta_*}}{S_0^{\delta_*}} \right) + \varepsilon.$$

Since the inequality holds for all $\varepsilon \in (0, 1)$, theorem follows. \square

Best Performance of NP also in the Short Term

Outperforming Expected Benchmarked Returns

$(*) \Rightarrow$ nonnegative \hat{S}^δ is supermartingale and $\hat{S}_t^{\delta*} = 1$

\implies

- expected benchmarked returns

$$E_t \left(\frac{\hat{S}_{t+h}^\delta - \hat{S}_t^\delta}{\hat{S}_t^\delta} \right) \leq E_t \left(\frac{\hat{S}_{t+h}^{\delta*} - \hat{S}_t^{\delta*}}{\hat{S}_t^{\delta*}} \right) = 0$$

$t > 0, h > 0, \hat{S}^\delta$ nonnegative

\implies No strictly positive expected benchmarked returns possible!

Outperforming Expected Growth

- *expected growth*

$$g_{t,h}^\delta = E_t \left(\ln \left(\frac{S_{t+h}^\delta}{S_t^\delta} \right) \right)$$

for $t, h \geq 0$.

invest in strictly positive portfolio S^δ and small fraction $\varepsilon \in (0, \frac{1}{12})$ in some nonnegative portfolio S^δ ,
which yields perturbed portfolio S^{δ_ε}

- **derivative of expected growth** in the direction of S^δ :

$$\frac{\partial g_{t,h}^{\delta_\varepsilon}}{\partial \varepsilon} \Big|_{\varepsilon=0} = \lim_{\varepsilon \rightarrow 0+} \frac{1}{\varepsilon} \left(g_{t,h}^{\delta_\varepsilon} - g_{t,h}^\delta \right)$$

Definition S^δ is called **growth optimal** if

$$\frac{\partial g_{t,h}^{\delta_\varepsilon}}{\partial \varepsilon} \Big|_{\varepsilon=0} \leq 0$$

for all t and $h \geq 0$.

- classical characterization: maximization of expected logarithmic utility from terminal wealth
Kelly (1956), Latané (1959), Breiman (1960), Hakansson (1971a), Thorp (1972), Merton (1973a), Roll (1973) and Markowitz (1976)
equivalent

Theorem

*The NP is growth optimal
if its expected growth is finite.*

NP outperforms expected growth!

Proof: Pl. (2011)

For $t \geq 0$, $h > 0$, and a nonnegative portfolio S^δ we introduce the portfolio ratio $A_t^{\delta,h} = \frac{S_{t+h}^\delta}{S_t^\delta}$, which is set to 1 for $S_t^\delta = 0$. For $\varepsilon \in (0, \frac{1}{2})$; and a nonnegative portfolio S^δ , with $S_t^\delta > 0$, consider the perturbed portfolio S^{δ_ε} with $S_t^{\delta_\varepsilon} = S_t^{\delta_*}$, yielding portfolio ratio $A_{t,h}^{\delta_\varepsilon} = \varepsilon A_{t,h}^\delta + (1 - \varepsilon) A_{t,h}^{\delta_*} > 0$. By the inequality $\ln(x) \leq x - 1$ for $x \geq 0$,

$$G_{t,h}^{\delta_\varepsilon} = \frac{1}{\varepsilon} \ln \left(\frac{A_{t,h}^{\delta_\varepsilon}}{A_{t,h}^{\delta_*}} \right) \leq \frac{1}{\varepsilon} \left(\frac{A_{t,h}^{\delta_\varepsilon}}{A_{t,h}^{\delta_*}} - 1 \right) = \frac{A_{t,h}^\delta - A_{t,h}^{\delta_*}}{A_{t,h}^{\delta_*}} - 1$$

and

$$G_{t,h}^{\delta_\varepsilon} = -\frac{1}{\varepsilon} \ln \left(\frac{A_{t,h}^{\delta_*}}{A_{t,h}^{\delta_\varepsilon}} \right) \geq -\frac{1}{\varepsilon} \left(\frac{A_{t,h}^{\delta_*}}{A_{t,h}^{\delta_\varepsilon}} - 1 \right) = \frac{A_{t,h}^\delta - A_{t,h}^{\delta_*}}{A_{t,h}^{\delta_\varepsilon}}.$$

Because of $A_{t,h}^{\delta_\varepsilon} > 0$ one obtains for $A_{t,h}^\delta - A_{t,h}^{\delta_*} \geq 0$

$$G_{t,h}^{\delta_\varepsilon} \geq 0,$$

and for $A_{t,h}^\delta - A_{t,h}^{\delta_*} < 0$ because of $\varepsilon \in (0, \frac{1}{2})$ and $A_{t,h}^\delta \geq 0$

$$G_{t,h}^{\delta_\varepsilon} \geq -\frac{A_{t,h}^{\delta_*}}{A_{t,h}^{\delta_\varepsilon}} = -\frac{1}{1 - \varepsilon + \varepsilon \frac{A_{t,h}^\delta}{A_{t,h}^{\delta_*}}} \geq -\frac{1}{1 - \varepsilon} \geq -2.$$

Summarizing yields

$$-2 \leq G_{t,h}^{\delta_\varepsilon} \leq \frac{A_{t,h}^\delta}{A_{t,h}^{\delta_*}} - 1,$$

where by $(*)$

$$E_t \left(\frac{A_{t,h}^\delta}{A_{t,h}^{\delta_*}} \right) \leq 1.$$

By the Dominated Convergence Theorem

$$\begin{aligned}
 \frac{\partial g_{t,h}^{\delta_\varepsilon}}{\partial \varepsilon} \Big|_{\varepsilon=0} &= \lim_{\varepsilon \rightarrow 0+} E_t \left(G_{t,h}^{\delta_\varepsilon} \right) = E_t \left(\lim_{\varepsilon \rightarrow 0+} G_{t,h}^{\delta_\varepsilon} \right) \\
 &= E_t \left(\frac{\partial}{\partial \varepsilon} \ln \left(\frac{A_{t,h}^{\delta_\varepsilon}}{A_{t,h}^{\delta_*}} \right) \Big|_{\varepsilon=0} \right) = E_t \left(\frac{A_{t,h}^\delta}{A_{t,h}^{\delta_*}} \right) - 1.
 \end{aligned}$$

This proves that NP S^{δ_*} is growth optimal. □

Systematic Outperformance

Definition A nonnegative portfolio S^δ systematically outperforms a strictly positive portfolio $S^{\tilde{\delta}}$ if

- (i) $S_{t_0}^\delta = S_{t_0}^{\tilde{\delta}}$;
- (ii) $P(S_t^\delta \geq S_t^{\tilde{\delta}}) = 1$ for some $t > 0$ and
- (iii) $P(S_t^\delta > S_t^{\tilde{\delta}}) > 0$.

- also called relative arbitrage in Fernholz & Karatzas (2005)

Theorem *The NP cannot be systematically outperformed by any nonnegative portfolio.*

- NP can systematically outperform other portfolios.

Proof: Consider nonnegative S^δ with $\hat{S}_t^\delta = 1$ where $\hat{S}_s^\delta \geq 1$ almost surely at $s \in [t, \infty)$. By the supermartingale property $(*)$

$$0 \geq E_t \left(\hat{S}_s^\delta - \hat{S}_t^\delta \right) = E_t \left(\hat{S}_s^\delta - 1 \right).$$

Since $\hat{S}_s^\delta \geq 1$ and $E_t(\hat{S}_s^\delta) \leq 1$, it can only follow $\hat{S}_s^\delta = 1$.
 $\implies S_s^\delta = S_s^{\delta_*}$. □

- **NP needs shortest expected time to reach a given wealth level**
Kardaras & Platen (2010)

Strong Arbitrage

- only market participants can exploit arbitrage with their total wealth
- **limited liability**

⇒ nonnegative total wealth of each market participant

Definition A nonnegative portfolio S^δ is a **strong arbitrage** if $S_0^\delta = 0$ and for some $T > 0$

$$P(S_T^\delta > 0) > 0.$$

- Loewenstein & Willard (2000) - same notion of arbitrage through economic arguments
- Pl. (2002)-mathematical motivation through supermartingale property (*)

Theorem *There is no strong arbitrage.*

Proof:

$$(*) \implies$$

\hat{S}^δ nonnegative supermartingale

$$\implies$$

$$0 = \hat{S}_0^\delta \geq E\left(\hat{S}_T^\delta \mid \mathcal{A}_0\right) = E(\hat{S}_T^\delta)$$

$$\implies$$

$$P(S_T^\delta > 0) = P(\hat{S}_T^\delta > 0) = 0. \quad \square$$

Classical notion of arbitrage is too restrictive

- Delbaen & Schachermayer (1998)

no free lunches with vanishing risk \triangleq existence of risk neutral measure
 \triangleq APT

- Loewenstein & Willard (2000)

free snacks & cheap thrills \triangleq some free lunch with vanishing risk

- exploiting weak forms of classical arbitrage requires to allow negative total wealth
- pricing via hedging by avoiding “strong arbitrage” makes no sense since there is no “strong arbitrage” under (*).

How to approximate the NP in the real market?

- diversification leads asymptotically to the NP (Pl. 2005)
- confirm empirically by visualizing diversification effect
- simplest diversification:
equal value weighting = naive diversification

Pl. & Rendek (2012)

- **investment universe:**

market capitalization weighted indices generating the MSCI

- **Industry Classification Benchmark** (Datastream, Reuters ...):

54 countries with country index

10 industry indices

19 supersector indices

41 sector indices

114 country subsector indices

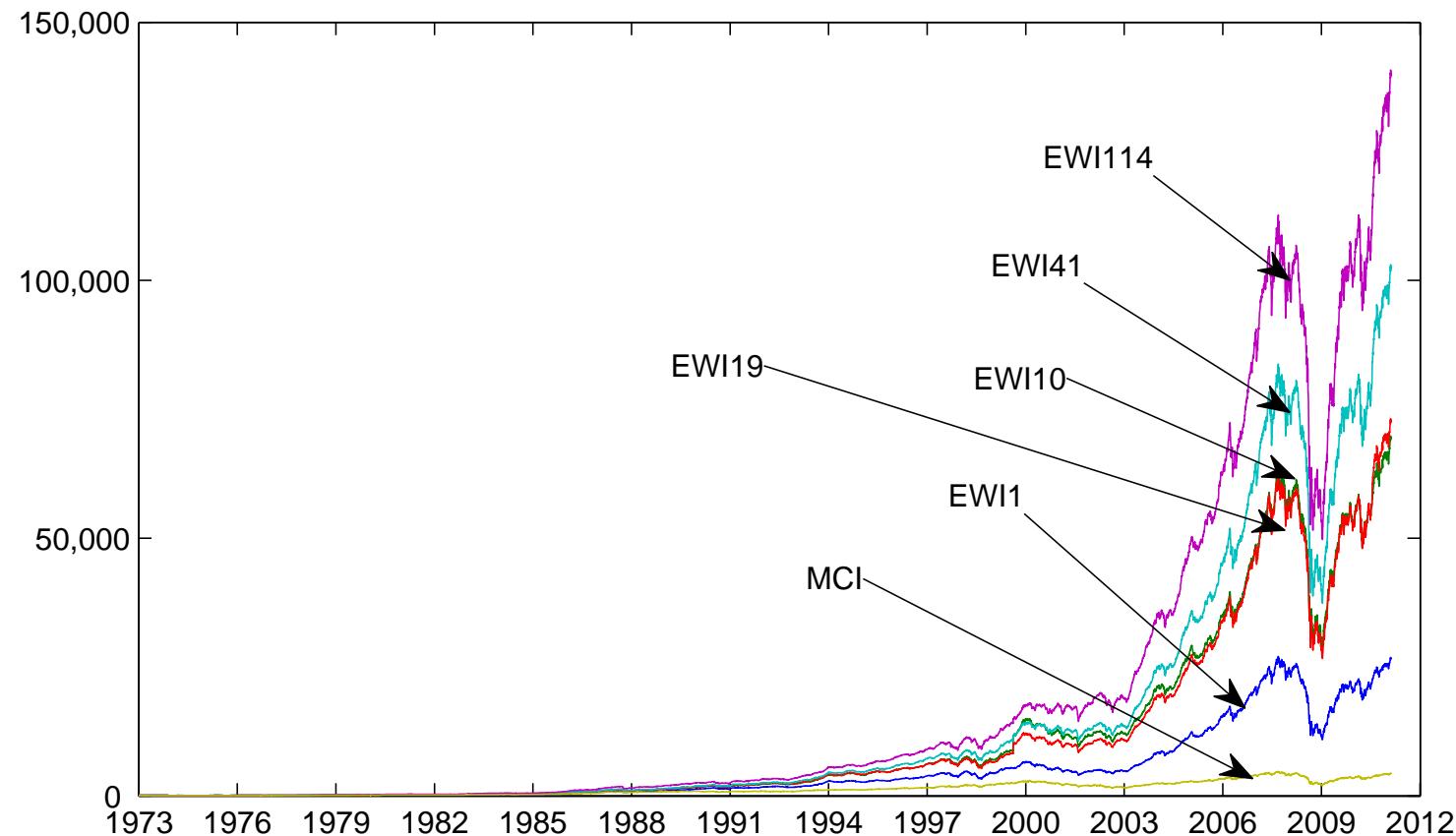
all market capitalization weighted

Equi-Weighted Index (EWId)

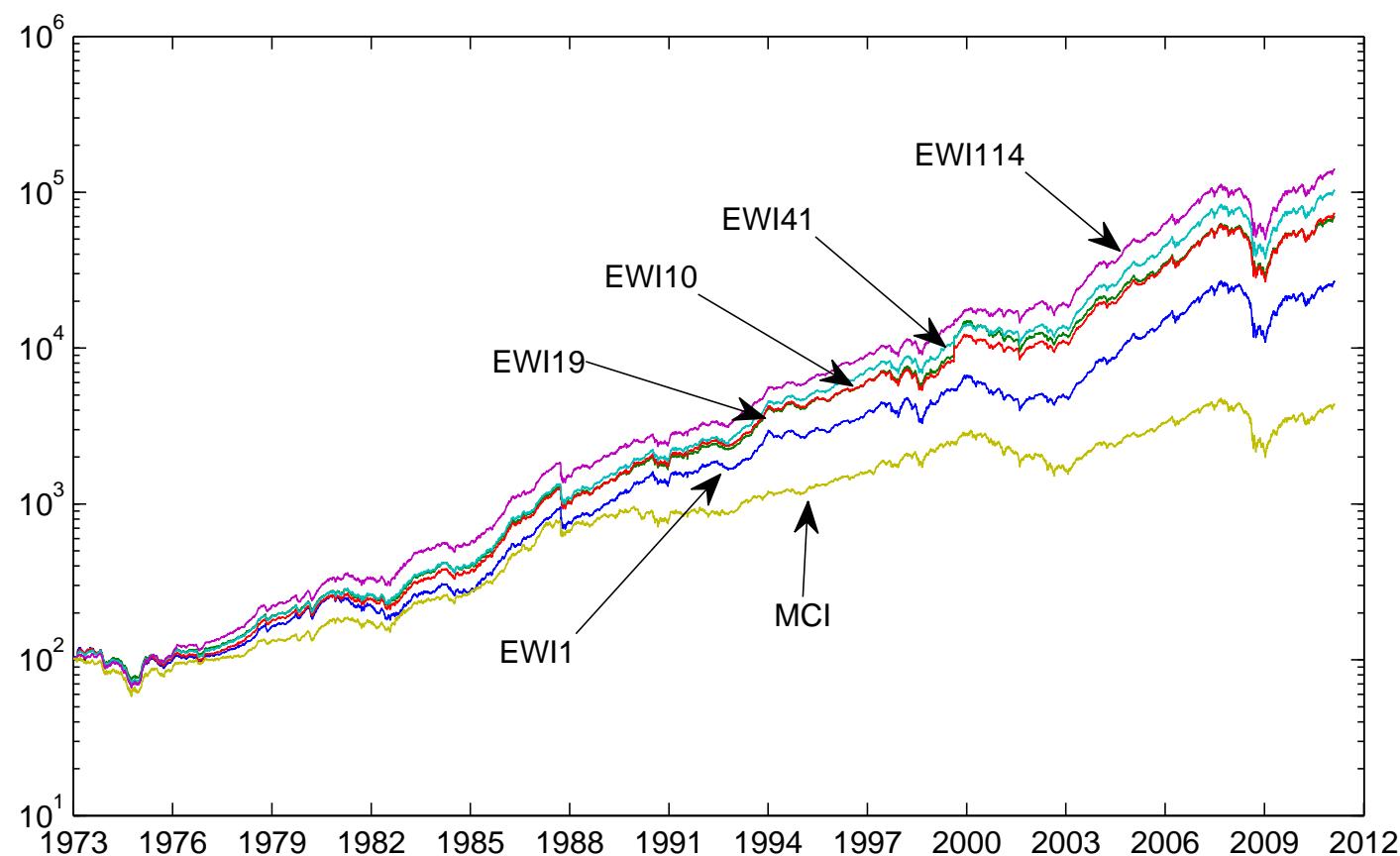
d constituents

$$S_{t_n}^{\delta_{EWId}} = S_{t_{n-1}}^{\delta_{EWId}} \left(1 + \sum_{j=1}^d \sum_{k=1}^{\ell_{d,j}} \pi_{\delta_{EWId}, t_{n-1}}^{j,k} \frac{S_{t_n}^{j,k} - S_{t_{n-1}}^{j,k}}{S_{t_{n-1}}^{j,k}} \right),$$

- Naive diversification over $\ell_{d,j}$ countries
and then over $d = 10, 19, 41, 114$ world industries
- equal value weighting uncertainties of different economic activities



The MCI and five equi-weighted indices: EWI1 (market), EWI10 (industry), EWI19 (supersector), EWI41 (sector), EWI114 (subsector).



The MCI and five equi-weighted indices in log-scale: EWI1 (market), EWI10 (industry), EWI19 (supersector), EWI41 (sector), EWI114 (subsector).

- all based on the same country industry subsectors of the equity market
- all resulting indices mostly driven by nondiversifiable uncertainty
- Empirically: **Better performance through better diversification!**

Illustrating Diversification Phenomenon:

- return of benchmarked constituent over small period $[t, t + h]$

$$R^j = \frac{\hat{S}_{t+h}^j - \hat{S}_t^j}{\hat{S}_t^j}$$

$$j \in \{1, 2, \dots\}$$

Assume for illustration purposes: R^1, R^2, \dots independent, and

$$E_t(R^j) = 0$$

$$E_t((R^j)^2) = \sigma^2 h$$

$$h > 0, j \in \{1, 2, \dots\}$$

- return of portfolio equals
sum of weighted returns of constituents
- return of benchmarked EWI ℓ with ℓ constituents:

$$R^{\delta_{EWI\ell}} = \frac{\hat{S}_{t+h}^{\delta_{EWI\ell}} - \hat{S}_t^{\delta_{EWI\ell}}}{\hat{S}_t^{\delta_{EWI\ell}}} = \frac{1}{\ell} \sum_{j=1}^{\ell} R^j$$

\implies

$$E_t(R^{\delta_{EWI\ell}}) = 0$$

$$E_t((R^{\delta_{EWI\ell}})^2) = \frac{1}{\ell^2} \sum_{j=1}^{\ell} \sigma^2 h = \frac{1}{\ell} \sigma^2 h$$

\implies

$$\lim_{\ell \rightarrow \infty} E_t((R^{\delta_{EWI\ell}})^2) = 0$$

Strong Law of Large Numbers (Kolmogorov)

\Rightarrow

$$\lim_{\ell \rightarrow 0} R^{\delta_{EWI\ell}} = 0$$

almost surely

for $\ell \gg 1$

$$\implies R^{\delta_{EWI\ell}} \approx 0$$

$$\implies \hat{S}_t^{\delta_{EWI\ell}} \approx 1$$

$$\implies S_t^{\delta_{EWI\ell}} = \hat{S}_t^{\delta_{EWI\ell}} S_t^{\delta*} \approx S_t^{\delta*}$$

a.s. for all t

Diversification yields approximation of the NP !

- diversification holds more generally
- model independent phenomenon
- arises naturally
- can be systematically exploited

Naive Diversification Theorem

Pl. & Rendek (2012)

- continuous market, Merton (1972)

- benchmarked constituents

$$\frac{d\hat{S}_t^j}{\hat{S}_t^j} = \sum_k \sigma_t^{j,k} dW_t^k$$

W^1, W^2, \dots - Brownian motions

- fraction of wealth invested

$$\pi_{\delta,t}^j = \frac{\delta_t^j \hat{S}_t^j}{\hat{S}_t^\delta} = \frac{\delta_t^j S_t^j}{S^{\delta_t}}$$

- benchmarked portfolio

$$\frac{d\hat{S}_t^\delta}{\hat{S}_t^\delta} = \sum_j \pi_{\delta,t}^j \frac{d\hat{S}_t^j}{\hat{S}_t^j} = \sum_j \pi_{\delta,t}^j \sum_k \sigma_t^{j,k} dW_t^k$$

- return process of a benchmarked portfolio \hat{S}_t^δ

$$d\hat{Q}_t^\delta = \frac{1}{\hat{S}_t^\delta} d\hat{S}_t^\delta$$

$$t \geq 0 \text{ with } \hat{Q}_0^\delta = 0$$

- Quadratic Variation of Return Process

$$\langle \hat{Q}^{\delta_\ell} \rangle_t = \int_0^t \sum_k \left(\sum_j \pi_{\delta_\ell, s}^j \sigma_s^{j,k} \right)^2 ds$$

- Tracking rate

$$T_{\delta_\ell}(t) = \frac{d}{dt} \langle \hat{Q}^{\delta_\ell} \rangle_t = \sum_k \left(\sum_j \pi_{\delta_\ell, t}^j \sigma_t^{j,k} \right)^2$$

- **benchmarked NP equals constant one**

return process of benchmarked NP equals zero

$$\frac{d\hat{S}_t^{\delta_*}}{\hat{S}_t^{\delta_*}} = \sum_k \sum_j \pi_{\delta_*, t}^j \sigma_t^{j,k} dW_t^k = 0$$

\implies tracking rate

$$T_{\delta_*}(t) = \frac{d}{dt} \langle \hat{Q}^{\delta_*} \rangle_t = 0$$

- **ℓ th equi-weighted index (EWI ℓ)**

$$\pi_{\delta_{EWI\ell}, t}^j = \begin{cases} \frac{1}{\ell} & \text{for } j \in \{1, 2, \dots, \ell\} \\ 0 & \text{otherwise.} \end{cases}$$

Definition:

A sequence $(\hat{S}^{\delta_\ell})_{\ell \in \{1, 2, \dots\}}$ of strictly positive benchmarked portfolios, with initial value equal to one, is called a sequence of **benchmarked approximate NPs** if for each $\varepsilon > 0$ and $t \geq 0$ one has

$$\lim_{\ell \rightarrow \infty} P \left(T_{\delta_\ell}(t) = \frac{d}{dt} \langle \hat{Q}^{\delta_\ell} \rangle_t > \varepsilon \right) = 0.$$

- benchmark captures general, systematic or non-diversifiable market uncertainty
- benchmarked primary security accounts capture specific or idiosyncratic uncertainty
- different types of economic activity yield naturally different specific uncertainties
- specific uncertainty can be diversified

Particular specific uncertainty drives in reality usually only
returns of restricted number of benchmarked primary security accounts:

Definition:

A market is **well-securitized** if there exists $q > 0$ and square integrable $\underline{\sigma}^2 = \{\underline{\sigma}_t^2, t \geq 0\}$ with finite mean for all $\ell, k \in \{1, 2, \dots\}$ and $t \geq 0$

$$\frac{1}{\ell} \left| \sum_{j=1}^{\ell} \sigma_t^{j,k} \right|^2 \leq \frac{1}{\ell^q} \underline{\sigma}_t^2$$

P -almost surely.

- **k -th uncertainty affects not too many constituents.**

Naive Diversification Theorem:

In a well-securitized market the sequence of benchmarked equi-weighted indices is a sequence of benchmarked approximate NPs.

Pl. & Rendek (2012)

- return process of ℓ th benchmarked EWI

$$\hat{Q}_t^{\delta_{EWI\ell}} = \sum_{j=1}^{\ell} \frac{1}{\ell} \sum_k \int_0^t \sigma_s^{j,k} dW_s^k$$

- quadratic variation of return process

$$\left\langle \hat{Q}^{\delta_{EWI\ell}} \right\rangle_t = \frac{1}{\ell} \int_0^t \sum_k \left| \frac{1}{\sqrt{\ell}} \sum_{j=1}^{\ell} \sigma_s^{j,k} \right|^2 ds$$

- tracking rate

$$T_{\delta_{EWI\ell}}(t) = \frac{d}{dt} \left\langle \hat{Q}^{\delta_{EWI\ell}} \right\rangle_t = \frac{1}{\ell} \sum_k \left| \frac{1}{\sqrt{\ell}} \sum_{j=1}^{\ell} \sigma_t^{j,k} \right|^2$$

Lemma Assume that for all $\varepsilon > 0$ and $t \geq 0$

$$\lim_{\ell \rightarrow \infty} P \left(\frac{1}{\ell} \sum_k^{\ell} \left| \frac{1}{\sqrt{\ell}} \sum_{j=1}^{\ell} \sigma_t^{j,k} \right|^2 > \varepsilon \right) = 0,$$

then the sequence of equi-weighted indices is a sequence of benchmarked approximate NPs.

Proof:

For $\varepsilon > 0$ and $t > 0$

$$\begin{aligned} 0 &= \lim_{\ell \rightarrow \infty} P \left(\frac{1}{\ell} \sum_k \left| \frac{1}{\sqrt{\ell}} \sum_{j=1}^{\ell} \sigma_t^{j,k} \right|^2 > \epsilon \right) \\ &= \lim_{\ell \rightarrow \infty} P \left(\frac{d}{dt} \langle \hat{Q}^{\delta_{EWI\ell}} \rangle_t > \epsilon \right) \\ &= \lim_{\ell \rightarrow \infty} P (T_{\delta_{EWI\ell}}(t) > \epsilon) \end{aligned}$$

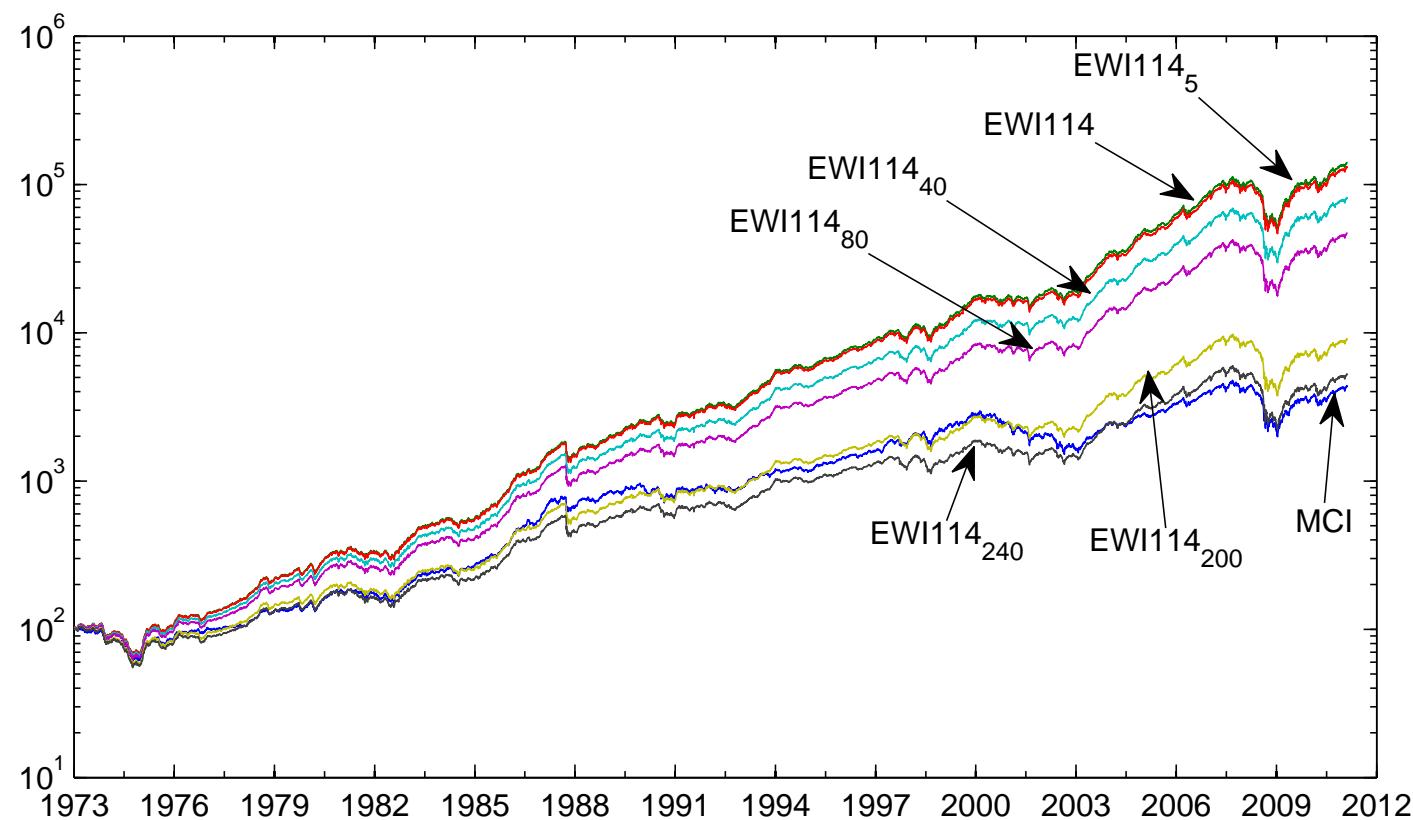
□

Proof of Naive Diversification Theorem:

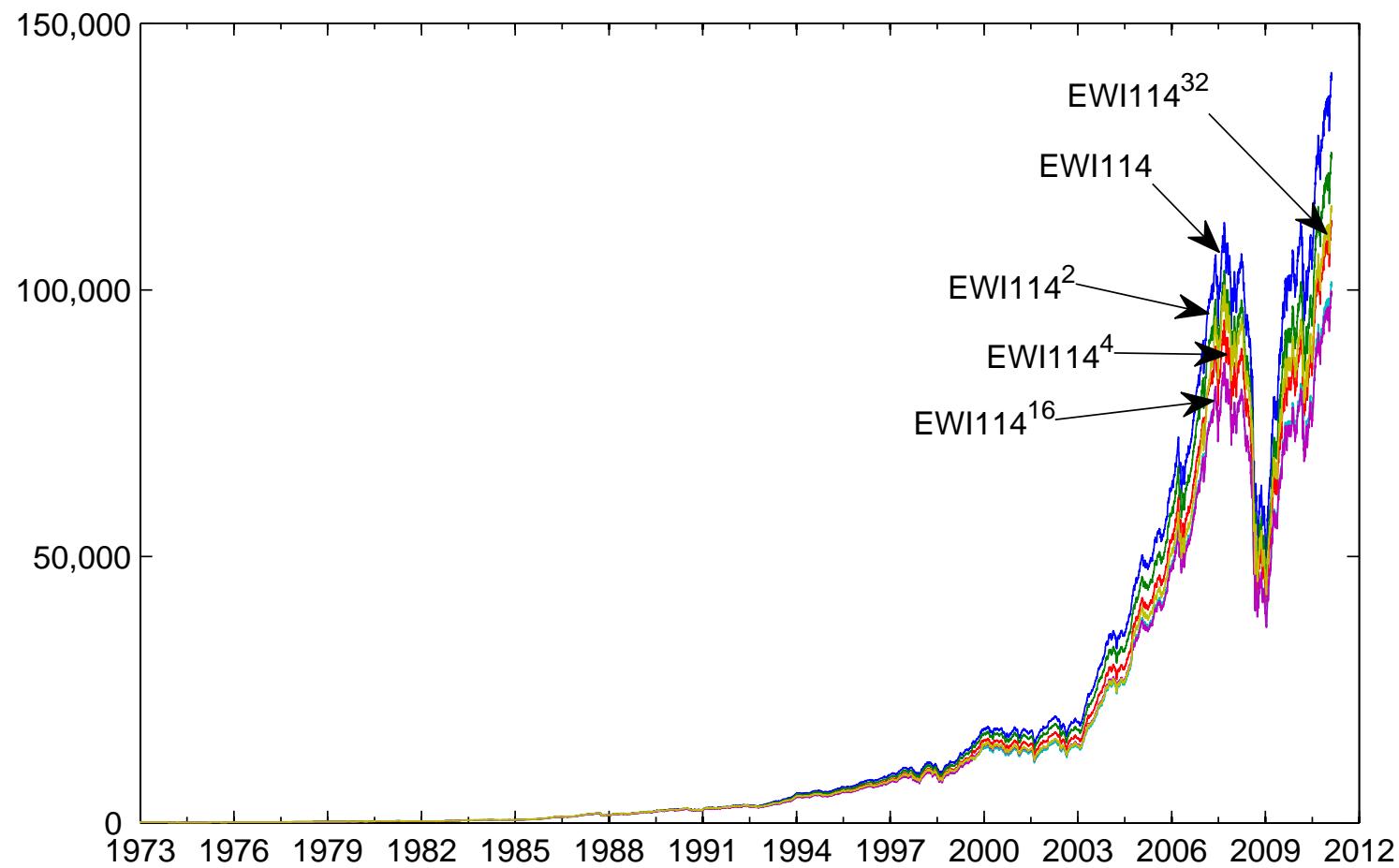
By Markov inequality in a well-securitized market

$$\begin{aligned} & \lim_{\ell \rightarrow \infty} P \left(\frac{1}{\ell} \sum_k \left| \frac{1}{\sqrt{\ell}} \sum_{j=1}^{\ell} \sigma_t^{j,k} \right|^2 > \varepsilon \right) \\ &= \lim_{\ell \rightarrow \infty} P \left(\frac{1}{\ell^q} \underline{\sigma}_t^2 > \varepsilon \right) \leq \lim_{\ell \rightarrow \infty} \frac{1}{\varepsilon} \frac{1}{\ell^q} E(\underline{\sigma}_t^2) = 0, \end{aligned}$$

NDT follows from previous Lemma. □



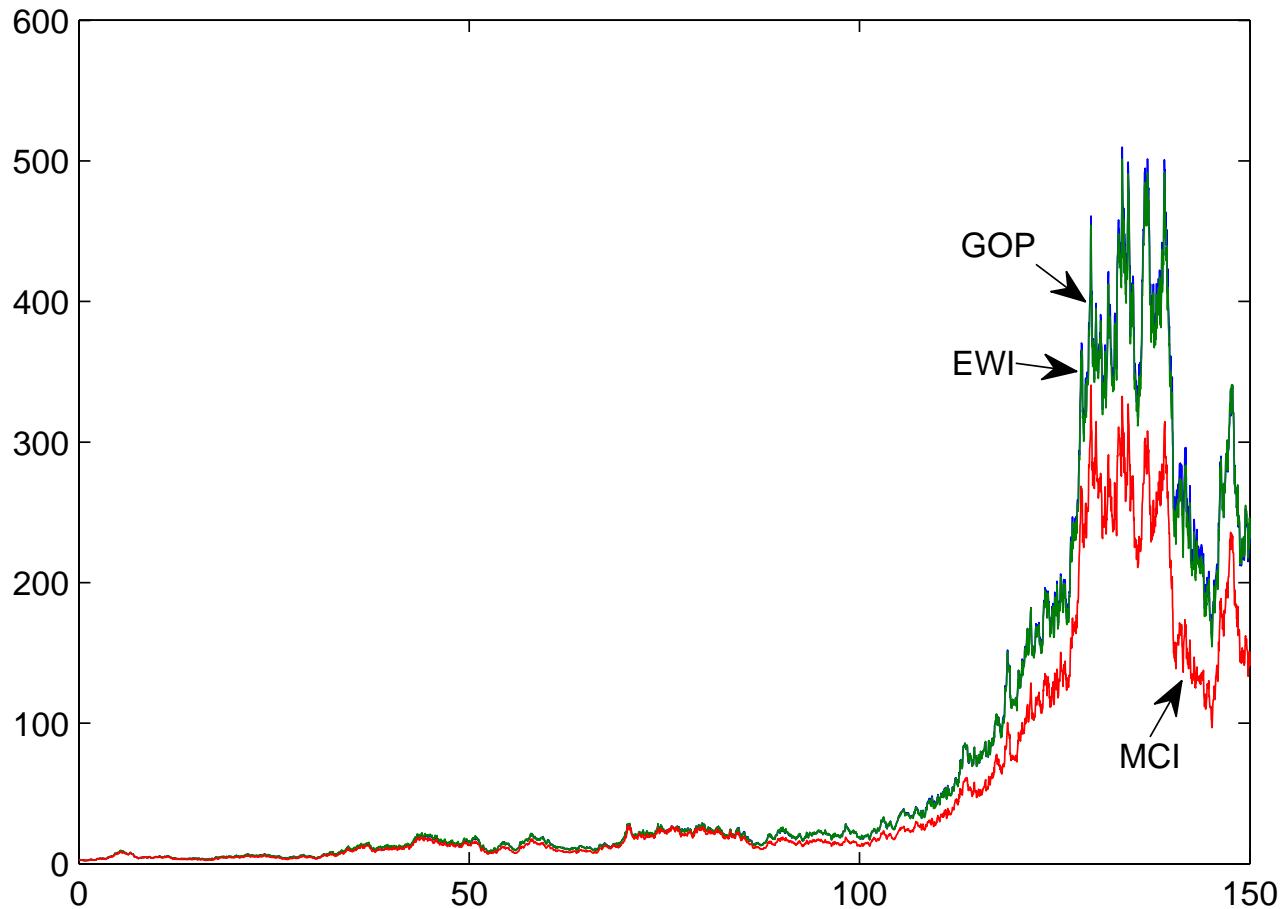
Logarithms of MCI, EWI114 without transaction cost and EWI114_ξ with transaction costs of 5,40,80,200 and 240 basis points.



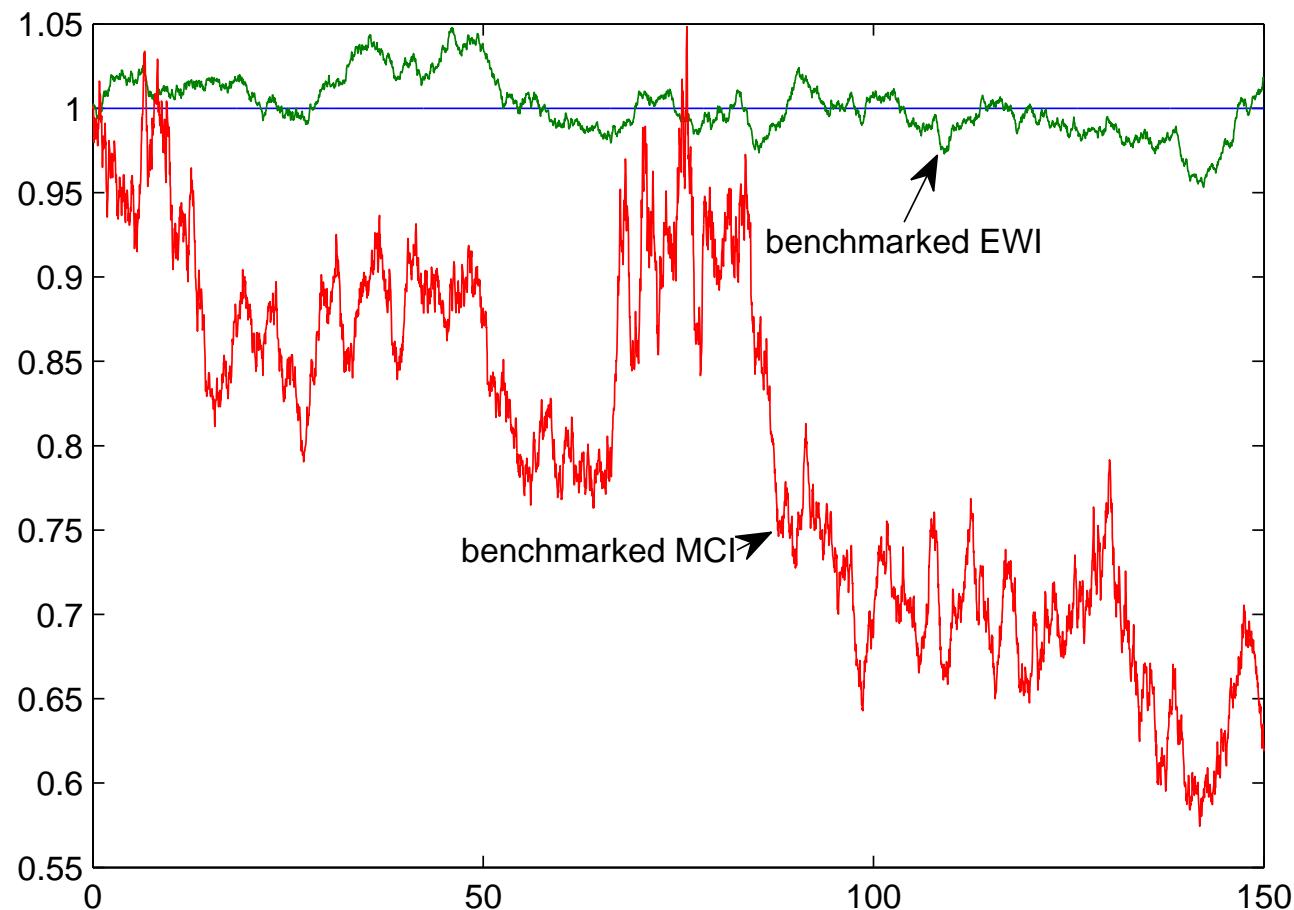
EWI114^{*m*} reallocated daily and every 2, 4, 8, 16 and 32 days.

Transaction cost	0	5	40	80	200	240
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Reallocation terms	1					
Final value	139338.64	130111.93	80543.07	46555.04	8988.23	5194.46
Annualised average return	0.1979	0.1961	0.1834	0.1689	0.1254	0.1109
Annualised volatility	0.1135	0.1135	0.1135	0.1135	0.1134	0.1134
Sharpe ratio	1.4205	1.4046	1.2930	1.1654	0.7822	0.6544
<hr/>						
Reallocation terms	2					
Final value	124542.04	119369.00	88697.63	63166.73	22808.64	16240.40
Annualised average return	0.1949	0.1938	0.1859	0.1770	0.1500	0.1411
Annualised volatility	0.1134	0.1134	0.1134	0.1134	0.1135	0.1136
Sharpe ratio	1.3955	1.3856	1.3163	1.2369	0.9987	0.9193
<hr/>						
Reallocation terms	4					
Final value	111899.82	108230.16	85698.25	65628.82	29467.48	22562.42
Annualised average return	0.1921	0.1912	0.1850	0.1780	0.1568	0.1497
Annualised volatility	0.1135	0.1135	0.1134	0.1134	0.1134	0.1135
Sharpe ratio	1.3699	1.3622	1.3080	1.2459	1.0591	0.9967

Transaction cost	0	5	40	80	200	240
<hr/>						
Reallocation terms	8					
Final value	100505.37	97963.66	81881.57	66705.62	36055.10	29367.02
Annualised average return	0.1892	0.1885	0.1837	0.1783	0.1621	0.1566
Annualised volatility	0.1127	0.1127	0.1127	0.1127	0.1128	0.1128
Sharpe ratio	1.3531	1.3471	1.3051	1.2569	1.1119	1.0634
<hr/>						
Reallocation terms	16					
Final value	98775.24	96892.29	84677.43	72588.14	45711.97	39177.91
Annualised average return	0.1887	0.1882	0.1847	0.1806	0.1684	0.1643
Annualised volatility	0.1130	0.1130	0.1130	0.1130	0.1131	0.1131
Sharpe ratio	1.3463	1.3418	1.3102	1.2740	1.1647	1.1281
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Reallocation terms	32					
Final value	114592.50	112929.09	101939.59	90678.85	63804.84	56744.28
Annualised average return	0.1927	0.1923	0.1896	0.1865	0.1772	0.1741
Annualised volatility	0.1131	0.1131	0.1131	0.1131	0.1133	0.1133
Sharpe ratio	1.3797	1.3763	1.3522	1.3245	1.2408	1.2127



Simulated GOP, EWI and MCI under the Black-Scholes model



Simulated benchmarked GOP, EWI and MCI under the Black-Scholes
model

Example of a Multi-Asset BS Model

- j th primary security account

$$dS_t^j = S_t^j \left[\left(r + \sigma^2 \left(1 + \frac{1}{\sqrt{d}} \right) \right) dt + \frac{\sigma}{\sqrt{d}} \sum_{k=1}^d dW_t^k + \sigma dW_t^j \right]$$

$$t \in [0, T], j \in \{1, 2, \dots, d\}, \sigma > 0$$

general market risk, specific market risk

- **NP fractions**

$$\pi_{\delta_*, t}^j = \left(\sqrt{d} \left(1 + \sqrt{d} \right) \right)^{-1}$$

$$j \in \{1, 2, \dots, d\}$$

$$\pi_{\delta_*, t}^0 = \left(1 + \sqrt{d} \right)^{-1}$$

- **NP SDE**

$$dS_t^{\delta_*} = S_t^{\delta_*} \left((r + \sigma^2) dt + \frac{\sigma}{\sqrt{d}} \sum_{k=1}^d dW_t^k \right)$$

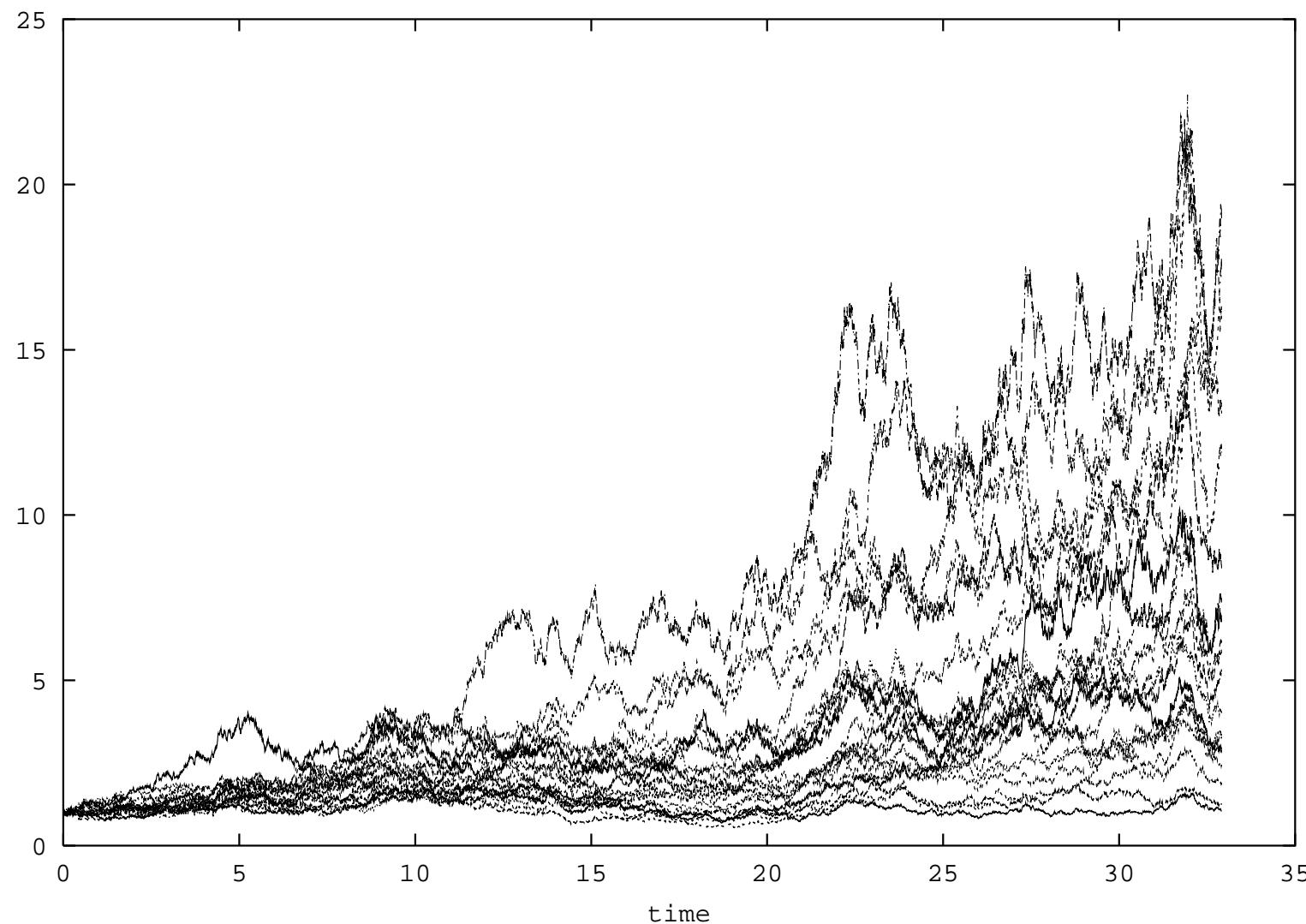
- simulate over $T = 32$ years $d = 50$ risky primary security accounts

$$\sigma = 0.15, r = 0.05$$

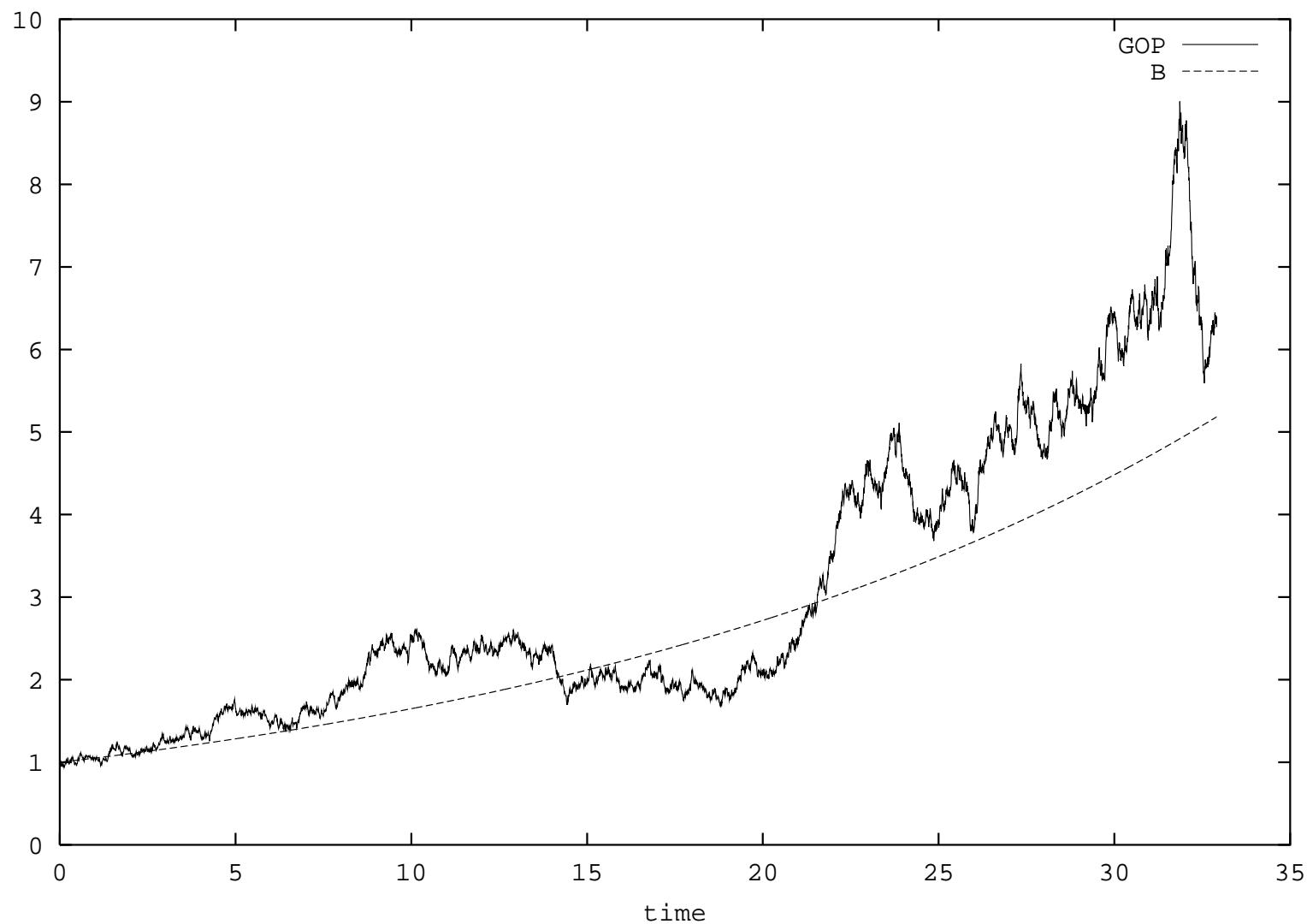
- Benchmarked primary security account
BS - model

$$d\hat{S}_t^j = d \left(\frac{S_t^j}{S_t^{\delta*}} \right) = \hat{S}_t^j \sigma dW_t^j$$

martingale



Simulated primary security accounts.



Simulated GOP and savings account.

Equal Weighted Index

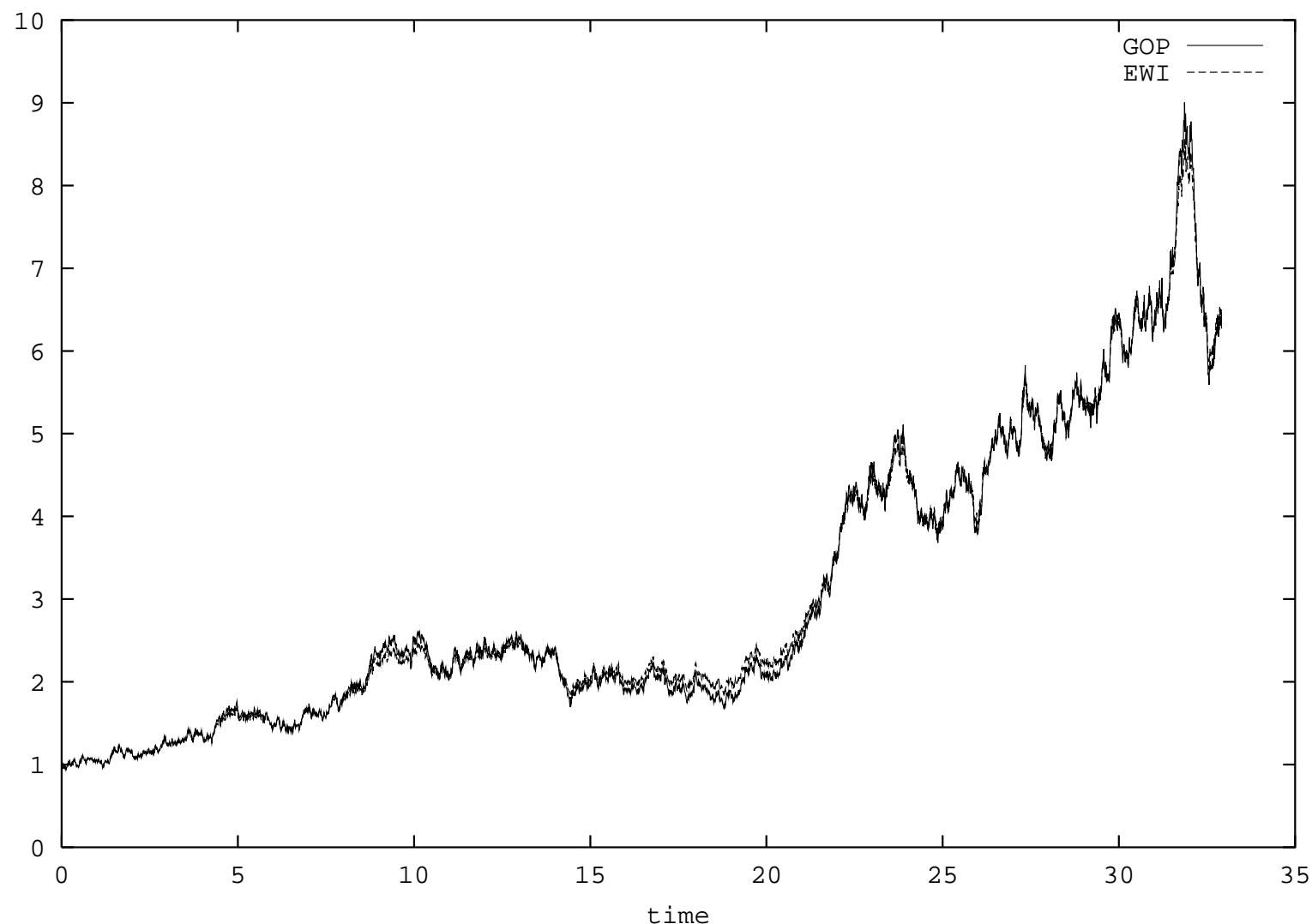
$S^{\delta_{\text{EWId}}}$ holds equal fractions

$$\pi_{\delta_{\text{EWId}}, t}^j = \frac{1}{d+1}$$

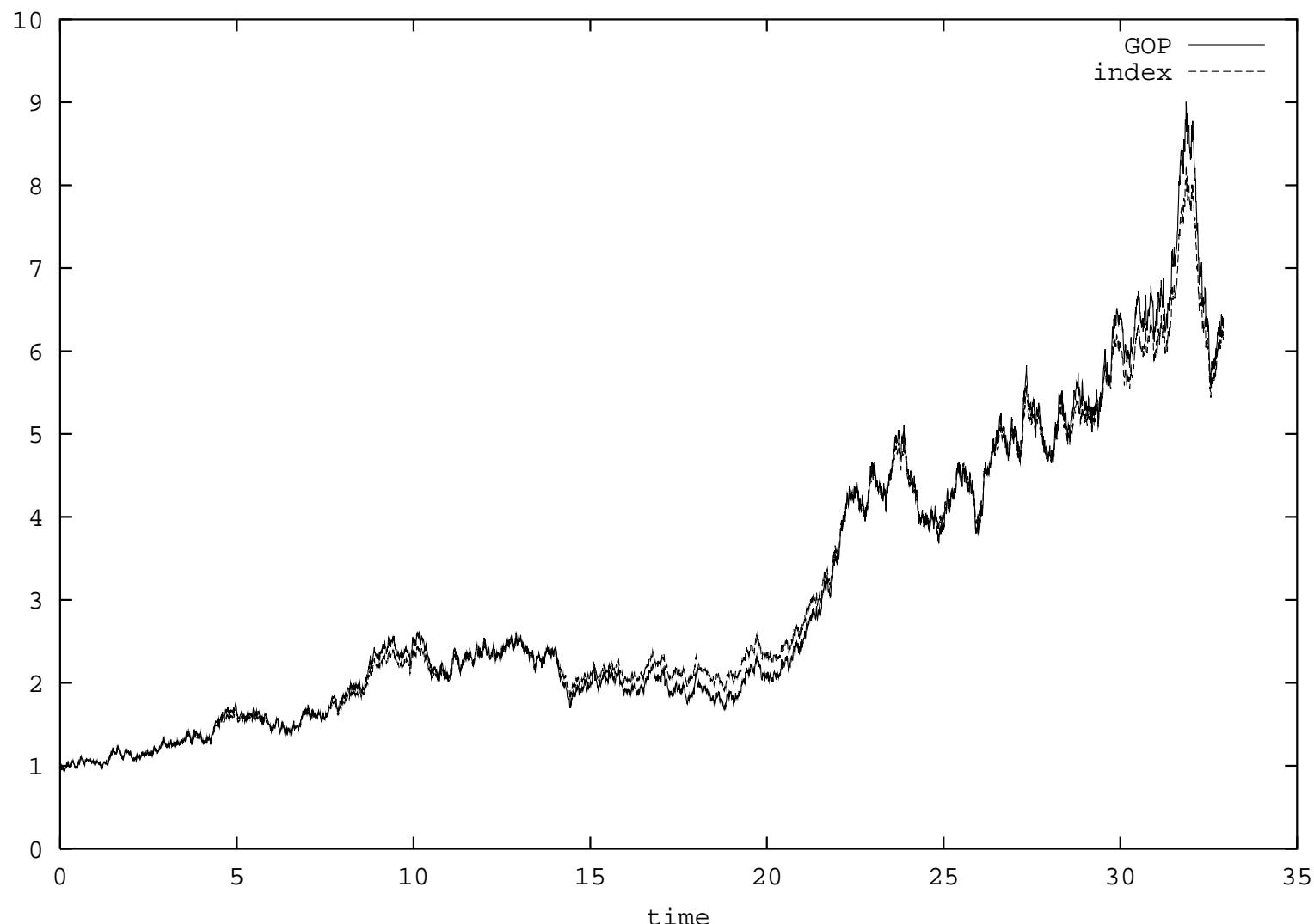
$$j \in \{0, 1, \dots, d\}$$

- tracking rate for BS example:

$$\begin{aligned} T_{\delta_{\text{EWId}}}^d(t) &= \sum_{k=1}^d \left(\frac{\sigma}{d+1} \left(\frac{1}{\sqrt{d}} - 1 \right) \right)^2 \\ &= d \left(\frac{\sigma}{d+1} \left(\frac{1}{\sqrt{d}} - 1 \right) \right)^2 \leq \frac{\sigma^2}{(d+1)} \rightarrow 0 \end{aligned}$$



Simulated NP and EWI.



Simulated diversified accumulation index and NP.

Diversification in an MMM Setting

Pl. (2001), minimal market model (MMM)

- **savings account**

$$S_{(d)}^0(t) = \exp\{r t\}$$

- **discounted GOP drift**

$$\alpha_t^{\delta_*} = \alpha_0 \exp\{\eta t\}$$

- **j th benchmarked primary security account**

$$\hat{S}_{(d)}^j(t) = \frac{1}{Y_t^j \alpha_t^{\delta_*}}$$

- square root process

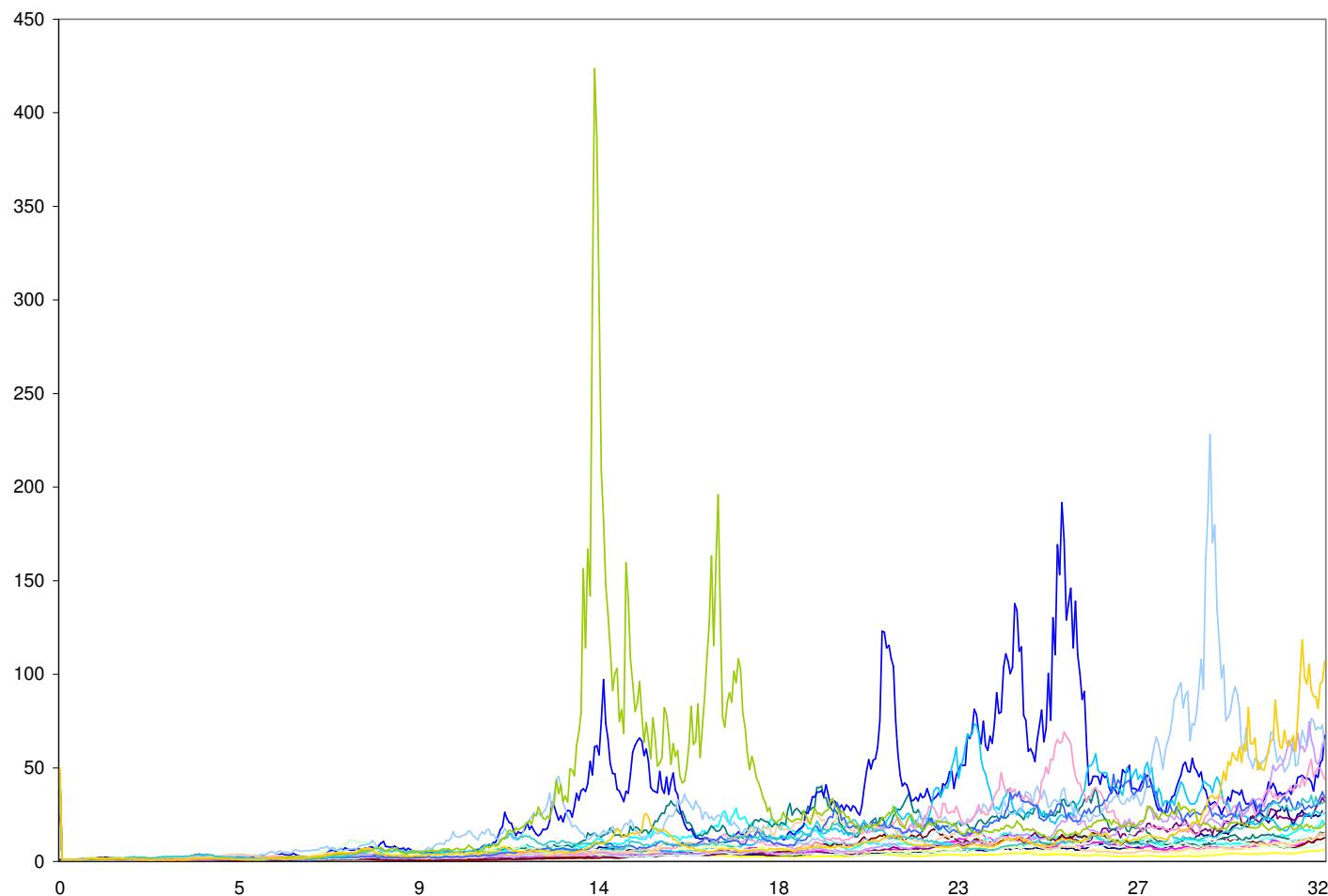
$$dY_t^j = \left(1 - \eta Y_t^j\right) dt + \sqrt{Y_t^j} dW_t^j$$

- NP

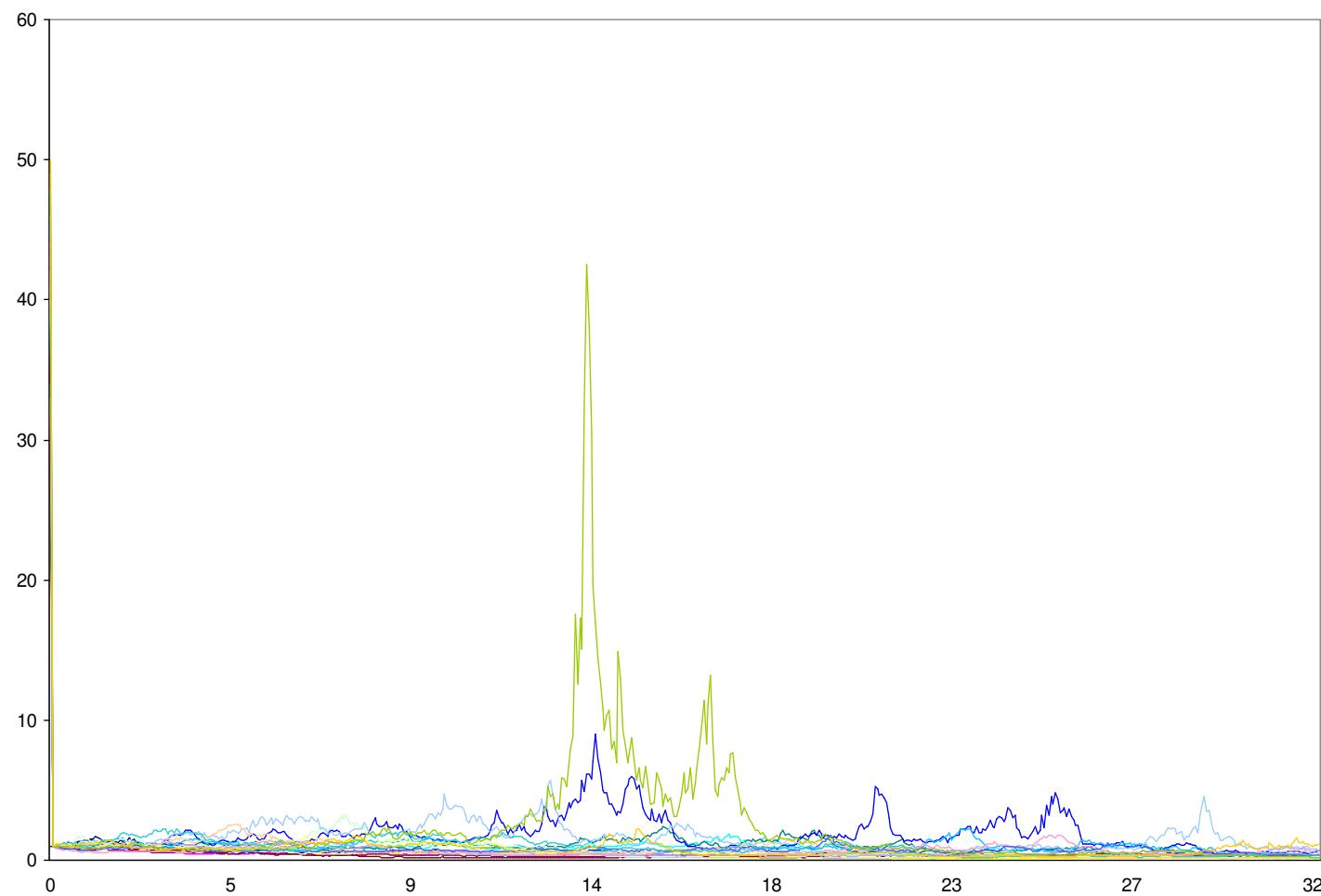
$$S_{(d)}^{\delta_*}(t) = \frac{S_{(d)}^0(t)}{\hat{S}_{(d)}^0(t)}$$

- j th primary security account

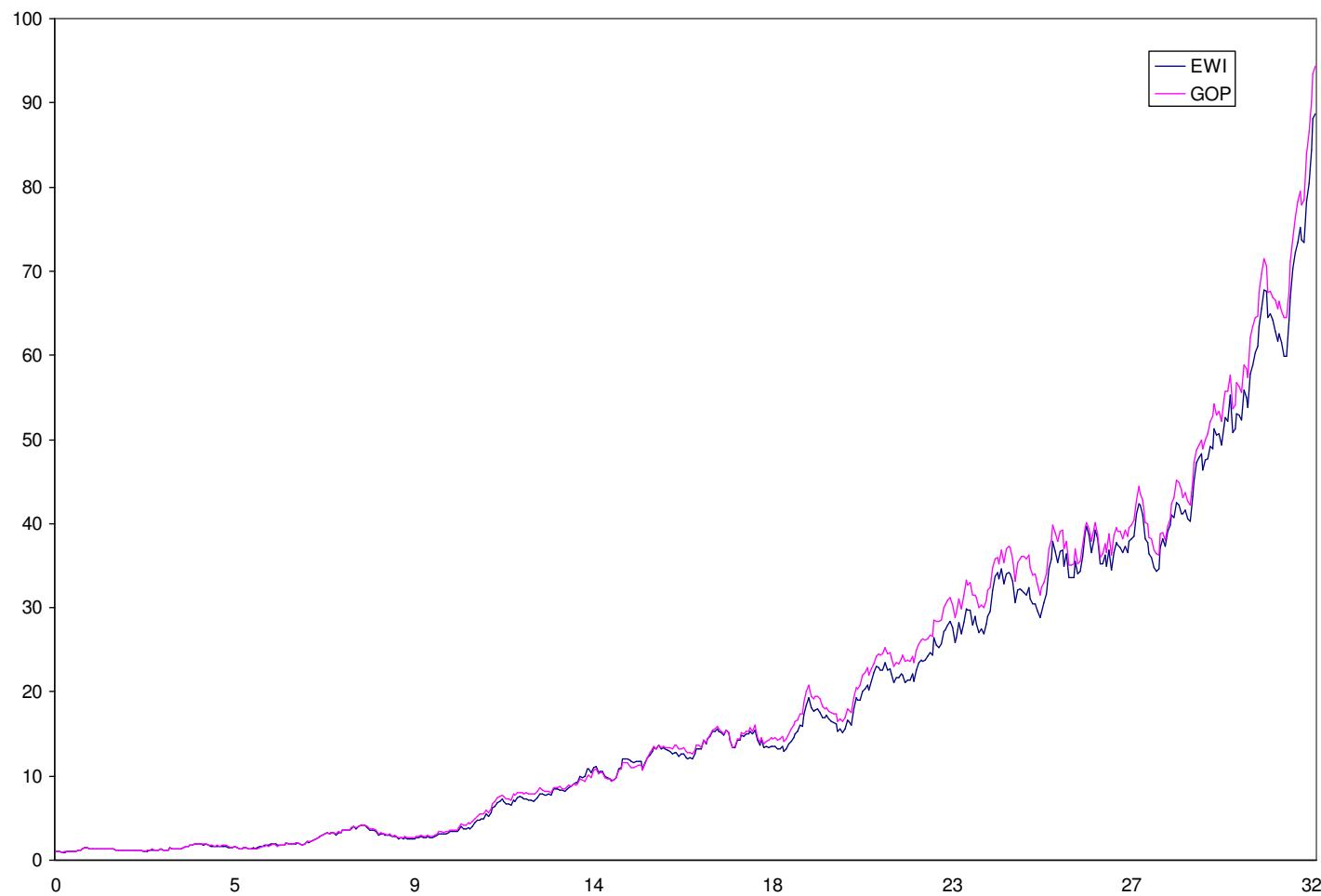
$$S_{(d)}^j(t) = \hat{S}_{(d)}^j(t) S_{(d)}^{\delta_*}(t)$$



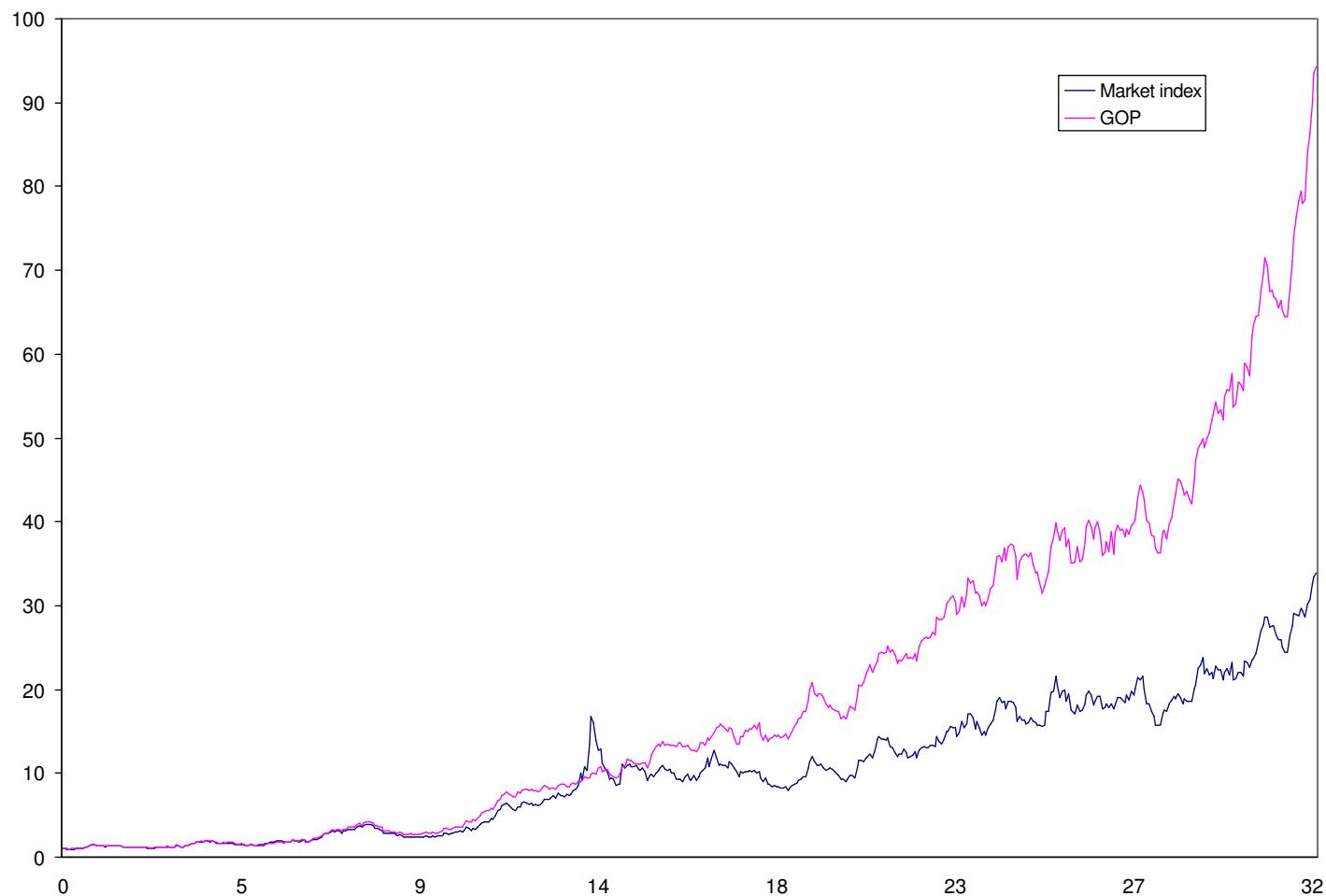
Primary security accounts under the MMM



Benchmarked primary security accounts



NP and EWI



NP and market index

Diversified Portfolios

Fundamental phenomenon of **diversification** leads naturally to **NP** also for more general diversified portfolios than the EWI:

Definition A sequence of strictly positive portfolios $(S^{\delta_d})_{d \in \{1, 2, \dots\}}$ is a sequence of **diversified portfolios** if

$$|\pi_{\delta_d, t}^j| \leq \begin{cases} \frac{K_2}{d^{\frac{1}{2} + \kappa_1}} & \text{for } j \in \{0, 1, \dots, d\} \\ 0 & \text{otherwise} \end{cases}$$

a.s. for $t \in [0, T]$.

- still small fractions but more general than EWIs

- j th benchmarked primary security account

$$\hat{S}_t^j = \frac{S_t^j}{S_t^{\delta_*}}$$

$$d\hat{S}_t^j = -\hat{S}_t^j \sum_k s_t^{j,k} dW_t^k$$

- $s_t^{j,k}$ (j, k) th **specific volatility**
measures **specific market risk** of S^j
with respect to W^k diversifiable uncertainty
- volatility of NP measures **general market risk**
 \triangleq nondiversifiable uncertainty

- **k th total specific volatility**

$$\hat{s}_t^k = \sum_j |s_t^{j,k}|$$

Definition *A market is called **regular** if*

$$E \left((\hat{s}_t^k)^2 \right) \leq K_5$$

for all $t \in [0, \infty)$, $k \in \{1, 2, \dots\}$.

Tracking Rate

$$T_\delta(t) = \frac{d}{dt} \langle \hat{Q}^\delta \rangle_t = \sum_k \left(\sum_{j=0}^d \pi_{\delta,t}^j s_t^{j,k} \right)^2$$

$T_\delta(t)$ - quantifies “distance” between S^δ and S^{δ_*}

$$T_{\delta_*}(t) = 0$$

Approximate NP

Definition *A strictly positive portfolio S^{δ_d} is an **approximate NP** if for all $t \in [0, T]$ and $\varepsilon > 0$*

$$\lim_{d \rightarrow \infty} P(T_{\delta_d}(t) > \varepsilon) = 0.$$

Diversification Theorem (Pl. 2005)

In a regular market

any diversified portfolio is an approximate NP.

- **model independent**

Proof:

$$\begin{aligned}
E(T_{\delta_d}(t)) &\leq \sum_{k=1}^d E\left(\left(\sum_{j=0}^d |\pi_{\delta,t}^j| |s_t^{j,k}|\right)^2\right) \\
&\leq \frac{(K_2)^2}{d^{(1+2K_1)}} \sum_{k=1}^d E\left(\left(\hat{s}_t^k\right)^2\right) \leq \frac{(K_2)^2 K_5}{d^{(1+2K_1)}} d \rightarrow 0
\end{aligned}$$

Markov inequality \implies

$$\lim_{d \rightarrow \infty} P(T_{\delta_d}(t) > \varepsilon) \leq \lim_{d \rightarrow \infty} \frac{1}{\varepsilon} E(T_{\delta_d}(t)) = 0. \quad \square$$

Summary on Diversification

- extremely robust approach to asset management
- does not rely on estimating some expected returns
- diversification can be refined
- diversified market indices are in reality similar in their dynamics
 ⇒ proxies for the NP
 e.g. MSCI, S&P 500, EWI114, EWI142 ...

Continuous Financial Market

- forces us to be more specific

$(\Omega, \mathcal{A}, \underline{\mathcal{A}}, P)$ - filtered probability space

$\underline{\mathcal{A}} = (\mathcal{A}_t)_{t \geq 0}$ - filtration

\mathcal{A}_t - information at time t

$E_t(\cdot) = E(\cdot | \mathcal{A}_t)$ - conditional expectation

- d sources of continuous **traded uncertainty**

$$W^1, W^2, \dots, W^d, \quad d \in \{1, 2, \dots\}$$

independent standard Brownian motions

Primary Security Accounts

S_t^j - j th primary security account at time t ,

$$j \in \{0, 1, \dots, d\}$$

cum-dividend share value or savings account value,

dividends or interest are reinvested

- vector process of primary security accounts

$$S = \{S_t = (S_t^0, \dots, S_t^d)^\top, t \in [0, T]\}$$

- assume **unique strong solution** of SDE

$$dS_t^j = S_t^j \left(a_t^j dt + \sum_{k=1}^d b_t^{j,k} dW_t^k \right)$$

$$t \in [0, \infty), \quad S_0^j > 0, \quad j \in \{1, 2, \dots, d\}$$

- predictable, integrable appreciation rate processes a^j
- predictable, square integrable volatility processes $b^{j,k}$

- **savings account**

$$S_t^0 = \exp \left\{ \int_0^t r_s \, ds \right\}$$

predictable, integrable short rate process

$$r = \{r_t, t \in [0, \infty)\}$$

- Market price of risk

$$\theta_t = (\theta_t^1, \theta_t^2, \dots, \theta_t^d)^\top$$

unique invariant of the market

solution of relation

$$b_t \theta_t = a_t - r_t \mathbf{1}$$

where

$$a_t = (a_t^1, a_t^2, \dots, a_t^d)^\top$$

$$\mathbf{1} = (1, 1, \dots, 1)^\top$$

Assumption:

Volatility matrix $b_t = [b_t^{j,k}]_{j,k=1}^d$ is **invertible**

\implies

- market price of risk

$$\theta_t = b_t^{-1} (a_t - r_t \mathbf{1})$$

- can rewrite j th primary security account SDE

$$dS_t^j = S_t^j \left\{ r_t dt + \sum_{k=1}^d b_t^{j,k} (\theta_t^k dt + dW_t^k) \right\}$$

$$t \in [0, \infty), j \in \{1, 2, \dots, d\}$$

- **strategy**

$$\delta = \{\delta_t = (\delta_t^0, \dots, \delta_t^d)^\top, t \in [0, T]\}$$

predictable and S -integrable,

δ_t^j number of units of j th primary security account

- **portfolio**

$$S_t^\delta = \sum_{j=0}^d \delta_t^j S_t^j$$

- S^δ **self-financing** \iff

$$dS_t^\delta = \sum_{j=0}^d \delta_t^j dS_t^j$$

- changes in the portfolio value only due to gains from trading,
- self-financing in each denomination

- j th fraction

$$\pi_{\delta,t}^j = \delta_t^j \frac{S_t^j}{S_t^\delta}$$

$$j \in \{0,1,\dots,d\}$$

$$\sum_{j=0}^d \pi_{\delta,t}^j = 1$$

- portfolio SDE

$$dS_t^\delta = S_t^\delta \left(r_t dt + \sum_{k=1}^d b_{\delta,t}^k (\theta_t^k dt + dW_t^k) \right)$$

with k th portfolio volatility

$$b_{\delta,t}^k = \sum_{j=0}^d \pi_{\delta,t}^j b_t^{j,k}$$

- **logarithm** of strictly positive portfolio

$$d \ln(S_t^\delta) = g_t^\delta dt + \sum_{k=1}^d b_{\delta,t}^k dW_t^k$$

with **growth rate**

$$g_t^\delta = r_t + \sum_{k=1}^d b_{\delta,t}^k \left(\theta_t^k - \frac{1}{2} b_{\delta,t}^k \right)$$

\implies maximize the growth rate

$$0 = \frac{\partial g_t^\delta}{\partial b_{\delta,t}^k} = \theta_t^k - b_{\delta_*,t}^k$$

\implies for each $k \in \{1, 2, \dots, d\}$

$$\theta_t^k = b_{\delta_*,t}^k = \sum_{\ell=1}^d \pi_{\delta_*,t}^\ell b_t^{\ell,k}$$

\implies

$$\theta_t^\top = \pi_{\delta_*,t}^\top b_t$$

Since b_t invertible \implies

$$\begin{aligned}\pi_{\delta_*, t} &= (\pi_{\delta_*, t}^1, \dots, \pi_{\delta_*, t}^d)^\top \\ &= (b_t^{-1})^\top \theta_t\end{aligned}$$

\implies Kelly portfolio, growth optimal portfolio, GOP,
log-optimal portfolio, NP

$$dS_t^{\delta_*} = S_t^{\delta_*} \left(\left[r_t + \sum_{k=1}^d (\theta_t^k)^2 \right] dt + \sum_{k=1}^d \theta_t^k dW_t^k \right)$$

finite

\implies

- GOP

$$dS_t^{\delta_*} = S_t^{\delta_*} \left(r_t dt + \sum_{k=1}^d \theta_t^k (\theta_t^k dt + dW_t^k) \right)$$

- general portfolio

$$dS_t^\delta = S_t^\delta \left(r_t dt + \sum_{k=1}^d \sum_{j=0}^d \pi_{\delta,t}^j b_t^{j,k} (\theta_t^k dt + dW_t^k) \right)$$

Benchmarked Portfolio

$$\hat{S}_t^\delta = \frac{S_t^\delta}{S_t^{\delta_*}}$$

Itô formula \implies

$$d\hat{S}_t^\delta = -\hat{S}_t^\delta \sum_{k=1}^d \sum_{j=0}^d \pi_{\delta,t}^j (\theta_t^k - b_t^{j,k}) dW_t^k$$

driftless \implies local martingale,

“zero expected return locally”

Local Efficient Market Hypothesis

similar to Fama (1965)

- nonnegative local martingale is supermartingale
 \implies **Kelly portfolio (GOP) is the NP**

- NP **benchmark for investing**
- **numéraire for pricing**
- captures general market risk in **risk management**
- central in **benchmark approach**

Benchmark Approach

- provides much wider framework for modeling than classical APT
- can capture more phenomena relevant in real world
- contains as special cases classical approaches
e.g. APT, actuarial approach, CAPM, ...

Capital Asset Pricing Model

- risk premium of security $V(t)$

$$p_V(t) = \lim_{h \downarrow 0} \frac{1}{h} E_t \left(\frac{V(t+h) - V(t)}{V(t)} \right) - r_t$$

excess expected return

- NP

$$dS_t^{\delta_*} = S_t^{\delta_*} \left((r_t + p_{S^{\delta_*}}(t)) dt + \sum_{k=1}^d \theta_t^k dW_t^k \right)$$

with volatility

$$|\theta_t| = \sqrt{\sum_{k=1}^d (\theta_t^k)^2} = \sqrt{p_{S^{\delta_*}}(t)}$$

- underlying security

$$dS_t = S_t \left(r_t dt + \sum_{k=1}^d \sigma_t^k (\theta_t^k dt + dW_t^k) \right)$$

risk premium

$$\begin{aligned} p_S(t) &= \sum_{k=1}^d \sigma_t^k \theta_t^k \\ &= \frac{d}{dt} \left[\ln(S), \ln(S^{\delta_*}) \right]_t \\ &= \lim_{h \downarrow 0} \frac{1}{h} E_t \left(\left(\frac{S_{t+h} - S_t}{S_t} \right) \left(\frac{S_{t+h}^{\delta_*} - S_t^{\delta_*}}{S_t^{\delta_*}} \right) \right) \end{aligned}$$

- portfolio

$$dS_t^\delta = S_t^\delta \left(r_t dt + \pi_{\delta,t}^1 \sum_{k=1}^d \sigma_t^k (\theta_t^k dt + dW_t^k) \right)$$

risk premium

$$\begin{aligned} p_{S^\delta}(t) &= \pi_{\delta,t}^1 \sum_{k=1}^d \sigma_t^k \theta_t^k \\ &= \lim_{h \downarrow 0} \frac{1}{h} E_t \left(\left(\frac{S_{t+h}^\delta - S_t^\delta}{S_t^\delta} \right) \left(\frac{S_{t+h}^{\delta_*} - S_t^{\delta_*}}{S_t^{\delta_*}} \right) \right) \\ &= \frac{d}{dt} \left[\ln(S^\delta), \ln(S^{\delta_*}) \right]_t \end{aligned}$$

Assumption: *Market portfolio* $S^{\delta_{MP}}$ *is diversified and approximates NP*

$$S_t^{\delta_{MP}} \approx S_t^{\delta_*} .$$

- systematic risk parameter **beta**

$$\beta_{S^\delta}(t) = \frac{\frac{d}{dt} [\ln(S^\delta), \ln(S^{\delta_{\text{MP}}})]_t}{\frac{d}{dt} [\ln(S^{\delta_{\text{MP}}})]_t}$$

\implies

$$\beta_{S^\delta}(t) \approx \frac{\frac{d}{dt} [\ln(S^\delta), \ln(S^{\delta_*})]_t}{\frac{d}{dt} [\ln(S^{\delta_*})]_t} = \frac{p_{S^\delta}(t)}{p_{S^{\delta_*}}(t)} \approx \frac{p_{S^\delta}(t)}{p_{S^{\delta_{\text{MP}}}}(t)}$$

\implies

- fundamental CAPM relation follows directly

$$p_{S^\delta}(t) \approx \beta_{S^\delta}(t) p_{S^{\delta_{\text{MP}}}}(t)$$

- CAPM relationship still holds for $S^{\delta_{\text{MP}}}$ with SDE

$$dS_t^{\delta_{\text{MP}}} = S_t^{\delta_{\text{MP}}} \left(r_t dt + \pi_t \sum_{k=1}^d \theta_t^k (\theta_t^k dt + dW_t^k) \right)$$

combination of savings account and NP, two fund separation

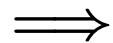
- $$\begin{aligned} \beta_{S^\delta}(t) &= \frac{\frac{d}{dt} [\ln(S^\delta), \ln(S^{\delta_{\text{MP}}})]_t}{\frac{d}{dt} [\ln(S^{\delta_{\text{MP}}})]_t} \\ &= \frac{\pi_{\delta,t}^1 \sum_{k=1}^d \sigma_t^k \theta_t^k \pi_t}{(\pi_t)^2 \sum_{k=1}^d (\theta_t^k)^2} \\ &= \frac{\pi_{\delta,t}^1 \sum_{k=1}^d \sigma_t^k \theta_t^k}{\pi_t \sum_{k=1}^d (\theta_t^k)^2} = \frac{p_{S^\delta}(t)}{p_{S^{\delta_{\text{MP}}}}(t)} \end{aligned}$$

no utility function involved,

no principal agent,

no equilibrium assumption

follows simply by Itô calculus and BA



- **capital asset pricing model (CAPM)**

Sharpe (1964), Lintner (1965),

Mossin (1966) and Merton (1973b)

Examples for Beta:

- savings account

$$\beta_{S^0}(t) = 0$$

- underlying security

$$\beta_S(t) = \frac{\sum_{k=1}^d \sigma_t^k \theta_t^k}{|\theta_t|^2}$$

- NP

$$\beta_{S^{\delta_*}}(t) = 1$$

- portfolio

$$\beta_{S^\delta}(t) = \frac{\pi_{\delta,t}^1 \sum_{k=1}^d \sigma_t^k \theta_t^k}{|\theta_t|^2}$$

- If S^δ involves only diversifiable uncertainty, which is not in NP,
then it has zero beta

\implies

$$\beta_{S^\delta}(t) = 0$$

Portfolio Optimization

- **Kelly portfolio (GOP, NP) best long term investment**

Kelly (1956), Latané (1959), Breiman (1961),
Hakansson (1971b), Thorp (1972)

- **optimal portfolio separated into two funds**

NP and savings account
Tobin (1958), Sharpe (1964)

- **mean-variance efficient portfolio**

Markowitz (1959)

- **intertemporal capital asset pricing model**

Merton (1973a)

Discounted Portfolio

- strictly positive portfolio $S^\delta \in \mathcal{V}^+$

- discounted value

$$\bar{S}_t^\delta = \frac{S_t^\delta}{S_t^0}$$

\implies

- SDE

$$d\bar{S}_t^\delta = \sum_{k=1}^d \psi_{\delta,t}^k (\theta_t^k dt + dW_t^k)$$

with k th diffusion coefficient

$$\psi_{\delta,t}^k = \sum_{j=1}^d \delta_t^j \bar{S}_t^j b_t^{j,k}$$

- discounted drift

$$\alpha_t^\delta = \sum_{k=1}^d \psi_{\delta,t}^k \theta_t^k$$

link to macro economy

- fluctuations

$$\bar{M}_t = \sum_{k=1}^d \int_0^t \psi_{\delta,s}^k dW_s^k$$

\implies

$$\bar{S}_t^\delta = \bar{S}_0^\delta + \int_0^t \alpha_s^\delta ds + \bar{M}_t$$

- aggregate diffusion coefficient (deviation)

$$\gamma_t^\delta = \sqrt{\sum_{k=1}^d (\psi_{\delta,t}^k)^2}$$

\implies aggregate volatility

$$b_t^\delta = \frac{\gamma_t^\delta}{\bar{S}_t^\delta}$$

Locally Optimal Portfolios

Definition

$S^{\tilde{\delta}} \in \mathcal{V}^+$ **locally optimal**, if for all $t \in [0, \infty)$ and $S^\delta \in \mathcal{V}^+$ with

$$\gamma_t^\delta = \gamma_t^{\tilde{\delta}}$$

almost surely:

$$\alpha_t^\delta \leq \alpha_t^{\tilde{\delta}}.$$

generalization of mean-variance optimality

Markowitz (1952, 1959)

Sharpe Ratio

Sharpe (1964, 1966)

- risk premium

$$p_{S^\delta}(t) = \frac{\alpha_t^\delta}{\bar{S}_t^\delta}$$

- aggregate volatility b_t^δ

- Sharpe ratio

$$s_t^\delta = \frac{p_{S^\delta}(t)}{b_t^\delta} = \frac{\alpha_t^\delta}{\gamma_t^\delta} = \frac{\text{mean}}{\text{deviation}}$$

- total market price of risk

$$|\theta_t| = \sqrt{\sum_{k=1}^d (\theta_t^k)^2}$$

Assumption:

$$0 < |\theta_t| < \infty$$

and

$$\pi_{\delta_*, t}^0 \neq 1$$

almost surely.

- **Portfolio Selection Theorem**

For any $S^\delta \in \mathcal{V}^+$ it follows Sharpe ratio

$$s_t^\delta \leq |\theta_t|,$$

where equality when S^δ **locally optimal**

\implies

$$d\bar{S}_t^\delta = \bar{S}_t^\delta \frac{b_t^\delta}{|\theta_t|} \sum_{k=1}^d \theta_t^k (\theta_t^k dt + dW_t^k),$$

\implies fractions

$$\pi_{\delta,t}^j = \frac{b_t^\delta}{|\theta_t|} \pi_{\delta_*,t}^j$$

Markowitz (1959), Sharpe (1964), Merton (1973a)

Khanna & Kulldorff (1999), Platen (2002)

Sharpe ratio maximization

Proof of Portfolio Selection Theorem

Lagrange multiplier λ

consider

$$\mathcal{L}(\psi_\delta^1, \dots, \psi_\delta^d, \lambda) = \sum_{k=1}^d \psi_\delta^k \theta^k + \lambda \left((\gamma^{\tilde{\delta}})^2 - \sum_{k=1}^d (\psi_\delta^k)^2 \right)$$

suppressing time dependence

first-order conditions

$$\frac{\partial \mathcal{L}(\psi_\delta^1, \dots, \psi_\delta^d, \lambda)}{\partial \psi_\delta^k} = \theta^k - 2\lambda \psi_\delta^k = 0$$

for all $k \in \{1, 2, \dots, d\}$ as well as

$$\frac{\partial \mathcal{L}(\psi_\delta^1, \dots, \psi_\delta^d, \lambda)}{\partial \lambda} = \left(\gamma^{\tilde{\delta}}\right)^2 - \sum_{k=1}^d (\psi_\delta^k)^2 = 0$$

\implies locally optimal portfolio $S^{(\tilde{\delta})}$ must satisfy

$$\psi_{\tilde{\delta}}^k = \frac{\theta^k}{2\lambda}$$

for all $k \in \{1, 2, \dots, d\}$.

Then

$$\sum_{k=1}^d (\psi_{\tilde{\delta}}^k)^2 = \left(\gamma^{\tilde{\delta}}\right)^2$$

\implies

$$\left(\gamma^{\tilde{\delta}}\right)^2 = \sum_{k=1}^d \left(\psi_{\tilde{\delta}}^k\right)^2 = \frac{\sum_{k=1}^d (\theta^k)^2}{4 \lambda^2}$$

we have $|\theta| = \sqrt{\sum_{k=1}^d (\theta^k)^2} > 0$,

then

$$\psi_{\tilde{\delta}}^k = \frac{\gamma^{\tilde{\delta}}}{|\theta|} \theta^k$$

for all $k \in \{1, 2, \dots, d\}$

\implies

$$\alpha_t^{\tilde{\delta}} = \gamma_t^{\tilde{\delta}} \frac{|\theta_t|^2}{|\theta_t|} = \gamma_t^{\tilde{\delta}} |\theta_t|$$

$$d\bar{S}_t^{\tilde{\delta}} = \gamma_t^{\tilde{\delta}} \sum_{k=1}^d \frac{\theta_t^k}{|\theta_t|} (\theta_t^k dt + dW_t^k)$$

$$\psi_{\tilde{\delta},t}^k = \sum_{j=1}^d \tilde{\delta}_t^j \bar{S}_t^j b_t^{j,k} = \bar{S}_t^{(\tilde{\delta})} \sum_{j=1}^d \pi_{\tilde{\delta},t}^j b_t^{j,k} = \frac{\gamma_t^{\tilde{\delta}}}{|\theta_t|} \theta_t^k = \bar{S}_t^{(\tilde{\delta})} b_t^{\tilde{\delta}} \frac{\theta_t^k}{|\theta_t|}$$

invertibility of volatility matrix

\implies

$$\pi_{\tilde{\delta},t}^j = \frac{b_t^{\tilde{\delta}}}{|\theta_t|} \sum_{k=1}^d \theta_t^k b_t^{-1,j,k} \quad \square$$

Two Fund Separation

- locally optimal portfolio

fraction of wealth in the NP

$$\frac{b_t^\delta}{|\theta_t|} = \frac{1 - \pi_{\delta,t}^0}{1 - \pi_{\delta_*,t}^0}$$

remainder in savings account

$$\pi_{\delta,t}^0 = 1 - \frac{b_t^\delta}{|\theta_t|} (1 - \pi_{\delta_*,t}^0)$$

also known as **fractional Kelly strategy**:

Kelly (1956), Latané (1959), Thorp (1972), Hakansson & Ziemba (1995)

Risk Aversion Coefficient

$$J_t^\delta = \frac{1 - \pi_{\delta^*, t}^0}{1 - \pi_{\delta, t}^0} = \frac{|\theta_t|}{b_t^\delta}$$

Pratt (1964), Arrow (1965)

- discounted locally optimal portfolio

$$d\bar{S}_t^\delta = \bar{S}_t^\delta \frac{1}{J_t^\delta} |\theta_t| (|\theta_t| dt + dW_t),$$

where

$$dW_t = \sum_{k=1}^d \frac{\theta_t^k}{|\theta_t|} dW_t^k$$

$\frac{1}{J_t^\delta}$ fraction in the NP

Capital Market Line

- expected rate of return

$$a_t^\delta = r_t + p_{S^\delta}(t)$$

- for a locally optimal portfolio

$$a_t^\delta = r_t + |\theta_t| b_t^\delta$$

- portfolio process at the capital market line

has fraction $\frac{1}{J_t^\delta} = \frac{b_t^\delta}{|\theta_t|}$ in NP

fractional Kelly strategy

Markowitz Efficient Frontier

- aggregate volatility

$$b_t^\delta = \frac{1 - \pi_{\delta,t}^0}{1 - \pi_{\delta^*,t}^0} |\theta_t|$$

Definition

Markowitz efficient portfolio has expected rate of return on efficient frontier if

$$a_t^\delta = r_t + \sqrt{(b_t^\delta)^2} |\theta_t|.$$

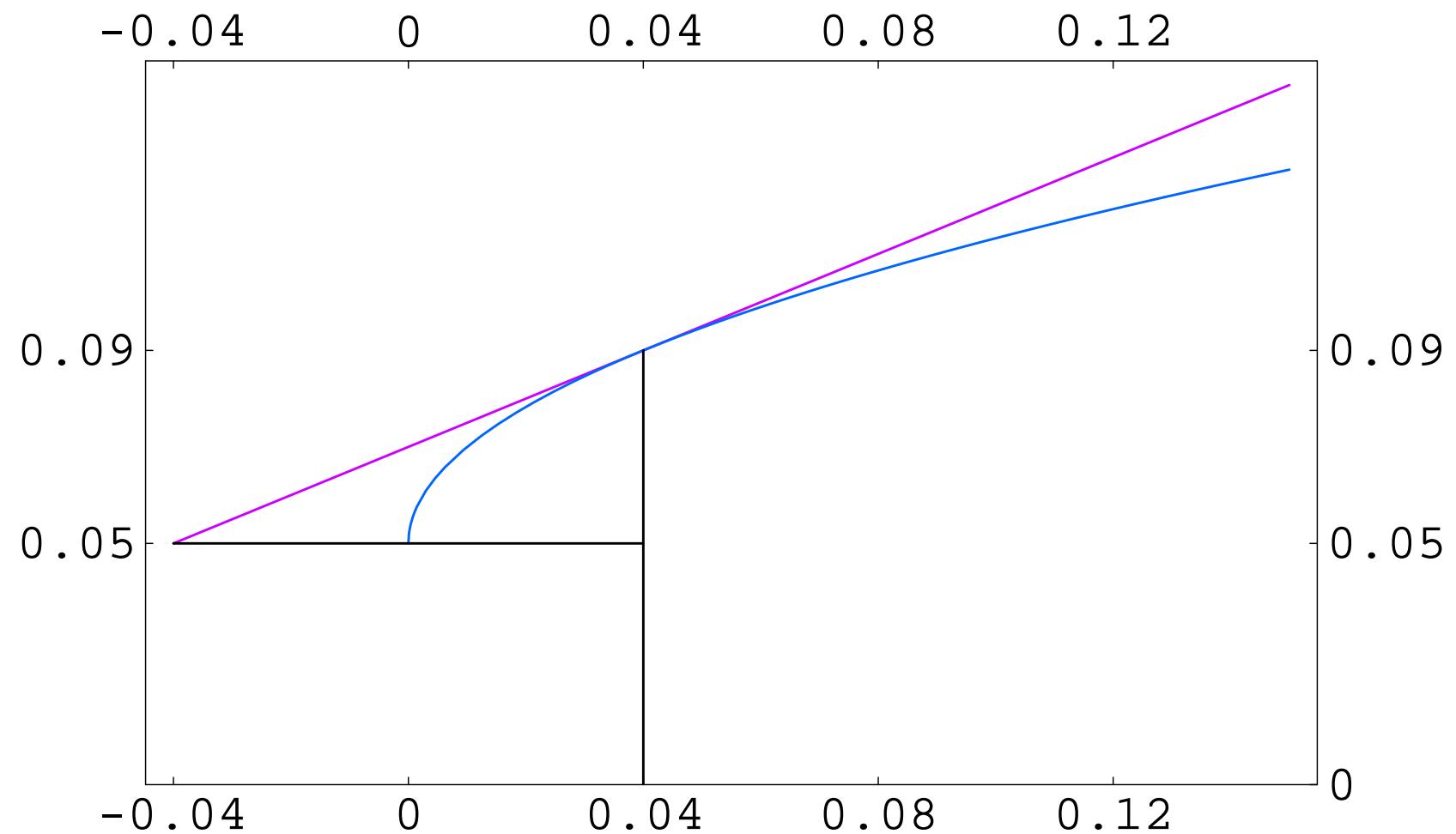
Theorem *Any locally optimal portfolio S^δ is Markowitz efficient portfolio.*

$$p_{S^\delta}(t) = \frac{1 - \pi_{\delta,t}^0}{1 - \pi_{\delta_*,t}^0} |\theta_t|^2 = b_t^\delta |\theta_t|$$

Sharpe ratio maximization

risk attitude expressed via $J_t^\delta = \frac{|\theta_t|}{b_t^\delta}$

short term view



Efficient frontier

Efficient Growth Rates

- efficient growth rate frontier

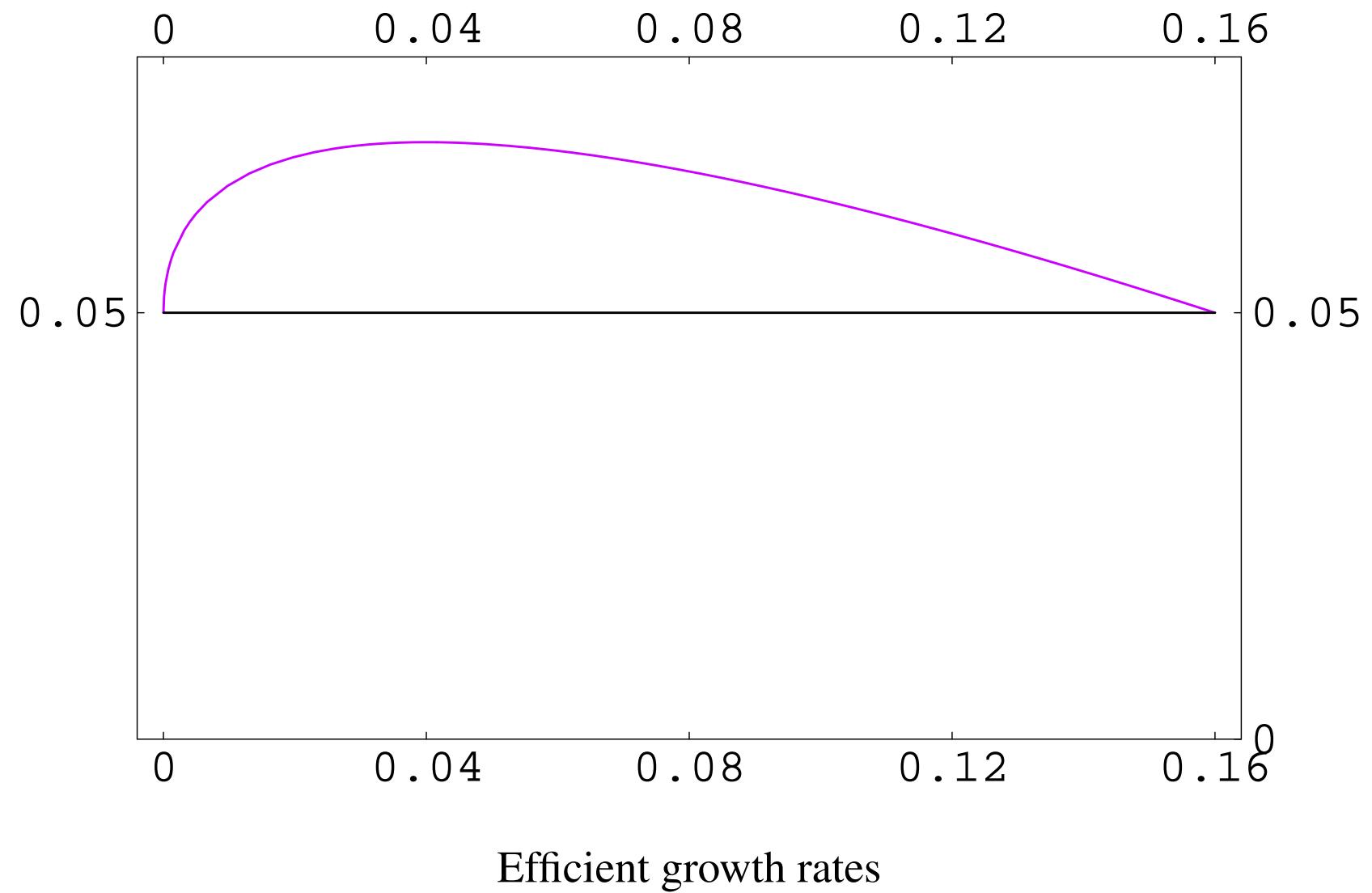
of locally optimal portfolio

$$g_t^\delta = r_t + \sqrt{|b_t^\delta|^2} |\theta_t| - \frac{1}{2} |b_t^\delta|^2 = r_t + \frac{|\theta_t|^2}{J_t^\delta} \left(1 - \frac{1}{2 J_t^\delta} \right)$$

- maximum

$$g_t^{\delta*} = r_t + \frac{1}{2} |\theta_t|^2$$

- long term view



Market Portfolio

$$S_t^{\delta_{\text{MP}}} = \sum_{\ell=1}^n S_t^{\delta_\ell}$$

portfolio of tradable wealth of ℓ th investor S^{δ_ℓ}

Remark:

If each investor forms a nonnegative, locally optimal portfolio with her or his total tradable wealth, then **MP is locally optimal** .

$$\begin{aligned}
d\bar{S}_t^{\delta_{\text{MP}}} &= \sum_{\ell=1}^n d\bar{S}_t^{\delta_\ell} \\
&= \sum_{\ell=1}^n \frac{\left(\bar{S}_t^{\delta_\ell} - \delta_\ell^0\right)}{\left(1 - \pi_{\delta_*, t}^0\right)} \sum_{k=1}^d \theta_t^k (\theta_t^k dt + dW_t^k) \\
&= \bar{S}_t^{\delta_{\text{MP}}} \frac{\left(1 - \pi_{\delta_{\text{MP}}, t}^0\right)}{\left(1 - \pi_{\delta_*, t}^0\right)} \sum_{k=1}^d \theta_t^k (\theta_t^k dt + dW_t^k)
\end{aligned}$$

- CAPM relationship still holds when MP locally optimal

Practical Feasibility of Sample Based Markowitz Mean-Variance Approach:

- NP theoretically known for given model
- Markowitz (1952)
- Best & Grauer (1991)
- DeMiguel, Garlappi & Uppal (2009)

practical application

requires for 50 assets observation window of about 500 years

- not realistic to be applied
- theoretically can be reconciled under strong assumptions
(MP locally optimal)

Maximum Drawdown Constrained Portfolios

\mathcal{X} - set of nonnegative continuous discounted portfolios

- **running maximum**

for $X = \{X_t, t \geq 0\} \in \mathcal{X}$

$$X_t^* = \sup_{u \in [0, t]} X_u$$

- **drawdown**

$$X_t^* - X_t$$

- relative drawdown

$$\frac{X_t}{X_t^*}$$

- maximum relative drawdown
- express attitude towards risk by restricting to $X \in {}^\alpha \mathcal{X}$, where

$$\frac{X_t}{X_t^*} \geq \alpha, \alpha \in [0, 1)$$

pathwise criterion

- **drawdown**

for $X \in {}^\alpha \mathcal{X}$

$$X_t \geq \alpha X_t^*$$

\implies

$$X_t^* - X_t \leq (1 - \alpha)X_t^*$$

- **maximum drawdown constrained portfolio**

$$\alpha \in [0, 1), X \in \mathcal{X}$$

$$\begin{aligned} {}^\alpha X_t &= \alpha(X_t^*)^{1-\alpha} + (1 - \alpha)X_t(X_t^*)^{-\alpha} \\ &= 1 + \int_0^t (1 - \alpha)(X_s^*)^{-\alpha} dX_s \in [\alpha X_t^*, (X_t^*)^{1-\alpha}] \end{aligned}$$

Grossman & Zhou (1993), Cvitanic & Karatzas (1994),
 Kardaras, Obloj & Pl. (2012)

\implies SDE

$$\frac{d^\alpha X_t}{{}^\alpha X_t} = {}^\alpha \pi_t \frac{dX_t}{X_t}$$

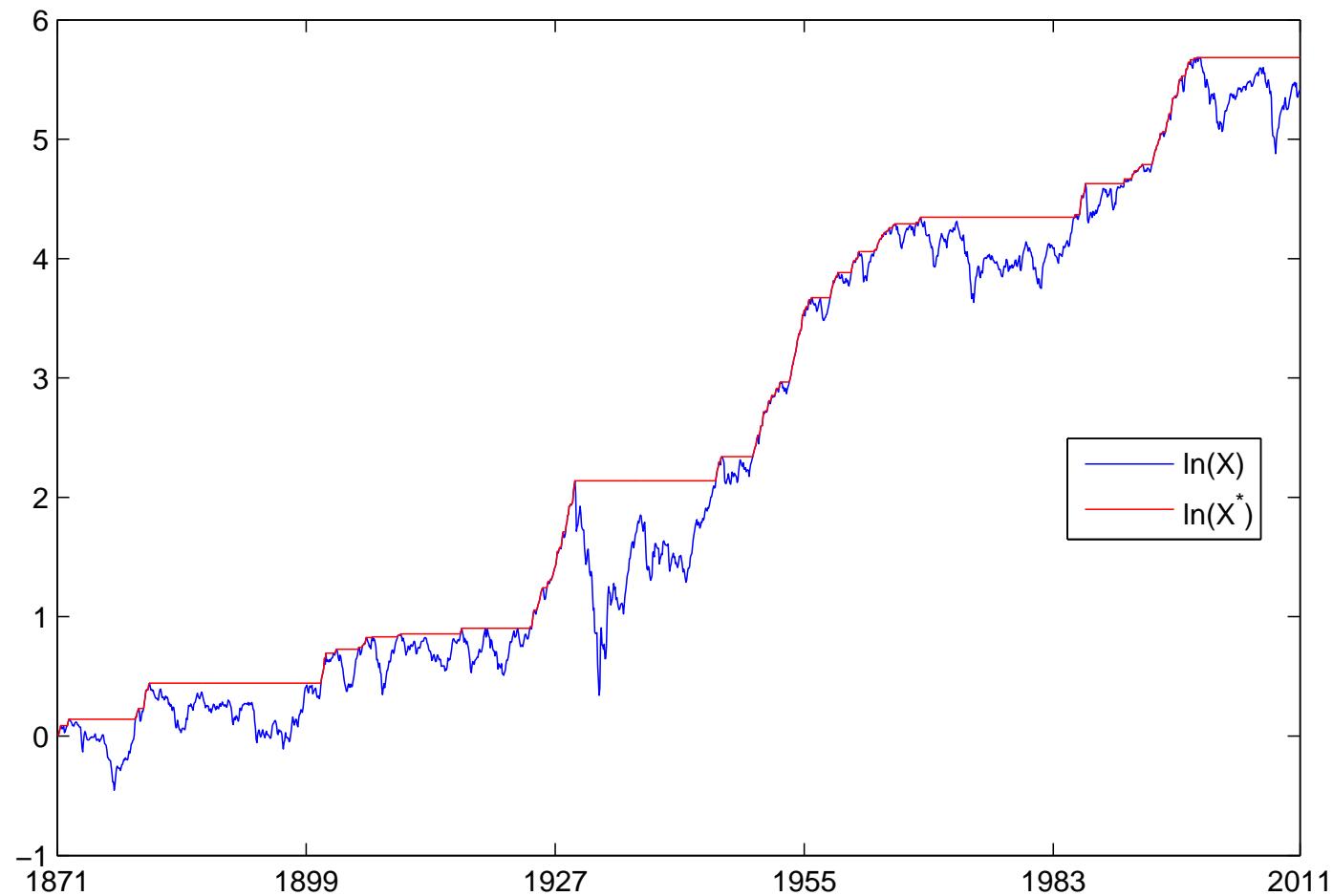
fraction

$${}^\alpha \pi_t = 1 - \frac{(1 - \alpha) \frac{X_t}{X_t^*}}{\alpha + (1 - \alpha) \frac{X_t}{X_t^*}}$$

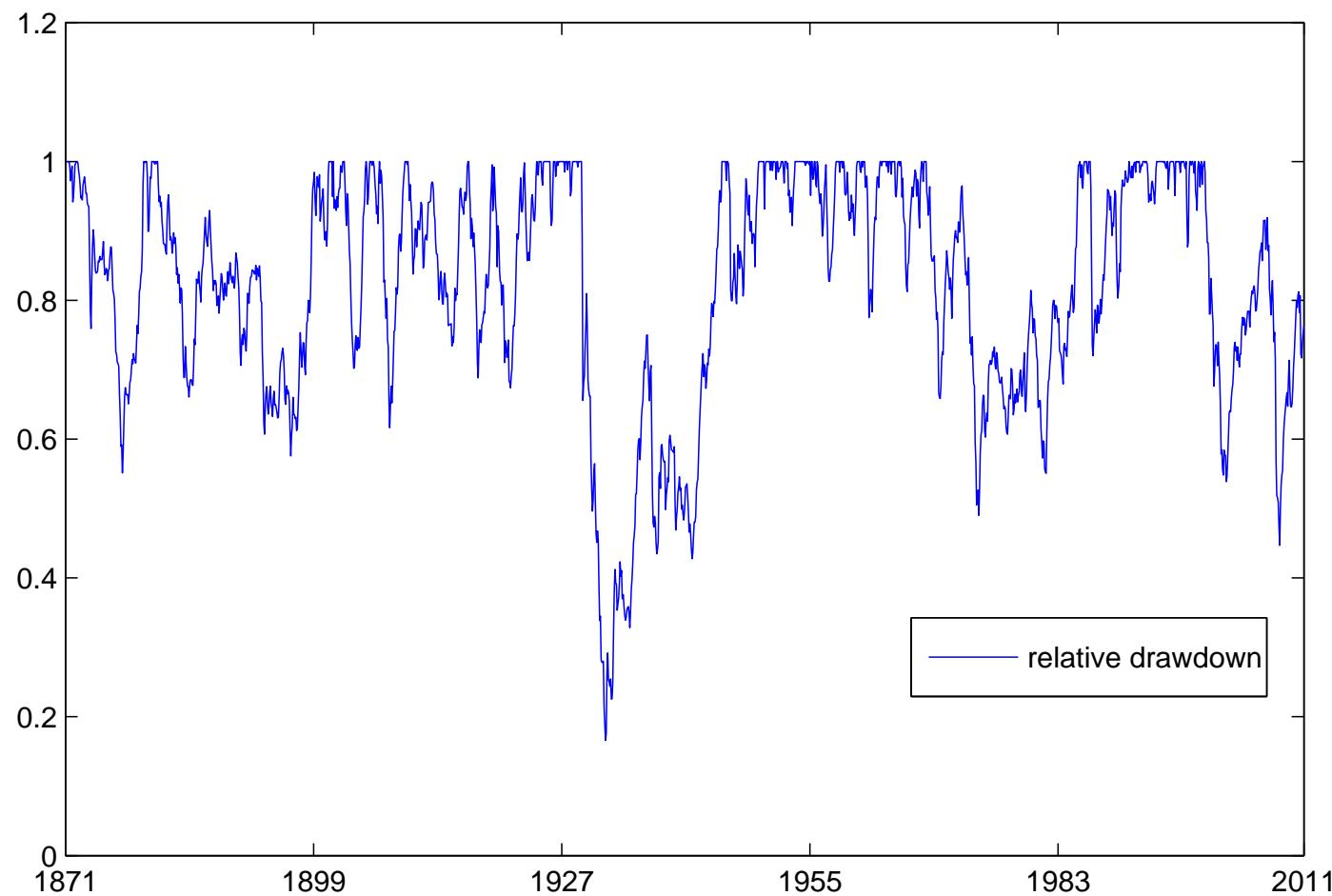
model independent

depends only on $\frac{X_t}{X_t^*}$ and α

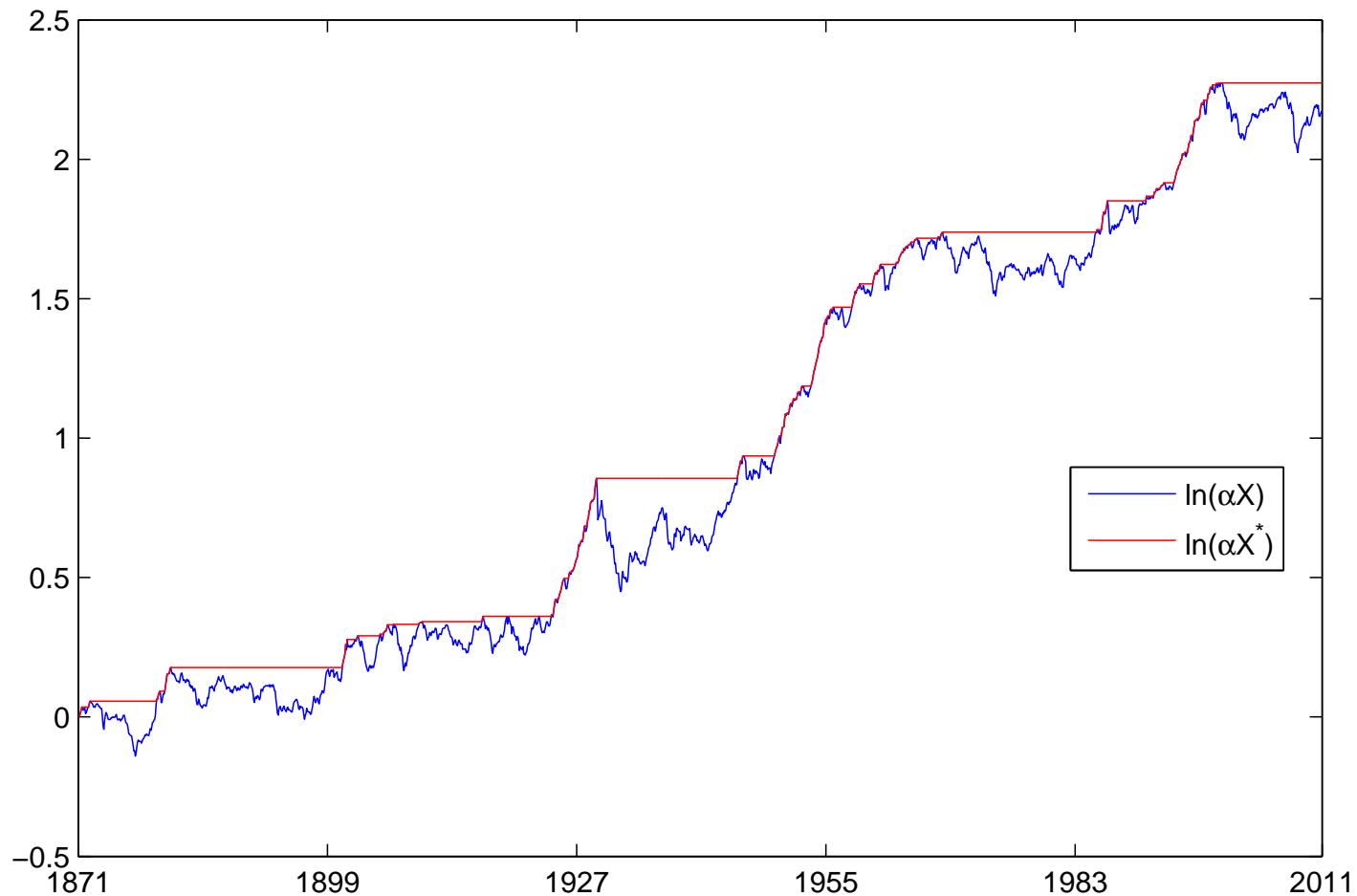
${}^\alpha \bar{S}_t^{\delta*}$ - locally optimal portfolio



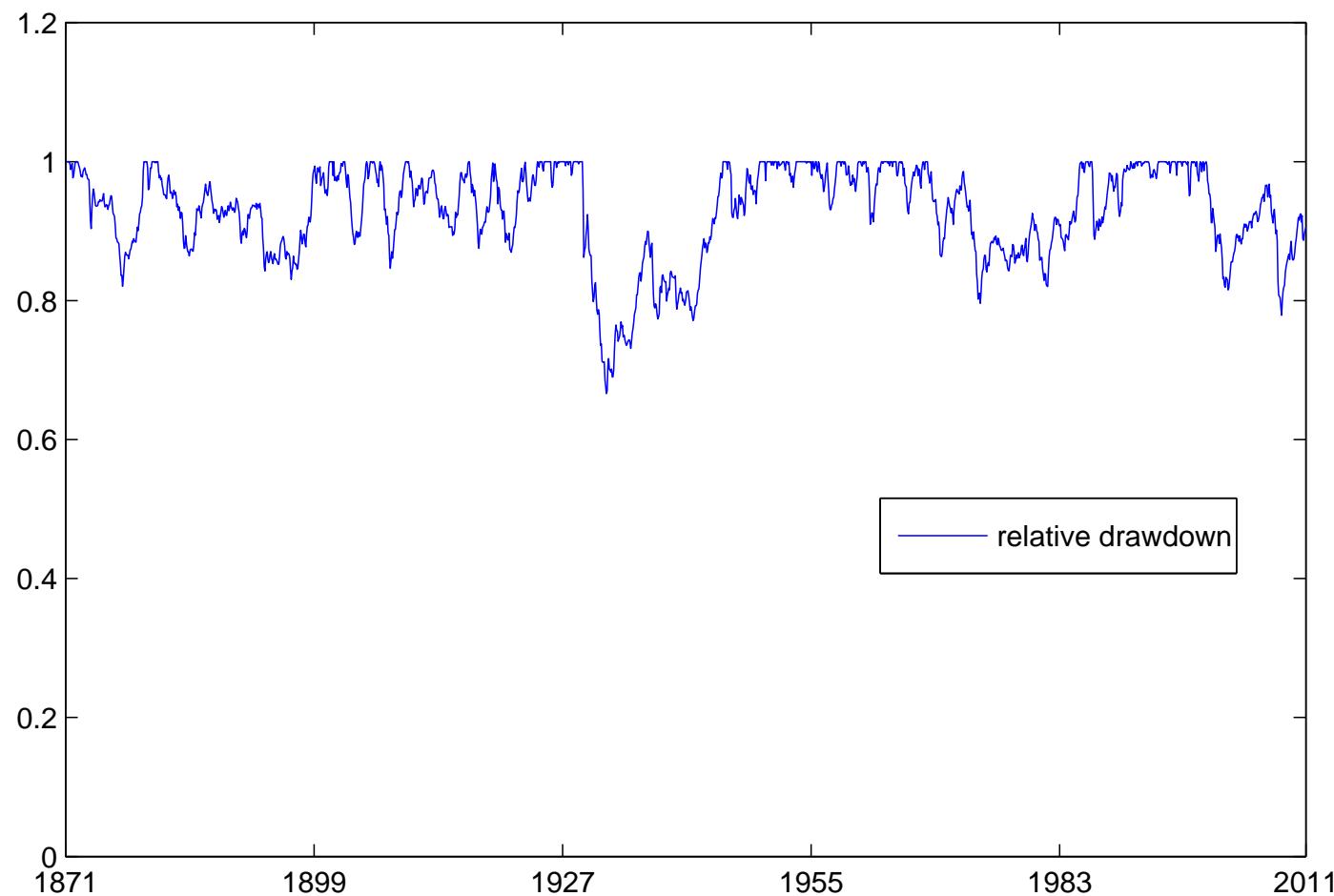
logarithm of discounted S&P500 and its running maximum



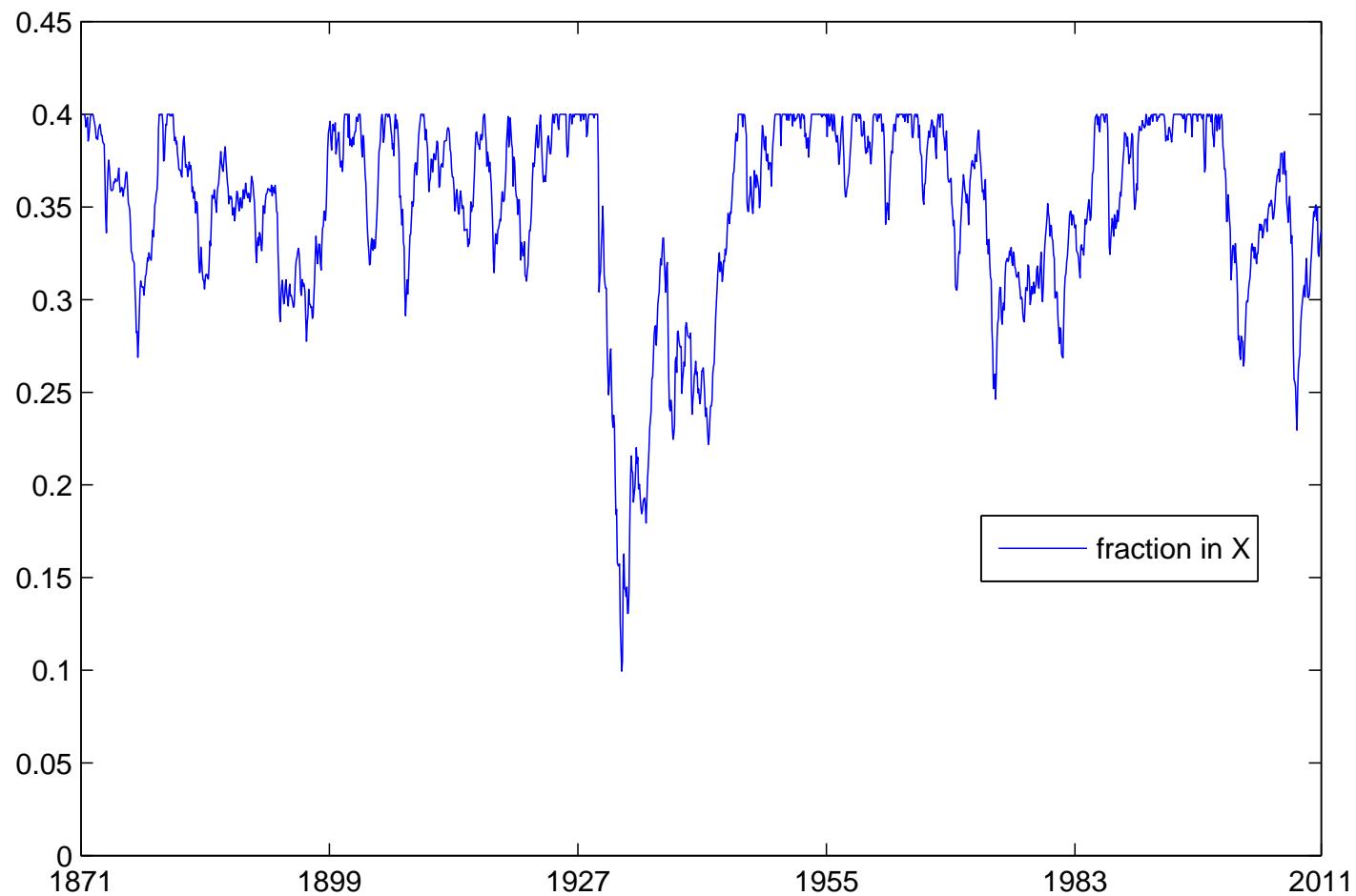
Relative drawdown of discounted S&P500



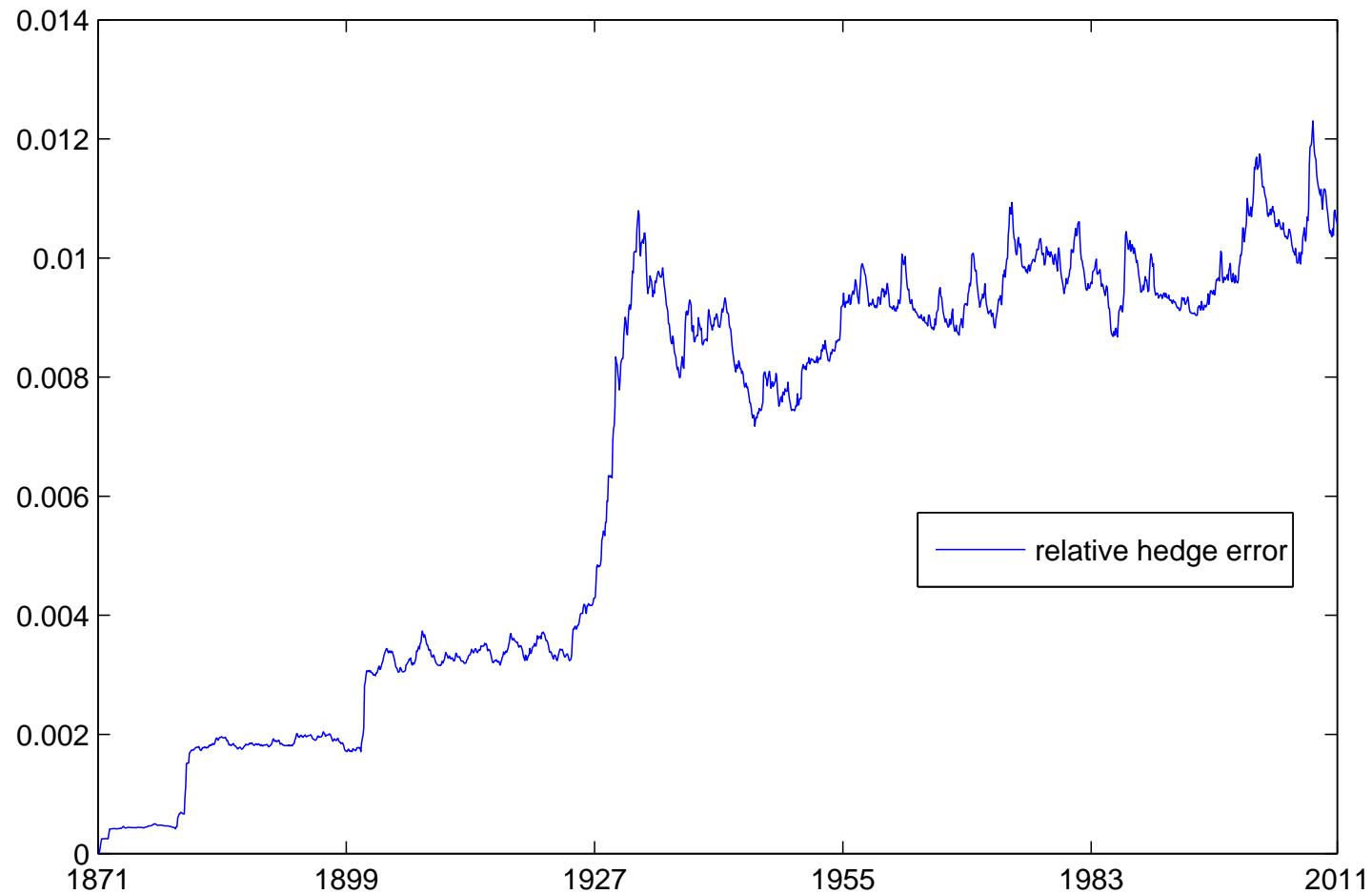
Logarithm of maximum drawdown constrained portfolio, $\alpha = 0.6$



Relative drawdown for drawdown constrained portfolio, $\alpha = 0.6$



Fraction of wealth in S&P500 for $\alpha = 0.6$



Relative hedge error, $\alpha = 0.6$, monthly hedging

- **Long term growth rate**

For $\alpha \in [0, 1)$, $X \in {}^\alpha \mathcal{X}$

$$\lim_{t \rightarrow \infty} \left(\frac{\log(X_t)}{G_t} \right) \leq 1 - \alpha = \lim_{t \rightarrow \infty} \left(\frac{\log({}^\alpha \bar{S}_t^{\delta*})}{G_t} \right)$$

$$\log(\bar{S}_t^{\delta*}) = G_t + \int_0^t |\theta_s| dW_s$$

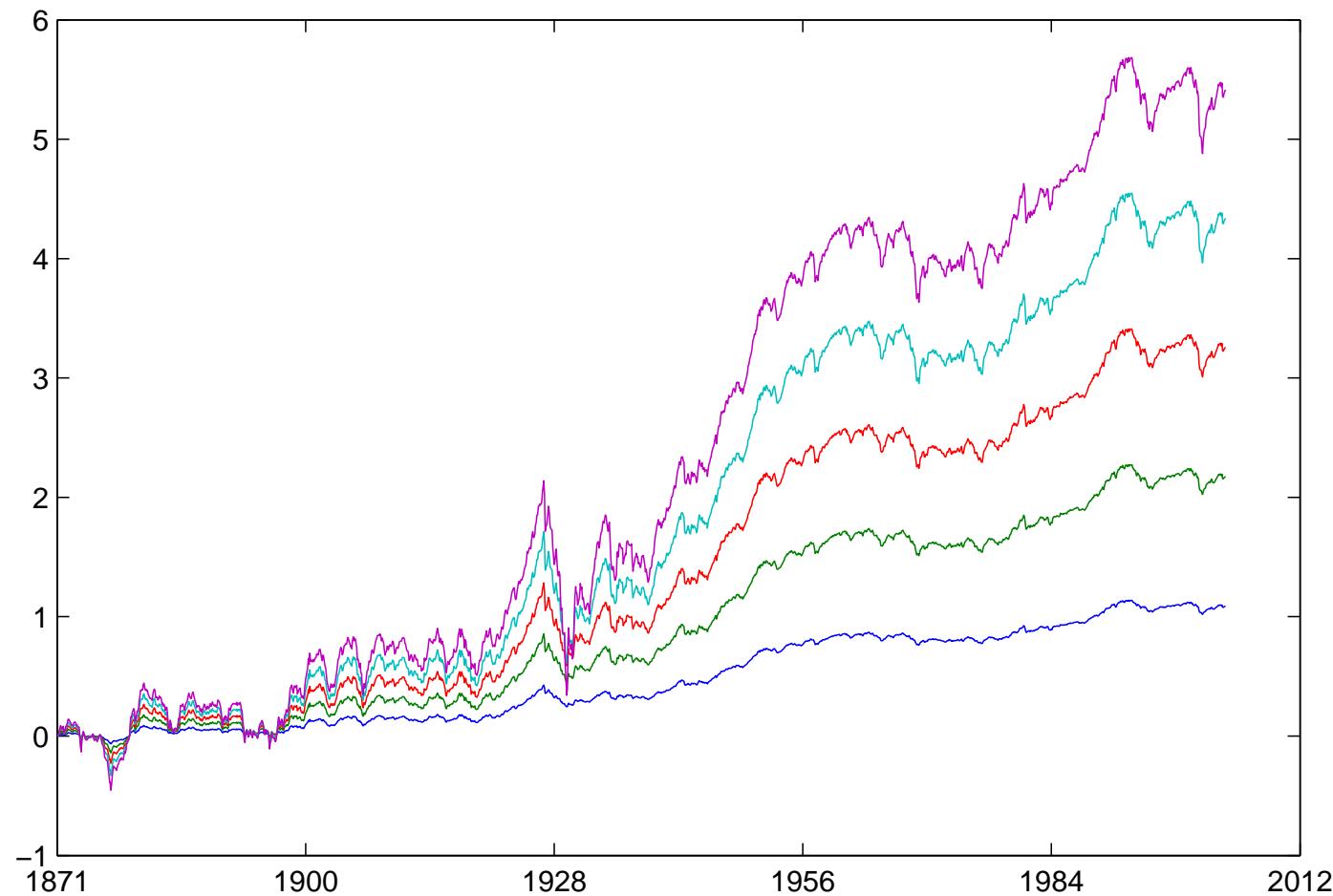
Kardaras, Obloj & Pl. (2012)

- asymptotic maximum long term growth rate

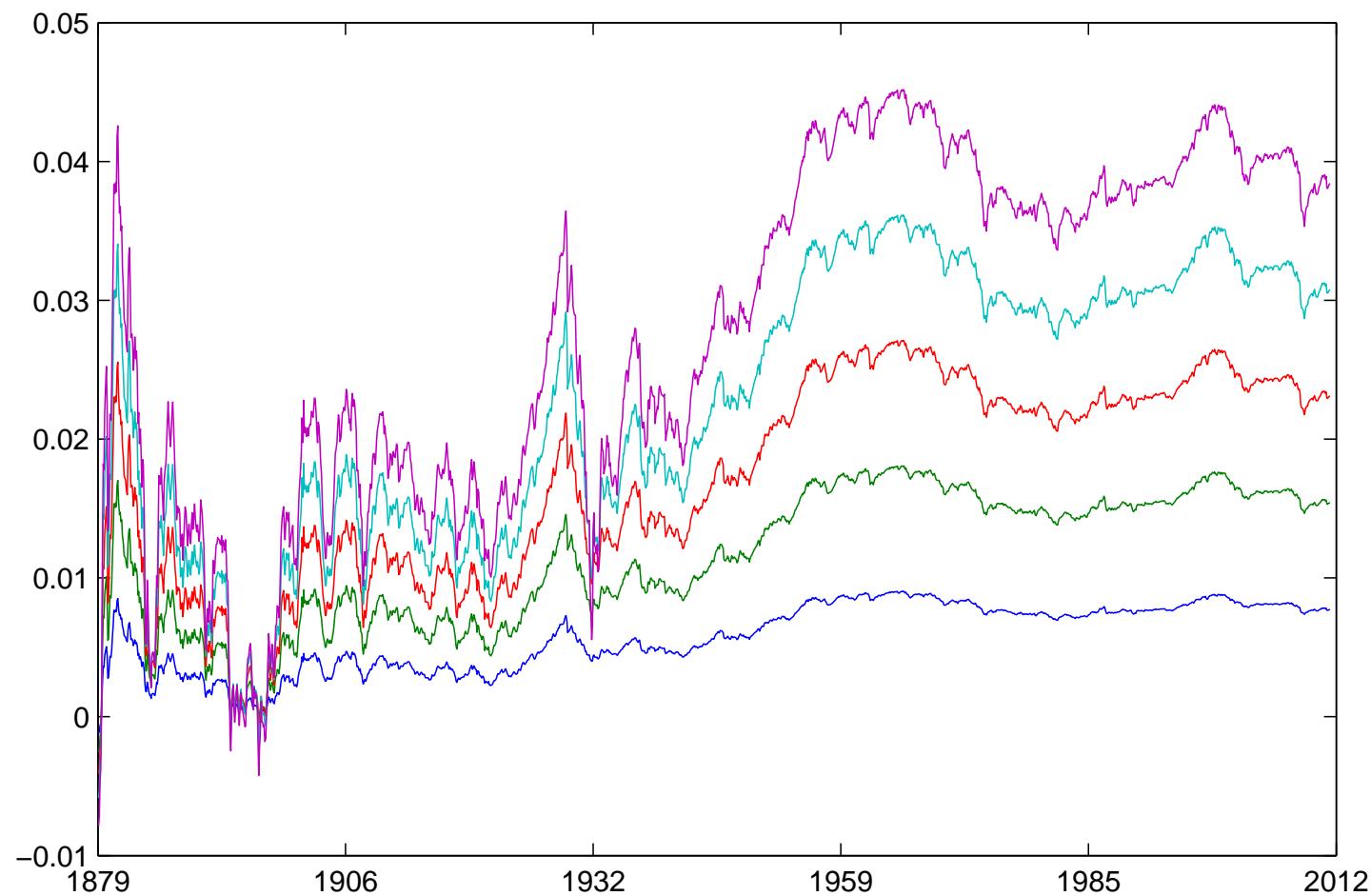
$$\lim_{t \rightarrow \infty} \frac{1}{t} \log({}^\alpha S_t^{\delta*}) = (1 - \alpha) \lim_{t \rightarrow \infty} \frac{1}{t} \log(S_t^{\delta*})$$

$${}^\alpha g = (1 - \alpha)g$$

restricted drawdown \implies reduced maximum growth rate
 long term view with short term attitude towards risk
 realistic alternative to Markowitz mean-variance approach
 and utility maximization



Logarithm of drawdown constrained portfolios



Long term growth of drawdown constrained portfolios

Expected Utility Maximization

- utility function $U(\cdot)$, $U'(\cdot) > 0$ $U''(\cdot) < 0$

- examples

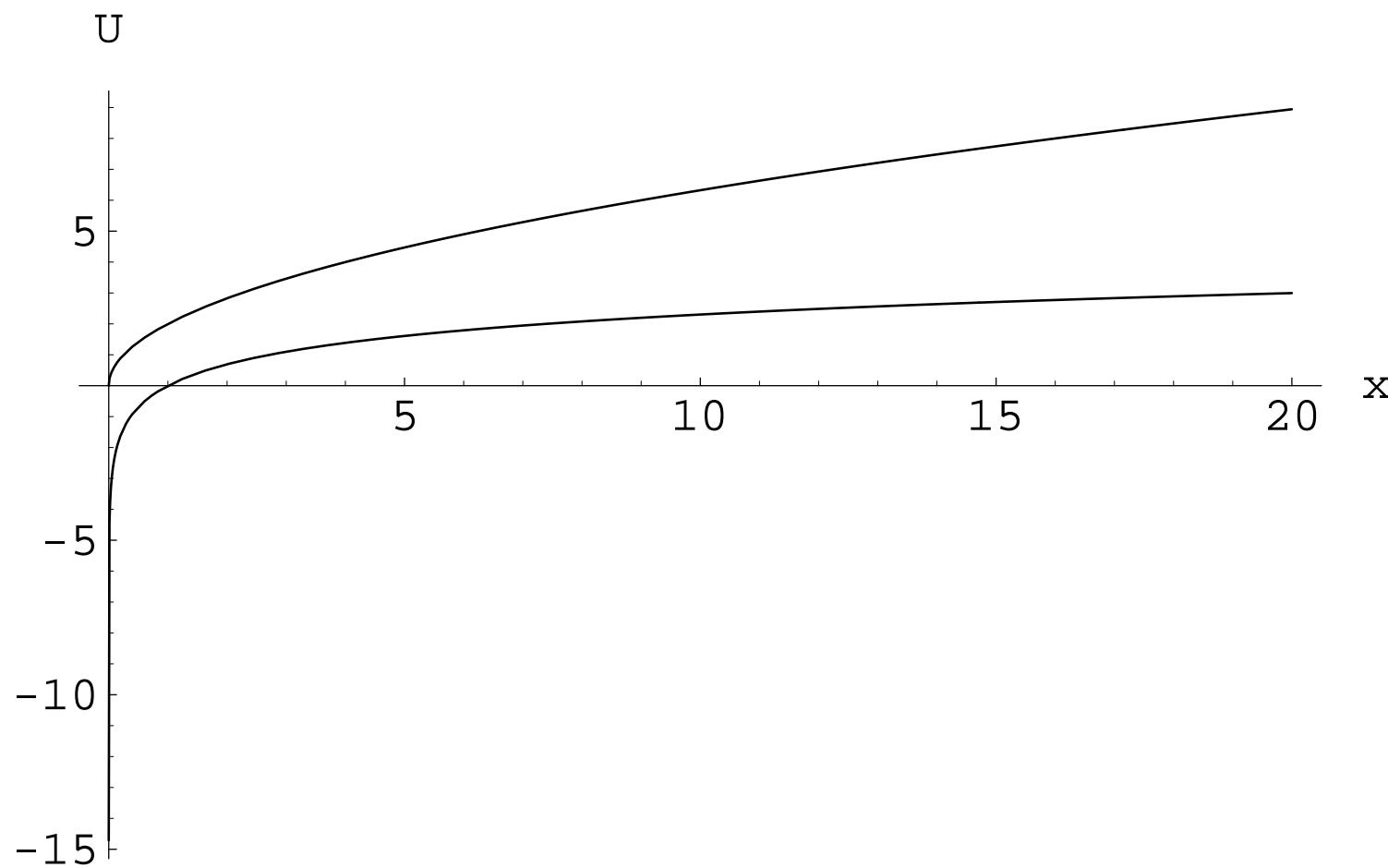
power utility

$$U(x) = \frac{1}{\gamma} x^\gamma$$

for $\gamma \neq 0$ and $\gamma < 1$

log-utility

$$U(x) = \ln(x)$$



Examples for power utility (upper graph) and log-utility (lower graph)

- assume \hat{S}^0 is **scalar Markov process**

- **maximize expected utility**

$$v^{\tilde{\delta}} = \max_{\bar{S}^{\delta} \in \bar{\mathcal{V}}_{S_0}^+} E \left(U \left(\bar{S}_T^{\delta} \right) \mid \mathcal{A}_0 \right)$$

$\bar{\mathcal{V}}_{S_0}^+$ set of strictly positive, discounted, fair portfolios

$$\bar{S}_0^{\delta} = S_0 > 0$$

Theorem

Benchmarked, expected utility maximizing portfolio:

$$\hat{S}_t^{\tilde{\delta}} = \hat{u}(t, \hat{S}_t^0) = E \left(U'^{-1} \left(\lambda \hat{S}_T^0 \right) \hat{S}_T^0 \mid \mathcal{A}_t \right),$$

$$\lambda \quad s.t. \quad S_0 = \hat{S}_0^{\tilde{\delta}} S_0^{\delta_*} = \hat{u}(0, \hat{S}_0) S_0^{\delta_*}$$

two fund separation with *risk aversion coefficient*

$$J_t^{\tilde{\delta}} = \frac{1}{1 - \frac{\hat{S}_t^0}{\hat{u}(t, \hat{S}_t^0)} \frac{\partial \hat{u}(t, \hat{S}_t^0)}{\partial \hat{S}_t^0}}.$$

\implies locally optimal portfolio

Proof:

$$\begin{aligned} v^\delta &= E(U(\bar{S}_T^\delta)) - \lambda \left(E(\hat{S}_T^\delta) - \hat{S}_0^\delta \right) \\ &= E(F(\bar{S}_T^\delta)) \end{aligned}$$

$$\begin{aligned} F(\bar{S}_T^\delta) &= U(\bar{S}_T^\delta) - \lambda \left(\frac{\bar{S}_T^\delta}{\bar{S}_T^{\delta*}} - \hat{S}_0^\delta \right) \\ F'(\bar{S}_T^\delta) &= U'(\bar{S}_T^\delta) - \frac{\lambda}{\bar{S}_T^{\delta*}} = 0 \end{aligned}$$

\implies

$$\begin{aligned} U'(\bar{S}_T^\delta) &= \frac{\lambda}{\bar{S}_T^{\delta*}} \\ \bar{S}_T^\delta &= U'^{-1}\left(\frac{\lambda}{\bar{S}_T^{\delta*}}\right) = U'^{-1}(\lambda \hat{S}_T^0) \\ \hat{S}_T^\delta &= U'^{-1}(\lambda \hat{S}_T^0) \hat{S}_T^0 \end{aligned}$$

□

- **log-utility function** $U(x) = \ln(x)$

$$U'(x) = \frac{1}{x}$$

$$U'^{-1}(y) = \frac{1}{y}$$

$$U''(x) = -\frac{1}{x^2}$$

utility concave

$$\hat{u}(t, \hat{S}_t^0) = E \left(U'^{-1} \left(\lambda \hat{S}_T^0 \right) \hat{S}_T^0 \mid \mathcal{A}_t \right) = E \left(\frac{1}{\lambda \hat{S}_T^0} \hat{S}_T^0 \mid \mathcal{A}_t \right) = \frac{1}{\lambda}$$

Lagrange multiplier

$$\lambda = \frac{S_0^{\delta_*}}{S_0}$$

risk aversion coefficient

$$J_t^{\tilde{\delta}} = 1$$

- expected log-utility

$$\begin{aligned} v^{\tilde{\delta}} &= E \left(\ln \left(\bar{S}_T^{\delta_*} \right) \mid \mathcal{A}_0 \right) \\ &= \ln(\lambda) + \ln(S_0) + \frac{1}{2} \int_0^T E \left((\theta(s, \bar{S}_s^{\delta_*}))^2 \mid \mathcal{A}_0 \right) ds \end{aligned}$$

if local martingale part forms a martingale

- power utility

$$U(x) = \frac{1}{\gamma} x^\gamma$$

for $\gamma < 1, \gamma \neq 0$

$$U'(x) = x^{\gamma-1}$$

$$U'^{-1}(y) = y^{\frac{1}{\gamma-1}}$$

$$U''(x) = (\gamma - 1) x^{\gamma-2}$$

concave

$$\hat{u}(t, \hat{S}_t^0) = E \left(\left(\frac{\lambda}{\bar{S}_T^{\delta_*}} \right)^{\frac{1}{\gamma-1}} \frac{1}{\bar{S}_T^{\delta_*}} \mid \mathcal{A}_t \right) = \lambda^{\frac{1}{\gamma-1}} E \left(\left(\bar{S}_T^{\delta_*} \right)^{\frac{\gamma}{1-\gamma}} \mid \mathcal{A}_t \right)$$

\bar{S}^{δ_*} geometric Brownian motion

$$\begin{aligned}\hat{u}(t, \hat{S}_t^0) &= \lambda^{\frac{1}{\gamma-1}} (\bar{S}_t^{\delta_*})^{\frac{\gamma}{1-\gamma}} \\ &\quad \times E \left(\exp \left\{ \frac{\gamma}{1-\gamma} \left(\frac{\theta^2}{2} (T-t) + \theta (W_T - W_t) \right) \right\} \middle| \mathcal{A}_t \right) \\ &= \lambda^{\frac{1}{\gamma-1}} (\bar{S}_t^{\delta_*})^{\frac{\gamma}{1-\gamma}} \exp \left\{ \frac{\theta^2}{2} \frac{\gamma}{(1-\gamma)^2} (T-t) \right\}\end{aligned}$$

Lagrange multiplier

$$\lambda = S_0^{\gamma-1} \bar{S}_0^{\delta_*} \exp \left\{ \frac{\theta^2}{2} \frac{\gamma}{1-\gamma} T \right\}$$

$$\hat{S}_t^0 = (\bar{S}_t^{\delta_*})^{-1}$$

$$\frac{\partial \hat{u}(t, \hat{S}_t^0)}{\partial \hat{S}^0} = \frac{\hat{u}(t, \hat{S}_t^0)}{\hat{S}_t^0} \frac{\gamma}{\gamma - 1}$$

risk aversion coefficient

$$J_t^{\tilde{\delta}} = 1 - \gamma$$

expected utility

$$v^{\tilde{\delta}} = E \left(\frac{1}{\gamma} \left(\bar{S}_T^{\tilde{\delta}} \right)^\gamma \mid \mathcal{A}_0 \right) = \exp \left\{ - \frac{\theta^2}{2} \frac{\gamma}{1-\gamma} T \right\} (S_0)^\gamma$$

Various Approaches to Asset Pricing

- Actuarial Pricing Approach (APA)
- Capital Asset Pricing Model (CAPM)
- Arbitrage Pricing Theory (APT)
- Utility Indifference Pricing
- Benchmark Approach (BA)

Classical

Law of One Price

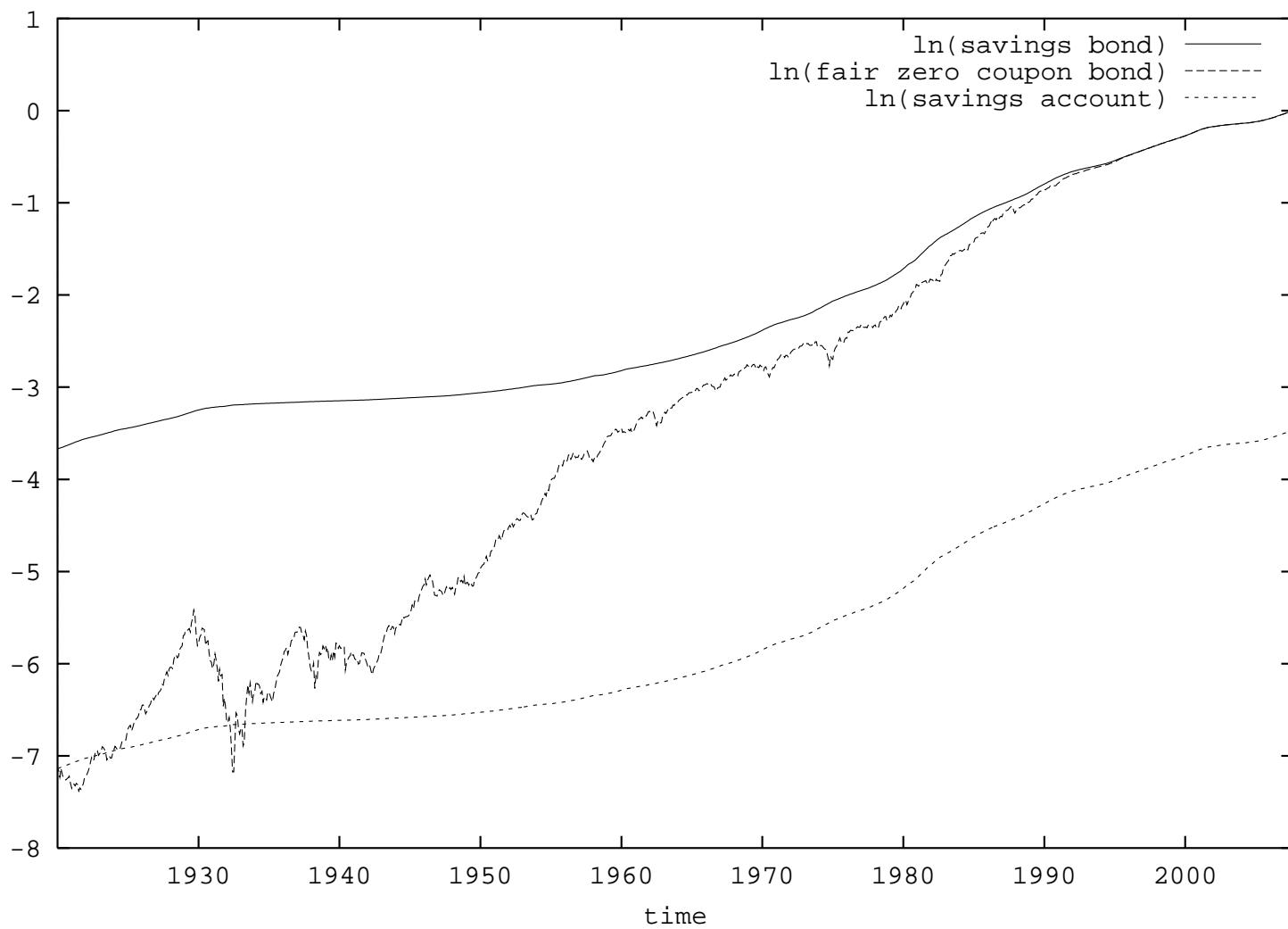
“All replicating portfolios of a payoff have the same price!”

Debreu (1959), Sharpe (1964), Lintner (1965),

Merton (1973a, 1973b), Ross (1976), Harrison & Kreps (1979),

Cochrane (2001), . . .

will be, in general, violated under the benchmark approach.



Logarithms of savings bond, fair zero coupon bond and savings account

- **benchmarked value**

$$\hat{S}_t^\delta = \frac{S_t^\delta}{S_t^{\delta_*}}$$

Supermartingale Property

For nonnegative S^δ

$$\hat{S}_t^\delta \geq E_t (\hat{S}_s^\delta)$$

$$0 \leq t \leq s < \infty$$

\hat{S}^δ supermartingale

- One needs consistent pricing concept
- All benchmarked nonnegative securities are supermartingales
(no upward trend)

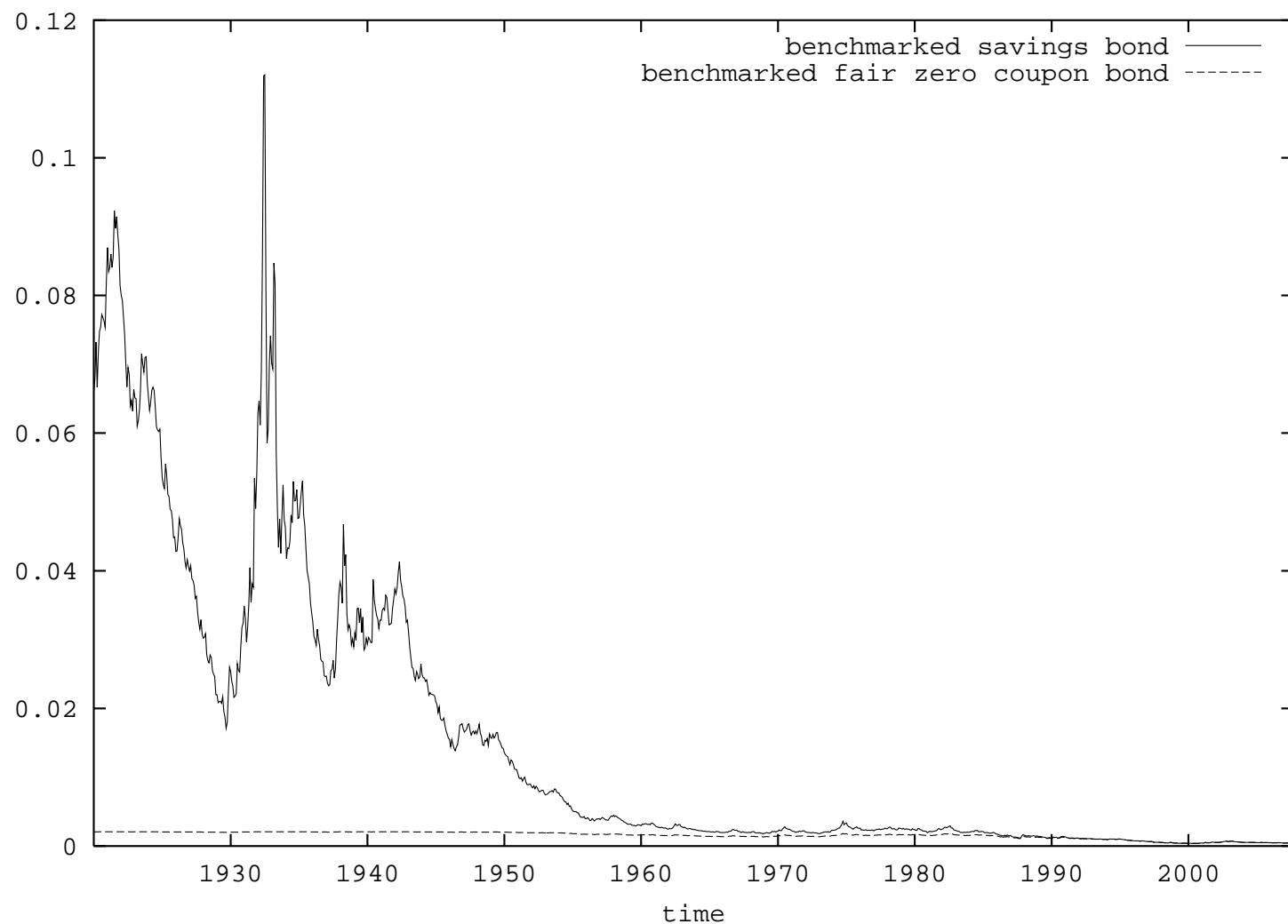
Definition *Price is fair if, when benchmarked, forms martingale
(no trend)*

$$\hat{S}_t^\delta = E_t \left(\hat{S}_s^\delta \right)$$

$$0 \leq t \leq s < \infty.$$

Lemma *The minimal nonnegative supermartingale that reaches a given benchmarked contingent claim is a martingale.*

see Revuz & Yor (1999)



Benchmarked savings bond and benchmarked fair zero coupon bond

Benchmark Approach:

Law of the Minimal Price

Pl. (2008)

Theorem *If a fair portfolio replicates a nonnegative payoff, then this represents the minimal replicating portfolio.*

- least expensive
- minimal possible hedge
- economically correct price in a competitive market

- contingent claim

$$H_T$$

$$E_0 \left(\frac{H_T}{S_T^{\delta_*}} \right) < \infty$$

- $S_t^{\delta_H}$ fair if

$$\hat{S}_t^{\delta_H} = \frac{S_t^{\delta_H}}{S_t^{\delta_*}} = E_t \left(\frac{H_T}{S_T^{\delta_*}} \right)$$

- real world expectation
best performing portfolio as numéraire
 \implies **Direct link with real world and economy !**

Law of the Minimal Price \implies

Corollary

*Minimal price for replicable H_T is **fair** and given by*

real world pricing formula

$$S_t^{\delta_H} = S_t^{\delta_*} E_t \left(\frac{H_T}{S_T^{\delta_*}} \right).$$

- Du & Pl. (2013) benchmarked risk minimization for nonreplicable claims

- most pricing concepts become unified and generalized by real world pricing

⇒

actuarial pricing

risk neutral pricing

pricing with stochastic discount factor

pricing with numeraire change

pricing with deflator

pricing with state pricing density

pricing with pricing kernel

utility indifference pricing

pricing with numeraire portfolio

classical risk minimization

benchmarked risk minimization

- Equivalent to risk neutral pricing are pricing via:
 - stochastic discount factor**, Cochrane (2001);
 - deflator**, Duffie (2001);
 - pricing kernel**, Constantinides (1992);
 - state price density**, Ingersoll (1987);
 - numeraire portfolio** as in Long (1990)

Actuarial Pricing

When H_T independent of $S_T^{\delta_*}$

\implies rigorous derivation of

- **actuarial pricing formula**

$$S_t^{\delta_H} = P(t, T) E_t(H_T)$$

with **zero coupon bond** as discount factor

$$P(t, T) = S_t^{\delta_*} E_t \left(\frac{1}{S_T^{\delta_*}} \right)$$

Risk Neutral Pricing

- real world pricing formula \implies

$$S_0^{\delta_H} = E_0 \left(\Lambda_T \frac{S_0^0}{S_T^0} H_T \right)$$

Radon-Nikodym derivative (density)

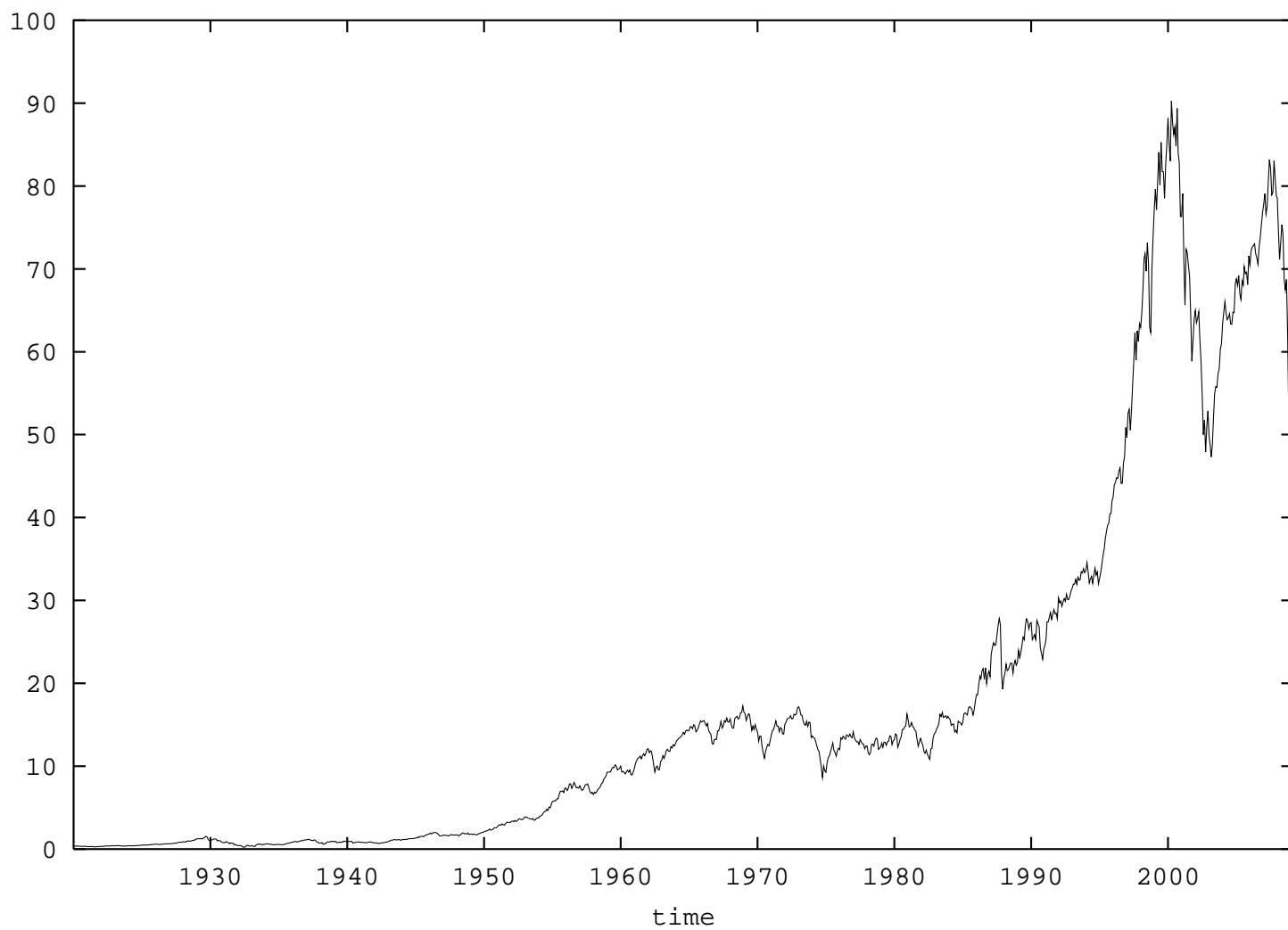
$$\Lambda_t = \frac{\hat{S}_t^0}{\hat{S}_0^0} \quad \text{supermartingale}$$

$$1 = \Lambda_0 \geq E_0(\Lambda_T)$$

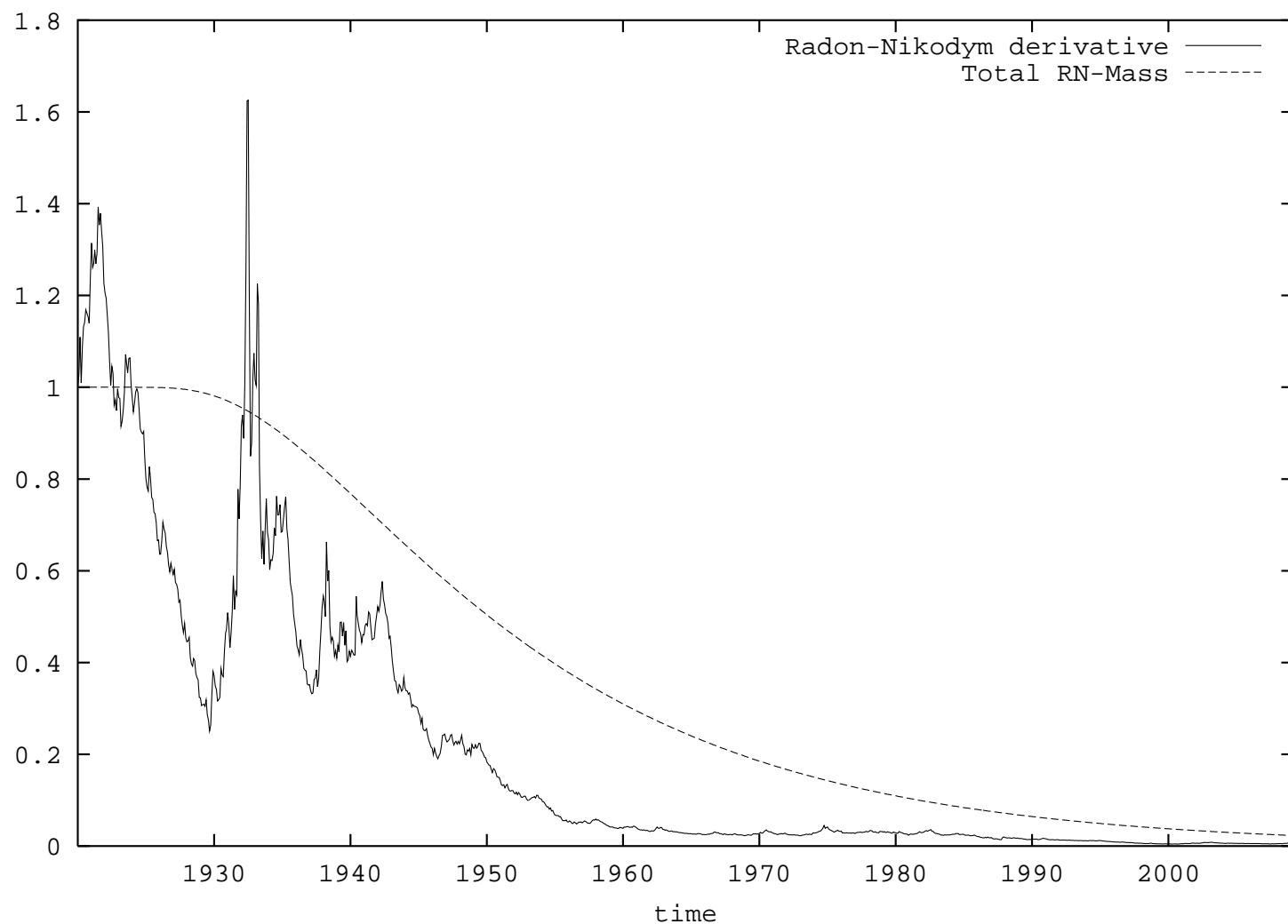
\implies

$$S_0^{\delta_H} \leq \frac{E_0 \left(\Lambda_T \frac{S_0^0}{S_T^0} H_T \right)}{E_0(\Lambda_T)}$$

similar for any numeraire (here S^0 savings account)



Discounted S&P500 total return index



Radon-Nikodym derivative and total mass of putative risk neutral measure

- special case when **savings account is fair**:

$$\implies \Lambda_T = \frac{dQ}{dP} \text{ forms martingale; } E_0(\Lambda_T) = 1;$$

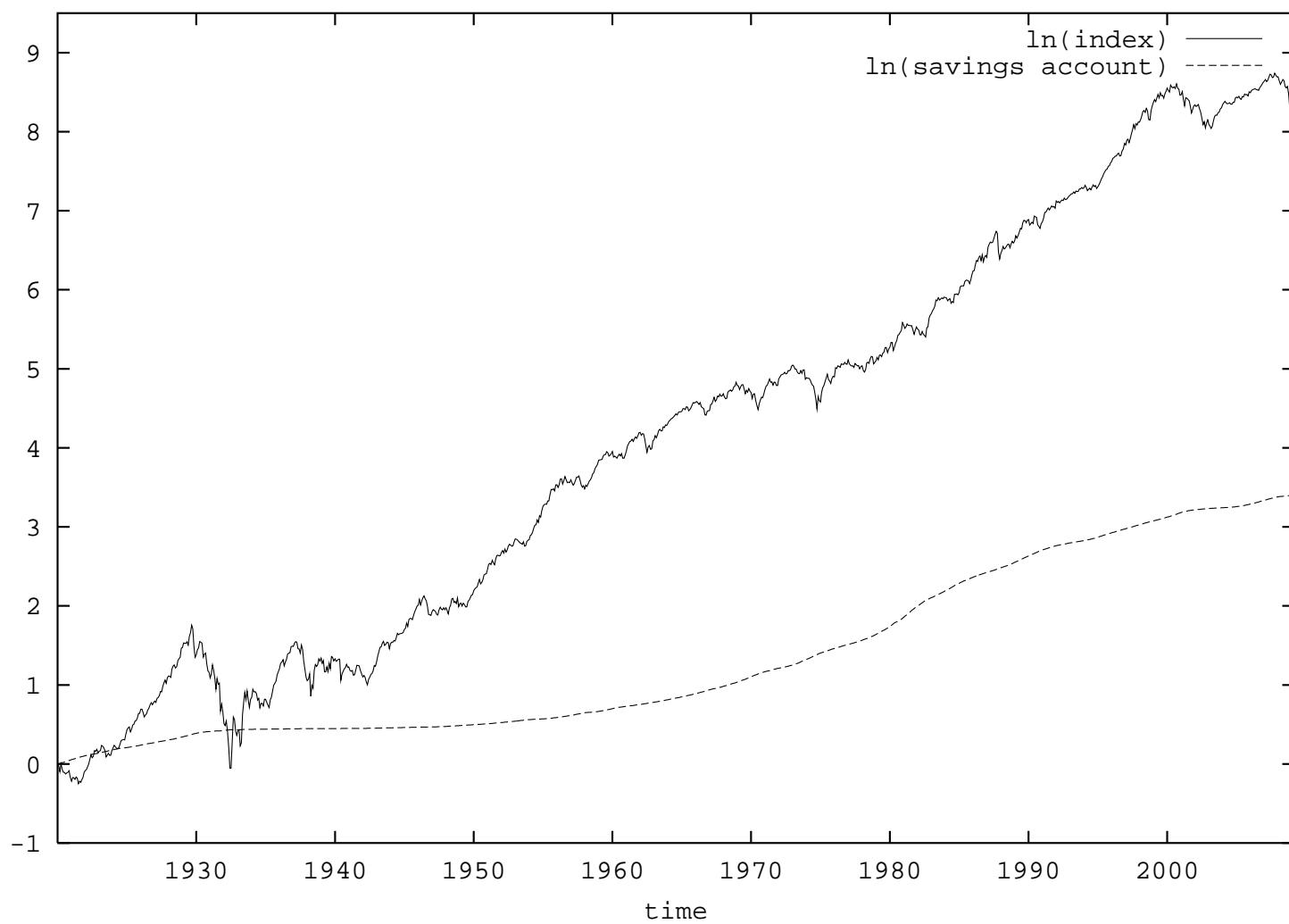
equivalent risk neutral probability measure Q exists;

Bayes' formula \implies
risk neutral pricing formula

$$S_0^{\delta_H} = E_0^Q \left(\frac{S_0^0}{S_T^0} H_T \right)$$

Harrison & Kreps (1979), Ingersoll (1987),
Constantinides (1992), Duffie (2001), Cochrane (2001), ...

- otherwise “risk neutral price” \geq real world price



$\ln(\text{S\&P500 accumulation index})$ and $\ln(\text{savings account})$

Risk Neutral Pricing under BS Model

$$\theta_t = \theta, \ a_t = a, \ r_t = r, \ \sigma_t = \sigma$$

Drifted Wiener Process

- drifted Wiener process W_θ

$$W_\theta(t) = W_t + \theta t$$

- market price of risk

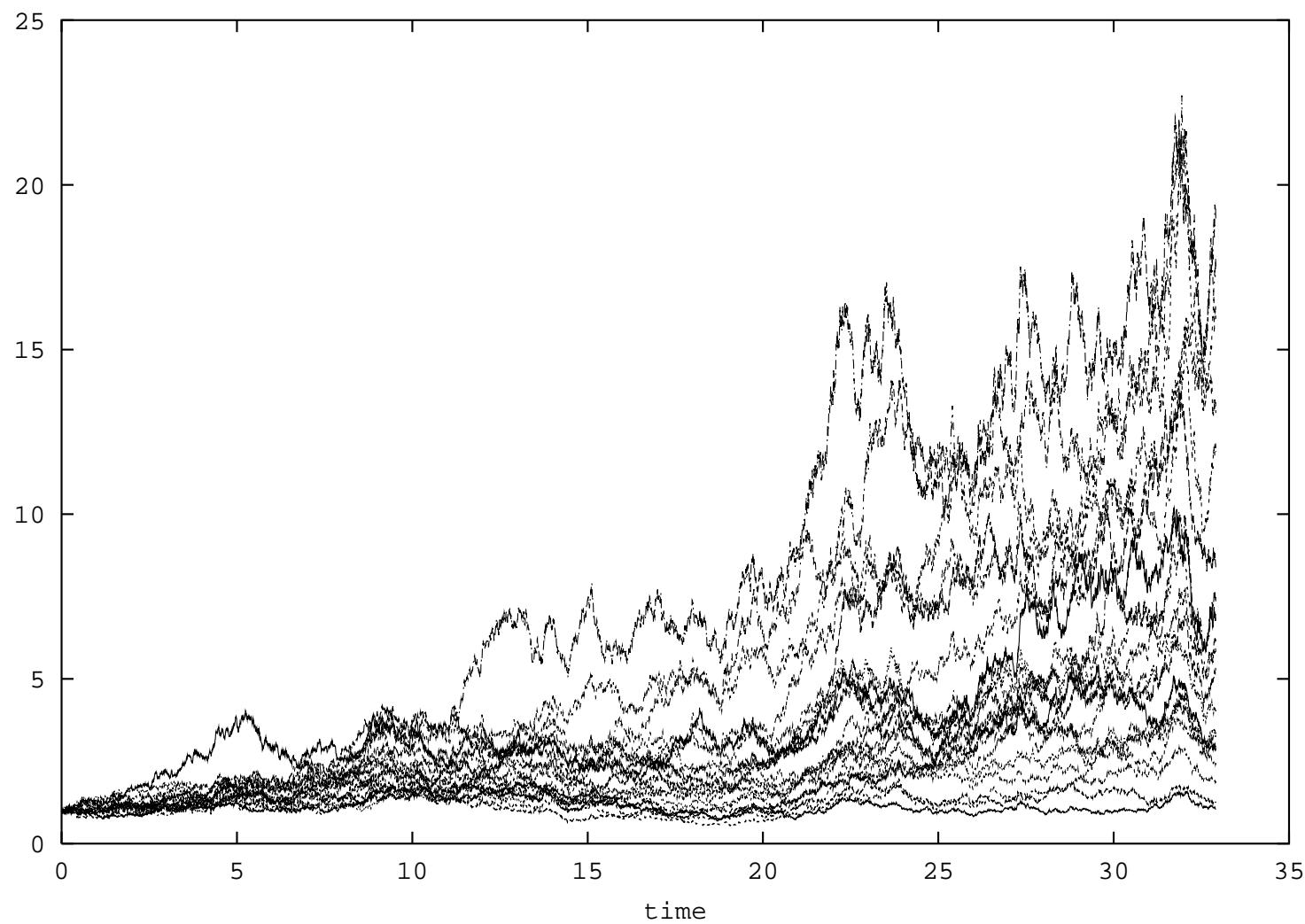
$$\theta = \frac{a - r}{\sigma}$$

- underlying security

$$\begin{aligned} dS_t &= (a - \sigma \theta) S_t dt + \sigma S_t (\theta dt + dW_t) \\ &= (a - \sigma \theta) S_t dt + \sigma S_t dW_\theta(t) \\ &= r S_t dt + \sigma S_t dW_\theta(t) \end{aligned}$$

$$S_t = S_0 \exp \left\{ \left(r - \frac{1}{2} \sigma^2 \right) t + \sigma W_\theta(t) \right\}$$

W_θ is **not** a Wiener process under the real world probability P ,
however, will be under some P_θ



Simulated primary security accounts.

- Radon-Nikodym derivative

$$\Lambda_\theta(T) = \frac{dP_\theta}{dP} = \frac{\hat{S}_T^0}{\hat{S}_0^0}$$

- risk neutral measure P_θ

$$\begin{aligned} P_\theta(A) &= \int_{\Omega} 1_{\{A\}} dP_\theta(\omega) = \int_A \frac{dP_\theta(\omega)}{dP(\omega)} dP(\omega) \\ &= \int_A \Lambda_\theta(T) dP(\omega) \end{aligned}$$

for all subsets $A \in \Omega$

Definition \tilde{P} is equivalent to \hat{P} if both have the same sets of events that have probability zero.

- equivalence of P and P_θ
fundamental for risk neutral pricing

Key Risk Neutral assumption:

Λ_θ is (\mathcal{A}, P) martingale

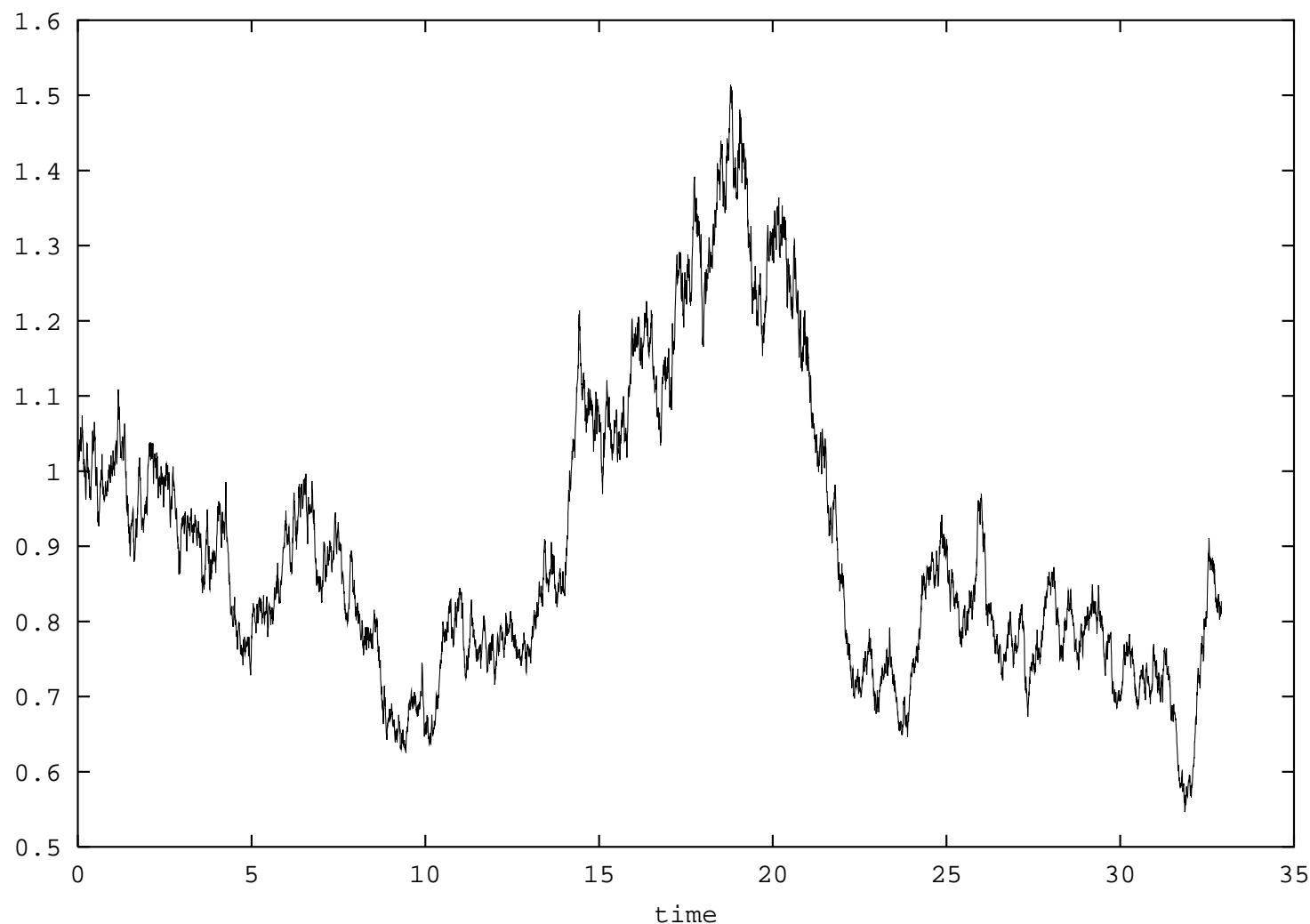
- under BS model

Radon-Nikodym derivative

$$\Lambda_\theta(t) = \frac{\hat{S}_t^0}{\hat{S}_0^0}$$

is $(\underline{\mathcal{A}}, P)$ -martingale,

\implies there exists equivalent risk neutral probability measure P_θ
for BS model



Benchmarked savings account under BS model, martingale.

Risk Neutral Measure Transformation

$$\frac{dP_\theta}{dP} \Big|_{\mathcal{A}_t} = \Lambda_\theta(t) = \exp \left\{ -\frac{1}{2} \theta^2 t - \theta W_t \right\}$$

- total risk neutral measure at $t = 0$:

$$\frac{dP_\theta}{dP} \Big|_{\mathcal{A}_0} = E(\Lambda_\theta(T) \mid \mathcal{A}_0) = \Lambda_\theta(0) = 1$$

$\implies P_\theta$ is a probability measure

Let us show that W_θ is Wiener process under P_θ :

For fixed $\tilde{y} \in \Re$,

$t \in [0, T]$ and $s \in [0, t]$

let A be the event

$$A = \{\omega \in \Omega : W_\theta(t, \omega) - W_\theta(s, \omega) < \tilde{y}\}$$

\implies

$$A = \{\omega \in \Omega : W_t(\omega) - W_s(\omega) < \tilde{y} - \theta(t - s)\}$$

E_θ denoting expectation under $P_\theta \implies$

$$\begin{aligned}
P_\theta(A) &= E_\theta(1_A | \mathcal{A}_0) = E \left(\frac{dP_\theta}{dP} \Big|_{\mathcal{A}_T} 1_A \Big| \mathcal{A}_0 \right) \\
&= E(\Lambda_\theta(T) 1_A | \mathcal{A}_0) = E \left(\Lambda_\theta(t) 1_A \frac{\Lambda_\theta(T)}{\Lambda_\theta(t)} \Big| \mathcal{A}_0 \right) \\
&= E \left(\Lambda_\theta(t) 1_A E \left(\frac{\Lambda_\theta(T)}{\Lambda_\theta(t)} \Big| \mathcal{A}_t \right) \Big| \mathcal{A}_0 \right) \\
&= E(\Lambda_\theta(t) 1_A | \mathcal{A}_0) = E \left(\Lambda_\theta(s) \frac{\Lambda_\theta(t)}{\Lambda_\theta(s)} 1_A \Big| \mathcal{A}_0 \right) \\
&= E(\Lambda_\theta(s) | \mathcal{A}_0) E \left(\frac{\Lambda_\theta(t)}{\Lambda_\theta(s)} 1_A \Big| \mathcal{A}_0 \right) \\
&= E \left(\frac{\Lambda_\theta(t)}{\Lambda_\theta(s)} 1_A \Big| \mathcal{A}_0 \right)
\end{aligned}$$

since $y = W_t - W_s$ is Gaussian $\sim \mathcal{N}(0, t-s)$ under $P \implies$

$$\begin{aligned}
P_\theta(A) &= E \left(\frac{\Lambda_\theta(t)}{\Lambda_\theta(s)} \mathbf{1}_A \right) \\
&= \int_{-\infty}^{\tilde{y}-\theta(t-s)} \exp \left\{ -\frac{\theta^2}{2} (t-s) - \theta y \right\} \frac{1}{\sqrt{2\pi(t-s)}} \\
&\quad \exp \left\{ -\frac{y^2}{2(t-s)} \right\} dy \\
&= \int_{-\infty}^{\tilde{y}-\theta(t-s)} \frac{1}{\sqrt{2\pi(t-s)}} \exp \left\{ -\frac{(y+\theta(t-s))^2}{2(t-s)} \right\} dy \\
&= \int_{-\infty}^{\tilde{y}} \frac{1}{\sqrt{2\pi(t-s)}} \exp \left\{ -\frac{z^2}{2(t-s)} \right\} dz
\end{aligned}$$

for $z = y + \theta(t-s)$

\implies

Theorem (Girsanov) *Under the above BS model the process*

$$W_\theta = \{W_\theta(t) = W_t + \theta t, t \geq 0\}$$

is a standard Wiener process in the filtered probability space $(\Omega, \mathcal{A}_T, \underline{\mathcal{A}}, P_\theta)$.

- W_θ is $(\underline{\mathcal{A}}, P_\theta)$ -Wiener process
- simply a transformation of variable
(not always possible, exploits here BS model)

Real World Pricing \implies **Risk Neutral Pricing**

$$\begin{aligned} V(0, S_0) &= S_0^{\delta_*} E \left(\frac{H(S_T)}{S_T^{\delta_*}} \mid \mathcal{A}_0 \right) \\ &= E \left(\frac{\hat{S}_T^0}{\hat{S}_0^0} \frac{H(S_T)}{S_T^0} \mid \mathcal{A}_0 \right) \\ &= E \left(\Lambda_{\theta}(T) \left(\frac{H(S_T)}{S_T^0} \right) \mid \mathcal{A}_0 \right) \\ &\leq \frac{E \left(\Lambda_{\theta}(T) \frac{H(S_T)}{S_T^0} \mid \mathcal{A}_0 \right)}{E \left(\Lambda_{\theta}(T) \mid \mathcal{A}_0 \right)} \end{aligned}$$

\implies **Risk Neutral Pricing** if Λ_θ is $(\underline{\mathcal{A}}, P)$ -martingale

$$\begin{aligned}
 V(0, S_0) &= E \left(\exp \left\{ -\frac{\theta^2}{2} T - \theta W_T \right\} \exp \{-r T\} H(S_T) \mid \mathcal{A}_0 \right) \\
 &= \exp \{-r T\} \int_{-\infty}^{\infty} H \left(S_0 \exp \left\{ \left(a - \frac{1}{2} \sigma^2 \right) T + \sigma \textcolor{red}{y} \right\} \right) \\
 &\quad \times \exp \left\{ -\frac{\theta^2}{2} T - \theta \textcolor{red}{y} \right\} \frac{1}{\sqrt{T}} N' \left(\frac{\textcolor{red}{y}}{\sqrt{T}} \right) d\textcolor{red}{y} \\
 &= \exp \{-r T\} \int_{-\infty}^{\infty} \left(H \left(S_0 \exp \left\{ \left(r - \frac{1}{2} \sigma^2 \right) T + \sigma \textcolor{blue}{z} \right\} \right) \right) \\
 &\quad \times \frac{1}{\sqrt{T}} N' \left(\frac{\textcolor{blue}{z}}{\sqrt{T}} \right) d\textcolor{blue}{z} \\
 &= \exp \{-r T\} E_\theta (H(S_T) \mid \mathcal{A}_0) = E_\theta \left(\frac{H(S_T)}{S_T^0} \mid \mathcal{A}_0 \right)
 \end{aligned}$$

$\textcolor{blue}{z} = \textcolor{red}{y} + \theta T \implies E_\theta$ - expectation with respect to P_θ

Risk Neutral SDEs under BS model

- discounted underlying security price

$$d\bar{S}_t = \sigma \bar{S}_t dW_\theta(t)$$

driftless under P_θ , $\sigma \bar{S} \in \mathcal{L}_T^2 \implies$

$(\underline{\mathcal{A}}, P_\theta)$ -martingale

- discounted option price

$$d\bar{V}(t, \bar{S}_t) = \frac{\partial \bar{V}(t, \bar{S}_t)}{\partial \bar{S}} \sigma \bar{S}_t dW_\theta(t)$$

driftless under P_θ , $\frac{\partial \bar{V}}{\partial \bar{S}} \sigma \bar{S} \in \mathcal{L}_T^2 \implies$

$(\underline{\mathcal{A}}, P_\theta)$ -martingale

- discounted savings account $\bar{S}_t^0 = 1 \implies (\underline{\mathcal{A}}, P_\theta)$ -martingale

- **discounted portfolio**

$$\begin{aligned}
 d\bar{S}_t^\delta &= d(\bar{S}_t^{\delta_*} \hat{S}_t^\delta) \\
 &= \bar{S}_t^\delta \pi_\delta^1(t) \sigma (\theta dt + dW_t) \\
 &= \bar{S}_t^\delta \pi_\delta^1(t) \sigma dW_\theta(t)
 \end{aligned}$$

$(\underline{\mathcal{A}}, P_\theta)$ -local martingale

- if $\bar{S}^\delta \pi_\delta^1 \delta = \delta^1 \bar{S} \sigma \in \mathcal{L}_T^2 \implies \bar{S}^\delta$ is an **$(\underline{\mathcal{A}}, P_\theta)$ -martingale**,
- \implies **risk neutral pricing formula** holds
- \implies
- $$\bar{S}_t^\delta = E_\theta \left(\bar{S}_T^\delta \mid \mathcal{A}_t \right) = E_\theta \left(\frac{S_T^\delta}{S_T^0} \mid \mathcal{A}_t \right)$$
- \bar{S}^{δ_*} is **$(\underline{\mathcal{A}}, P_\theta)$ -martingale** under BS model

- strict supermartingale portfolio under BS model

if Z_t strict $(\underline{\mathcal{A}}, P_\theta)$ -supermartingale, independent of $\bar{S}_t^{\delta_*}$ and

$$\bar{S}_t^\delta = \bar{S}_t^{\delta_*} Z_t$$

$\implies \bar{S}_t^\delta$ strict $(\underline{\mathcal{A}}, P_\theta)$ -supermartingale

- risk neutral pricing formula does **not** hold for such S_t^δ

$$\bar{S}_t^\delta > E_{\theta} (\bar{S}_T^\delta \mid \mathcal{A}_t) \text{ for } T > t$$

\implies **not** all discounted portfolios are $(\underline{\mathcal{A}}, P_\theta)$ -martingales

even when Radon-Nikodym derivative Λ_θ is $(\underline{\mathcal{A}}, P)$ -martingale

Change of Variables

- we performed a change of variables from W_t to $W_\theta(t)$

W Wiener processes under P

W_θ Wiener processes under P_θ

- relies on existence of an
equivalent risk neutral probability measure P_θ

General Girsanov Transformation

- **m -dimensional standard Wiener process**

$$W = \{W_t = (W_t^1, \dots, W_t^m)^\top, t \in [0, T]\}$$

$$(\Omega, \mathcal{A}_T, \underline{\mathcal{A}}, P)$$

- **$\underline{\mathcal{A}}$ -adapted, predictable m -dimensional stochastic process**

$$\theta = \{\theta_t = (\theta_t^1, \dots, \theta_t^m)^\top, t \in [0, T]\}$$

$$\int_0^T \sum_{i=1}^m (\theta_t^i)^2 dt < \infty$$

- assume **Radon-Nikodym derivative**

$$\begin{aligned}
 \Lambda_\theta(t) &= 1 - \sum_{i=1}^m \int_0^t \Lambda_\theta(s) \theta_s^i dW_s^i \\
 &= \exp \left\{ - \int_0^t \theta_s^\top dW_s - \frac{1}{2} \int_0^t \theta_s^\top \theta_s ds \right\} < \infty
 \end{aligned}$$

forms $(\underline{\mathcal{A}}, P)$ -martingale

- Λ_θ $(\underline{\mathcal{A}}, P)$ -martingale

\implies

$$E(\Lambda_\theta(t) \mid \mathcal{A}_s) = \Lambda_\theta(s)$$

for $t \in [0, T]$ and $s \in [0, t]$

\implies

$$E(\Lambda_\theta(t) \mid \mathcal{A}_0) = \Lambda_\theta(0) = 1$$

- define a candidate risk neutral measure P_θ with

$$\frac{dP_\theta}{dP} \Big|_{\mathcal{A}_T} = \Lambda_\theta(T)$$

by setting

$$\begin{aligned} P_\theta(A) &= E(\Lambda_\theta(T) \mathbf{1}_A) \\ &= E_\theta(\mathbf{1}_A) \end{aligned}$$

- since Λ_θ is an $(\underline{\mathcal{A}}, P)$ -martingale

$$P_\theta(\Omega) = E(\Lambda_\theta(T)) = E(\Lambda_\theta(T) \mid \mathcal{A}_0) = \Lambda_\theta(0) = 1$$

$\implies P_\theta$ is a probability measure

Girsanov Theorem

Theorem (Girsanov) *If a given strictly positive Radon-Nikodym derivative process Λ_θ is an $(\underline{\mathcal{A}}, P)$ -martingale, then the m -dimensional process $W_\theta = \{W_\theta(t), t \in [0, T]\}$, given by*

$$W_\theta(t) = W_t + \int_0^t \theta_s \, ds$$

for all $t \in [0, T]$, is an m -dimensional standard Wiener process on the filtered probability space $(\Omega, \mathcal{A}_T, \underline{\mathcal{A}}, P_\theta)$.

- W is $(\underline{\mathcal{A}}, P)$ -Wiener process

W_θ is $(\underline{\mathcal{A}}, P_\theta)$ -Wiener process

Novikov Condition

When is the strictly positive $(\underline{\mathcal{A}}, P)$ -local martingale Λ_θ an $(\underline{\mathcal{A}}, P)$ -martingale?

- a **sufficient** condition is the **Novikov condition**

Novikov (1972)

$$E \left(\exp \left\{ \frac{1}{2} \int_0^T |\theta_s^\top \theta_s| ds \right\} \right) < \infty$$

(is satisfied for the BS model)

- for other sufficient conditions see Revuz & Yor (1999)

Bayes's Theorem

Theorem (Bayes) *Assume that Λ_θ is an $(\underline{\mathcal{A}}, P)$ -martingale, then for any given stopping time $\tau \in [0, T]$ and any \mathcal{A}_τ -measurable random variable Y , satisfying the integrability condition*

$$E_\theta(|Y|) < \infty,$$

one has the Bayes rule

$$E_\theta(Y | \mathcal{A}_s) = \frac{E(\Lambda_\theta(\tau) Y | \mathcal{A}_s)}{E(\Lambda_\theta(\tau) | \mathcal{A}_s)}$$

for $s \in [0, \tau]$.

Proof of Bayes's Theorem (*)

For each \mathcal{A}_τ -measurable random variable Y and a set $A \in \mathcal{A}_s$ with $s \in [0, \tau]$, $\tau \leq T$

$$\begin{aligned}
1_A E_\theta(Y | \mathcal{A}_s) &= E_\theta(1_A Y | \mathcal{A}_s) = E(1_A Y \Lambda_\theta(T) | \mathcal{A}_s) \\
&= E(1_A Y \Lambda_\theta(\tau) | \mathcal{A}_s) = E(1_A E(Y \Lambda_\theta(\tau) | \mathcal{A}_s) | \mathcal{A}_s) \\
&= E\left(\Lambda_\theta(s) \left\{ \frac{1_A}{\Lambda_\theta(s)} E(Y \Lambda_\theta(\tau) | \mathcal{A}_s) \right\} \middle| \mathcal{A}_s\right) \\
&= E_\theta\left(\frac{1_A}{\Lambda_\theta(s)} E(Y \Lambda_\theta(\tau) | \mathcal{A}_s) \middle| \mathcal{A}_s\right) \\
&= E_\theta\left(1_A \frac{E(\Lambda_\theta(\tau) Y | \mathcal{A}_s)}{E(\Lambda_\theta(\tau) | \mathcal{A}_s)} \middle| \mathcal{A}_s\right) \\
&= 1_A \frac{E(\Lambda_\theta(\tau) Y | \mathcal{A}_s)}{E(\Lambda_\theta(\tau) | \mathcal{A}_s)}. \quad \square
\end{aligned}$$

Change of Numeraire Technique

Geman, El Karoui & Rochet (1995)

Jamshidian (1997)

- **numeraire**

$$S^{\bar{\delta}} = \{S_t^{\bar{\delta}}, t \in [0, T]\}$$

normalizes all other portfolios

- may lead to simplifications in calculations
- relative price of a portfolio

$$\frac{S_t^\delta}{S_t^{\bar{\delta}}}$$

Self-financing under Numeraire Change

$$S_t^\delta = \delta_t^0 S_t^0 + \delta_t^1 S_t$$

$$dS_t^\delta = \delta_t^0 dS_t^0 + \delta_t^1 dS_t$$

$$dS_t^{\bar{\delta}} = \bar{\delta}_t^0 dS_t^0 + \bar{\delta}_t^1 dS_t$$

Itô formula \implies

$$\begin{aligned}
d \left(\frac{S_t^\delta}{S_t^{\bar{\delta}}} \right) &= \frac{1}{S_t^{\bar{\delta}}} dS_t^\delta + S_t^\delta d \left(\frac{1}{S_t^{\bar{\delta}}} \right) + d \left[\frac{1}{S^{\bar{\delta}}}, S^\delta \right]_t \\
&= \delta_t^0 \left(\frac{1}{S_t^{\bar{\delta}}} dS_t^0 + S_t^0 d \left(\frac{1}{S_t^{\bar{\delta}}} \right) \right) \\
&\quad + \delta_t^1 \left(\frac{1}{S_t^{\bar{\delta}}} dS_t + S_t d \left(\frac{1}{S_t^{\bar{\delta}}} \right) + d \left[\frac{1}{S^{\bar{\delta}}}, S \right]_t \right)
\end{aligned}$$

$$[X, Y]_t \approx \sum_{i=0}^{i_t} (X_{t_{i+1}} - X_{t_i}) (Y_{t_{i+1}} - Y_{t_i}) \quad \text{covariation}$$

Itô formula for

$$\frac{S_t^0}{S_t^{\bar{\delta}}} \quad \text{and} \quad \frac{S_t}{S_t^{\bar{\delta}}}$$

\implies

$$d \left(\frac{S_t^\delta}{S_t^{\bar{\delta}}} \right) = \delta_t^0 d \left(\frac{S_t^0}{S_t^{\bar{\delta}}} \right) + \delta_t^1 d \left(\frac{S_t}{S_t^{\bar{\delta}}} \right)$$

\implies

- in denomination of another numeraire **always self-financing**

Numeraire Pairs

- numeraire pair $(S^{\bar{\delta}}, P_{\theta \bar{\delta}})$

real world pricing formula:

$$S_t = E \left(\frac{S_t^{\delta_*}}{S_T^{\delta_*}} S_T \mid \mathcal{A}_t \right)$$

existence of GOP

\implies numeraire pair (S^{δ_*}, P)

if Λ_θ is (\mathcal{A}, P) -martingale \implies

- **risk neutral** pricing formula:

$$S_t = E_\theta \left(\frac{S_t^0}{S_T^0} S_T \mid \mathcal{A}_t \right)$$

\implies numeraire pair (S^0, P_θ)

- for fair price process S
from real world pricing formula

$$\frac{S_t}{S_t^{\delta_*}} = E \left(\frac{S_T}{S_T^{\delta_*}} \mid \mathcal{A}_t \right)$$

\implies for any strictly positive numeraire $S^{\bar{\delta}}$

$$\begin{aligned} \frac{S_t}{S_t^{\bar{\delta}}} &= E \left(\frac{S_t^{\delta_*}}{S_t^{\bar{\delta}}} \frac{S_T^{\bar{\delta}}}{S_T^{\delta_*}} \frac{S_T}{S_T^{\bar{\delta}}} \mid \mathcal{A}_t \right) \\ &= E \left(\frac{\Lambda_{\theta_{\bar{\delta}}}(T)}{\Lambda_{\theta_{\bar{\delta}}}(t)} \frac{S_T}{S_T^{\bar{\delta}}} \mid \mathcal{A}_t \right) \end{aligned}$$

with **benchmarked normalized numeraire**

\implies

$$\Lambda_{\theta_{\bar{\delta}}}(t) = \frac{\hat{S}_t^{\bar{\delta}}}{\hat{S}_0^{\bar{\delta}}}$$

$$\begin{aligned} d\Lambda_{\theta_{\bar{\delta}}}(t) &= d \left(\frac{\hat{S}_t^{\bar{\delta}}}{\hat{S}_0^{\bar{\delta}}} \right) \\ &= \Lambda_{\theta_{\bar{\delta}}}(t) (\pi_{\bar{\delta}}^1(t) \sigma_t - \theta_t) dW_t \end{aligned}$$

$(\underline{\mathcal{A}}, P)$ -local martingale

\implies **Radon-Nikodym derivative** for candidate measure $P_{\theta_{\bar{\delta}}}$

$$\frac{dP_{\theta_{\bar{\delta}}}}{dP} = \Lambda_{\theta_{\bar{\delta}}}(T)$$

- drifted Wiener process

$$dW_{\theta_{\bar{\delta}}}(t) = dW_t + \theta_{\bar{\delta}}(t) dt,$$

where

$$\theta_{\bar{\delta}}(t) = \theta_t - \pi_{\bar{\delta}}^1(t) \sigma_t$$

- If $\Lambda_{\theta_{\bar{\delta}}}$ is true **(\underline{A}, P)-martingale**

$\implies P_{\theta_{\bar{\delta}}}$ equivalent to P and probability measure

allows to apply **Girsanov transformation**

$\implies W_{\theta_{\bar{\delta}}}$ standard Wiener process under $P_{\theta_{\bar{\delta}}}$

\implies **numeraire pair $(S^{\bar{\delta}}, P_{\theta_{\bar{\delta}}})$**

Some Examples

- **GOP** S^{δ_*} as numeraire $\implies \theta_{\delta_*}(t) = 0$

- **savings account** S^0 as numeraire

$$\pi_{\bar{\delta}}^1(t) = 0 \implies \theta_{\bar{\delta}}(t) = \theta_t$$

- **underlying security** S as numeraire

$$\pi_{\bar{\delta}}^1(t) = 1 \implies \theta_{\bar{\delta}}(t) = \theta_t - \sigma_t$$

- **unfair portfolio** $S_t^{\bar{\delta}} = S_t^{\delta_*} Z_t$ as numeraire
with Z an $(\underline{\mathcal{A}}, P)$ -strict supermartingale, independent of S^{δ_*}

$\implies \hat{S}^{\bar{\delta}}$ is **not** an $(\underline{\mathcal{A}}, P)$ -martingale

Girsanov Theorem **can not** be applied !

\implies **If savings account is not fair,**
then the change of numeraire approach is not applicable !

Pricing Formula for General Numeraire $S^{\underline{\delta}}$

$$\begin{aligned} V(0, S_0) &= E \left(\frac{S_0^{\delta_*}}{S_T^{\delta_*}} H(S_T) \mid \mathcal{A}_0 \right) \\ &= E \left(\Lambda_{\theta_{\bar{\delta}}}(T) \frac{H(S_T)}{S_T^{\bar{\delta}}} \mid \mathcal{A}_0 \right) \end{aligned}$$

the quantity

$$\frac{\Lambda_{\theta_{\bar{\delta}}}(T)}{S_T^{\bar{\delta}}} = \frac{S_0^{\delta_*}}{S_T^{\delta_*}}$$

remains **numeraire invariant**

- if $\Lambda_{\theta_{\bar{\delta}}}$ forms an $(\underline{\mathcal{A}}, P)$ -martingale

Bayes's Theorem \implies

$$E \left(\Lambda_{\theta_{\bar{\delta}}}(T) \frac{H(S_T)}{S_T^{\bar{\delta}}} \mid \mathcal{A}_0 \right) = E_{\theta_{\bar{\delta}}} \left(\frac{H(S_T)}{S_T^{\bar{\delta}}} \mid \mathcal{A}_0 \right)$$

\implies

pricing formula with **general numeraire**

$$V(0, S_0) = E_{\theta_{\bar{\delta}}} \left(\frac{H(S_T)}{S_T^{\bar{\delta}}} \mid \mathcal{A}_0 \right)$$

- represents a **transformation of variables** in integration
- may provide **computational advantages**

Utility Indifference Pricing

discounted payoff $\bar{H} = \frac{H}{S_T^0}$ (not perfectly hedgable)

$$v_{\epsilon, V}^{\tilde{\delta}} = E \left(U \left((S_0 - \epsilon V) \frac{\bar{S}_T^{\tilde{\delta}}}{S_0} + \epsilon \bar{H} \right) \mid \mathcal{A}_0 \right)$$

- assume \hat{S}^0 is **scalar diffusion**

Definition *Utility indifference price* V s.t.

$$\lim_{\epsilon \rightarrow 0} \frac{v_{\epsilon, V}^{\tilde{\delta}} - v_{0, V}^{\tilde{\delta}}}{\epsilon} = 0$$

e.g. Davis (1997)

- Taylor expansion

$$\begin{aligned}
v_{\varepsilon, V}^{\tilde{\delta}} &\approx E \left(U \left(\bar{S}_T^{\tilde{\delta}} \right) + \varepsilon U' \left(\bar{S}_T^{\tilde{\delta}} \right) \left(\bar{H} - \frac{V \bar{S}_T^{\tilde{\delta}}}{S_0} \right) \mid \mathcal{A}_0 \right) \\
&= v_{0, V}^{\tilde{\delta}} + \varepsilon E \left(U' \left(\bar{S}_T^{\tilde{\delta}} \right) \bar{H} \mid \mathcal{A}_0 \right) - V \varepsilon E \left(U' \left(\bar{S}_T^{\tilde{\delta}} \right) \frac{\bar{S}_T^{\tilde{\delta}}}{S_0} \mid \mathcal{A}_0 \right)
\end{aligned}$$

\implies

- utility indifference price

$$\begin{aligned}
 V &= \frac{E \left(U' \left(\bar{S}_T^{\tilde{\delta}} \right) \bar{H} \mid \mathcal{A}_0 \right)}{E \left(U' \left(\bar{S}_T^{\tilde{\delta}} \right) \frac{\bar{S}_T^{\tilde{\delta}}}{S_0} \mid \mathcal{A}_0 \right)} = \frac{E \left(U' \left(U'^{-1} \left(\frac{\lambda}{\bar{S}_T^{\delta_*}} \right) \right) \bar{H} \mid \mathcal{A}_0 \right)}{E \left(U' \left(U'^{-1} \left(\frac{\lambda}{\bar{S}_T^{\delta_*}} \right) \right) \frac{\bar{S}_T^{\tilde{\delta}}}{S_0} \mid \mathcal{A}_0 \right)} \\
 &= \frac{E \left(\frac{H}{S_T^{\delta_*}} \mid \mathcal{A}_0 \right)}{\frac{1}{S_0} \frac{S_0}{S_0^{\delta_*}}}
 \end{aligned}$$

independent of U and independent of particular model \implies

- real world pricing formula

$$V = S_0^{\delta_*} E \left(\frac{H}{S_T^{\delta_*}} \mid \mathcal{A}_0 \right)$$

Pricing via Hedging

- underlying security

$$S = \{S_t, t \in [0, T]\}$$

$$dS_t = a_t S_t dt + \sigma_t S_t dW_t$$

$$t \in [0, T], S_0 > 0$$

appropriate stochastic appreciation rate a_t

and volatility σ_t

Wiener process $W = \{W_t, t \in [0, T]\}$

- **savings account**

$$dS_t^0 = r_t S_t^0 dt$$

$t \in [0, T]$, $S_0^0 = 1$

appropriate stochastic short rate

- replicate payoff

$$f(S_T) \geq 0$$

assuming

$$E(f(S_T)) < \infty$$

- hedge portfolio

$$V(t, S_t) = \delta_t^0 S_t^0 + \delta_t^1 S_t$$

δ_t^0 units of savings account

δ_t^1 units of underlying security

- hedging strategy

$$\delta = \{\delta_t = (\delta_t^0, \delta_t^1)^\top, t \in [0, T]\}$$

$$\delta^0 = \{\delta_t^0, t \in [0, T]\}$$

$$\delta^1 = \{\delta_t^1, t \in [0, T]\}$$

predictable processes

- replicating portfolio

$$V(T, S_T) = f(S_T)$$

- SDE for hedge portfolio

$$\begin{aligned} dV(t, S_t) &= \delta_t^0 dS_t^0 + \delta_t^1 dS_t \\ &\quad + S_t^0 d\delta_t^0 + S_t d\delta_t^1 + d [\delta^1, S]_t \end{aligned}$$

- **self-financing**

portfolio changes are caused by gains from trade

$$dV(t, S_t) = \delta_t^0 dS_t^0 + \delta_t^1 dS_t$$

\implies

self-financing condition:

$$S_t^0 d\delta_t^0 + S_t d\delta_t^1 + d [\delta^1, S]_t = 0$$

- discounted underlying security

$$\bar{S}_t = \frac{S_t}{S_t^0}$$

$$t \in [0, T]$$

Itô formula \implies

$$d\bar{S}_t = (a - r) \bar{S}_t dt + \sigma_t \bar{S}_t dW_t$$

$$t \in [0, T], \bar{S}_0 = S_0$$

- discounted value function

$$\bar{V} : [0, T] \times [0, \infty) \rightarrow [0, \infty)$$

$$\bar{V}(t, \bar{S}_t) = \frac{V(t, S_t)}{S_t^0}$$

- profit and loss (P&L)

$$C_t = \text{price} - \text{gains from trade} - \text{initial price}$$

- discounted P&L

$$\begin{aligned}\bar{C}_t &= \frac{C_t}{S_t^0} \\ &= \bar{V}(t, \bar{S}_t) - I_{\delta^1, \bar{S}}(t) - \bar{V}(0, \bar{S}_0)\end{aligned}$$

- discounted gains from trade

$$\begin{aligned}
 I_{\delta^1, \bar{S}}(t) &= \int_0^t \delta_u^1 d\bar{S}_u \\
 &= \int_0^t \delta_u^1 (a - r) \bar{S}_u du \\
 &\quad + \int_0^t \delta_u^1 \sigma \bar{S}_u dW_u
 \end{aligned}$$

$$I_{\delta^0, \bar{S}^0}(t) = \int_0^t \delta_u^0 d\bar{S}_u^0 = 0$$

- initial payment

$$V(0, S_0) = \bar{V}(0, \bar{S}_0),$$

\implies

$$\bar{C}_0 = 0$$

- identify hedging strategy δ

for which discounted P&L zero

$$\bar{C}_t = 0$$

for all $t \in [0, T]$

\implies perfect hedge

$V(t, S_t)$ - replicating portfolio

- increments of discounted P&L

$$\begin{aligned}
0 = \bar{C}_t - \bar{C}_s &= \bar{V}(t, \bar{S}_t) - \bar{V}(s, \bar{S}_s) - \int_s^t \delta_u^1 d\bar{S}_u \\
&= \int_s^t \left[\frac{\partial \bar{V}(u, \bar{S}_u)}{\partial t} + \frac{1}{2} \sigma_u^2 \bar{S}_u^2 \frac{\partial^2 \bar{V}(u, \bar{S}_u)}{\partial \bar{S}^2} \right. \\
&\quad \left. + (a_u - r_u) \bar{S}_u \left(\frac{\partial \bar{V}(u, \bar{S}_u)}{\partial \bar{S}} - \delta_u^1 \right) \right] du \\
&\quad + \int_s^t \sigma_u \bar{S}_u \left(\frac{\partial \bar{V}(u, \bar{S}_u)}{\partial \bar{S}} - \delta_u^1 \right) dW_u \\
&\implies
\end{aligned}$$

- hedge ratio

$$\delta_u^1 = \frac{\partial \bar{V}(u, \bar{S}_u)}{\partial \bar{S}}$$

- discounted PDE

$$\frac{\partial \bar{V}(u, \bar{S})}{\partial t} + \frac{1}{2} \sigma_u^2 \bar{S}^2 \frac{\partial^2 \bar{V}(u, \bar{S})}{\partial \bar{S}^2} = 0$$

$$u \in [0, T), \bar{S} \in (0, \infty)$$

with **terminal condition**

$$\bar{V}(T, \bar{S}) = \frac{f(\bar{S} S_T^0)}{S_T^0} = \frac{f(S)}{S_T^0}$$

\implies

- **PDE**

$$\frac{\partial V(u, S)}{\partial t} + r_u S \frac{\partial V(u, S)}{\partial S}$$

$$+ \frac{1}{2} \sigma_u^2 S^2 \frac{\partial^2 V(u, S)}{\partial S^2} - r_u V(u, S) = 0$$

$$u \in [0, T), S \in (0, \infty)$$

with **terminal condition**

$$V(T, S) = f(S)$$

- special boundary condition not specified
PDE may **not** provide minimal possible price

\implies

- units in savings account

$$\delta_t^0 = \frac{V(t, S_t) - \delta_t^1 S_t}{S_t^0} = \bar{V}(t, \bar{S}_t) - \delta_t^1 \bar{S}_t$$

- option price

$$\begin{aligned} V(t, S_t) &= \bar{V}(t, \bar{S}_t) S_t^0 \\ &= \delta_t^0 S_t^0 + \delta_t^1 S_t \end{aligned}$$

$$t \in [0, T]$$

\implies **strategy**

$$\delta_t = (\delta_t^0, \delta_t^1)^\top$$

such that

$$0 = d\bar{C}_t = d\bar{V}(t, \bar{S}_t) - \delta_t^1 d\bar{S}_t$$

\implies

$$d\bar{V}(t, \bar{S}_t) = \delta_t^0 d\bar{S}_t^0 + \delta_t^1 d\bar{S}_t$$

\implies **self-financing**

P&L vanishes

$$C_t = \bar{C}_t S_t^0 = 0$$

for all $t \in [0, T]$

If $V(t, S_t)$ is not fair, then there may exist less expensive hedge portfolio!

Numeraire Invariance

\bar{V} self-financing

$$\begin{aligned} dV(t, S_t) &= d(\bar{V}(t, \bar{S}_t) S_t^0) \\ &= S_t^0 d\bar{V}(t, \bar{S}_t) + \bar{V}(t, \bar{S}_t) dS_t^0 + d[S^0, \bar{V}(\cdot, \bar{S}_\cdot)]_t \\ &= S_t^0 \delta_t^1 d\bar{S}_t + (\delta_t^0 + \delta_t^1 \bar{S}_t) dS_t^0 + \delta_t^1 d[S^0, \bar{S}]_t \\ &= \delta_t^0 dS_t^0 + \delta_t^1 (S_t^0 d\bar{S}_t + \bar{S}_t dS_t^0 + d[S^0, \bar{S}]_t) \\ &= \delta_t^0 dS_t^0 + \delta_t^1 d(S_t^0 \bar{S}_t) \\ &= \delta_t^0 dS_t^0 + \delta_t^1 dS_t \end{aligned}$$

$\implies V$ also **self-financing**

The Black-Scholes Formula

- solution $c_{T,K}(t, S)$ of PDE

with call payoff function

$$f(S) = (S - K)^+$$

- $a_t = a, \sigma_t = \sigma, r_t = r$

Black & Scholes (1973)

Merton (1973a)

Black-Scholes Formula

$$c_{T,K}(t, S_t) = S_t N(d_1(t)) - K \frac{S_t^0}{S_T^0} N(d_2(t))$$

with

$$d_1(t) = \frac{\ln\left(\frac{S_t}{K}\right) + \left(r + \frac{1}{2}\sigma^2\right)(T-t)}{\sigma\sqrt{T-t}}$$

$$d_2(t) = d_1(t) - \sigma\sqrt{T-t}$$

$$t \in [0, T)$$

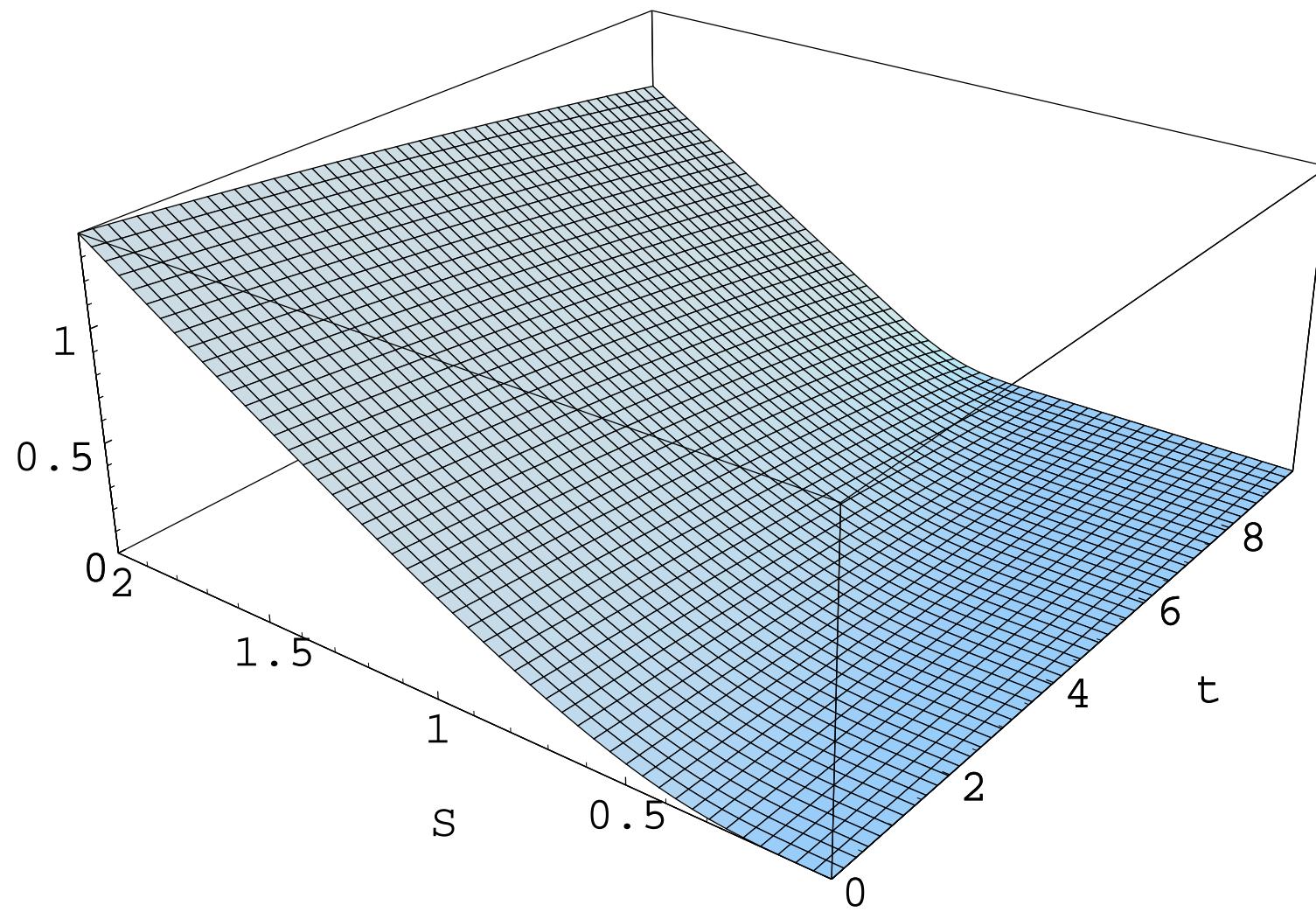
$$\hat{c}_{T,K}(t, S_t) = \frac{c_{T,K}(t, S_t)}{S_t^{\delta*}} \quad (P, \mathcal{A}) - \text{martingale}$$

minimal possible price

- $N(\cdot)$ standard **Gaussian distribution function**

with density

$$N'(x) = \frac{1}{\sqrt{2\pi}} \exp\left\{-\frac{x^2}{2}\right\}$$



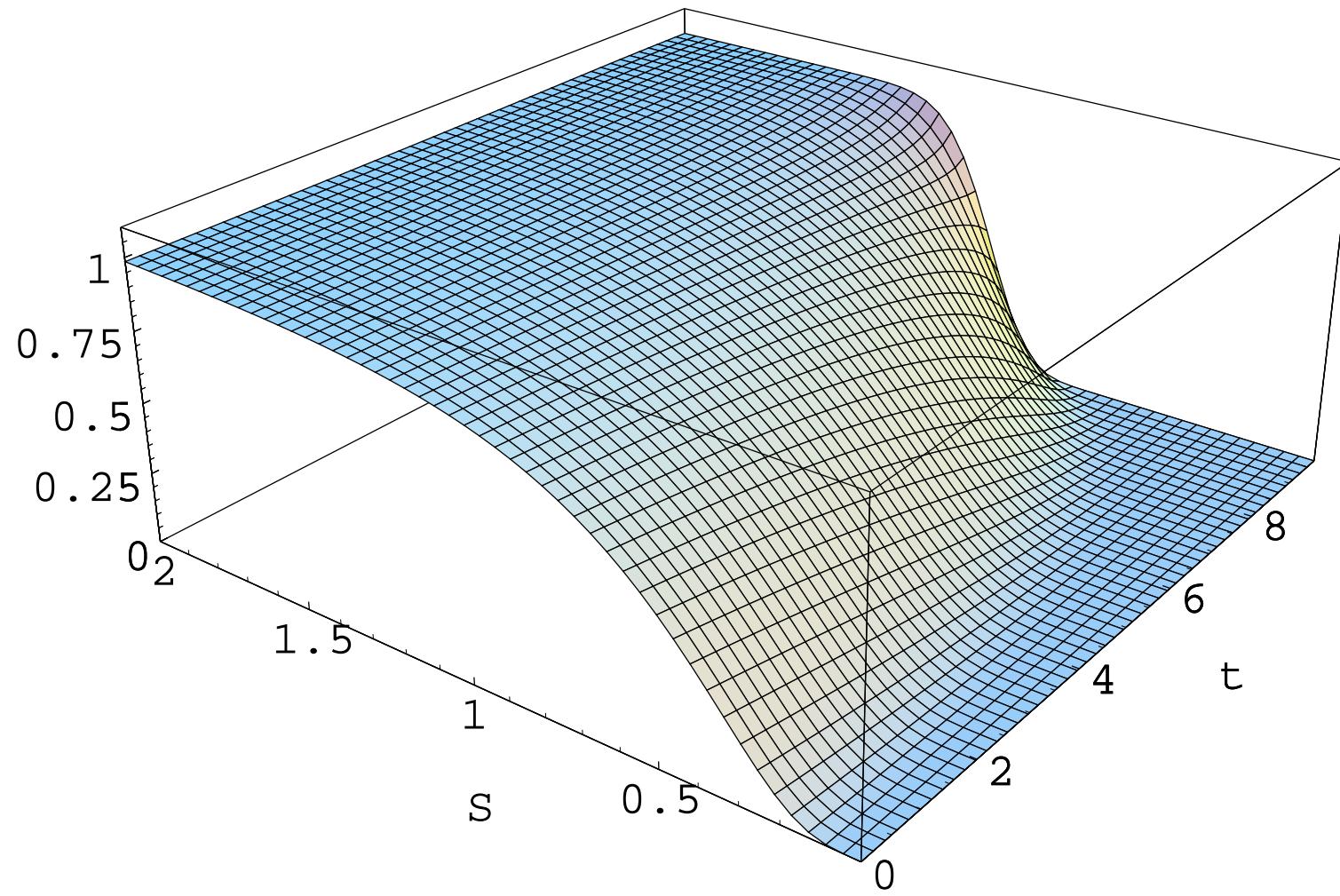
Black-Scholes European call option price.

Delta

- sensitivity with respect to the underlying S_t

$$\Delta = \frac{\partial V(t, S_t)}{\partial S} = \delta_t^1$$

$$\Delta = N(d_1(t))$$



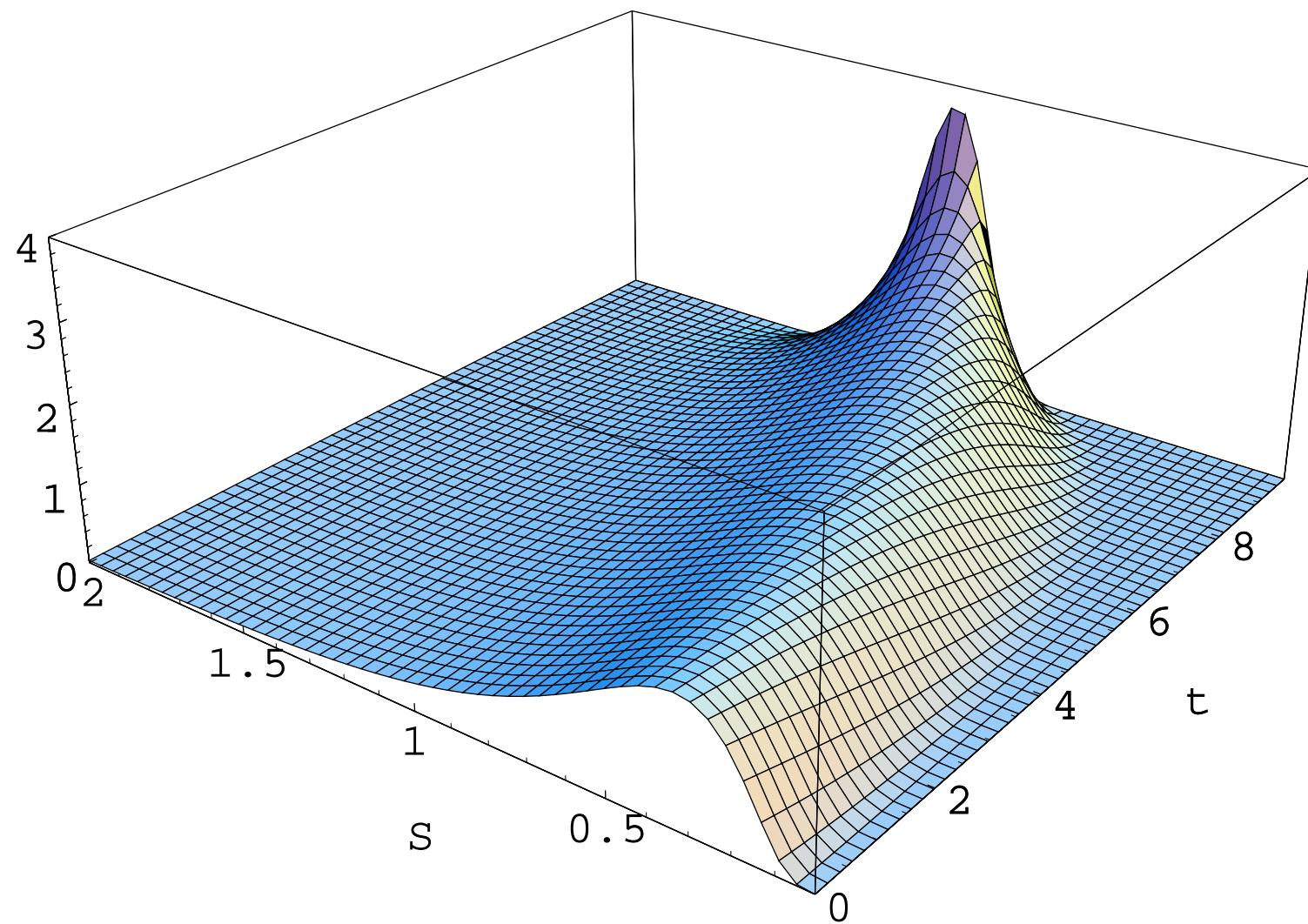
Delta as a function of t and S_t .

Gamma

- sensitivity of the hedge ratio with respect to underlying

$$\Gamma = \frac{\partial \Delta}{\partial S} = \frac{\partial^2 V(t, S_t)}{\partial S^2}$$

$$\Gamma = N'(d_1(t)) \frac{1}{S_t \sigma \sqrt{T-t}}$$



Gamma as a function of t and S_t .

Simulation of Delta Hedging

- simulate scenario along an equidistant time discretization
- price of call option

$$c_{T,K}(t, S_t) = \delta_t^1 S_t + \delta_t^0 S_t^0$$

- hedging strategy

$$\delta = \{\delta_t = (\delta_t^0, \delta_t^1)^\top, t \in [0, T]\}$$

- hedge ratio, units in the underlying

$$\begin{aligned}
 \delta_t^1 &= N(d_1(t)) \\
 &= N\left(\frac{\ln\left(\frac{S_t}{K}\right) + \left(r + \frac{1}{2}\sigma^2\right)(T-t)}{\sigma\sqrt{T-t}} \right)
 \end{aligned}$$

- units in domestic savings account

$$\begin{aligned}
 \delta_t^0 &= -\frac{K}{S_T^0} N(d_2(t)) \\
 &= -\frac{K}{S_T^0} N\left(\frac{\ln\left(\frac{S_t}{K}\right) + \left(r - \frac{1}{2}\sigma^2\right)(T-t)}{\sigma\sqrt{T-t}}\right)
 \end{aligned}$$

- discounted P&L

$$\bar{C}_t = \bar{V}(t, \bar{S}_t) - \int_0^t \delta_s^1 d\bar{S}_s - \bar{V}(0, \bar{S}_0)$$

$$\bar{V}(t, \bar{S}_t) = \bar{V}(0, \bar{S}_0) + \int_0^t \delta_s^1 d\bar{S}_s$$

⇒ theoretically under continuous hedging

$$\bar{C}_t = 0$$

for all $t \in [0, T]$

In-the-Money Scenario

- security price ends up in-the-money $S_T > K$

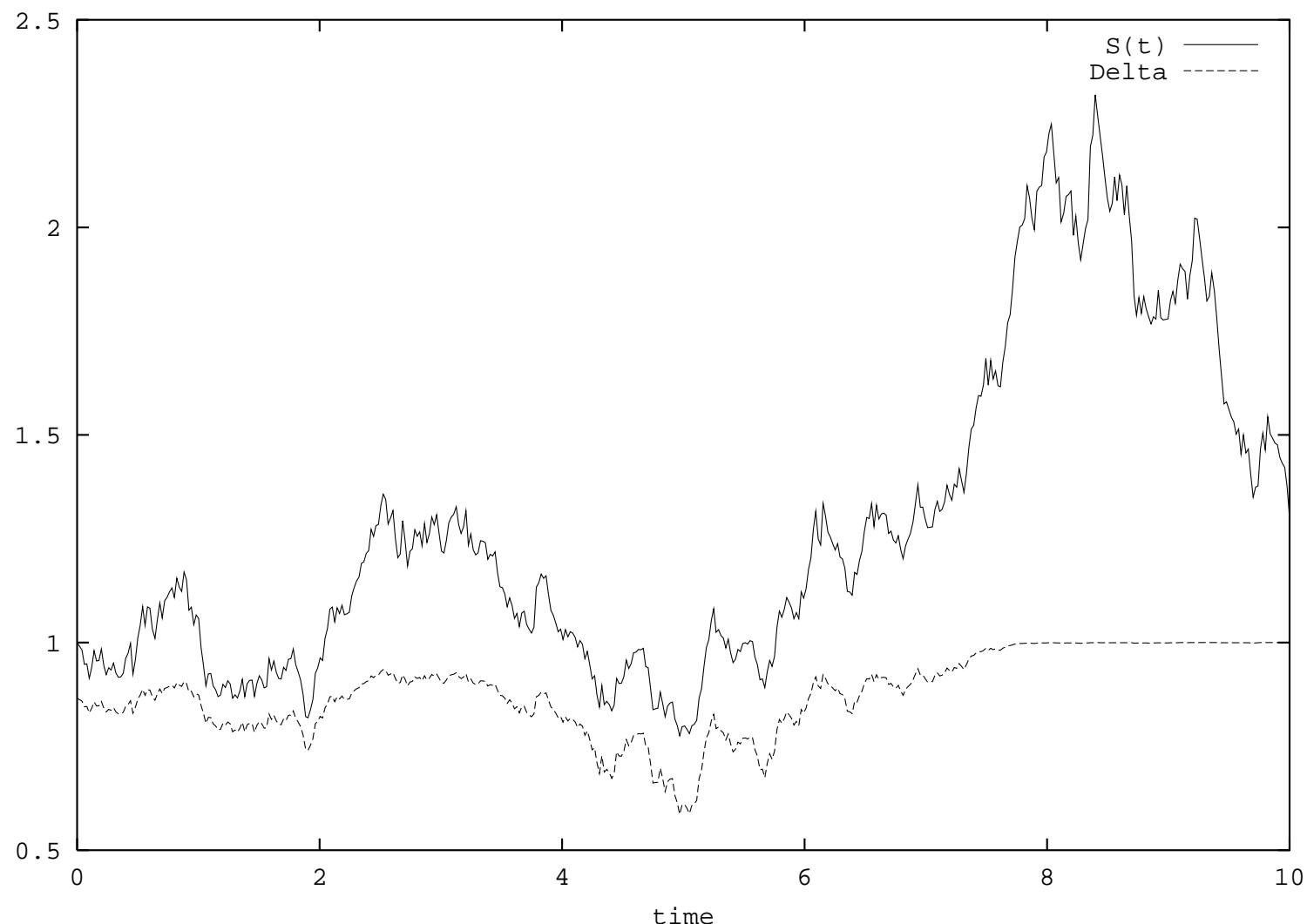
\implies hedge ratio converges to $\delta_T^1 = 1$

- self-financing strategy $\delta_t = (\delta_t^0, \delta_t^1)^\top$

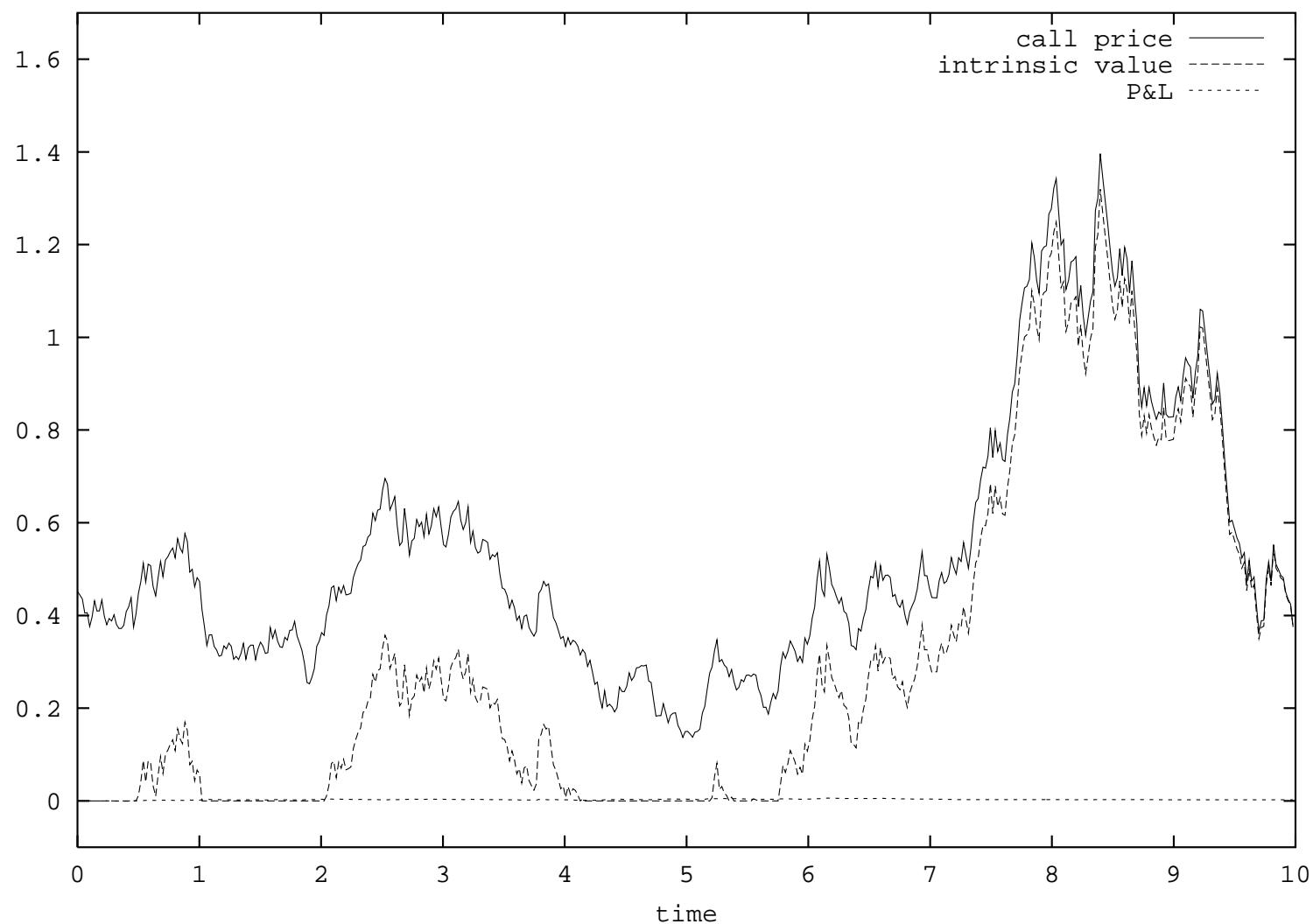
- **intrinsic value**

$$f(S_t) = (S_t - K)^+$$

- discounted P&L \bar{C}_t remains almost perfectly zero

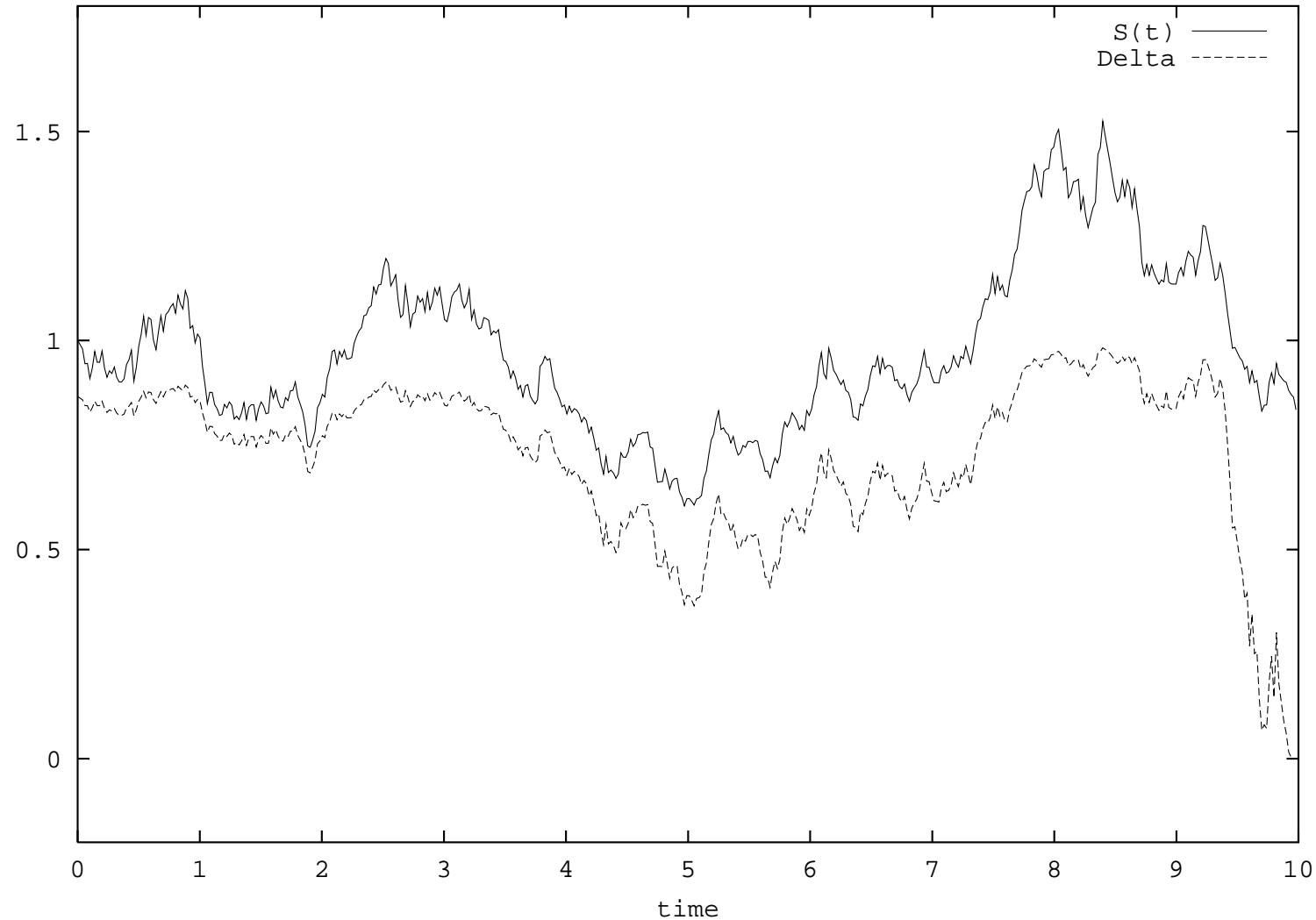


Underlying security and hedge ratio for in-the-money call.

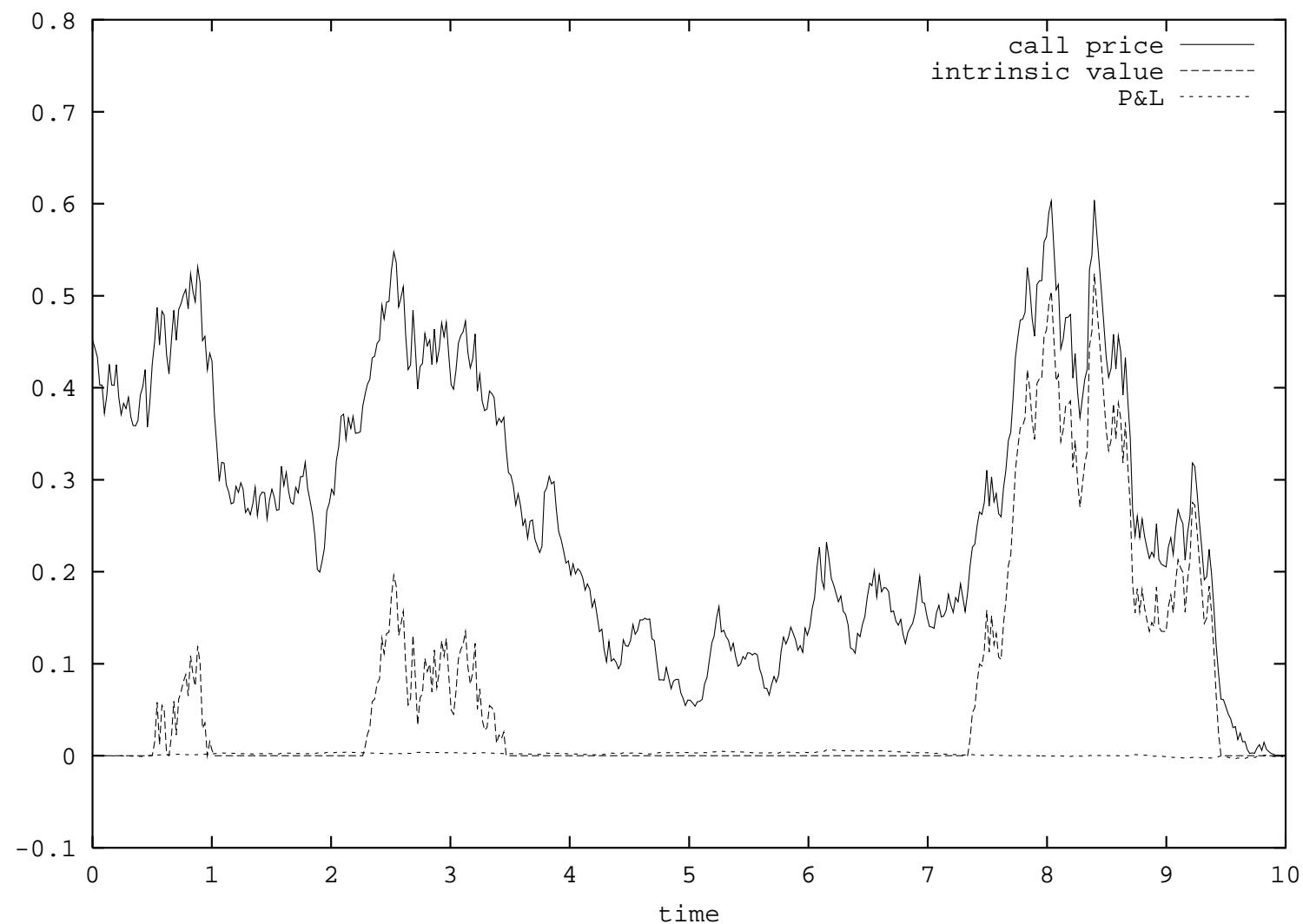


Price, intrinsic value and P&L for in-the-money call.

Out-of-the Money Scenario

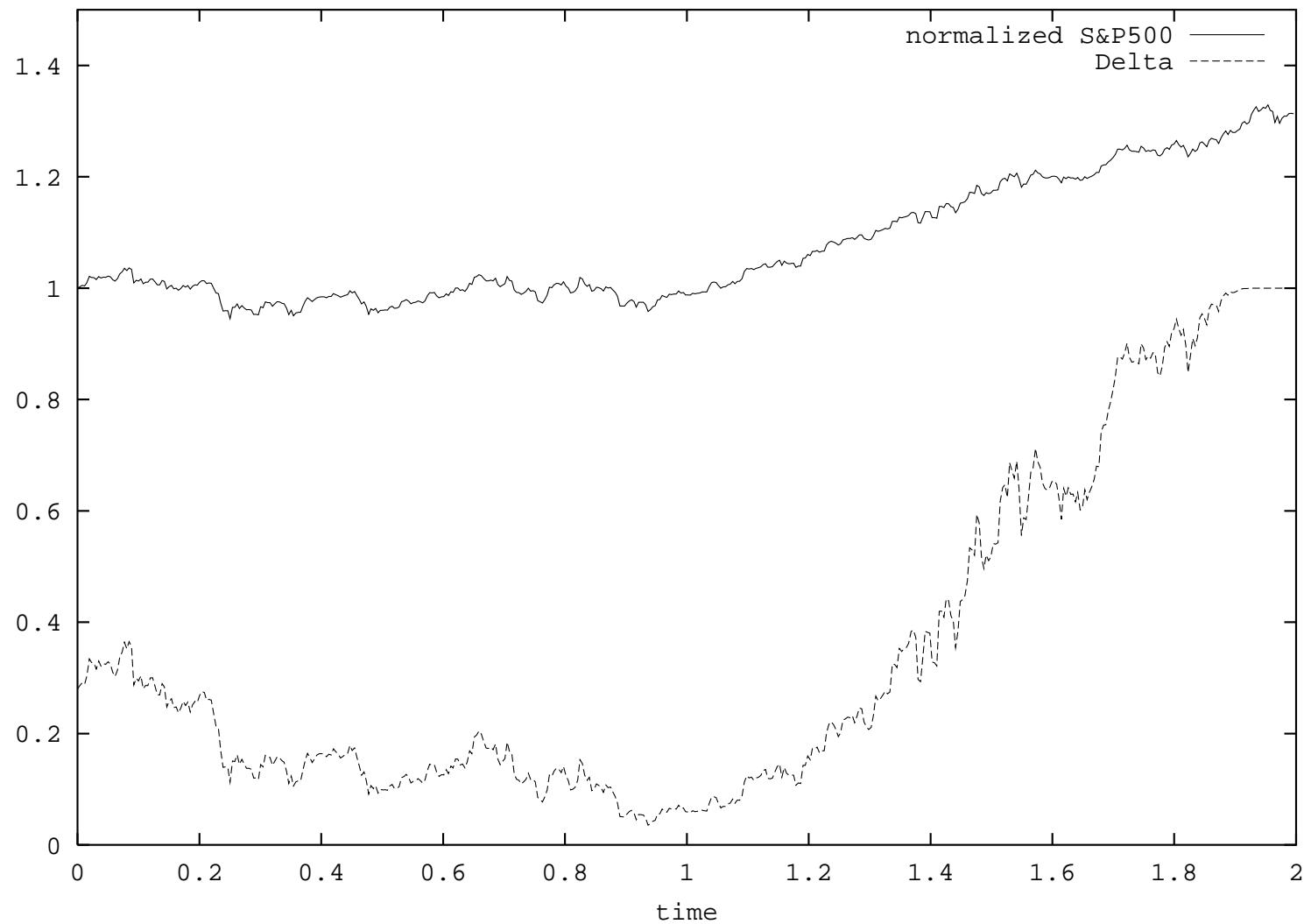


Underlying security and hedge ratio for out-of-the-money call. 266

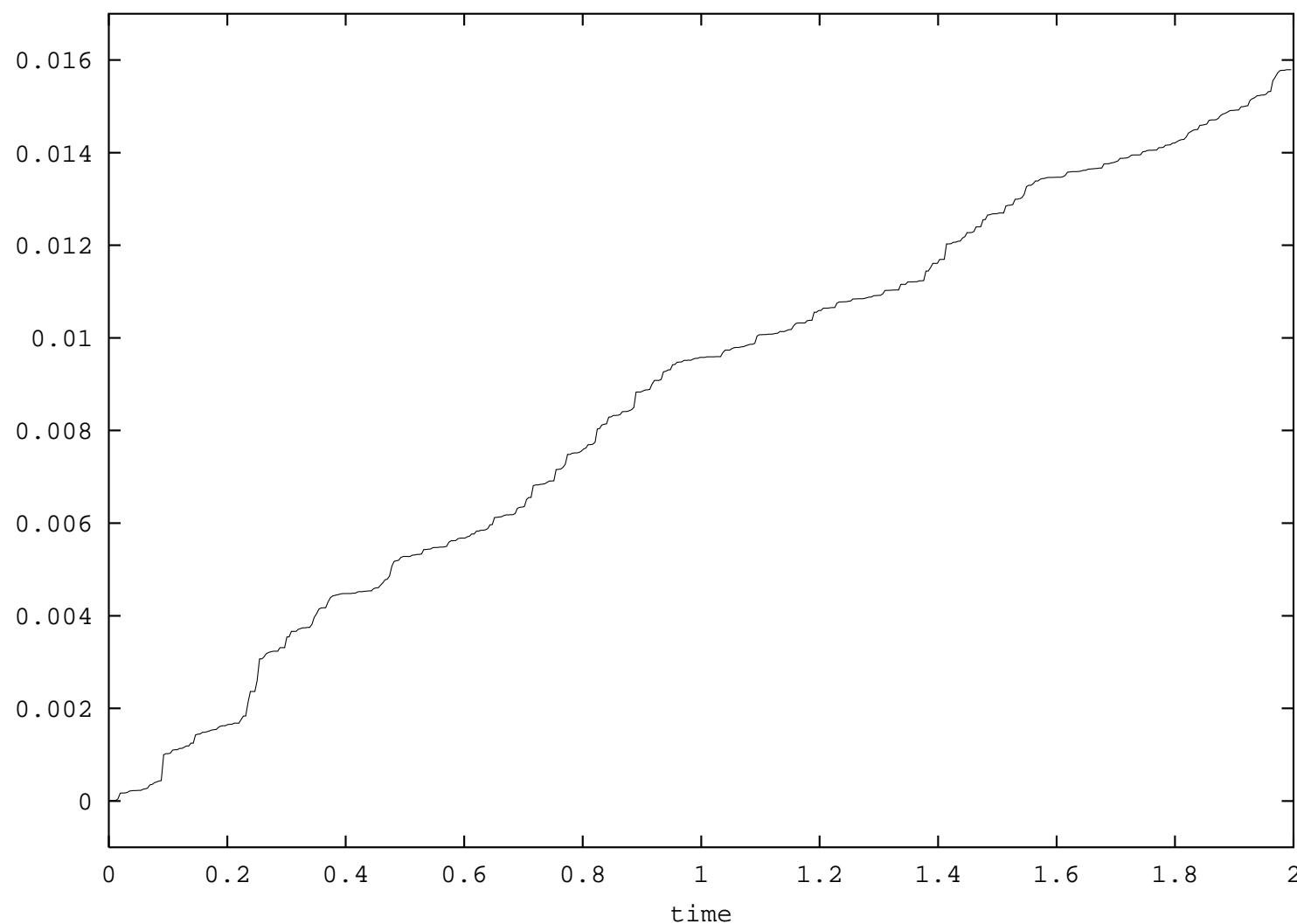


Price, intrinsic value and P&L for out-of-the-money call.

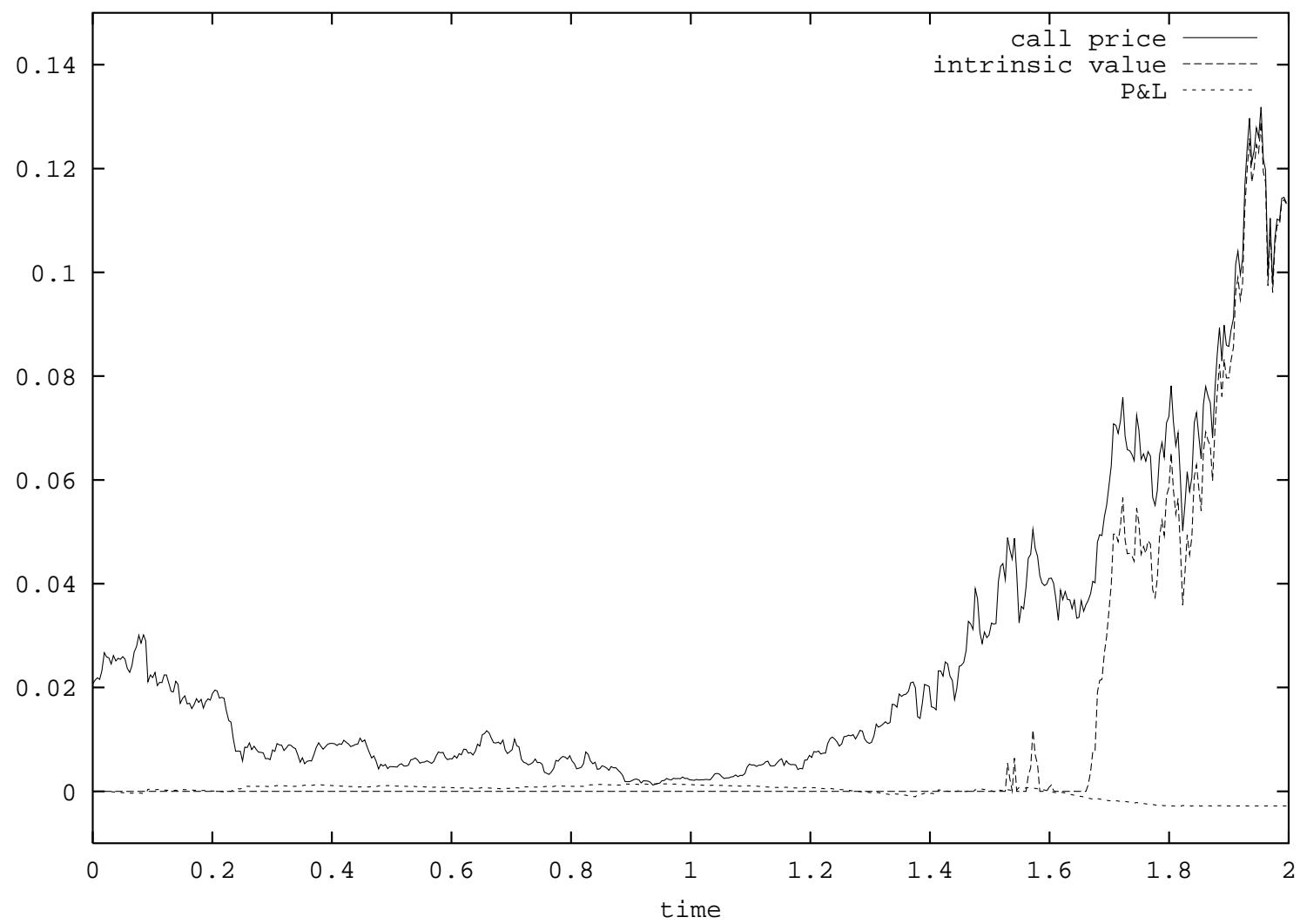
Hedging a European Call Option on the S&P500



Normalized S&P500 and hedge ratio for $K = 1.2$.



Quadratic variation of log-S&P500.



Call price on S&P500, intrinsic value and P&L.

Benchmarked Risk Minimization

Risk Minimization

- Föllmer and Sondermann (1986)
- Föllmer and Schweizer (1989)
- Schweizer (1991, 2001)
- Biagini et al. (1999)
- Biagini and Cretarola (2009)
- Du & Pl. (2012)

Financial Market

- d primary security accounts
- $S_t^{i,j}$ - j th primary security account in i th security denomination

- locally riskless savings account

$$S_t^{i,i} = S_0^{i,i} \exp \left\{ \int_0^t r_s^i ds \right\}$$

- portfolio

$$S_t^{i,\delta} = \sum_{j=1}^d \delta_t^j S_t^{i,j}$$

- self-financing portfolio

$$S_t^{i,\delta} = S_0^{i,\delta} + \sum_{j=1}^d \int_0^t \delta_s^j dS_s^{i,j}$$

Numéraire Portfolio

- $S_t^{i,\delta*}$ numéraire portfolio

$$\hat{S}_t^\delta = \frac{S_t^{i,\delta}}{S_t^{i,\delta*}} \geq E_t(\hat{S}_s^\delta)$$

$$0 \leq t \leq s < \infty$$

- supermartingale property

- benchmarked j th primary security account

$$\hat{S}_t^j = \frac{S_t^{i,j}}{S_t^{i,\delta*}}$$

local martingale; supermartingale

- matrix valued optional covariation process

$$[\hat{\mathbf{S}}] = \{ [\hat{\mathbf{S}}]_t = ([\hat{S}^i, \hat{S}^j]_t)_{i,j=1}^d, t \in [0, \infty) \}$$

- benchmarked primary security accounts

$$\hat{\mathbf{S}} = \{ \hat{\mathbf{S}}_t = (\hat{S}_t^1, \dots, \hat{S}_t^d)^\top, t \in [0, \infty) \}$$

Definition A dynamic trading strategy v , initiated at time $t = 0$, $v = \{v_t = (\eta_t, \vartheta_t^1, \dots, \vartheta_t^d)^\top, t \in [0, \infty)\}$, where $\vartheta = \{\vartheta_t = (\vartheta_t^1, \dots, \vartheta_t^d)^\top, t \in [0, \infty)\}$ forms benchmarked self-financing part $\vartheta_t^\top \hat{S}_t$ of benchmarked price process

$$\hat{V}_t^v = \vartheta_t^\top \hat{S}_t + \eta_t$$

ϑ is predictable

$$\int_0^t \vartheta_u^\top d[\hat{S}]_u \vartheta_u < \infty$$

$\eta = \{\eta_t, t \in [0, \infty)\}$ adapted with $\eta_0 = \mathbf{0}$ monitors the benchmarked non-self-financing part of

$$\hat{V}_t^v = \hat{V}_0^v + \int_0^t \vartheta_s^\top d\hat{S}_s + \eta_t$$

\hat{V}_t^v forms a supermartingale.

In general, capital has to be added or removed to match desired price

$$\hat{V}_t^v = \hat{S}_t^\delta = \sum_{j=1}^d \delta_t^j \hat{S}_t^j$$

with

$$\delta_t^j = \vartheta_t^j + \eta_t \delta_{*,t}^j \ .$$

- benchmarked contingent claim \hat{H}_T
dynamic trading strategy v *delivers* \hat{H}_T if

$$\hat{V}_T^v = \hat{S}_T^\delta = \hat{H}_T$$

\implies *replicable* if self-financing

Proposition *If for \hat{H}_T a self-financing benchmarked portfolio $\hat{S}^{\delta_{\hat{H}_T}}$ exists, satisfying*

$$\hat{S}_t^{\delta_{\hat{H}_T}} = E(\hat{H}_T | \mathcal{F}_t)$$

for all $t \in [0, T]$ P -a.s., then this provides least expensive hedge for \hat{H}_T .

Real World Pricing

- least expensive replication by a self-financing benchmarked portfolio

\implies

real world pricing formula

$$\hat{S}_t^{\delta_{\hat{H}_T}} = E_t(\hat{H}_T)$$

minimal possible price process

Definition For a dynamic trading strategy v which delivers

$$\hat{S}_t^\delta = \sum_{j=1}^d \delta_t^j \hat{S}_t^j$$

benchmarked P&L

$$\hat{C}_t^\delta = \hat{S}_t^\delta - \sum_{j=1}^d \int_0^t \vartheta_u^j d\hat{S}_u^j - \hat{S}_0^\delta$$

Corollary $\hat{C}_t^\delta = \eta_t$ for $t \in [0, \infty)$.

- usually fluctuating benchmarked P&L
- intrinsic risk

What criterion would be most natural?

- symmetric view with respect to all primary security accounts, including the domestic savings account
- pooling in large trading book
 \implies vanishing total hedge error

Proposition $\hat{H}_{T,l}$, \hat{V}^{v_l} with

\hat{C}^{v_l} independent square integrable martingales with

$$E \left(\left(\frac{\hat{C}_t^{v_l}}{\hat{V}_0^{v_l}} \right)^2 \right) \leq K_t < \infty \text{ for } l \in \{1, 2, \dots\}, t \in [0, T], T \in [0, \infty).$$

at initial time well diversified trading book holds equal fractions

- total benchmarked wealth $\hat{U}_t = \frac{\hat{U}_0}{m} \sum_{l=1}^m \frac{\hat{V}_t^{v_l}}{\hat{V}_0^{v_l}}$
- total benchmarked P&L

$$\hat{R}_m(t) = \frac{\hat{U}_0}{m} \sum_{l=1}^m \frac{\hat{C}_t^{v_l}}{\hat{V}_0^{v_l}},$$

\Rightarrow

$$\lim_{m \rightarrow \infty} \hat{R}_m(t) = 0$$

P -a.s.

Definition A dynamic trading strategy v , initiated at time zero, is called locally real-world mean-self-financing if its adapted process $\eta_t = \hat{C}_t^\delta$ forms an local martingale.

- **classical risk minimization**

Föllmer and Sondermann (1986)

Föllmer and Schweizer (1989)

Schweizer (1991, 1995)

Schweizer (1999)

Föllmer-Schweizer decomposition assuming a risk neutral probability measure

Buckdahn (1993) Schweizer (1994) Stricker (1996)

Delbaen et al. (1997) Pham et al. (1998)

- **local risk minimization**

Bouleau and Lamberton (1989)

Duffie and Richardson (1991)

Schweizer (1994)

Biagini et al. (1999)

Schweizer (2001)

- Benchmarked P&L is *orthogonal* if

$$\eta_t \hat{S}_t$$

is vector local martingale.

- above product has no trend
- benchmarked hedge error orthogonal to any benchmarked traded wealth

- $\mathcal{V}_{\hat{H}_T}$ set of locally real-world mean-self-financing dynamic trading strategies v initiated at time zero with

$$\hat{V}_t^v = \hat{S}_t^\delta, \quad t \geq 0$$

which deliver \hat{H}_T

with orthogonal benchmarked P&L

- market participants prefer **more for less**

\Rightarrow **Benchmarked Risk Minimization**

For \hat{H}_T strategy $\tilde{v} \in \mathcal{V}_{\hat{H}_T}$

benchmarked risk minimizing (BRM) if for all $v \in \mathcal{V}_{\hat{H}_T}$ with $\hat{V}_t^v = \hat{S}_t^\delta$

price $\hat{S}_t^{\tilde{\delta}}$ is minimal

$$\hat{S}_t^{\tilde{\delta}} \leq \hat{S}_t^\delta$$

P -a.s. for all $t \in [0, T]$.

Regular Benchmarked Contingent Claims

\hat{H}_T is called *regular* if

$$\hat{H}_T = E_t(\hat{H}_T) + \sum_{j=1}^d \int_t^T \vartheta_{\hat{H}_T}^j(s) d\hat{S}_s^j + \eta_{\hat{H}_T}(T) - \eta_{\hat{H}_T}(t)$$

$\vartheta_{\hat{H}_T}$ - predictable

$\eta_{\hat{H}_T}$ - local martingale, orthogonal to \hat{S}

Theorem A regular \hat{H}_T has a BRM strategy $v = \{v_t = (\eta_t, \vartheta_t^1, \dots, \vartheta_t^d)^\top, t \in [0, T]\} \in \mathcal{V}_{\hat{H}_T}$ with

$$\hat{S}_t^\delta = \hat{V}_t^v = E_t(\hat{H}_T) ,$$

$$\hat{S}_T^\delta = \hat{H}_T ,$$

and orthogonal benchmarked P&L: $\hat{C}_t^\delta = \eta_{\hat{H}_T}(t)$

\implies

\hat{C}_t^δ - local martingale

$\hat{C}_t^\delta \hat{\mathbf{S}}_t$ - vector local martingale

orthogonal

hedgeable part:

$$E_t(\hat{H}_T) - \hat{C}_t^\delta = E_0(\hat{H}_T) + \int_0^t \vartheta_{\hat{H}_T}^\top(s) d\hat{S}_s .$$

local martingale

- benchmarked price $\hat{V}_t^v = E_t(\hat{H}_T)$ minimal
- do not request square integrability
- no equivalent risk neutral probability measure required
- need only to check zero drift of $\eta_{\hat{H}_T}$
- and $\eta_{\hat{H}_T}(t), \hat{S}_t$ orthogonality

Hedging a Regular Claim

- **benchmarked contingent claim** \hat{H}_T driven by continuous local martingales W^1, W^2, \dots, W^d orthogonal to each other
- **benchmarked primary security account** $\hat{S}_t^j, j \in \{1, \dots, d\}$

$$d\hat{S}_t^j = -\hat{S}_t^j \sum_{k=1}^{d-1} \theta_t^{j,k} dW_t^k$$

- $d \times d$ matrix $\Phi_t = [\Phi_t^{i,k}]_{i,k=1}^d$

$$\Phi_t^{i,k} = \begin{cases} \theta_t^{i,k} & \text{for } k \in \{1, \dots, d-1\} \\ 1 & \text{for } k = d \end{cases}$$

for $t \in [0, T]$.

Proposition assume Φ_t invertible and $\hat{V}_t = E(\hat{H}_T | \mathcal{F}_t)$ has representation

$$\hat{V}_t = \hat{V}_0 + \sum_{k=1}^d \int_0^t x_s^k dW_s^k + \int_0^t x_s^d dW_s^d .$$

Then \hat{H}_T is a regular benchmarked contingent claim

$$\begin{aligned}\hat{V}_t^{v_{\hat{H}_T}} &= \hat{V}_t \quad \text{for all } t \in [0, T] \\ \vartheta_{\hat{H}_T}(t) &= \text{diag}(\hat{S}_t)^{-1} (\Phi_t^\top)^{-1} \xi_t \\ \eta_{\hat{H}_T}(t) &= \int_0^t x_s^d dW_s^d\end{aligned}$$

with $\xi_t = (x_t^1, \dots, x_t^{d-1}, \hat{V}_0 + \sum_{k=1}^{d-1} \int_0^t x_s^k dW_s^k)^\top$.

$$\bar{S}^{i,j}_t = \frac{\hat{S}^j_t}{\hat{S}^i_t}$$

$$d\bar{S}^{i,j}_t=\bar{S}^{i,j}_t\sum_{k=1}^{d-1}(\theta^{i,k}_t-\theta^{j,k}_t)(\theta^{i,k}_tdt+dW^k_t)$$

$$b^{i,j,k}_t=(\theta^{i,k}_t-\theta^{j,k}_t)~.$$

$$b^d_t=[b^{d,j,k}_t]_{j,k=1}^{d-1,d-1}$$

$$301\\$$

Proposition *The matrix Φ_t is invertible if and only if b_t^d is invertible.*

Proof: elementary transform of b_t^d

$$\Phi_t \longleftrightarrow \left(\begin{array}{ccc|c} & b_t^d & & 0 \\ \hline \theta_t^{d,1} & \dots & \theta_t^{d,d-1} & 1 \end{array} \right)$$

Φ_t has full rank if and only if b_t^d has full rank

Comparison with Quadratic Criterion

- $\int_0^t \vartheta_{\hat{H}_T}^\top(s) d\hat{S}_s$ and $\eta_{\hat{H}_T}(t)$ independent, square integrable martingales
- $\hat{S}_t^1, \dots, \hat{S}_t^d$ and $\eta_{\hat{H}_T}(t)$ independent, square integrable martingales
- $\eta_{\hat{H}_T}$ orthogonal to benchmarked primary security accounts
- \hat{H}_T square integrable
- $\hat{V}_t^v = \hat{S}_t^\delta$ delivers \hat{H}_T

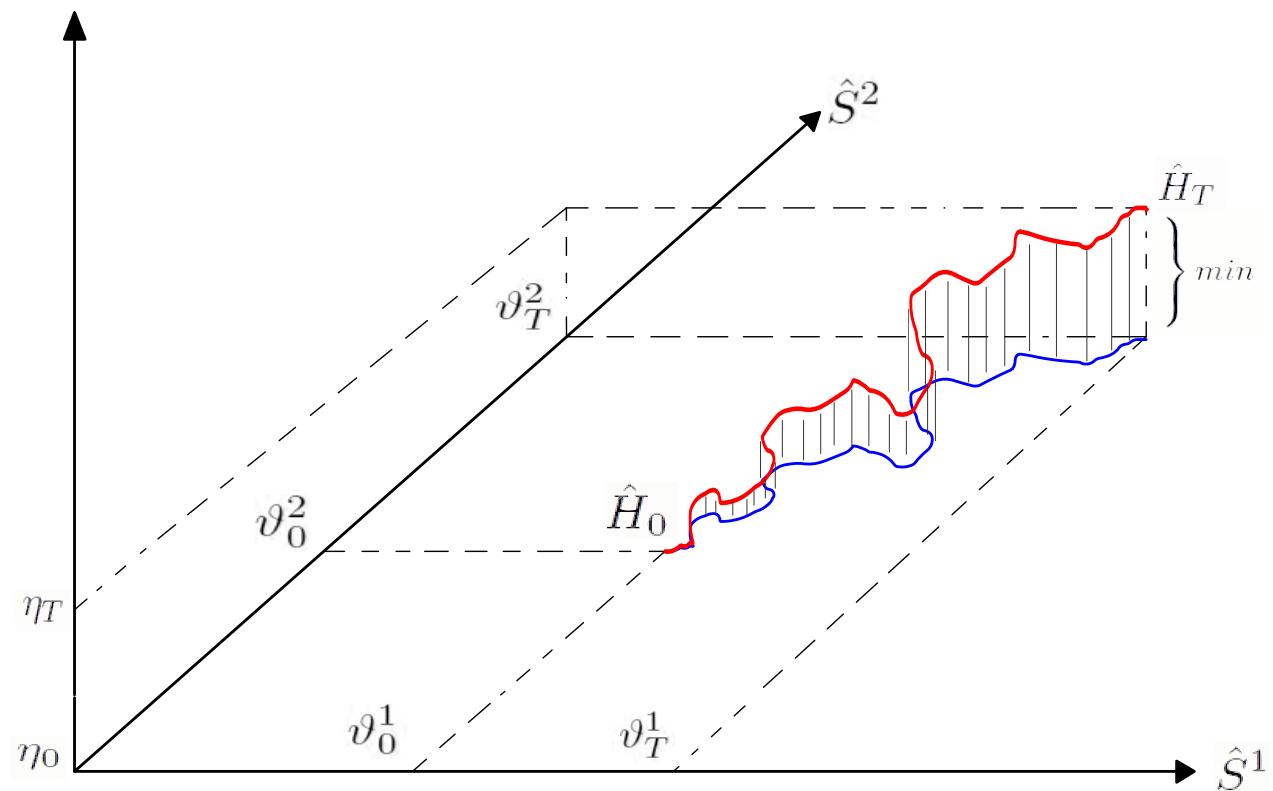
Quadratic Criterion

$$E \left((\hat{C}_T^\delta)^2 \right) = E \left((\hat{H}_T - \int_0^T \vartheta_s^\top d\hat{S}_s - \hat{S}_0^\delta)^2 \right) \Rightarrow \min$$

$$\begin{aligned}
& E \left((\hat{C}_T^\delta)^2 \right) \\
&= E \left(\left(E(\hat{H}_T) - \hat{S}_0^\delta + \int_0^T (\vartheta_{\hat{H}_T}^\top(s) - \vartheta_s^\top) d\hat{S}_s + \eta_{\hat{H}_T}(T) \right)^2 \right) \\
&= E \left(E(\hat{H}_T) - \hat{S}_0^\delta \right)^2 \\
&\quad + E \left(\int_0^T (\vartheta_{\hat{H}_T}^\top(s) - \vartheta_s^\top)^2 d[\hat{S}]_s \right) + E((\eta_{\hat{H}_T}(T))^2)
\end{aligned}$$

\implies

- $\hat{S}_0^\delta = E(\hat{H}_T)$
- $\vartheta_t = \vartheta_{\hat{H}_T}(t), t \in [0, T]$
- $\eta_{\hat{H}_T}(T) = \hat{C}_T^\delta$
- second moment of benchmarked P&L becomes minimal



Nonhedgeable Contingent Claim

- savings account $S_t^{1,1} = 1$
- numéraire portfolio $S_t^{1,\delta*}$
- $H_t = E_t(H_T), t \in [0, T]$
 - independent from $S_t^{1,\delta*}$
 - continuous

Benchmarked risk minimization

- real world pricing formula

$$\hat{S}_t^{\delta_{\hat{H}_T}} = E_t(\hat{H}_T) = E_t(H_T)E_t(\hat{S}_T^1) = H_t \hat{P}(t, T)$$

$$\hat{P}(t, T) = E_t(\hat{S}_T^1)$$

$$d\hat{P}(t, T) = \frac{\partial \hat{P}(t, T)}{\partial \hat{S}^1} d\hat{S}_t^1$$

- benchmarked NP $\hat{S}_t^{\delta^*} = \frac{S_t^{1,\delta^*}}{S_t^{1,\delta^*}} = 1$
- benchmarked P&L

$$\begin{aligned} d\hat{C}_t^{\delta_{\hat{H}_T}} &= d\hat{S}_t^{\delta_{\hat{H}_T}} - \vartheta_{\hat{H}_T}^1(t) d\hat{S}_t^1 \\ &= \left(H_t \frac{\partial \hat{P}(t, T)}{\partial \hat{S}^1} - \vartheta_{\hat{H}_T}^1(t) \right) d\hat{S}_t^1 + \hat{P}(t, T) dH_t \end{aligned}$$

- Assume orthogonal benchmarked primary security accounts
- product $\hat{C}^{\delta_{\hat{H}_T}} \hat{S}$ has zero drift if

$$0 = \frac{d[\hat{S}, \hat{C}^{\delta_{\hat{H}_T}}]_t}{dt} = \left(H_t \frac{\partial \hat{P}(t, T)}{\partial \hat{S}^1} - \vartheta_{\hat{H}_T}^1(t) \right) \frac{d[\hat{S}^1]_t}{dt}.$$

\implies

$$\vartheta_{\hat{H}_T}^1(t) = H_t \frac{\partial \hat{P}(t, T)}{\partial \hat{S}^1}$$

Depends on H_t !

- Benchmarked P&L

$$\eta_{\hat{H}_T}(t) = \hat{C}_t^{\delta_{\hat{H}_T}} = \int_0^t \hat{P}(s, T) dH_s$$

local martingale, orthogonal

- self-financing benchmarked hedgeable part

$$\hat{S}_t^{\delta_{\hat{H}_T}} - \eta_{\hat{H}_T}(t) = \hat{S}_0^{\delta_{\hat{H}_T}} + \int_0^t \vartheta_{\hat{H}_T}^1(s) d\hat{S}_s^1$$

- $\vartheta_{\hat{H}_T}^1(t)$ units in \hat{S}_t^1
- remaining wealth in NP
- minimizing $\frac{d}{dt} [\hat{C}^{\delta_{\hat{H}_T}}]_t$

- Regular benchmarked contingent claim

$$\begin{aligned}
\hat{H}_T &= \frac{H_T}{S_T^{1,\delta*}} \\
&= \hat{S}_0^{\delta_{\hat{H}_T}} + \int_0^T H_t \frac{\partial \hat{P}(t, T)}{\partial \hat{S}^1} d\hat{S}_t^1 + \int_0^T \hat{P}(t, T) dH_t \\
\vartheta_{\hat{H}_T}^1(t) &= H_t \frac{\partial \hat{P}(t, T)}{\partial \hat{S}^1} \\
\eta_{\hat{H}_T}(t) &= \int_0^t \hat{P}(s, T) dH_s
\end{aligned}$$

Evolving Information

- consider special case when \hat{S}^1 true martingale
- classical risk minimization can be applied
- real world pricing formula yields

$$\bar{S}_t^{\delta_{\hat{H}_T}} = \frac{\hat{S}_t^{\delta_{\hat{H}_T}}}{\hat{S}_t^1} = E_t \left(\frac{\hat{S}_T^1}{\hat{S}_t^1} \frac{H_T}{S_T^{1,1}} \right) = E_t \left(\frac{\Lambda_T}{\Lambda_t} \frac{H_T}{S_T^{1,1}} \right) = E_t^Q \left(\frac{H_T}{S_T^{1,1}} \right)$$

- equivalent minimal martingale measure Q

$$\Lambda_T = \frac{dQ}{dP}|_{\mathcal{F}_T} = \frac{\hat{S}_T^1}{\hat{S}_t^1}$$

- same initial price as benchmarked risk minimization

$$\bar{S}_0^{\delta_{\hat{H}_T}} = H_0 = \hat{S}_0^{\delta_{\hat{H}_T}} (\hat{S}_0^1)^{-1}$$

- classical risk minimization invests total initial value in the savings account

$$\vartheta_{CRM}^1(t) = H_0$$

- BRM strategy different
- $\hat{P}(t, T) = \hat{S}_t^1, \frac{\partial \hat{P}(t, T)}{\partial \hat{S}_t^1} = 1$
 \implies

$$\vartheta_{\hat{H}_T}^1(t) = H_t = E_t(H_T)$$

- intuitive and practically appealing
- under classical risk minimization, evolving information ignored
- provides less fluctuations for benchmarked P&L
- easier diversified

Modeling the Dynamics of Diversified Indices

Pl. & Rendek (2012)

- conjecture for normalized aggregate wealth dynamics
time transformed square root process
- Naive Diversification Theorem \Rightarrow equity index = proxy of numéraire portfolio
- empirical stylized facts \Rightarrow falsify models

- \Rightarrow propose realistic one factor, two component long term index model
- realistic model outside classical theory \Rightarrow benchmark approach
- exact, almost exact simulation \Rightarrow verify empirical facts, effects of estimation techniques etc.

The Affine Nature of Diversified Wealth Dynamics

Object: normalized units of wealth

Total wealth: $Y_{\tau_i}^{\Delta}$

Time steps: $\tau_i = i\Delta$

Wealth unit value: $\sqrt{\Delta}$

Economic activity: during $[\tau_i, \tau_{i+1})$ "projects"
generate each independent wealth with variance increment $v^2 \Delta^{3/2}$;
consume each $\eta \Delta$ fraction of wealth;
generate together on average $\beta \sqrt{\Delta}$ new units.

Mean for increment of aggregate wealth: $(\beta - \eta Y_{\tau_i}^\Delta) \Delta$

Variance for increment aggregate wealth: \triangleq sum of variances

- \implies proportional to number of wealth units: $\frac{Y_{\tau_i}^\Delta}{\sqrt{\Delta}}$
- \implies proportional to aggregate wealth
- \implies variance of increment of aggregate wealth is $v^2 Y_{\tau_i}^\Delta \Delta$

for $\Delta \rightarrow 0$

$$Y_{\tau_{i+1}}^\Delta - Y_{\tau_i}^\Delta = (\beta - \eta Y_{\tau_i}^\Delta) \Delta + v \sqrt{Y_{\tau_i}^\Delta} \Delta W_{\tau_i}$$

$$E(\Delta W_{\tau_i}) = 0, \quad E((\Delta W_{\tau_i})^2) = \Delta$$

conjectures drift and diffusion terms

\implies **Conjecture via
Weak convergence**

Kleoden & Pl. (1992), Alfonsi (2005), Diop (2003)

for parameters: $\beta = \eta = v = 1$

square root process:

$$dY_{\tau_t} = (1 - Y_{\tau_t}) d\tau_t + \sqrt{Y_{\tau_t}} dW_{\tau_t}$$

- Quadratic variation:

$$[Y_{\tau_\cdot}]_t = \int_0^t Y_{\tau_s} d\tau_s = \int_0^t Y_{\tau_s} M_s ds$$

- Market activity:

$$M_t = \frac{d\tau_t}{dt}$$

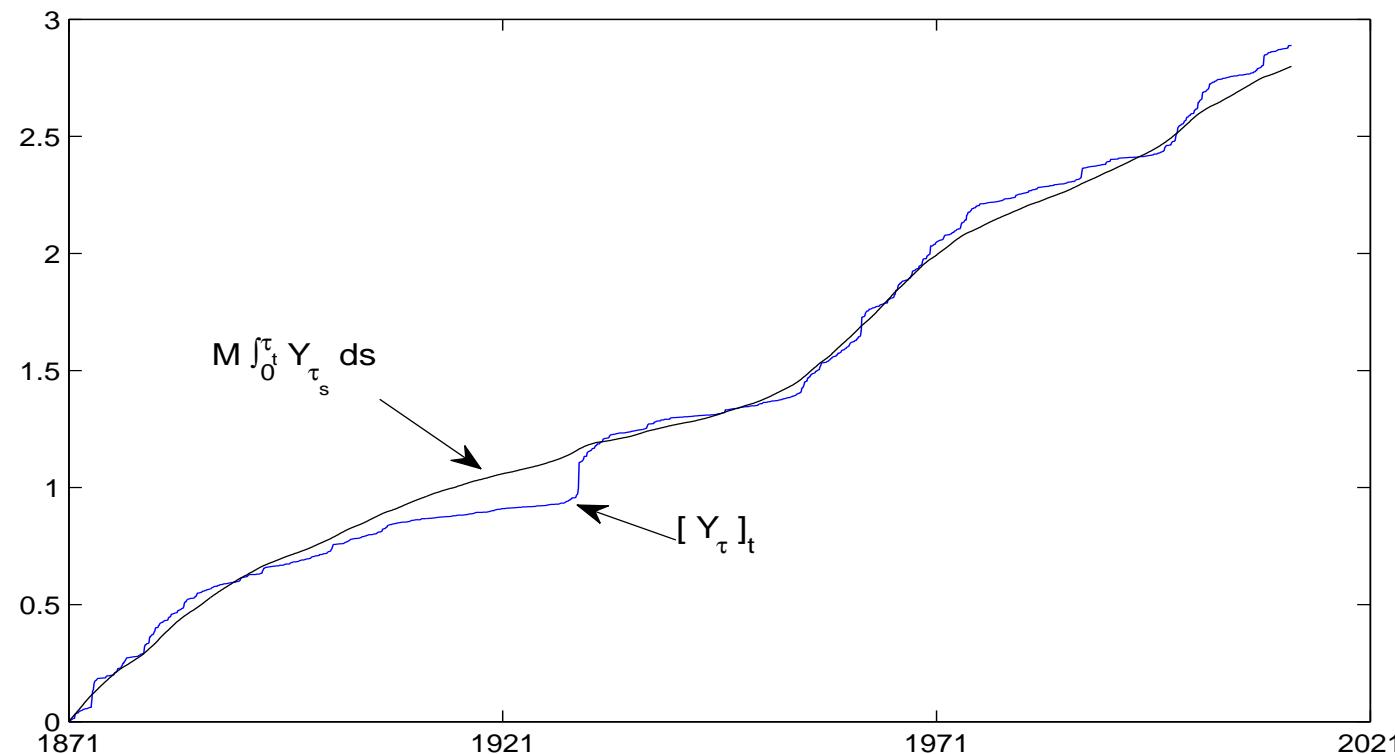
- integrate normalized index:

can we find $M = \text{const.}$ s.t.

$$M \int_0^t Y_{\tau_s} ds \approx [Y_{\tau_\cdot}]_t \quad ?$$

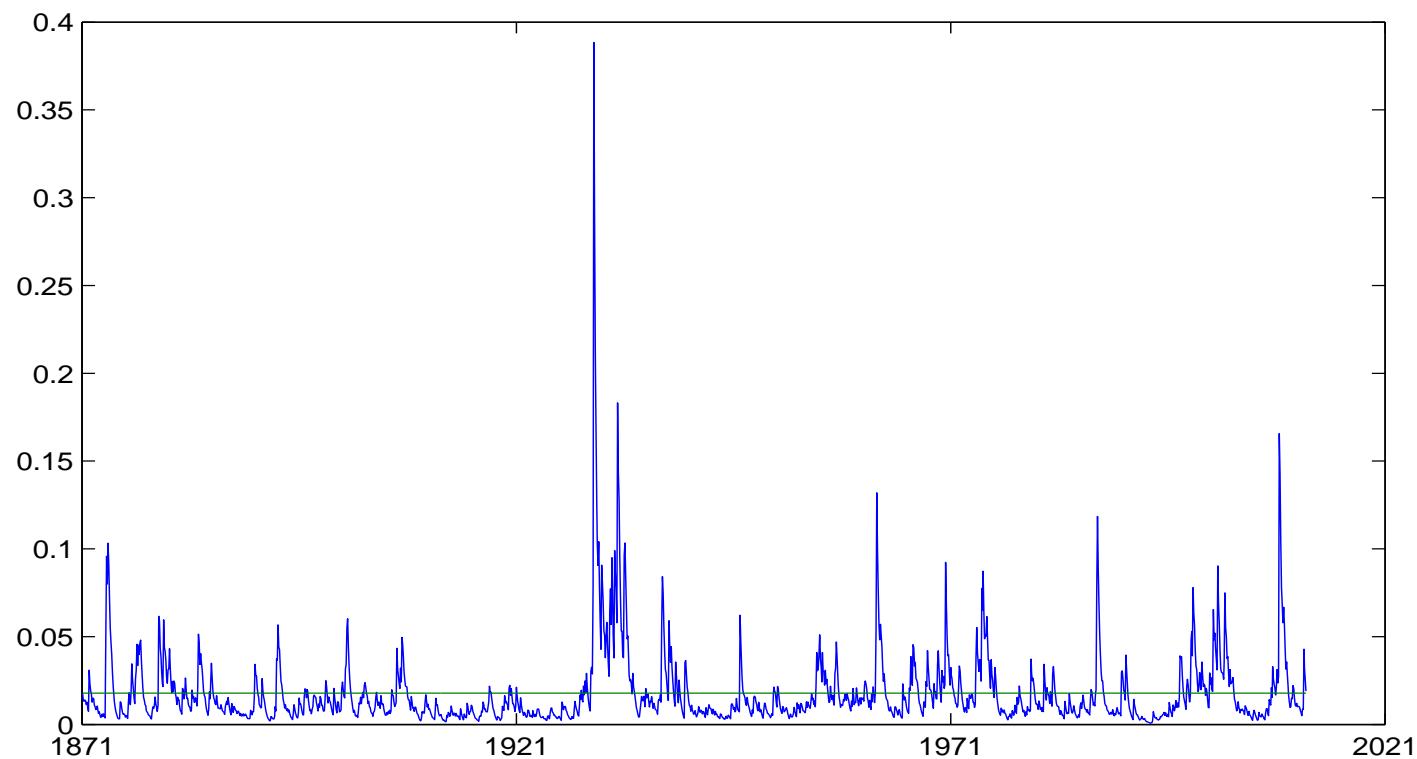
Quadratic variation and integrated normalized S&P500

monthly data, calendar time, Shiller data

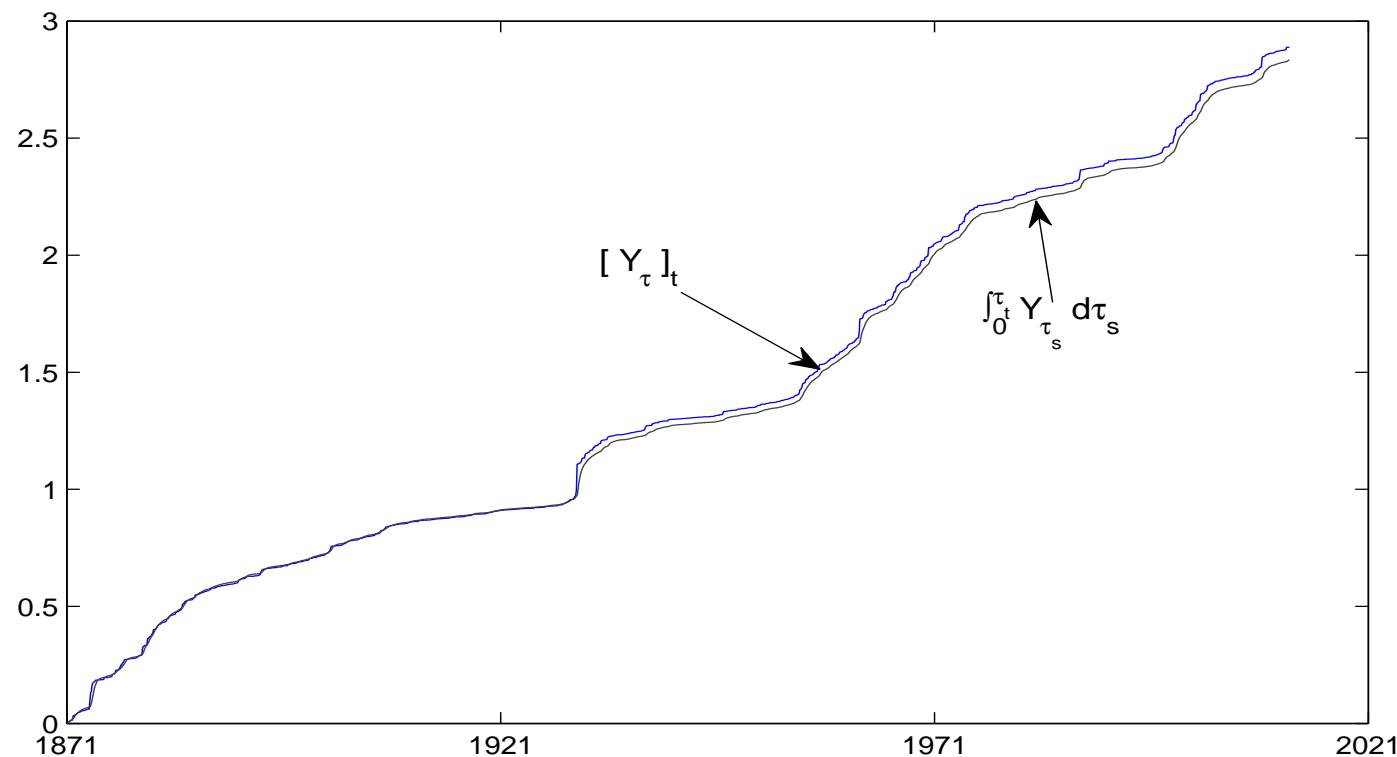


$M \approx 0.0178$
average long term fit

Market Activity: $M_t = \frac{d\tau_t}{dt}$ from model



Quadratic variation and integrated normalized S&P500 monthly data, τ -time

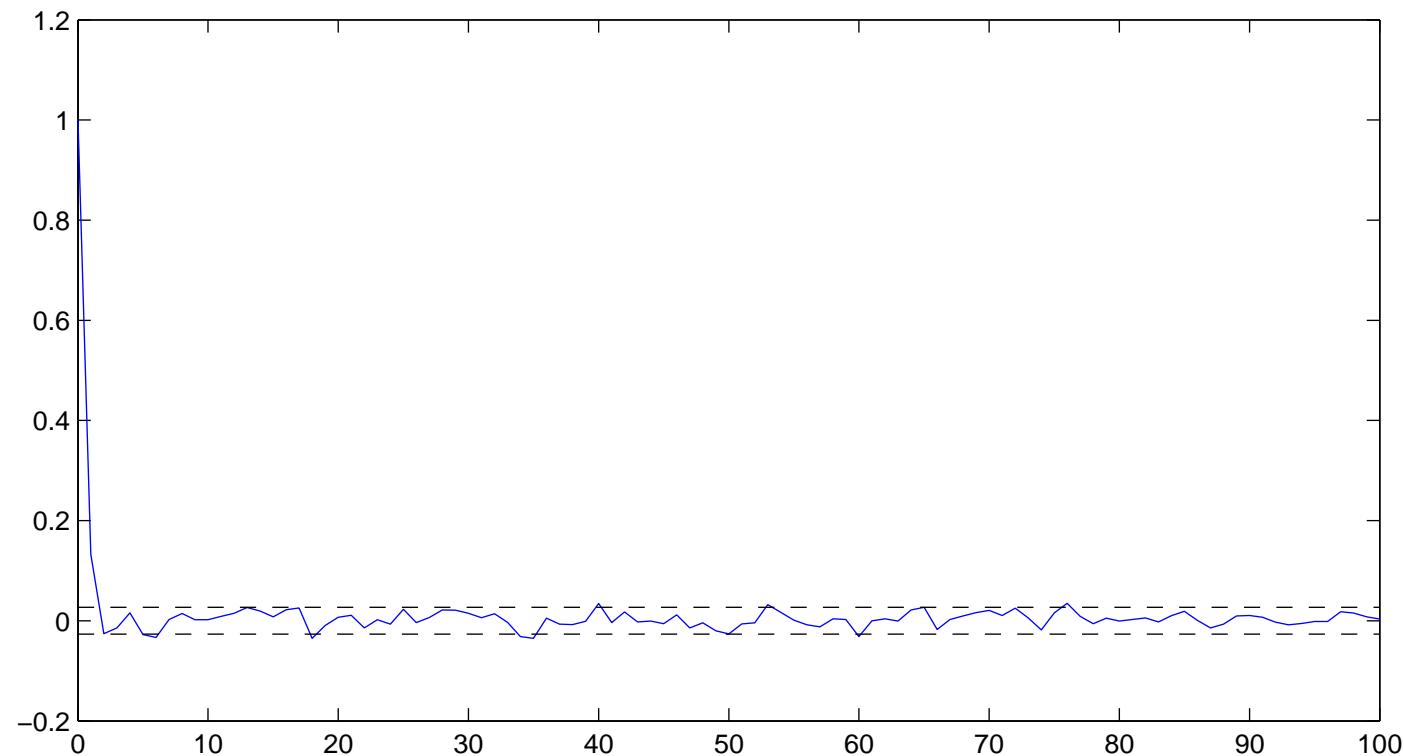


Eight Stylized Empirical Facts

- falsify potential models, Popper (1959)
- TOTMKWD in 26 currency denominations

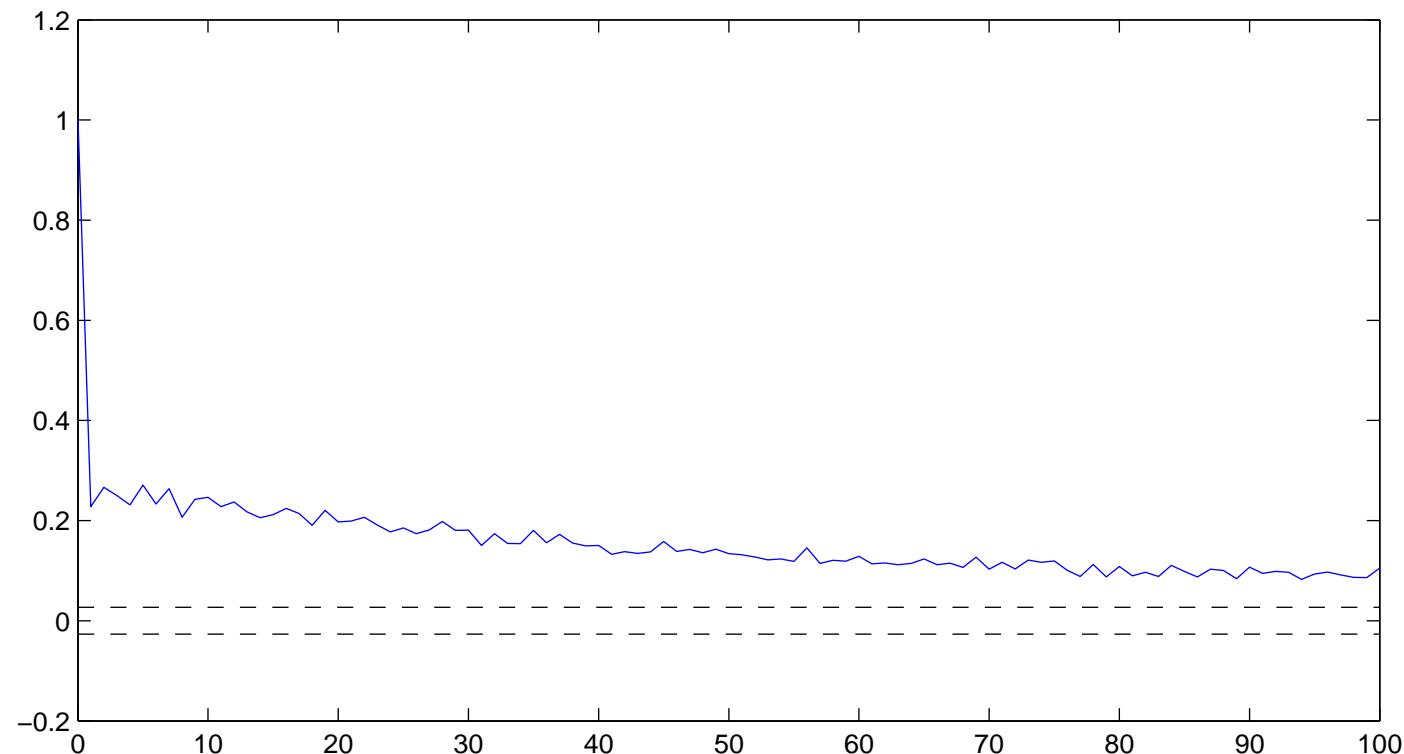
about 1000 years of daily data

(i) uncorrelated log-returns



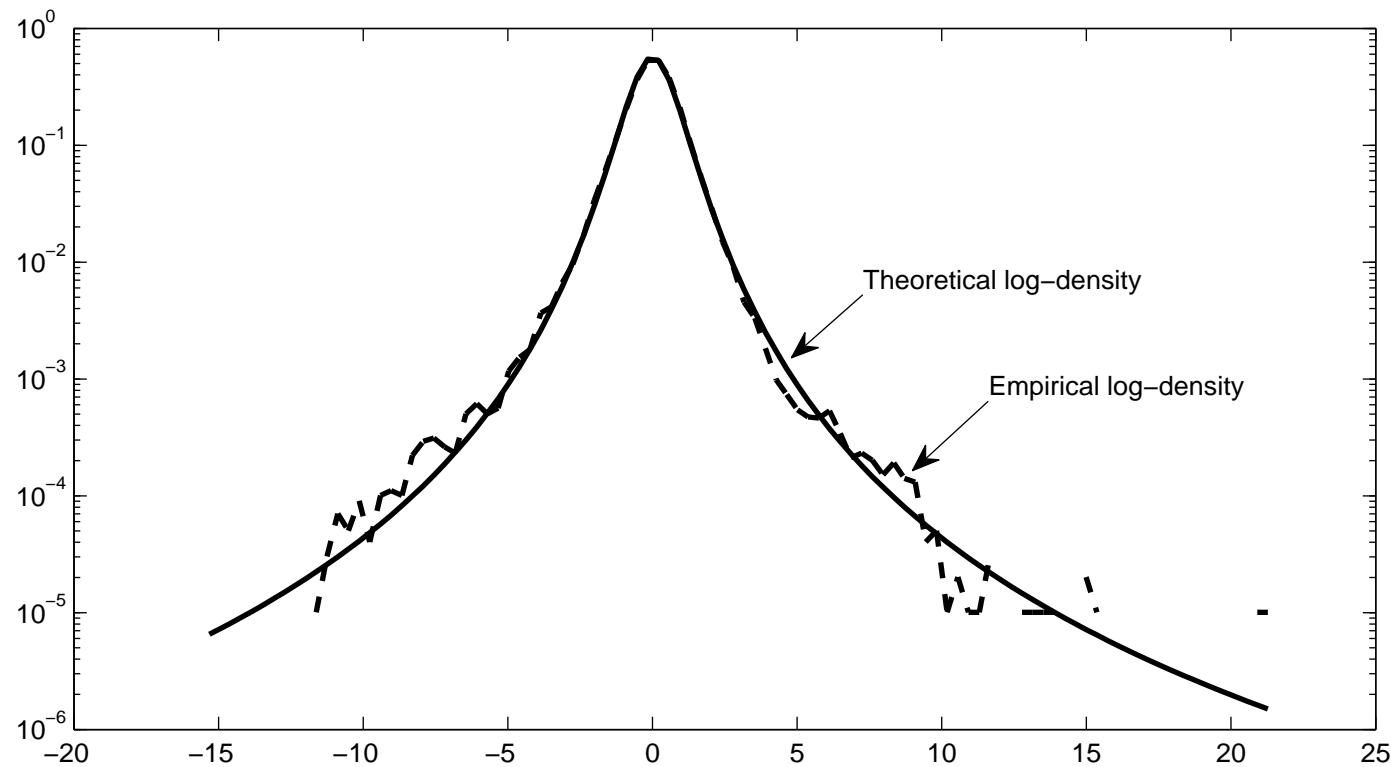
Average autocorrelation function for log-returns

(ii) correlated absolute log-returns



Average autocorrelation function for absolute log-returns

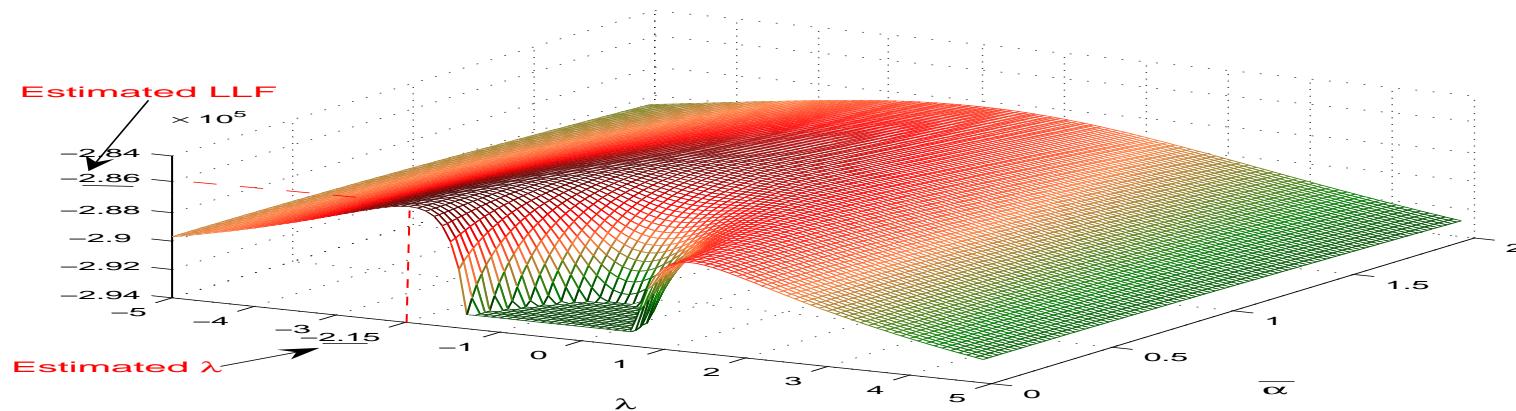
(iii) Student- t distributed log-returns



Logarithm of empirical density of normalized log-returns with Student- t density

Results for log-returns of the EWI104s

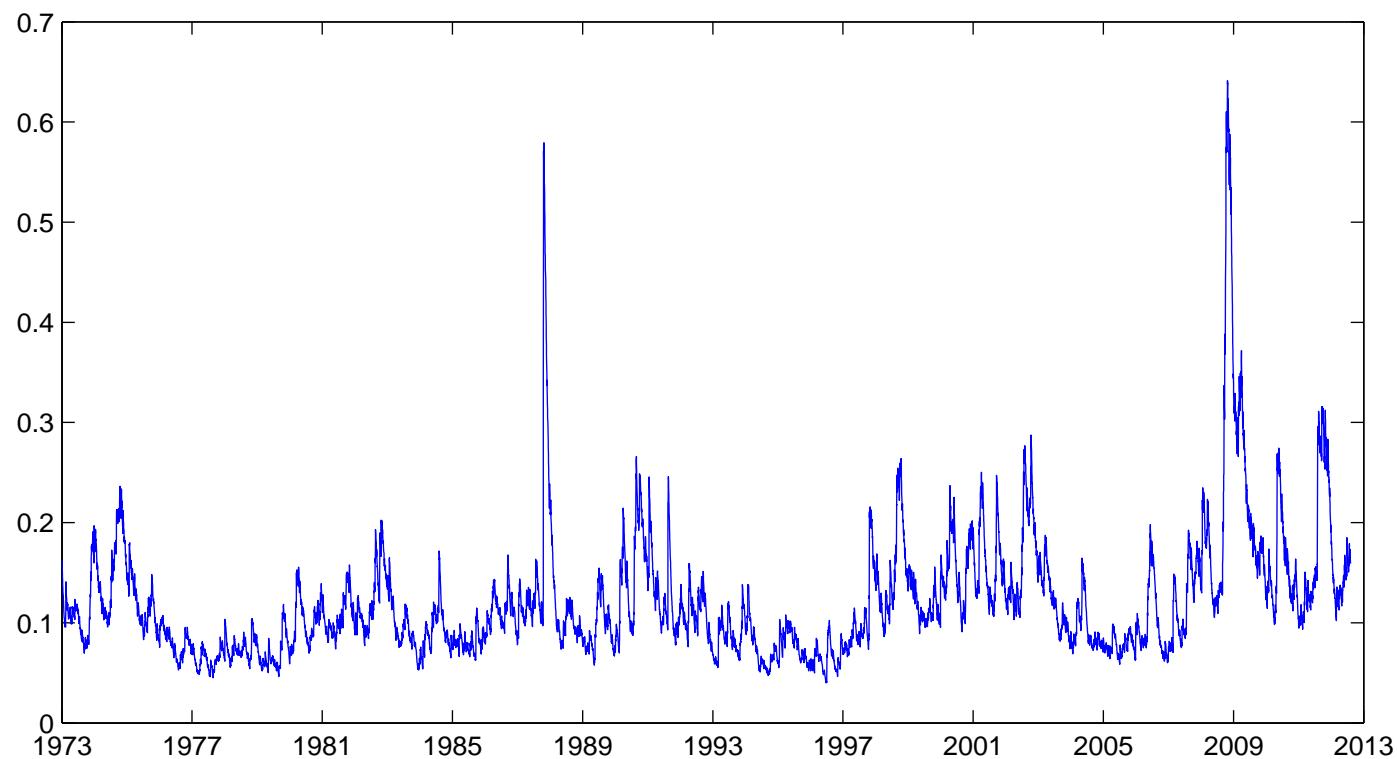
Pl. & Rendek (2008)



	SGH	Student- <i>t</i>	NIG	Hyperbolic	VG
σ	0.98	0.72	0.97	0.96	0.96
$\bar{\alpha}$	0.00		0.97	0.72	
λ	-2.16				1.49
ν		4.33			
$\ln(\mathcal{L}^*)$	-285796.39	-285796.39	-286448.94	-287152.08	-287499.83
L_n		0.0000004	1305.10	2711.38	3406.88

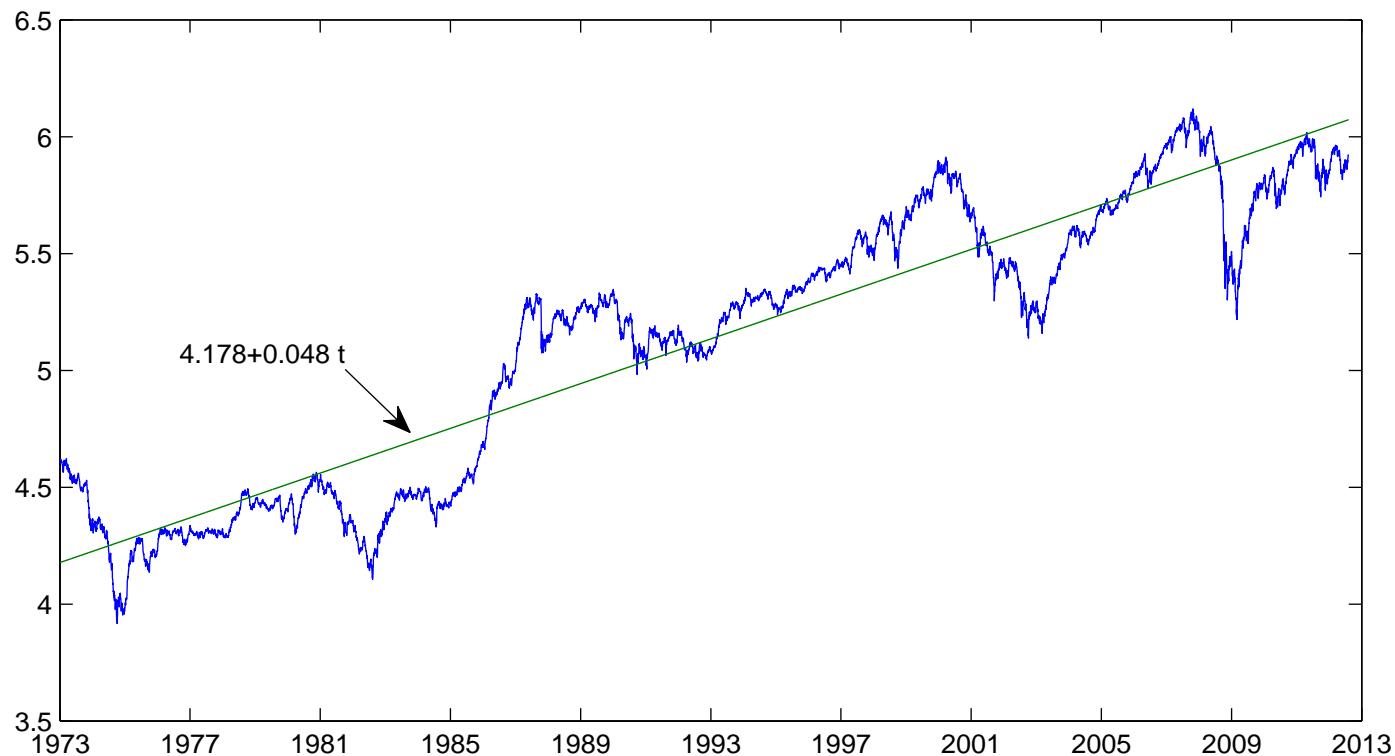
$$L_n = 0.0000004 < \chi^2_{0.001, 1} \approx 0.000002$$

(iv) volatility clustering



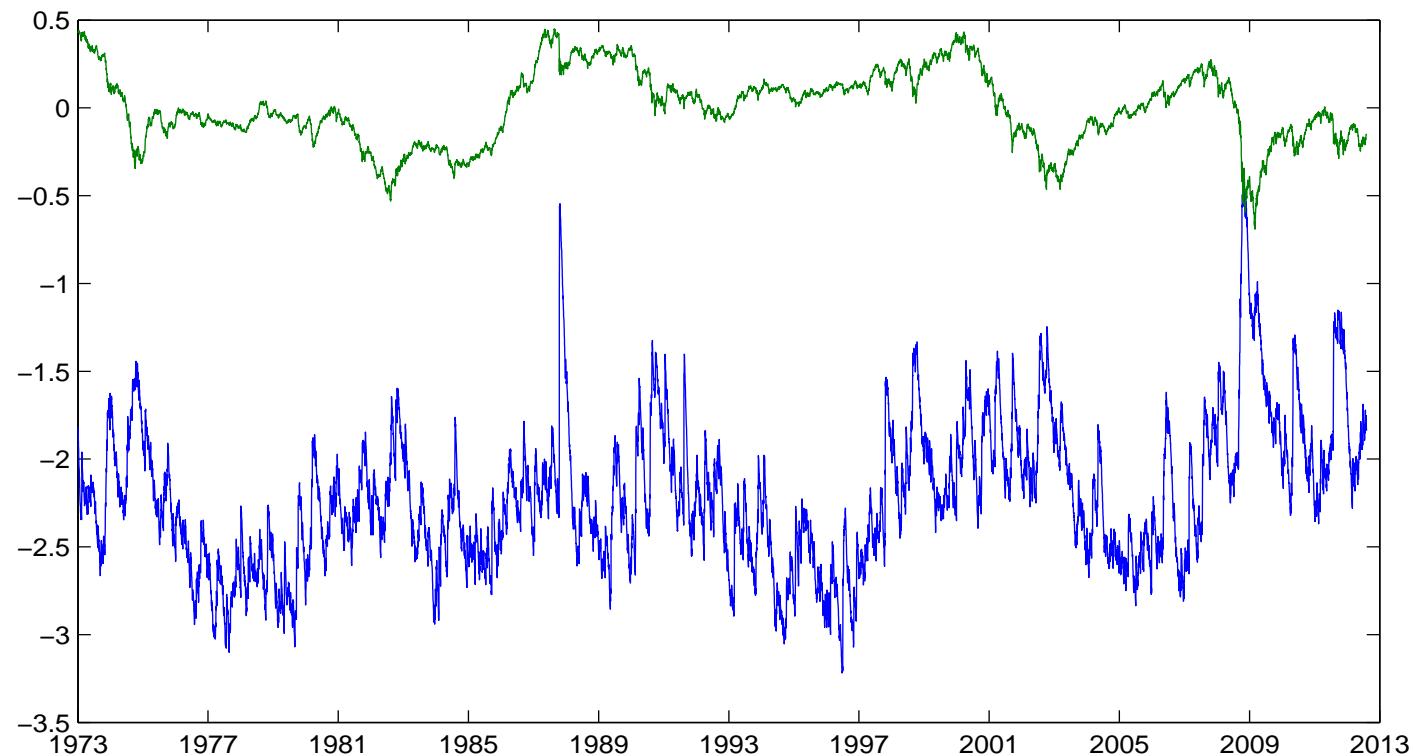
Estimated volatility

(v) long term exponential growth



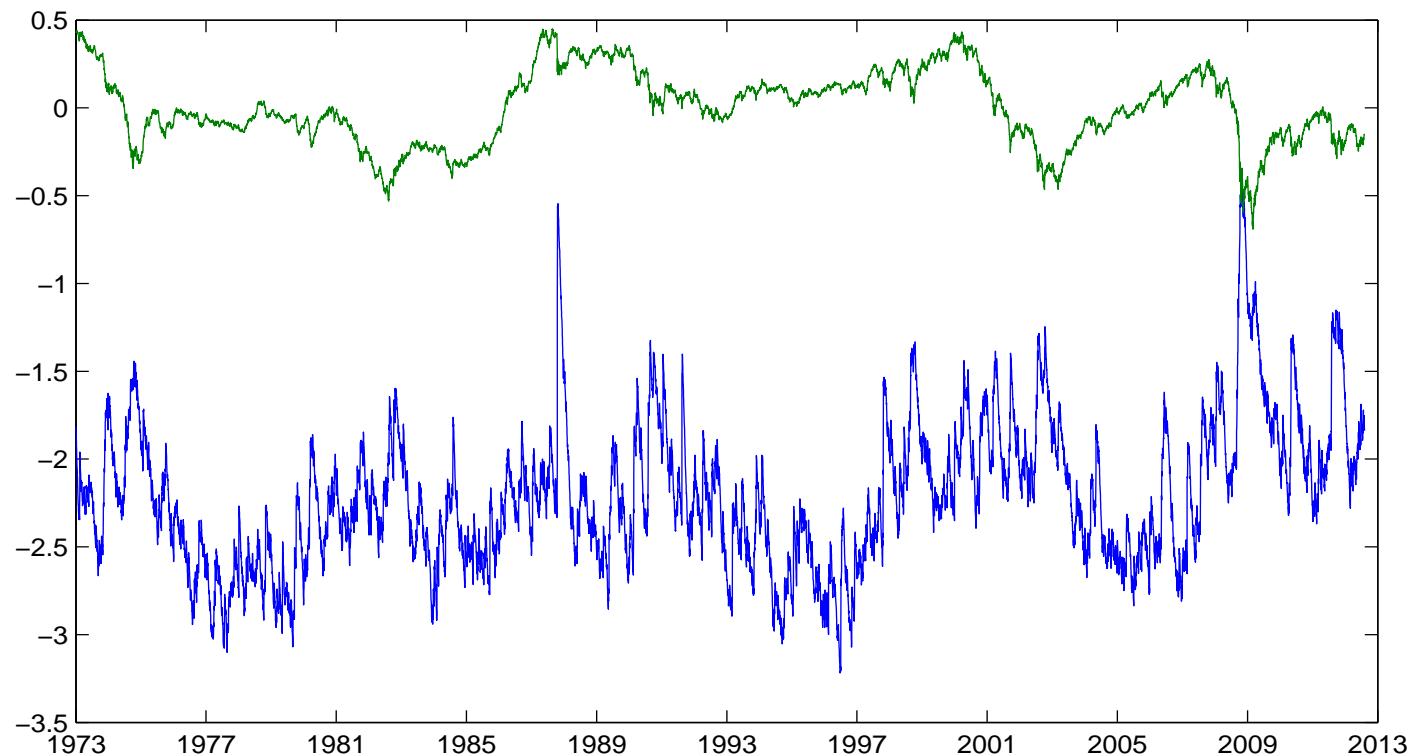
Logarithm of index with trend line

(vi) leverage effect



Logarithms of normalized index and its volatility

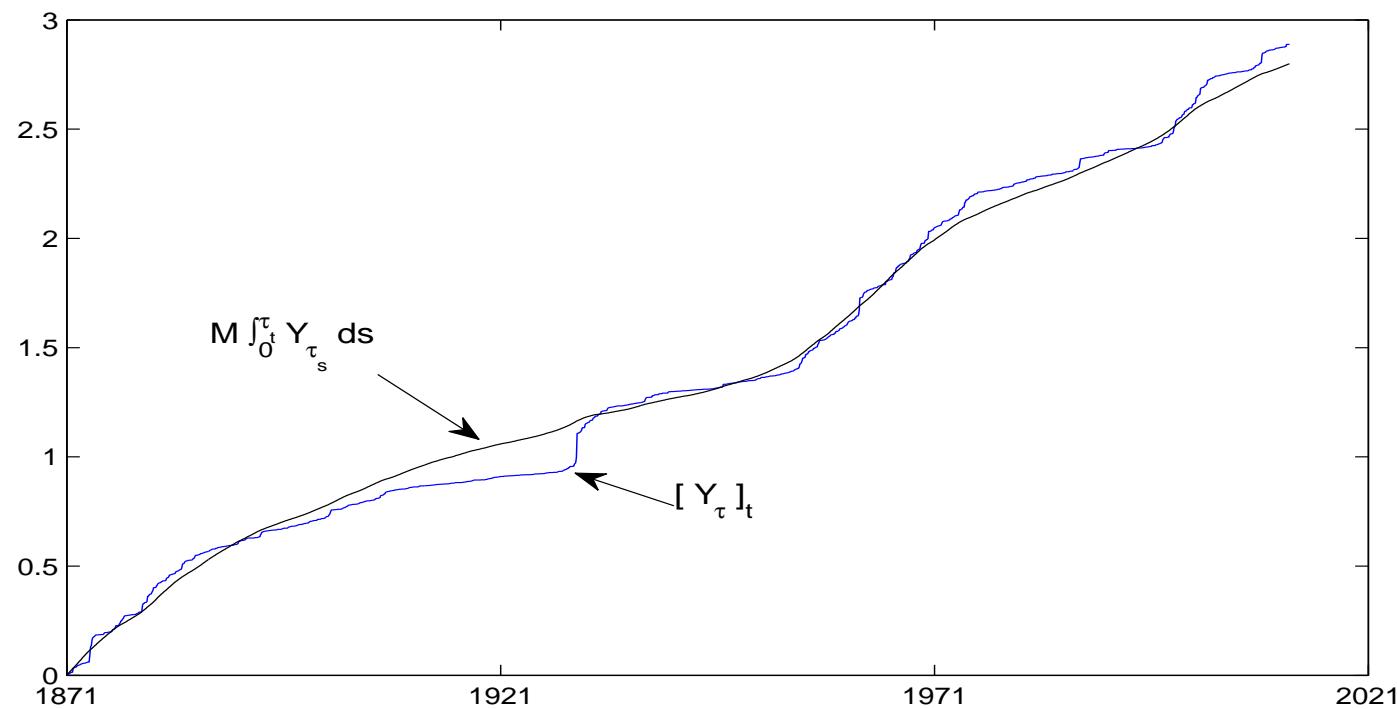
(vii) extreme volatility at major downturns



Logarithms of normalized index and its volatility

(viii) Integrated Normalized Index and its Quadratic Variation

monthly data, calendar time



$M \approx 0.0178$
average long term fit

\Rightarrow Discounted Index Model generalizing conjecture

$$S_t = A_{\tau_t} (Y_{\tau_t})^q,$$

$$A_{\tau_t} = A \exp\{a\tau_t\}$$

$$q > 0, a > 0, A > 0$$

conjecture expects $q = 1$

Normalized index: $(Y_{\tau_t})^q = \frac{S_t}{A_{\tau_t}}$

$$dY_\tau = \left(\frac{\delta}{4} - \frac{1}{2} \left(\frac{\Gamma(\frac{\delta}{2} + q)}{\Gamma(\frac{\delta}{2})} \right)^{\frac{1}{q}} Y_\tau \right) d\tau + \sqrt{Y_\tau} dW(\tau)$$

Long term mean: $\lim_{\tau \rightarrow \infty} \frac{1}{\tau} \int_0^\tau (Y_s)^q ds = 1$ P-a.s

Market activity time: $d\tau_t = M_t dt$

Inverse of market activity:

$$d\left(\frac{1}{M_t}\right) = \left(\frac{\nu}{4}\gamma - \epsilon\frac{1}{M_t}\right) dt + \sqrt{\frac{\gamma}{M_t}} dW_t,$$

where

$$dW(\tau_t) = \sqrt{\frac{d\tau_t}{dt}} dW_t = \sqrt{M_t} dW_t$$

- ◊ only one W_t
- ◊ two component model

Discounted index SDE:

$$dS_t = S_t (\mu_t dt + \sigma_t dW_t)$$

Expected rate of return:

$$\mu_t = \left(\frac{a}{M_t} - \frac{q}{2} \left(\frac{\Gamma(\frac{\delta}{2} + q)}{\Gamma(\frac{\delta}{2})} \right)^{\frac{1}{q}} + \left(\frac{\delta}{4}q + \frac{1}{2}q(q-1) \right) \frac{1}{M_t Y_{\tau_t}} \right) M_t$$

Volatility:

$$\sigma_t = q \sqrt{\frac{M_t}{Y_{\tau_t}}}$$

Pl. & Rendek (2012c)

Benchmark Approach

\hat{B}_t – benchmark savings account

$$d\hat{B}_t = \hat{B}_t ((-\mu_t + \sigma_t^2) dt - \sigma_t dW_t)$$

$\sigma_t^2 \leq \mu_t \Rightarrow \hat{B}_t$ is an $(\underline{\mathcal{A}}, P)$ -supermartingale
 \Rightarrow no strong arbitrage; Pl. (2011)

Assumptions:

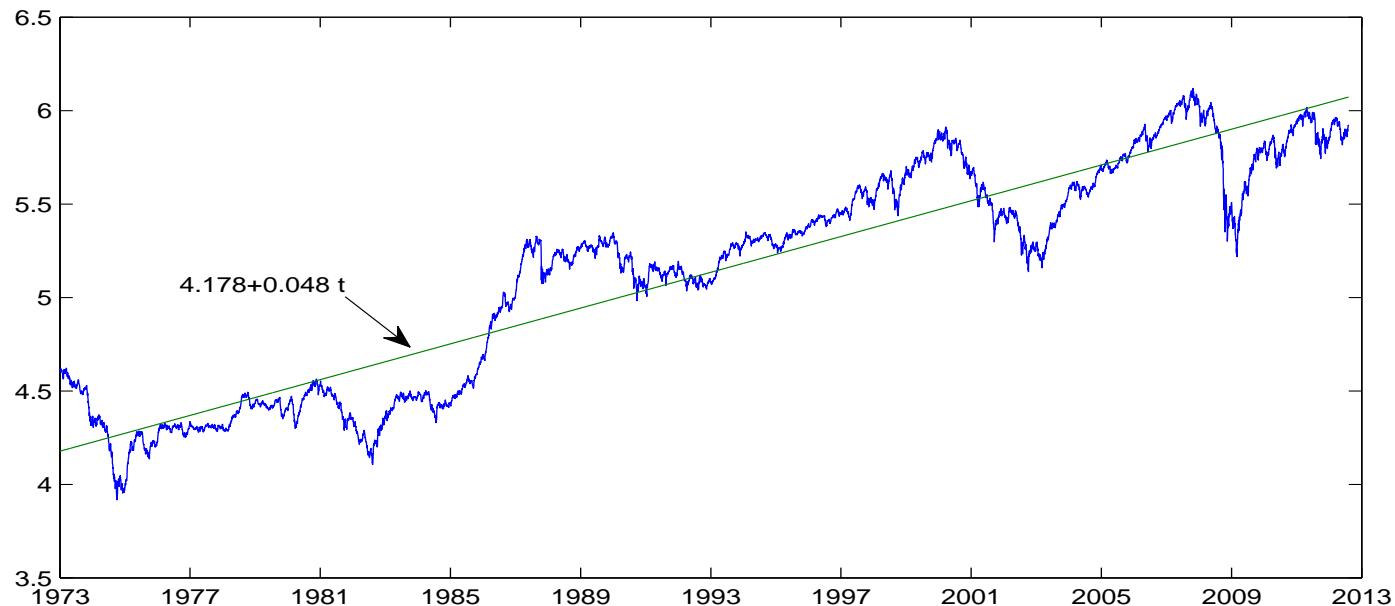
$$\text{A1. } \delta = 2(q+1) \quad \text{A2. } \frac{q}{2} \left(\frac{\Gamma(2q+1)}{\Gamma(q+1)} \right)^{\frac{1}{q}} \leq a$$

$$\implies \sigma_t^2 \leq \mu_t$$

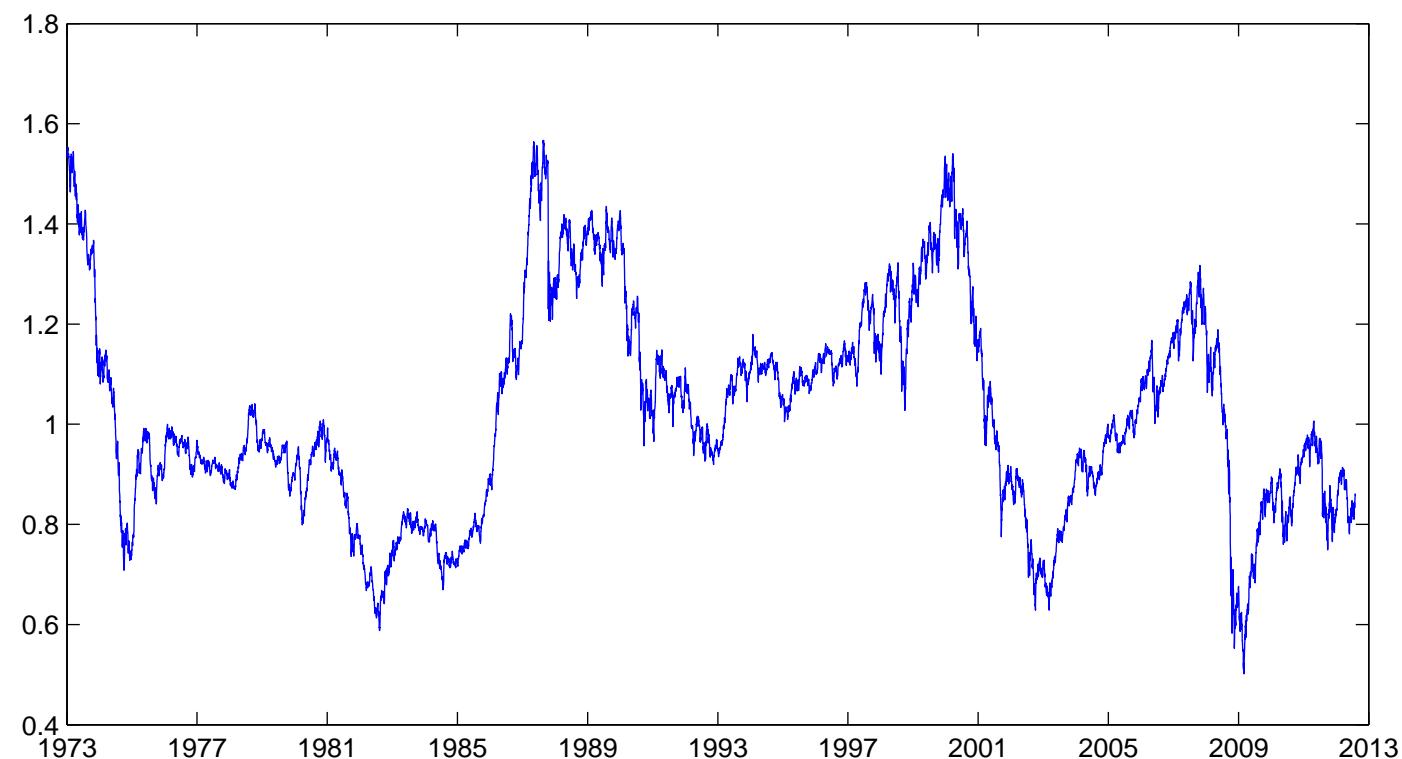
Fitting the model to TOTMKWD

Step 1: Normalization of Index

$$A_{\tau_t} \approx A \exp\left\{\frac{4a\epsilon}{\gamma(\nu-2)}t\right\} \Rightarrow A = 65.21, \frac{4a\epsilon}{\gamma(\nu-2)} \approx 0.048$$



Normalized TOTMKWD



Step 2: Power q:

Affine nature \implies conjecture: $q = 1 = \frac{\delta}{2} - 1 \implies \delta = 4$

- student-t with about four degree of freedom
is considered and cannot be falsified,
- \implies We set $q = 1$ and $\delta = 4$

Step 3: Observing Market Activity:

$$\frac{d[\sqrt{Y}]_{\tau_t}}{dt} = \frac{1}{4} \frac{d\tau_t}{dt} = \frac{M_t}{4}$$

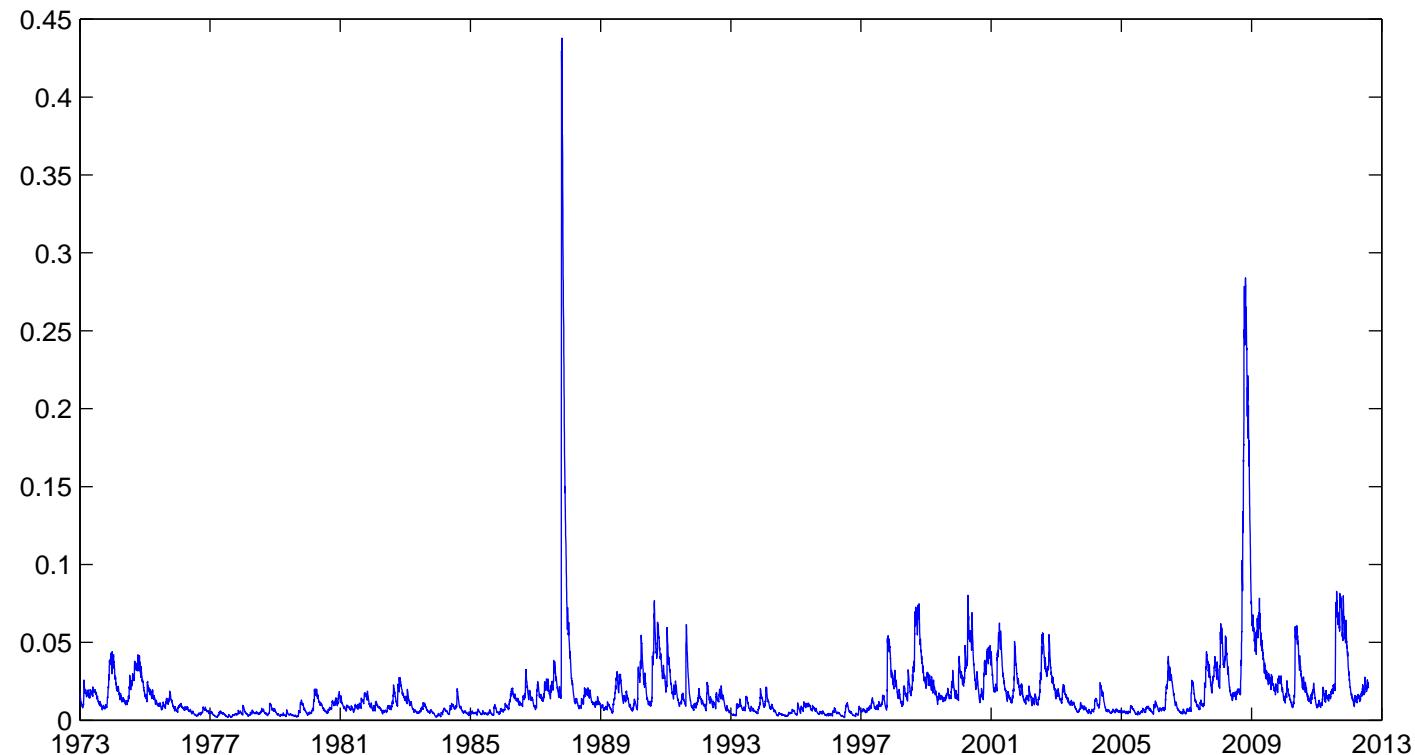
$$\hat{Q}_{t_i} \approx \frac{[\sqrt{Y}]_{\tau_{t_{i+1}}} - [\sqrt{Y}]_{\tau_{t_i}}}{t_{i+1} - t_i}$$

$$\tilde{Q}_{t_{i+1}} = \alpha \sqrt{t_{i+1} - t_i} \hat{Q}_{t_i} + (1 - \alpha \sqrt{t_{i+1} - t_i}) \tilde{Q}_{t_i}, \alpha = 0.92$$

exponential smooth

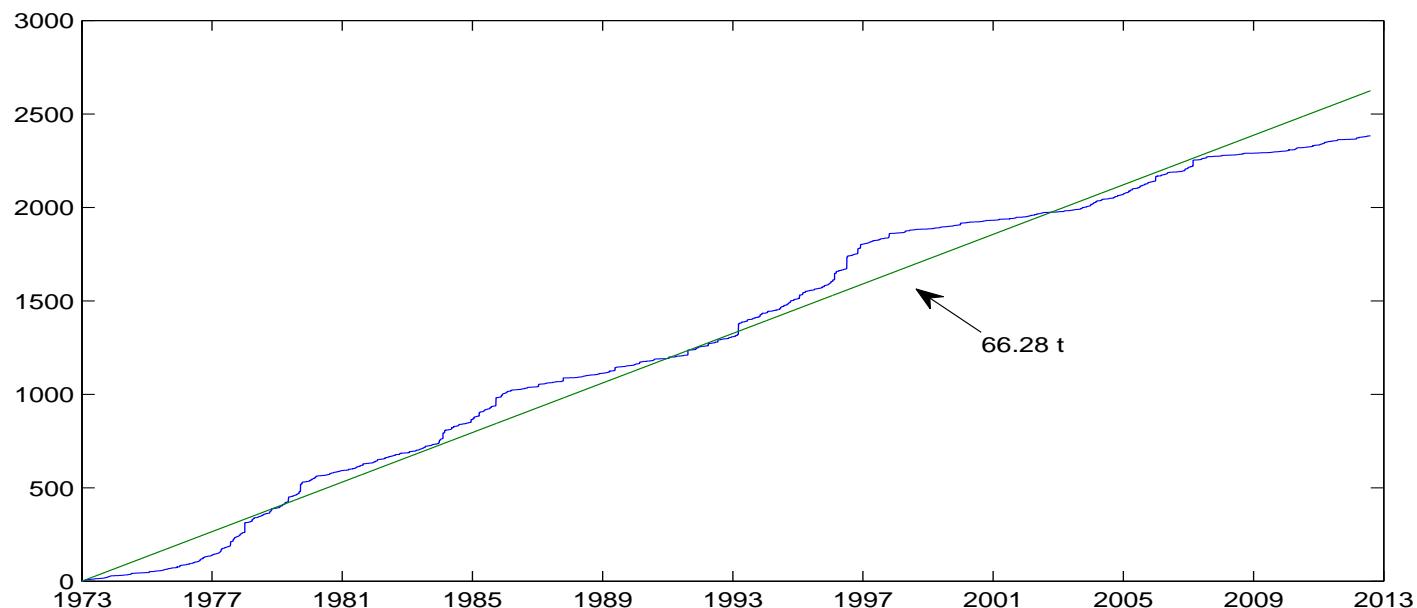
robust

Market activity: $M_t \approx 4\tilde{Q}_t$



$$M_0 = 0.0175$$

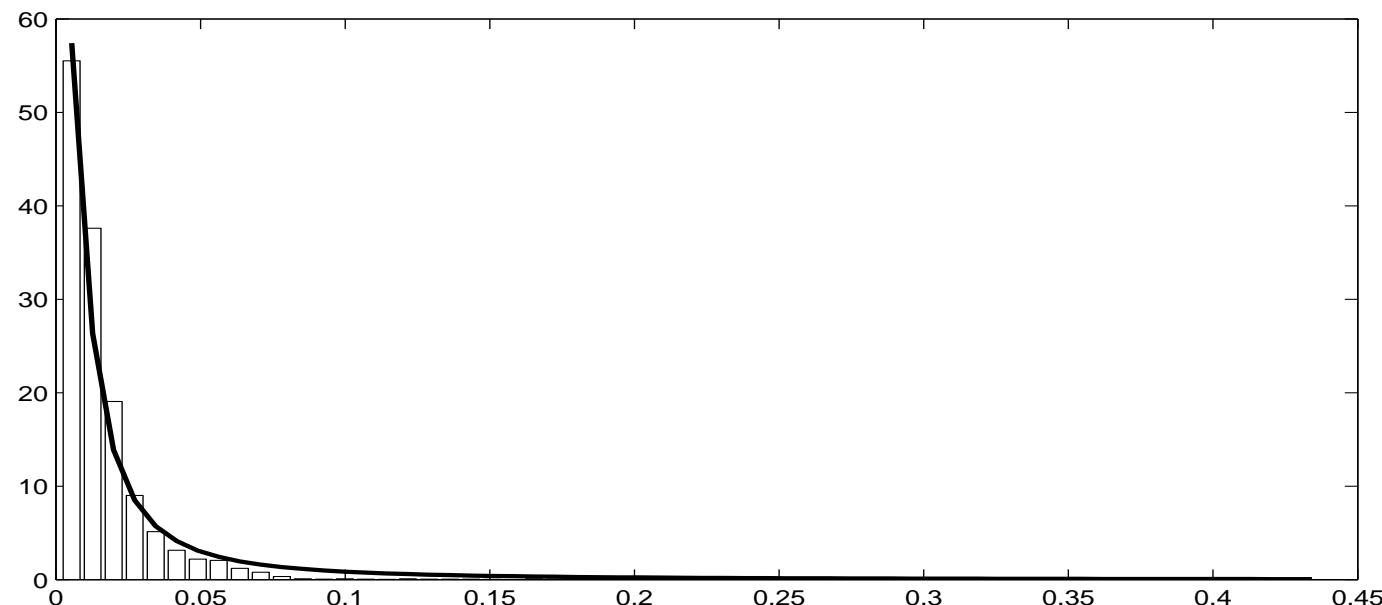
Step 4: Parameter γ :



$$\gamma = 265.12$$

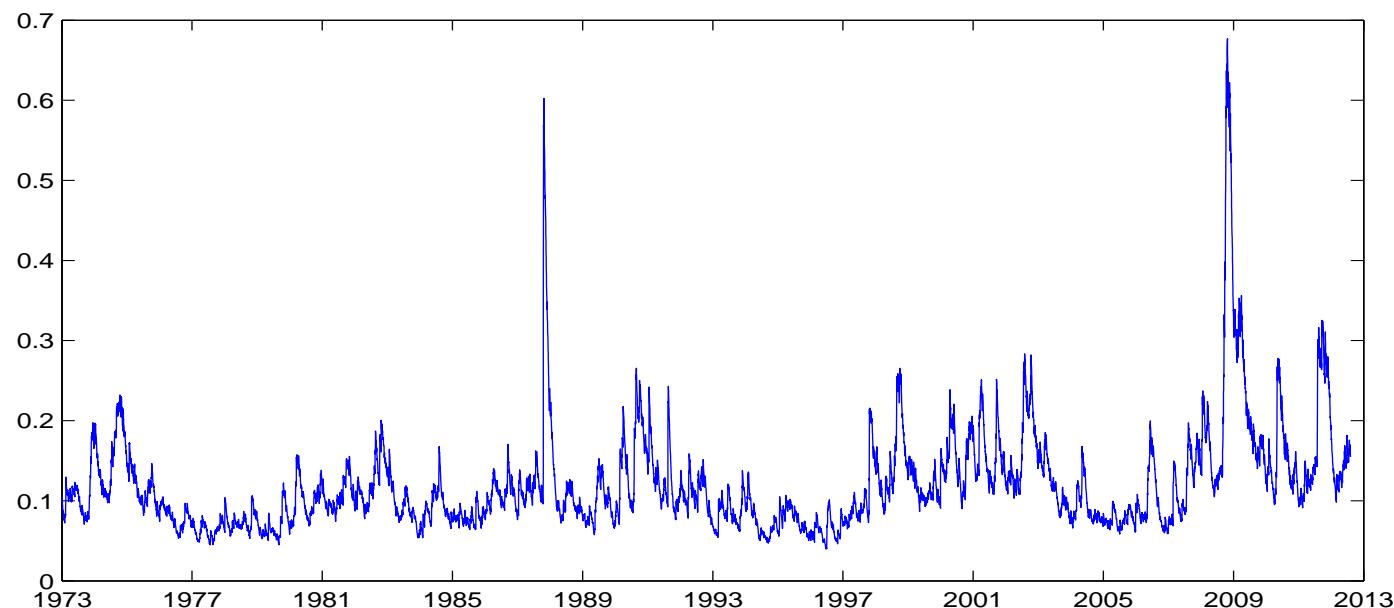
Step 5: Parameters ν and ϵ and

Step 6: Long Term Average Net Growth Rate a :



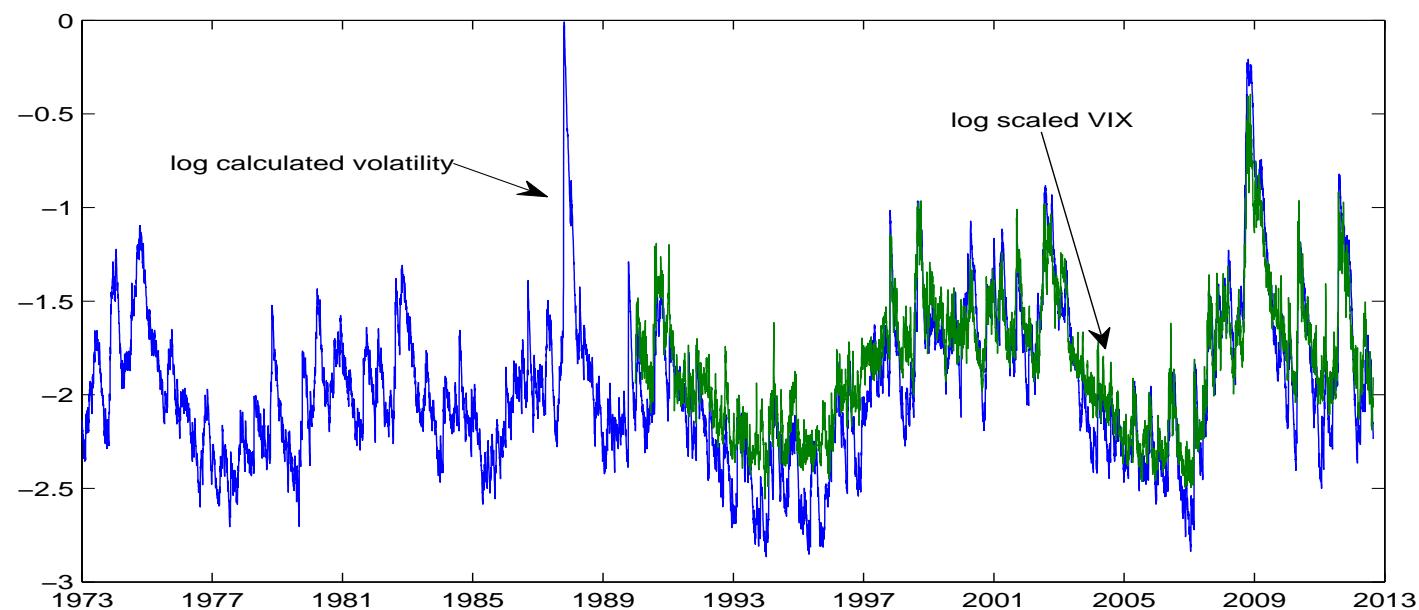
$\nu \approx 4, \epsilon \approx 2.18 \Rightarrow a = 2.55 \Rightarrow$ no strong arbitrage

Fitted model allows visualizing Calculated Volatility



$$\sigma_t \approx \sqrt{\frac{4\tilde{Q}_t}{Y_{\tau_t}}}, \text{ average volatility: } 11.9\%$$

**Model applies also to proxies of numéraire portfolio
S&P500 and VIX**

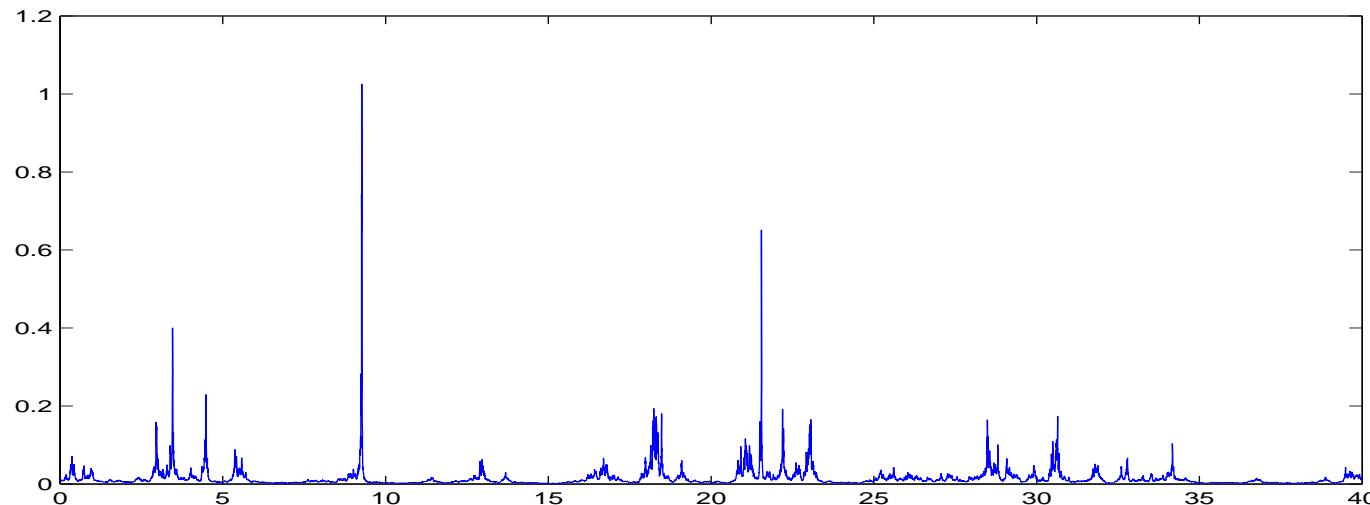


$$A = 52.09, \epsilon = 2.15, \gamma = 172.3, a = 1.5$$

Simulation Study

Step 1: Market activity:

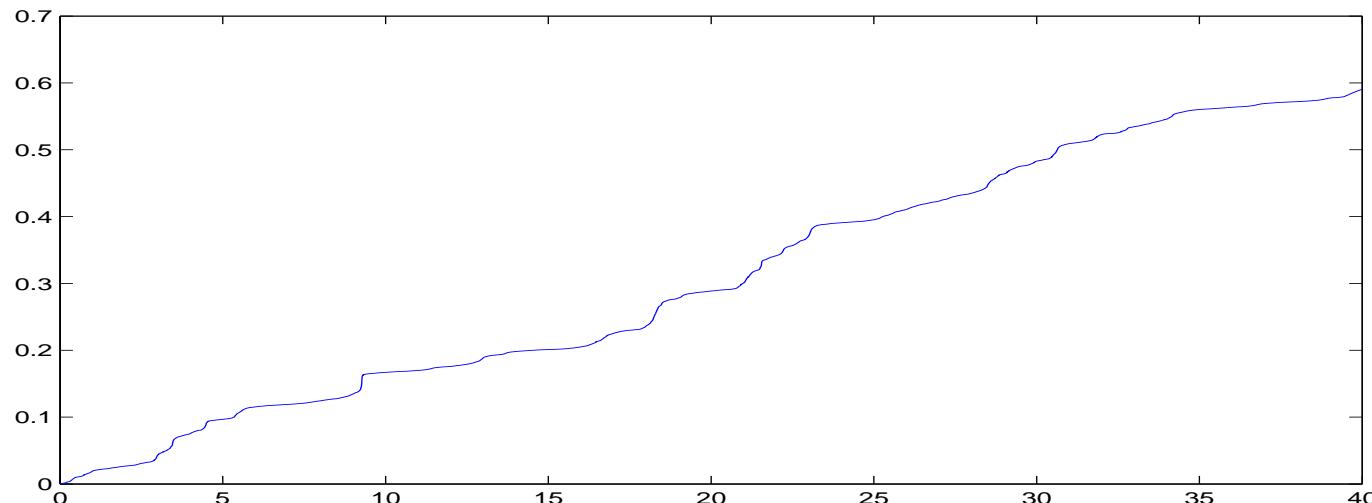
$$\frac{1}{M_{t_{i+1}}} = \frac{\gamma(1 - e^{-\epsilon(t_{i+1} - t_i)})}{4\epsilon} \left(\chi_{3,i}^2 + \left(\sqrt{\frac{4\epsilon e^{-\epsilon(t_{i+1} - t_i)}}{\gamma(1 - e^{-\epsilon(t_{i+1} - t_i)})}} \frac{1}{M_{t_i}} + Z_i \right)^2 \right)$$



exact simulation

Step 2: τ -time:

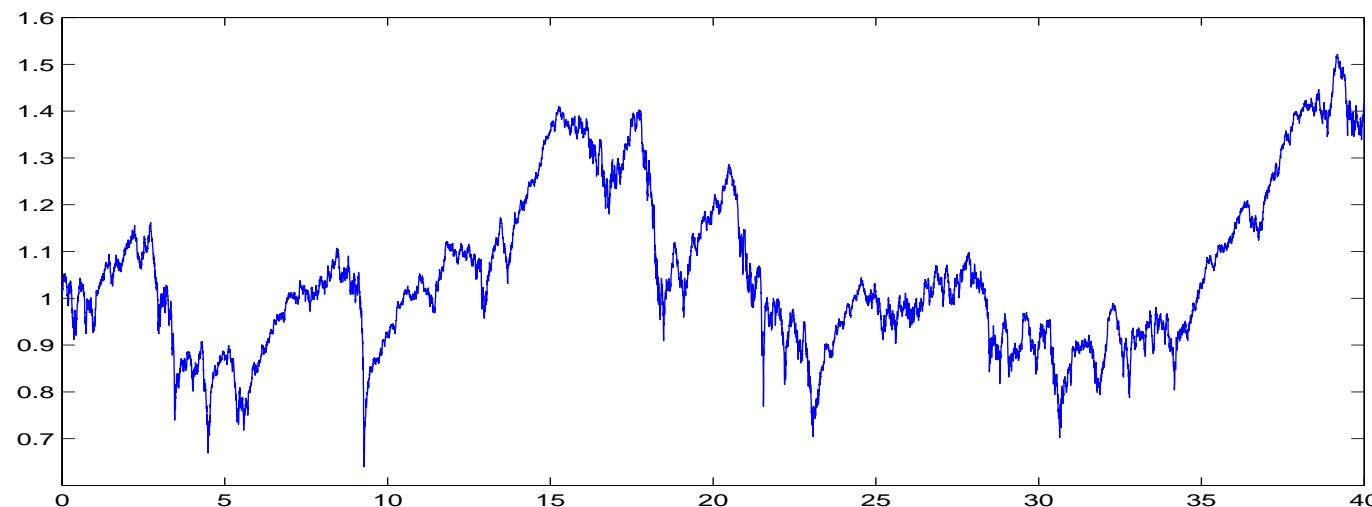
$$\tau_{t_{i+1}} - \tau_{t_i} = \int_{t_i}^{t_{i+1}} M_s ds \approx M_{t_i} (t_{i+1} - t_i)$$



almost exact simulation

Step 3: Normalized index:

$$Y_{\tau_{t_i+1}} = \frac{1 - e^{-(\tau_{t_{i+1}} - \tau_{t_i})}}{4} \left(\chi_{3,i}^2 + \left(\sqrt{\frac{4e^{-(\tau_{t_{i+1}} - \tau_{t_i})}}{1 - e^{-(\tau_{t_{i+1}} - \tau_{t_i})}} Y_{\tau_{t_i}} + Z_i} \right)^2 \right)$$

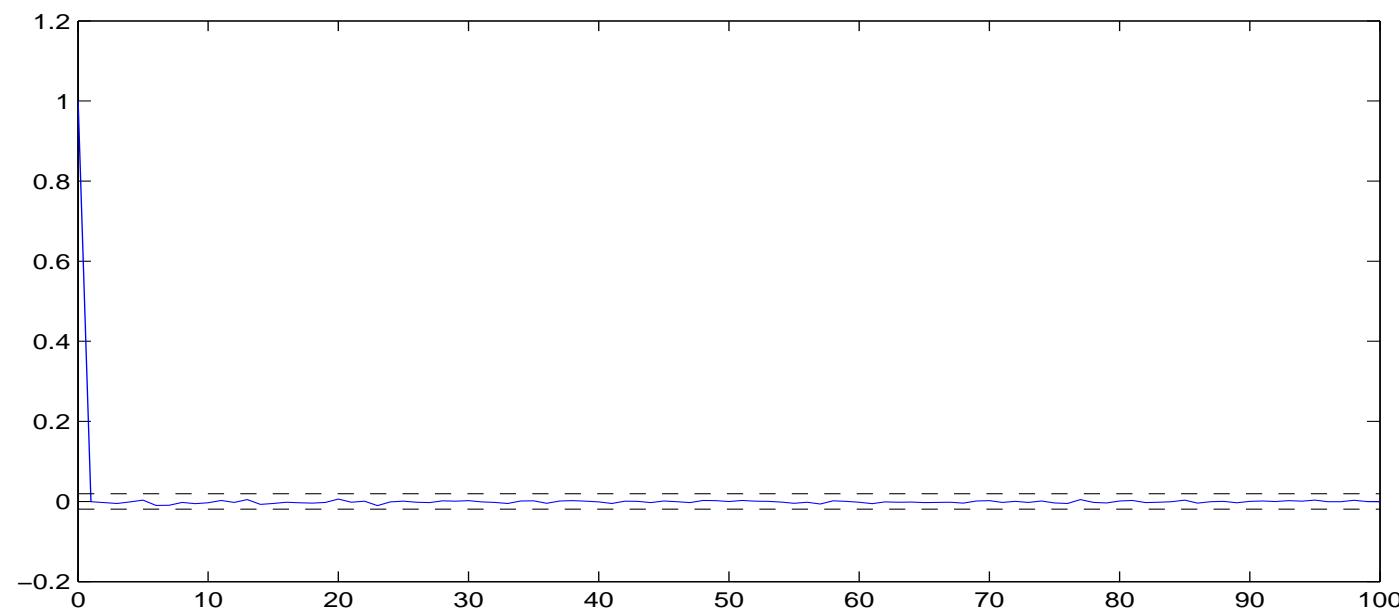


almost exact simulation

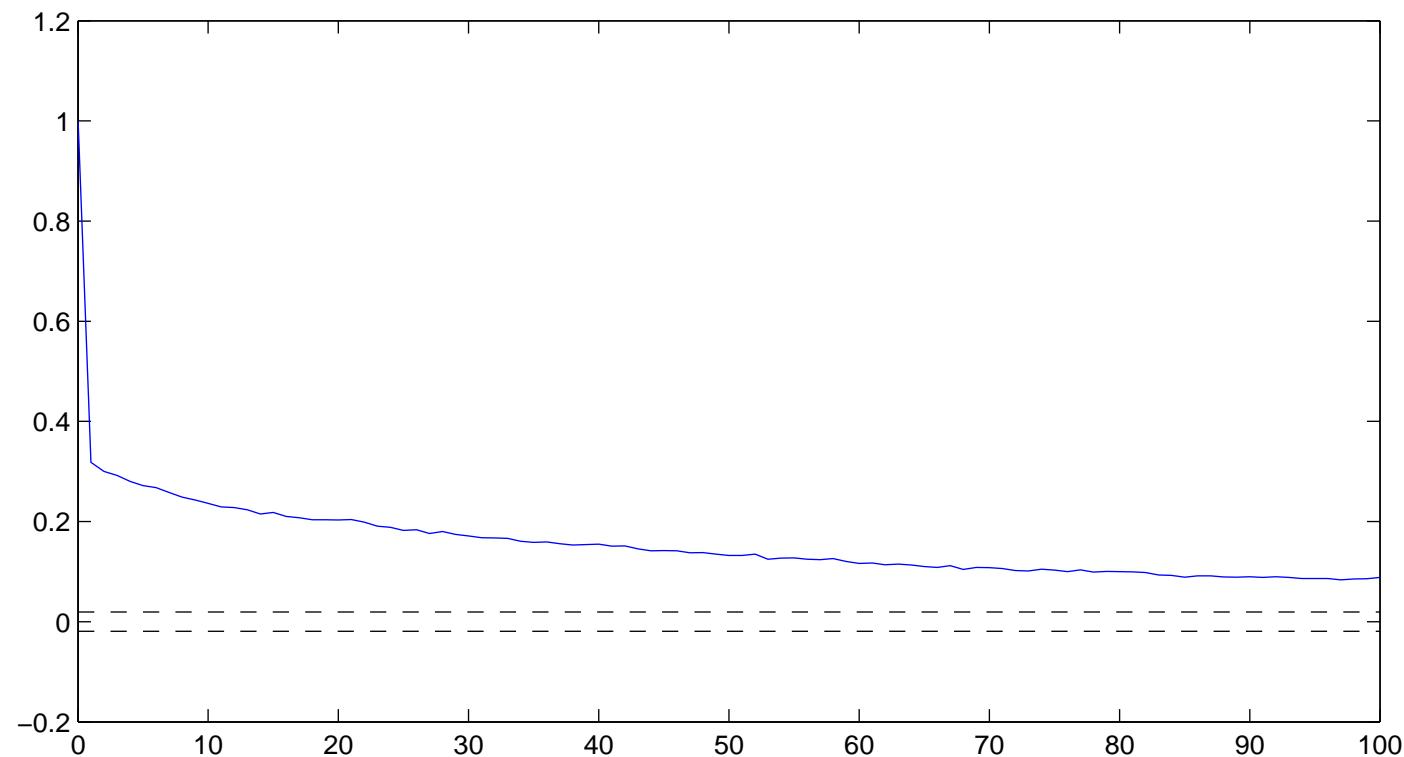
Model recovers stylized empirical facts:

Model is difficult to falsify: Popper (1934)

1. Uncorrelated returns



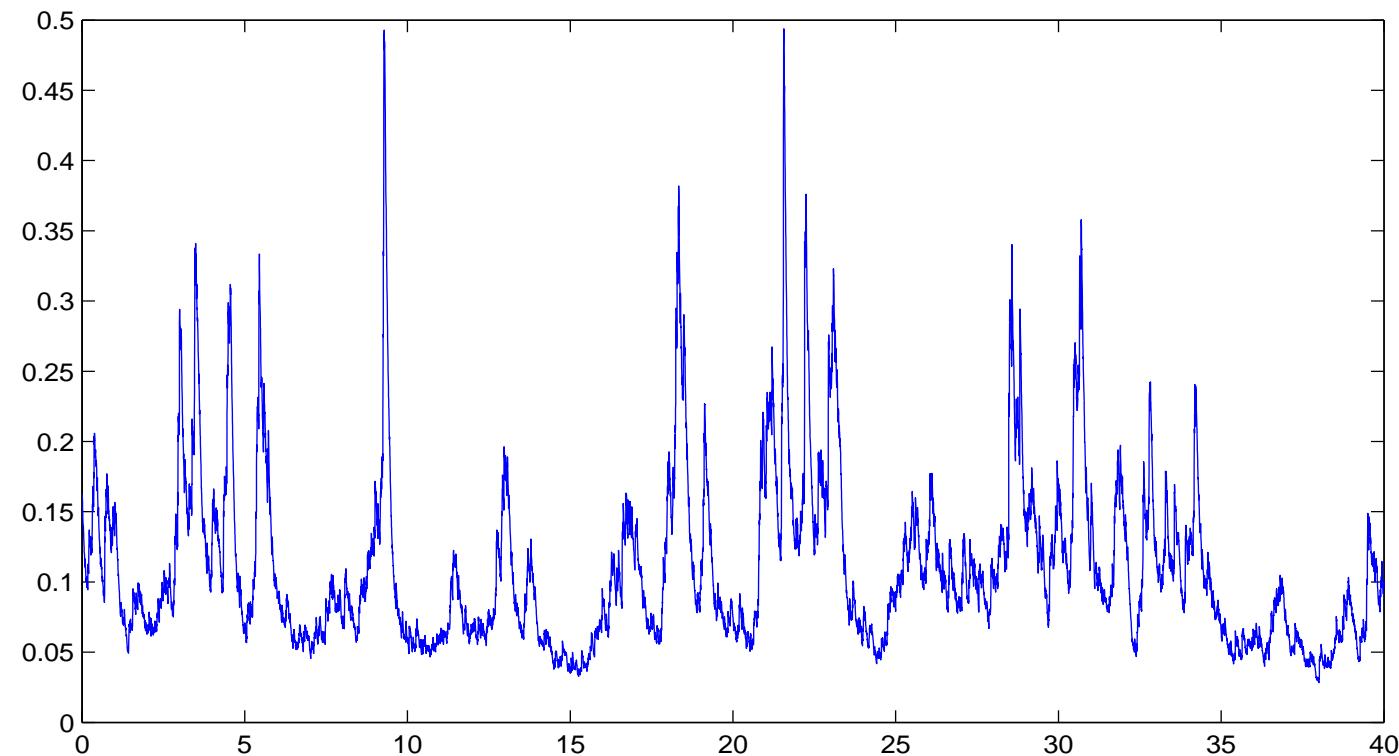
2. Correlated absolute returns



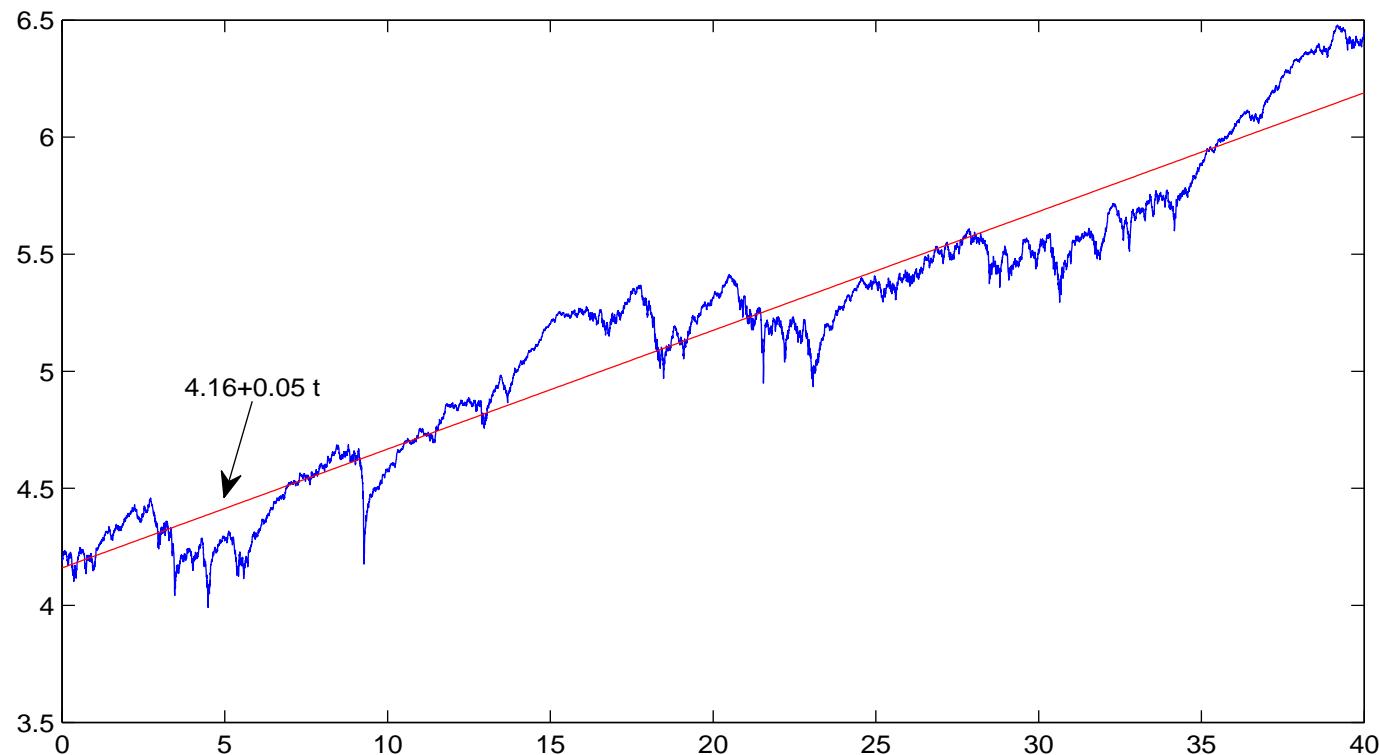
3. Student-*t* distributed returns

Simulation	Student-t	NIG	Hyperbolic	VG
1	0.008934	37.474149	102.719638	131.240780
2	11.485226	11.175028	96.457136	132.916256
3	0.000000	100.928524	244.190151	294.719960
4	9.002421	35.759464	347.060676	331.014904
5	8.767003	11.551178	121.190482	144.084964
6	0.401429	60.570898	205.788160	252.591737
7	12.239056	4.354888	46.411554	78.273485
8	1.693411	23.910523	94.408789	130.623174
9	1.232454	47.830407	202.073144	237.168411
10	0.000000	43.037206	128.807757	162.582353
estimated statistics	11	0.433645	47.782681	172.736397
	12	0.000000	56.019354	146.077121
	13	7.137154	48.219756	579.922931
	14	5.873948	16.515390	107.770531
	15	0.000000	54.718046	184.112794
	16	6.982560	3.991610	29.192198
	17	2.966916	22.914863	108.513143
	18	0.000000	52.066364	129.790856
	19	0.006909	39.568695	111.398645
	20	0.000001	56.845664	169.915512
	21	1.674578	17.681088	61.710576
	22	14.010840	3.279722	47.433693
	23	11.198940	12.074044	114.888817
	24	0.455557	27.676102	86.841704
				114.452947

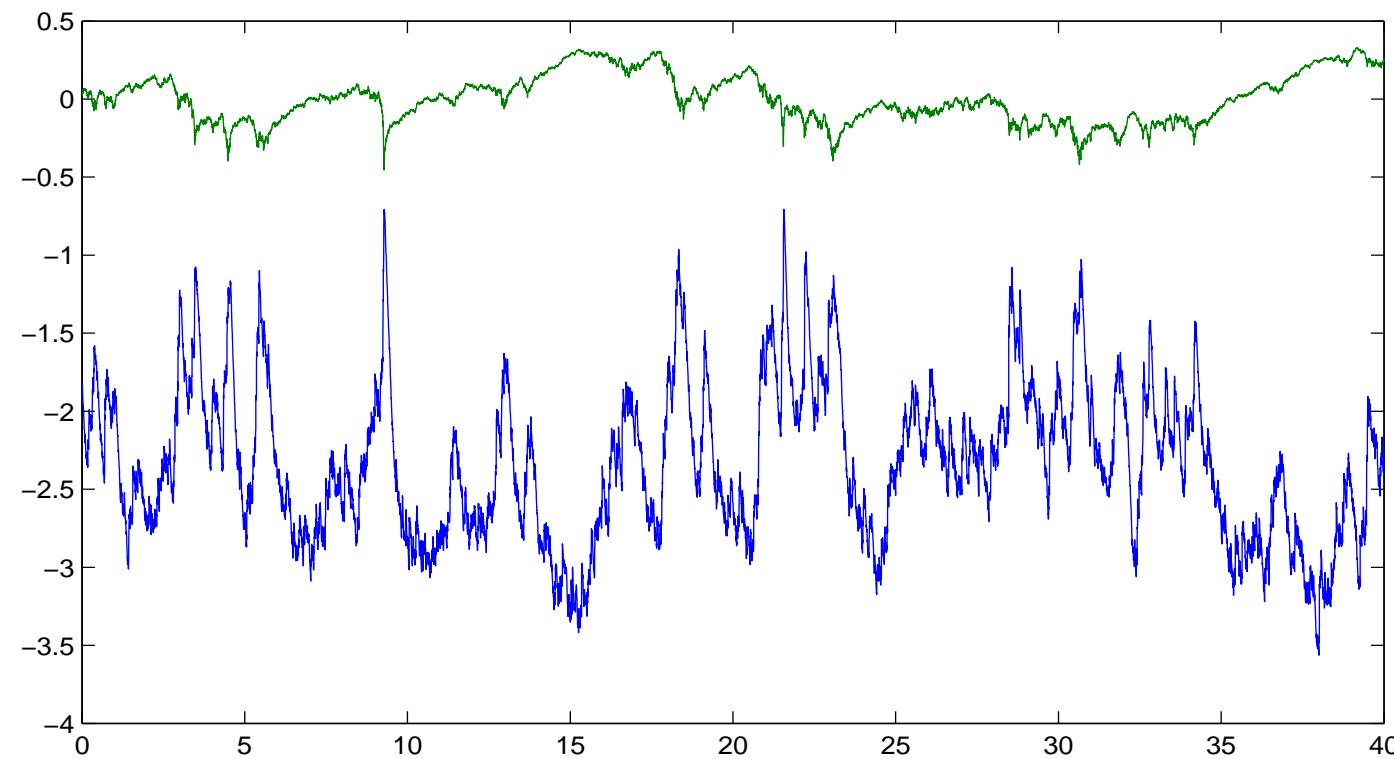
4. Volatility clustering



5. Long term exponential growth



6. Leverage effect and 7. Extreme volatility at major market downturns



Conclusions:

- ◊ equity index model: **3** initial parameters, **3** structural parameters and **1** Wiener process (nondiversifiable uncertainty)
- ◊ model recovers **7** stylized empirical facts
- ◊ allows long dated derivative pricing under benchmark approach
- ◊ leads outside classical theory
- ◊ if benchmarked savings account local martingale
 \implies only **2** structural parameters

Squared Bessel Processes (*)

Revuz & Yor (1999)

- squared Bessel process BESQ_x^δ

dimension $\delta \geq 0$

$$dX_\varphi = \delta d\varphi + 2 \sqrt{|X_\varphi|} dW_\varphi$$

$\varphi \in [0, \infty)$ with $X_0 = x \geq 0$

- scaling property

$$Z_\varphi = \frac{1}{a} X_{a\varphi}$$

$\text{BESQ}_{\frac{x}{a}}^\delta, a > 0$

- sum of squares

$$X_\varphi = \sum_{k=1}^{\delta} (w^k + W_\varphi^k)^2$$

W_φ^k Wiener process

\implies

$$dX_\varphi = \delta d\varphi + 2 \sum_{k=1}^{\delta} (w^k + W_\varphi^k) dW_\varphi^k$$

$$X_0 = \sum_{k=1}^{\delta} (w^k)^2 = x$$

- setting

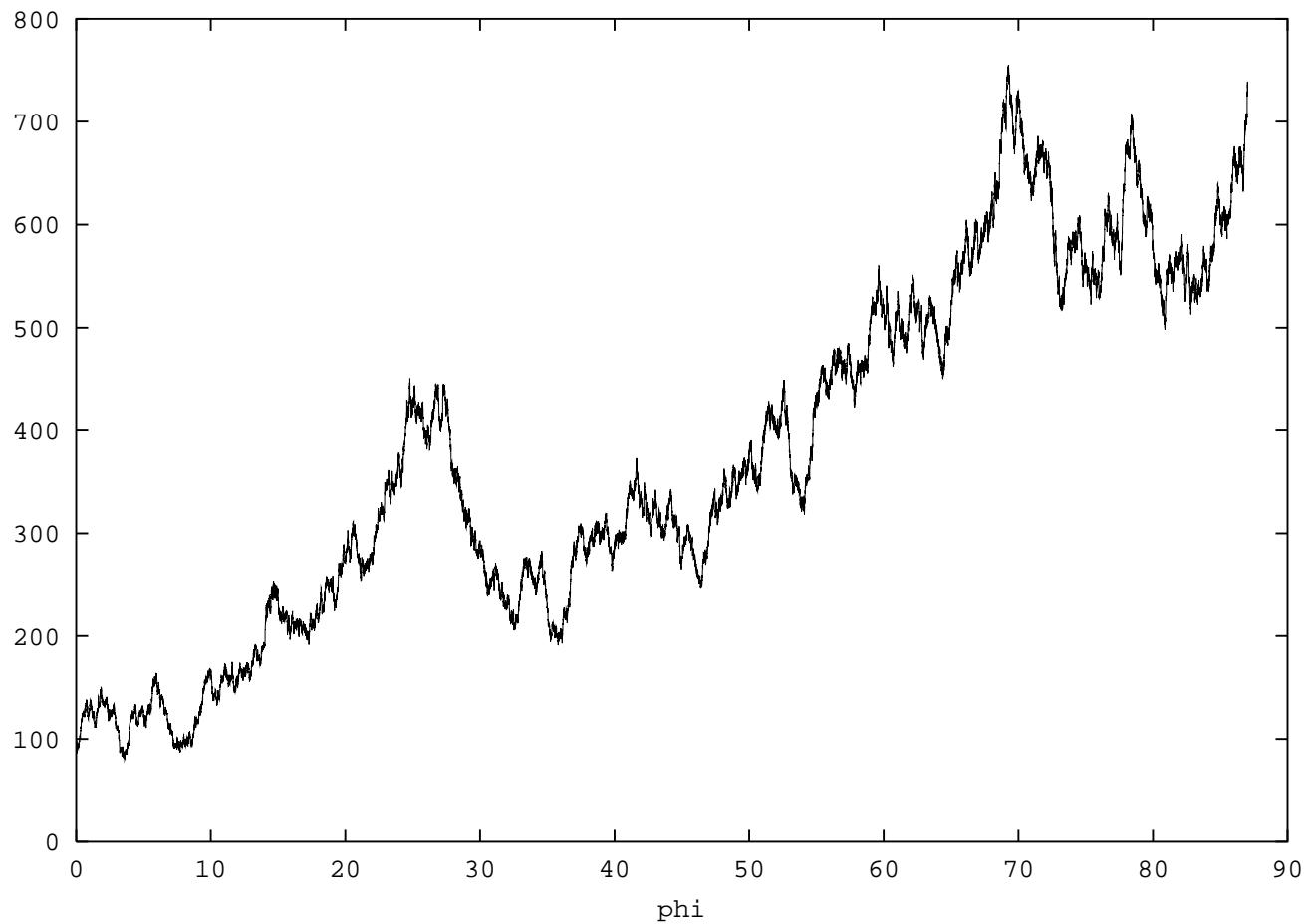
$$dW_\varphi = |X_\varphi|^{-\frac{1}{2}} \sum_{k=1}^{\delta} \left(w^k + W_\varphi^k \right) dW_\varphi^k$$

- quadratic variation

$$[W]_\varphi = \int_0^\varphi \frac{1}{X_s} \sum_{k=1}^{\delta} \left(w^k + W_s^k \right)^2 ds = \varphi$$

Lévy's theorem

$\implies W$ Wiener process in φ time



Squared Bessel process of dimension $\delta = 4$ in φ -time

- **additivity property**

Shiga & Watanabe (1973)

$$X = \{X_\varphi, \varphi \in [0, \infty)\} \quad \text{BESQ}_x^\delta$$

$$Y = \{Y_\varphi, \varphi \in [0, \infty)\} \quad \text{independent BESQ}_y^{\delta'}$$

$$\implies X_\varphi + Y_\varphi \text{ forms } \text{BESQ}_{x+y}^{\delta+\delta'}$$

$$x, y, \delta, \delta' \geq 0$$

- $\delta > 2$, strict positivity

with $X_0 = x > 0$

$$P \left(\inf_{0 \leq \varphi < \infty} X_\varphi > 0 \right) = 1$$

- $\delta = 2$, no strict positivity

$$P \left(\inf_{0 \leq \varphi < \infty} X_\varphi > 0 \right) = 0$$

- $\delta \in [0, 2)$, no strict positivity

$X_0 = x > 0$

$$P \left(\inf_{0 \leq \varphi \leq \varphi'} X_\varphi = 0 \right) > 0$$

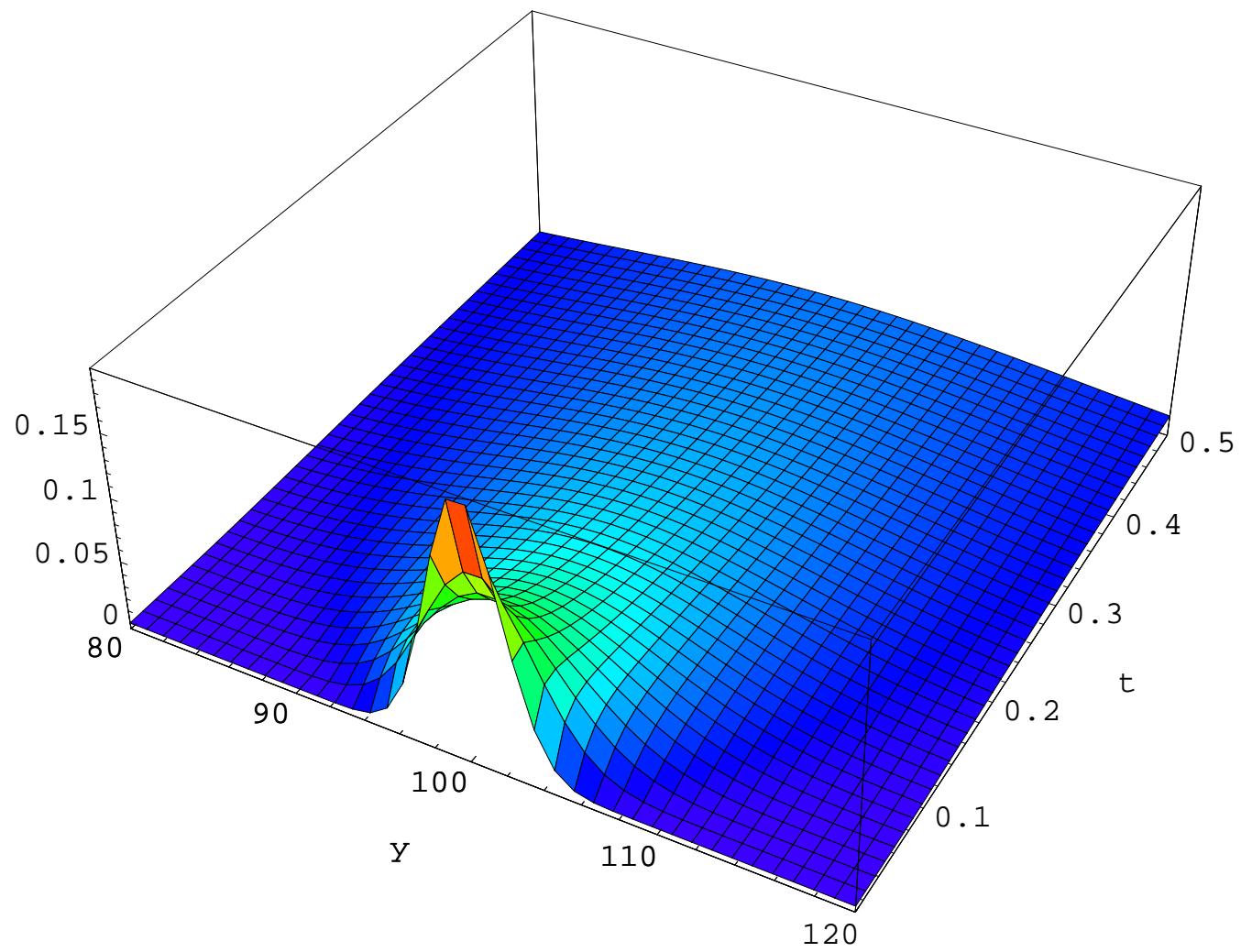
- transition density

$$\delta > 0, \quad x > 0$$

$$p_\delta(\varrho, x; \varphi, y) = \frac{1}{2(\varphi - \varrho)} \left(\frac{y}{x} \right)^{\frac{\nu}{2}} \exp \left\{ -\frac{x + y}{2(\varphi - \varrho)} \right\} I_\nu \left(\frac{\sqrt{xy}}{\varphi - \varrho} \right)$$

I_ν modified Bessel function of the first kind

$$\text{index } \nu = \frac{\delta}{2} - 1$$



Transition density of squared Bessel process for $\delta = 4$

- $Y = \frac{X_\varphi}{\varphi}$ non-central chi-square distributed

dimension δ

non-centrality parameter $\ell = \frac{x}{\varphi}$

$$P\left(\frac{X_\varphi}{\varphi} < u\right) = \sum_{k=0}^{\infty} \frac{\exp\left\{-\frac{\ell}{2}\right\} \left(\frac{\ell}{2}\right)^k}{k!} \left(1 - \frac{\Gamma\left(\frac{u}{2}; \frac{\delta+2k}{2}\right)}{\Gamma\left(\frac{\delta+2k}{2}\right)}\right)$$

$\Gamma(\cdot; \cdot)$ incomplete gamma function

- **Moments** for $\alpha > -\frac{\delta}{2}$, $\varphi \in (0, \infty)$ and $\delta > 2$

$$E(X_\varphi^\alpha | \mathcal{A}_0) = \begin{cases} (2\varphi)^\alpha \exp\left\{-\frac{X_0}{2\varphi}\right\} \sum_{k=0}^{\infty} \left(\frac{X_0}{2\varphi}\right)^k \frac{\Gamma(\alpha+k+\frac{\delta}{2})}{k! \Gamma(k+\frac{\delta}{2})} & \text{for } \alpha > -\frac{\delta}{2} \\ \infty & \text{for } \alpha \leq -\frac{\delta}{2} \end{cases}$$

\implies

$$E(X_\varphi | \mathcal{A}_0) = X_0 + \delta \varphi$$

- by monotonicity of the gamma function

$$\begin{aligned} E(X_\varphi^\alpha | \mathcal{A}_0) &\leq (2\varphi)^\alpha \exp\left\{\frac{-X_0}{2\varphi}\right\} \left(\frac{\Gamma(\alpha + \frac{\delta}{2})}{\Gamma(\frac{\delta}{2})} + \exp\left\{\frac{X_0}{2\varphi}\right\} \right) \\ &< \infty \end{aligned}$$

- for $\delta = 4$

$$E(X_\varphi^{-1} | \mathcal{A}_0) = X_0^{-1} \left(1 - \exp\left\{\frac{-X_0}{2\varphi}\right\} \right)$$

A Strict Local Martingale

$Z_\varphi = X_\varphi^{-1}$ inverse of a squared Bessel process X_φ of dimension 4

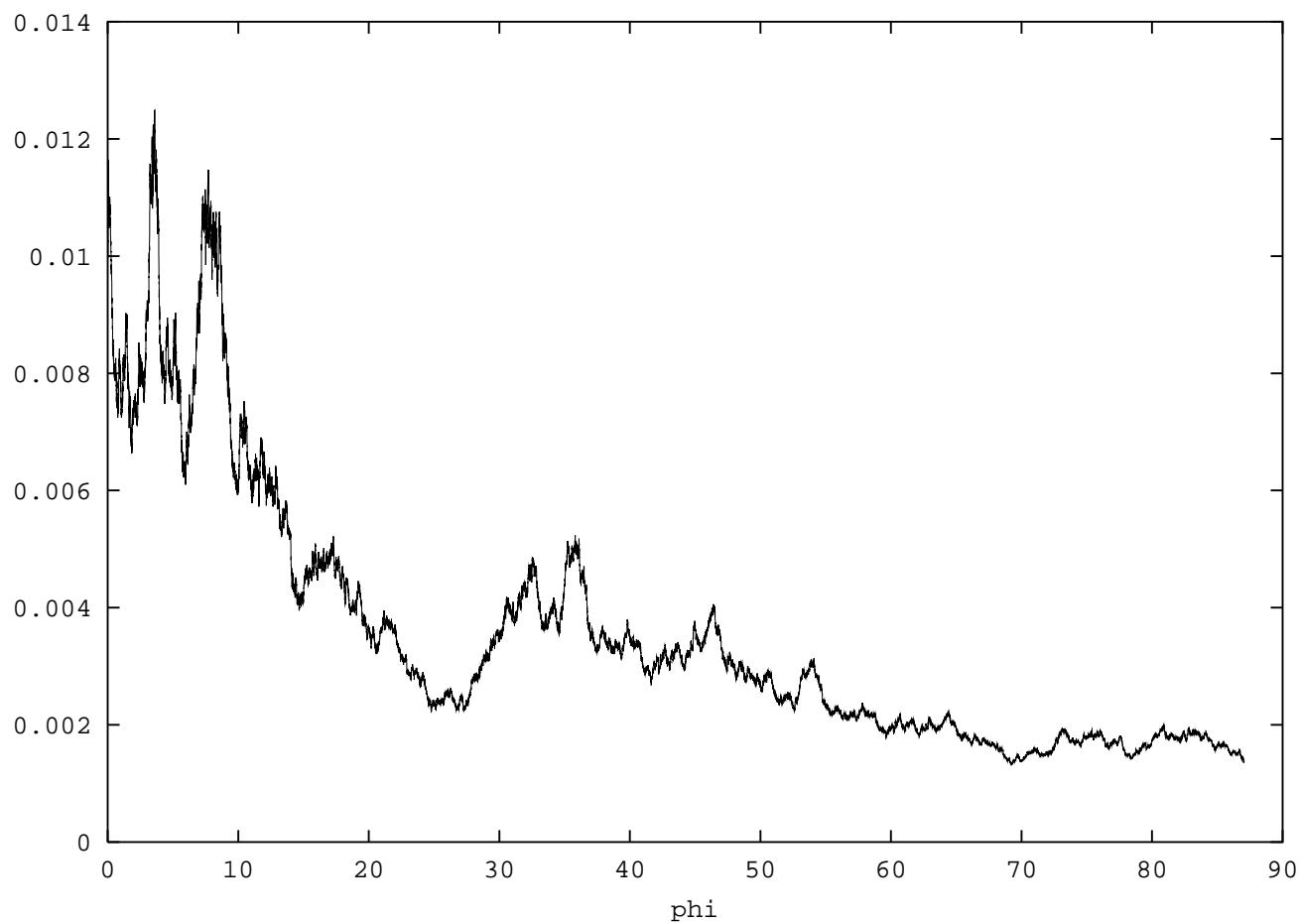
$$dZ_\varphi = -2 Z_\varphi^{\frac{3}{2}} dW_\varphi$$

\implies local martingale

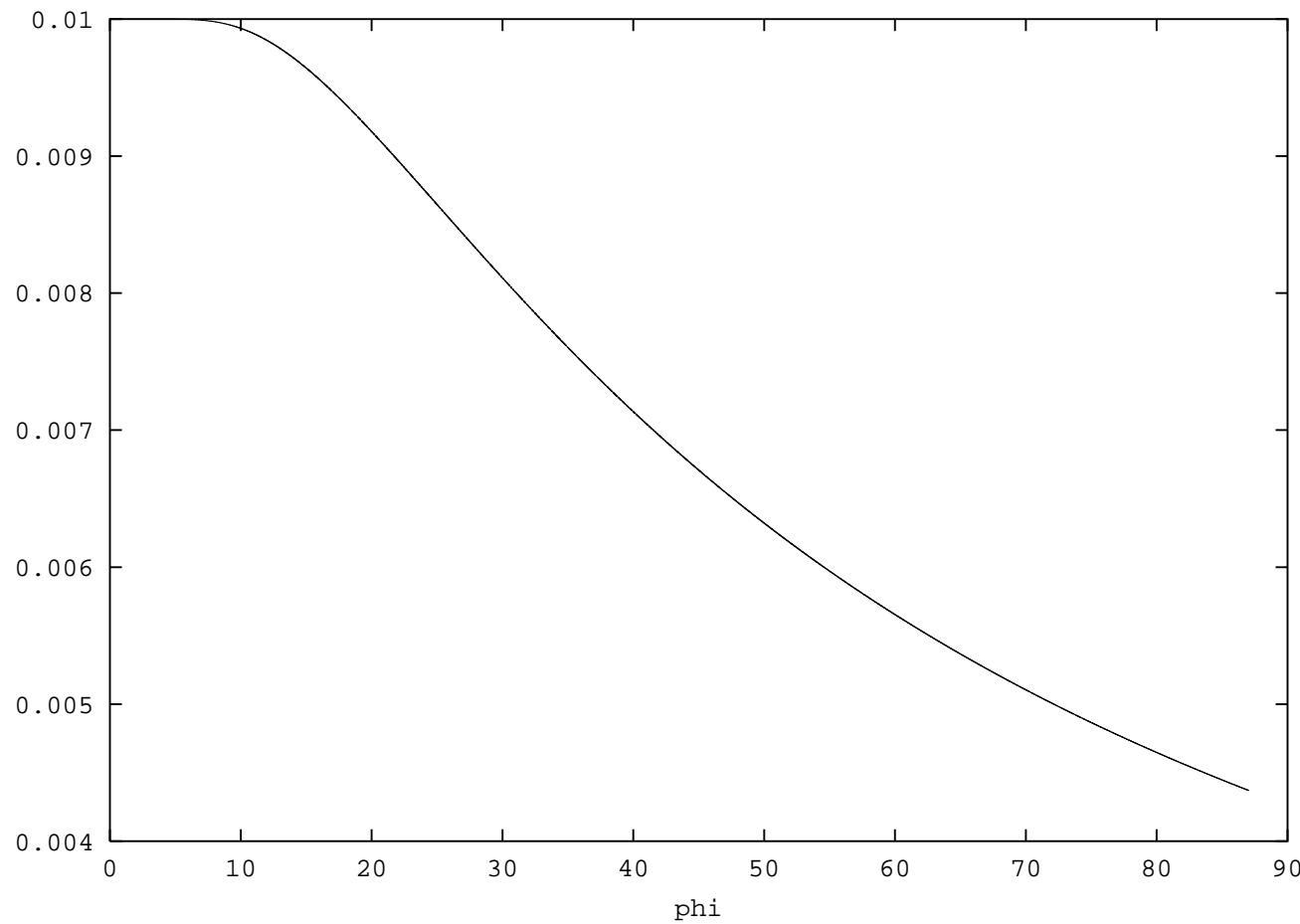
$$\begin{aligned} E(Z_\varphi \mid \mathcal{A}_0) &= E(X_\varphi^{-1} \mid \mathcal{A}_0) \\ &= Z_0 \left(1 - \exp \left\{ \frac{-1}{2 Z_0 \varphi} \right\} \right) < Z_0 \end{aligned}$$

\implies strict local martingale

\implies **strict supermartingale**



Inverse of a squared Bessel process of dimension $\delta = 4$ in φ -time



Expectation of the inverse of the squared Bessel process for $\delta = 4$ in
 φ -time

Time Transformation

- **φ -time**

$$\varphi(t) = \varphi(0) + \frac{1}{4} \int_0^t \frac{c_u^2}{s_u} du$$

$$t \in [0, \infty), s_0 > 0$$

with

$$s_t = s_0 \exp \left\{ \int_0^t b_u du \right\}$$

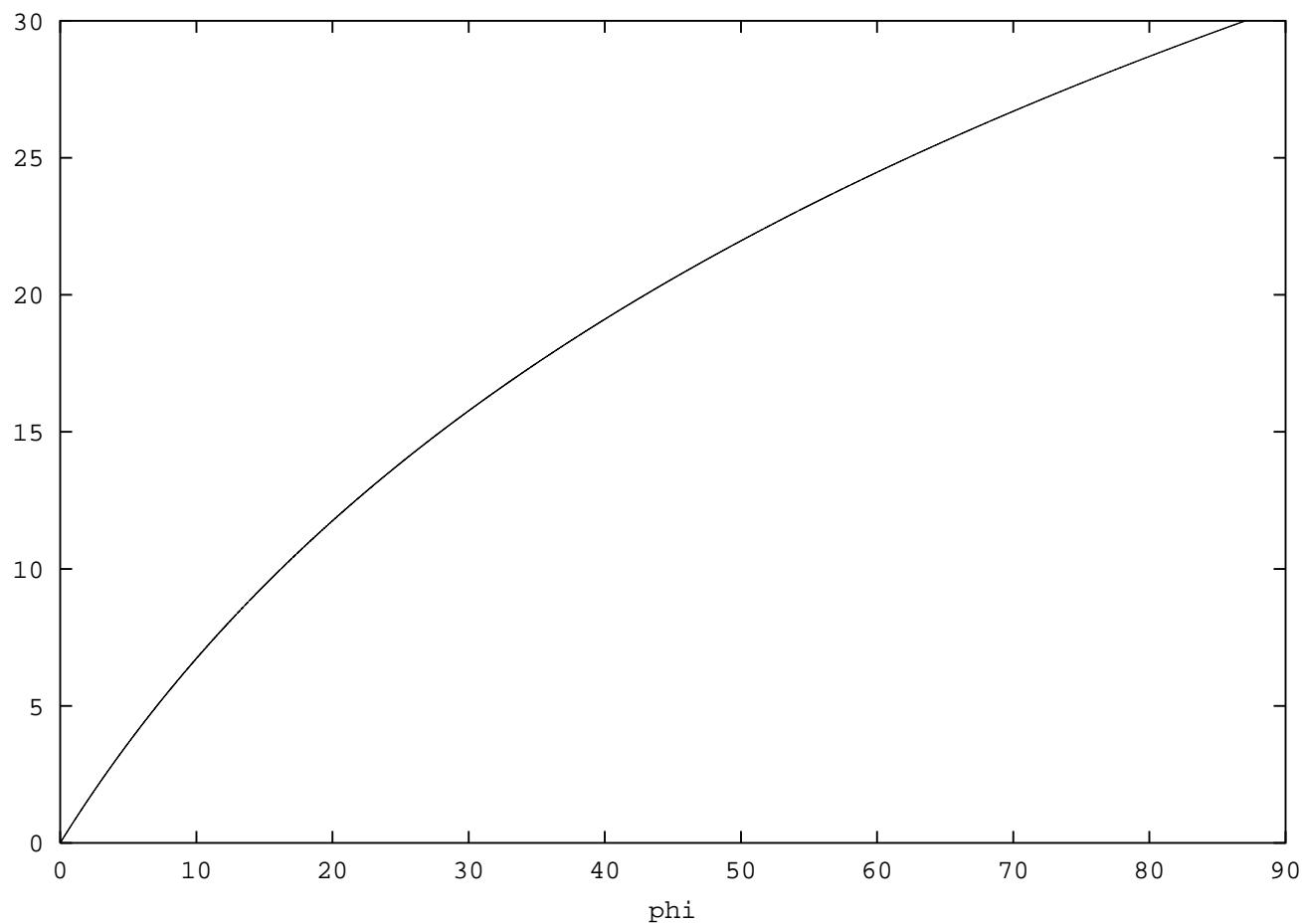
- **φ -time**

for constant $b < 0, c \neq 0$

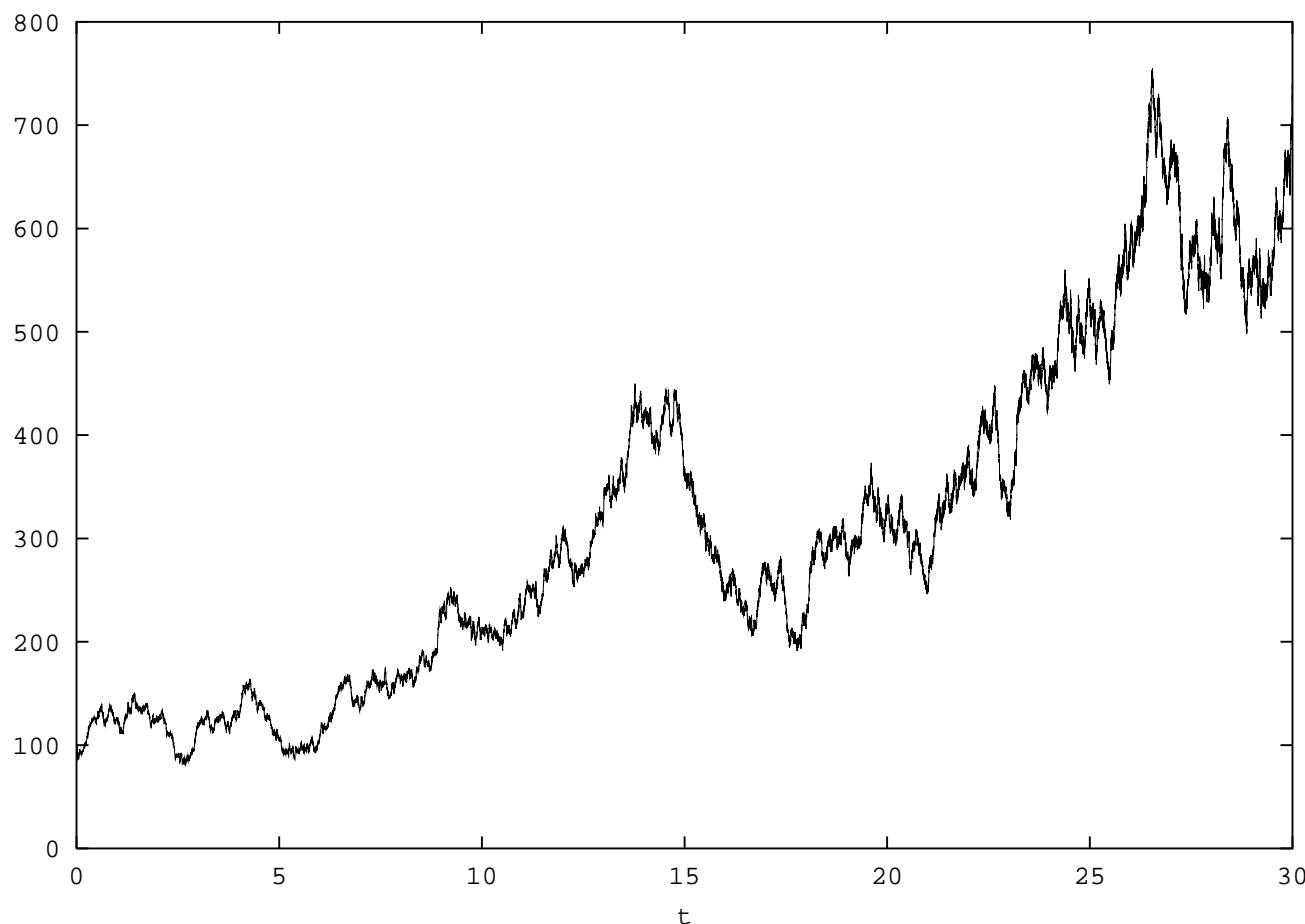
$$\varphi(t) = \varphi(0) + \frac{c^2}{4 b s_0} (1 - \exp\{-b t\})$$

- time

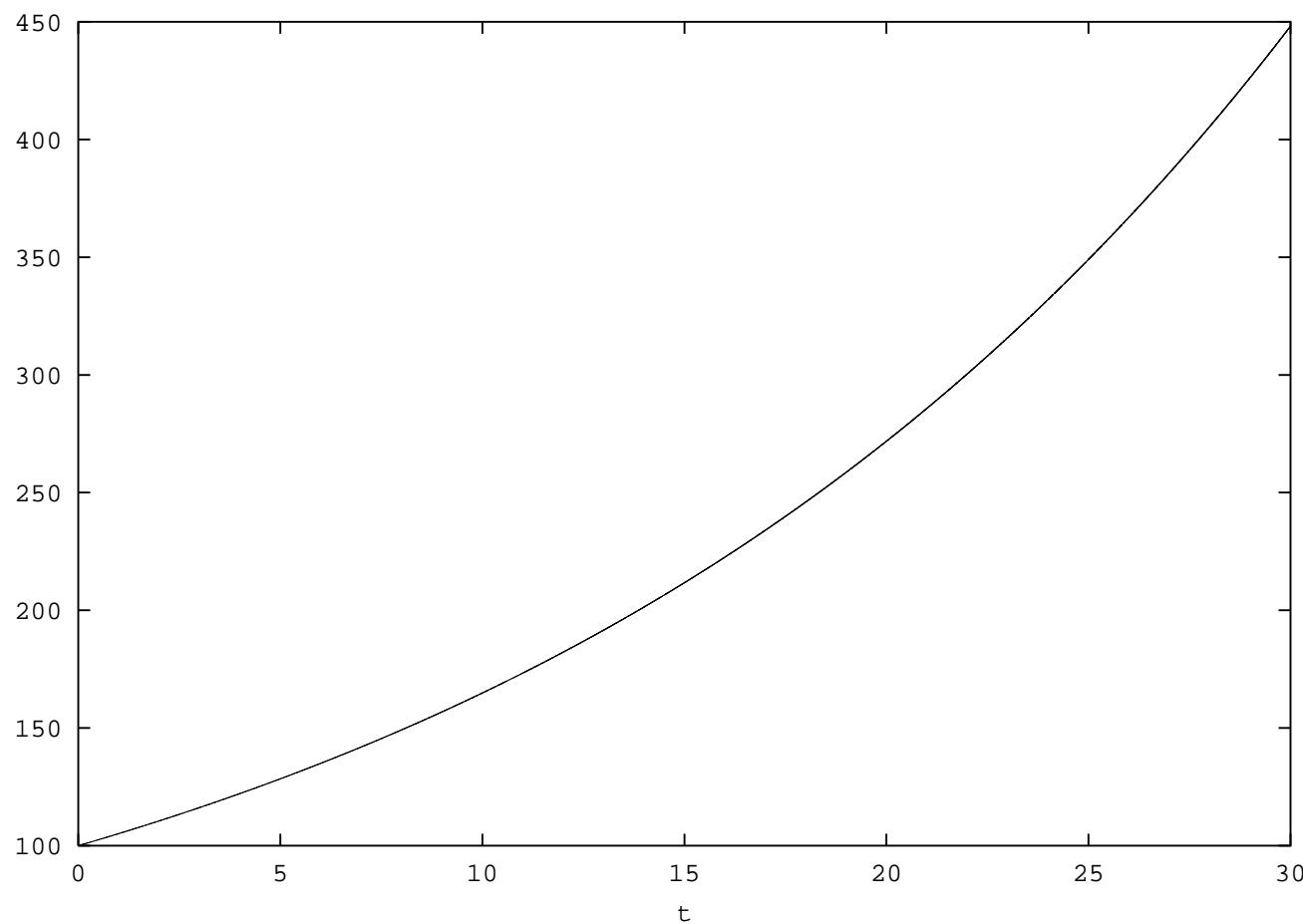
$$t(\varphi) = -\frac{1}{b} \ln \left(1 - \frac{4 b s_0}{c^2} (\varphi - \varphi(0)) \right)$$



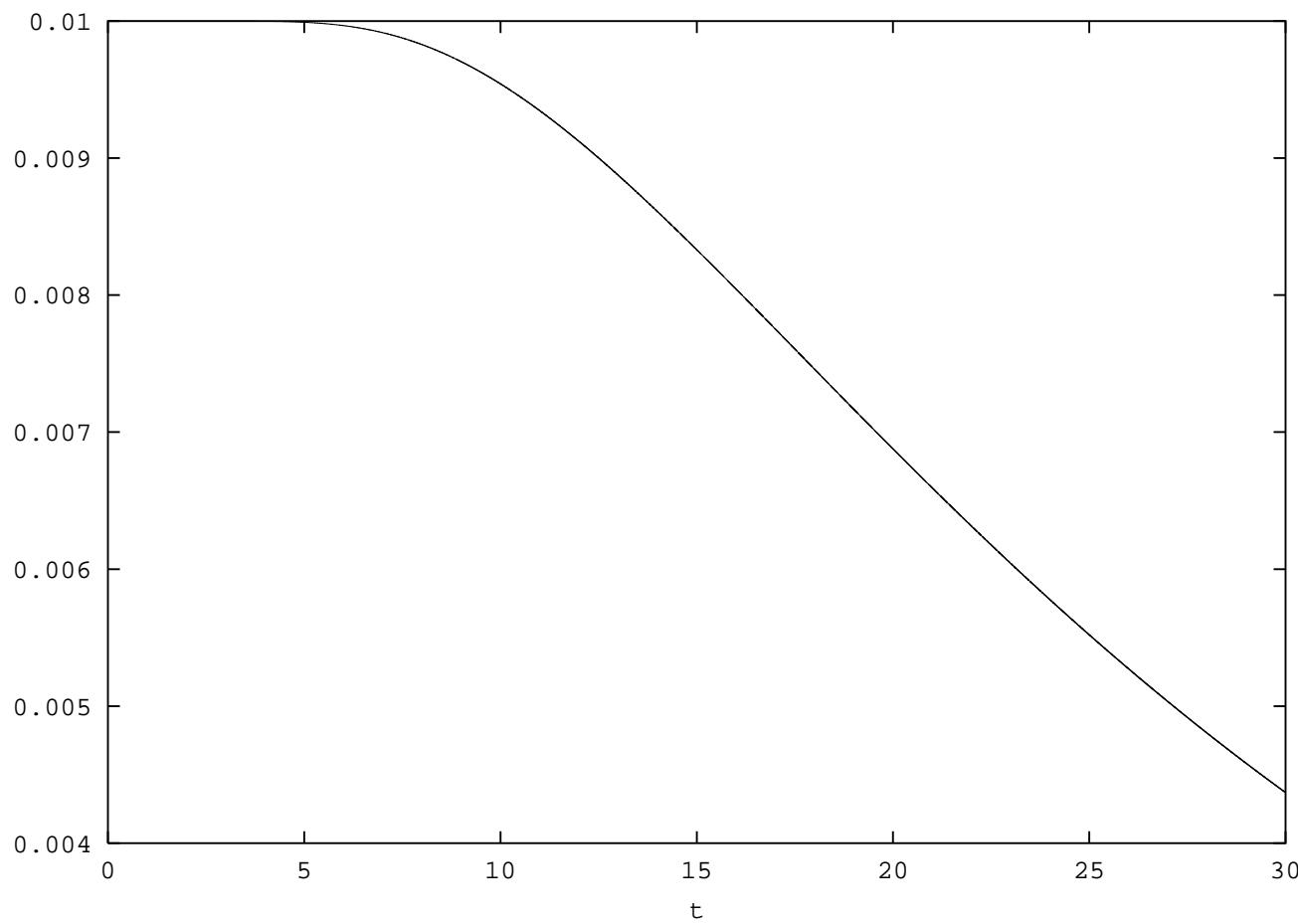
Time $t(\varphi)$ against φ -time



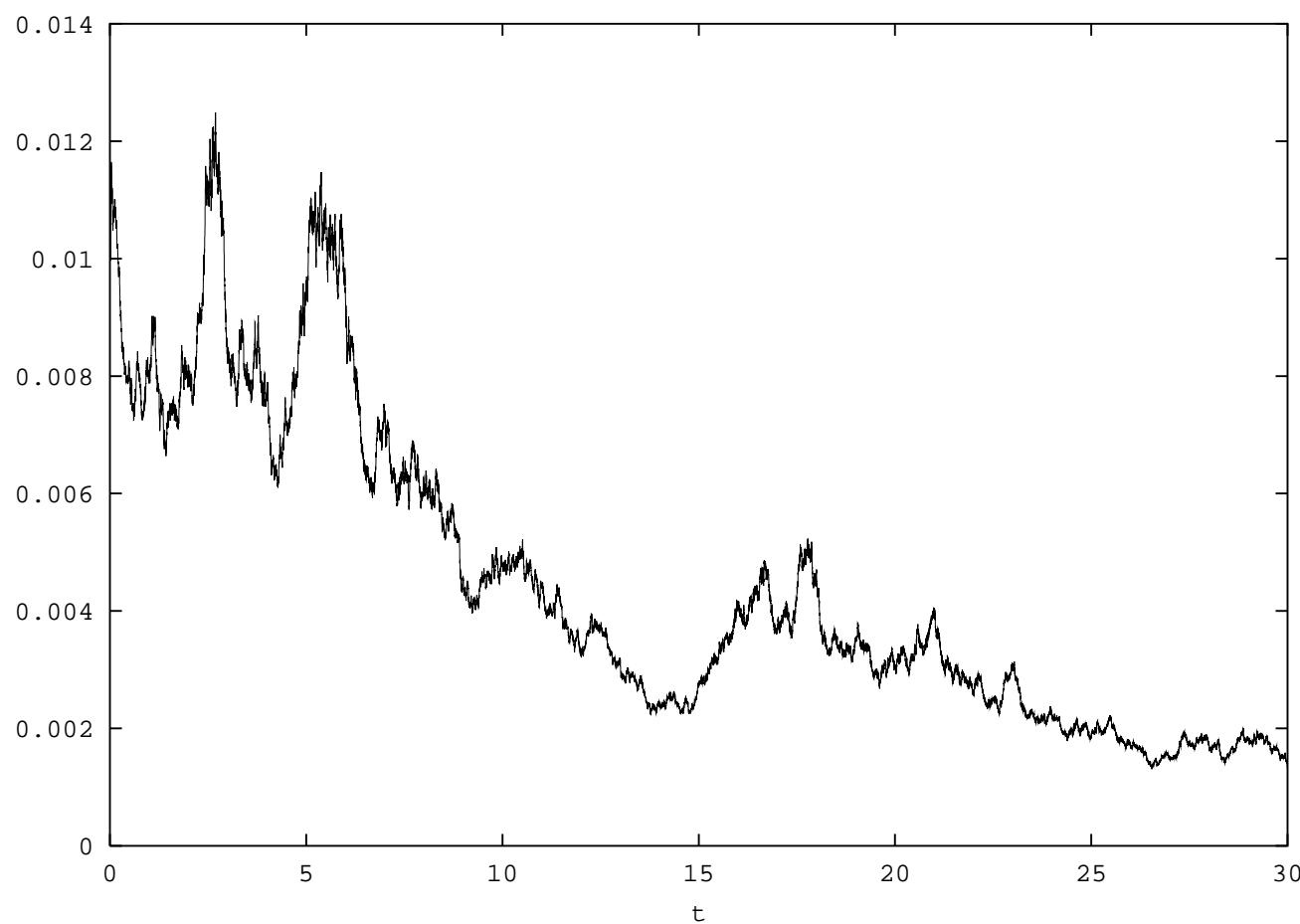
Squared Bessel process in dependence on time t



Expectation of a squared Bessel process in dependence on time t



Expectation of the inverse of a squared Bessel process in dependence on
time t



Inverse of squared Bessel process in dependence on time t

Square Root Process

$$Y_t = s_t X_{\varphi(t)}$$

$$\begin{aligned} dY_t &= s_t dX_{\varphi(t)} + X_{\varphi(t)} ds_t \\ &= s_t \delta d\varphi(t) + s_t 2 \sqrt{X_{\varphi(t)}} dW_{\varphi(t)} + X_{\varphi(t)} s_t b_t dt \\ &= \left(\frac{\delta}{4} c_t^2 + b_t Y_t \right) dt + c_t \sqrt{Y_t} \sqrt{\frac{4 s_t}{c_t^2}} dW_{\varphi(t)} \end{aligned}$$

$$dU_t = \sqrt{\frac{4 s_t}{c_t^2}} dW_{\varphi(t)}$$

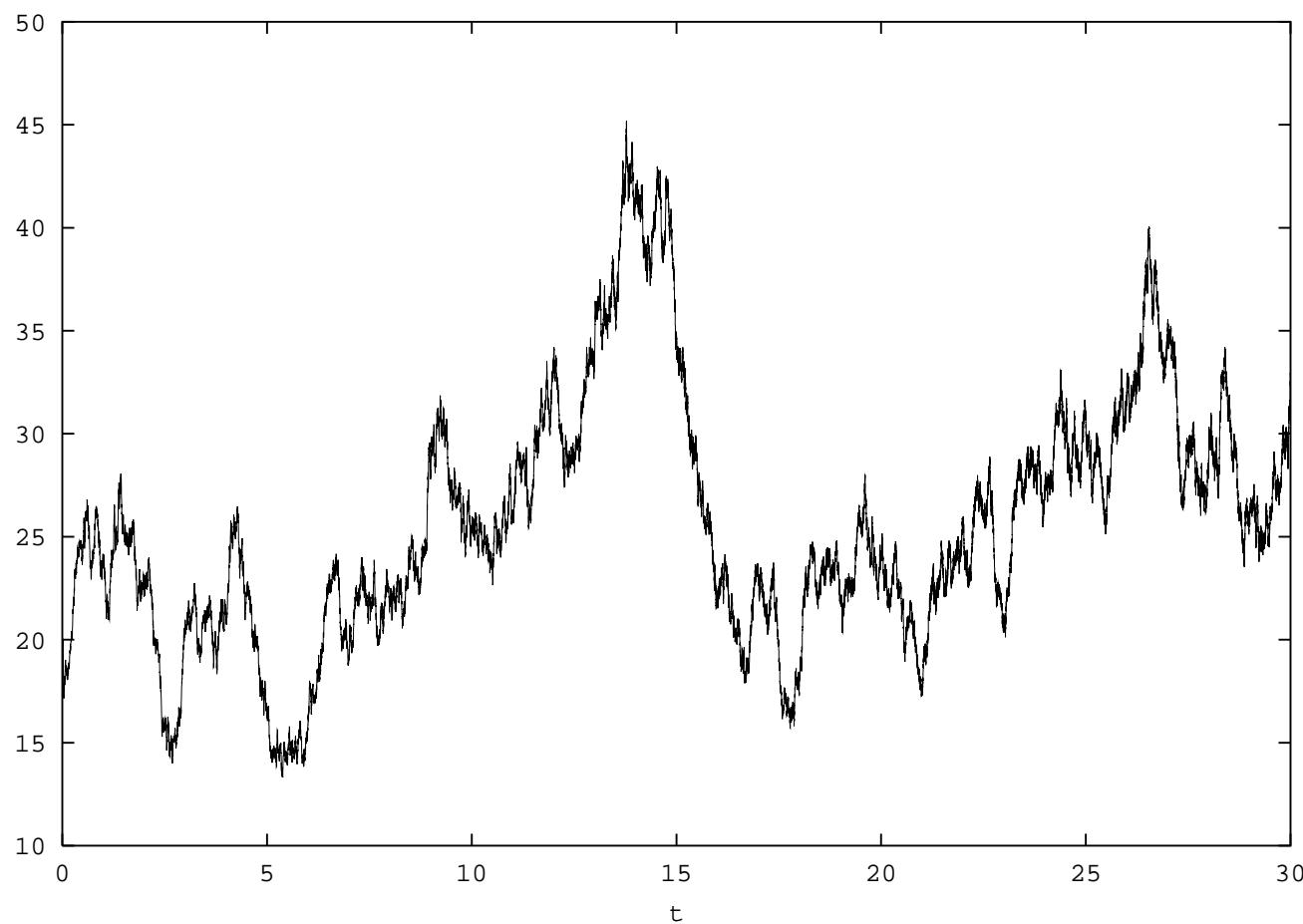
has quadratic variation

$$[U]_t = \int_0^t \frac{4 s_z}{c_z^2} d\varphi(z) = t$$

U Wiener process

\implies SR process

$$dY_t = \left(\frac{\delta}{4} c_t^2 + b_t Y_t \right) dt + c_t \sqrt{Y_t} dU_t$$



Sample path of a square root process in dependence on time t

- moment

$$E(Y_t^\alpha | \mathcal{A}_0) = (2\bar{\varphi}_t \bar{s}_t)^\alpha \exp\left\{-\frac{Y_0}{2\bar{\varphi}_t}\right\} \sum_{k=0}^{\infty} \left(\frac{Y_0}{2\bar{\varphi}_t}\right)^k \frac{\Gamma(\alpha + k + \frac{\delta}{2})}{k! \Gamma(k + \frac{\delta}{2})}$$

$$\delta > 2, \quad \alpha > -\frac{\delta}{2}$$

- moment estimate

$$E(Y_t^\alpha | \mathcal{A}_0) \leq (2\bar{\varphi}_t \bar{s}_t)^\alpha \exp\left\{-\frac{Y_0}{2\bar{\varphi}_t}\right\} \left(\frac{\Gamma(\alpha + \frac{\delta}{2})}{\Gamma(\frac{\delta}{2})} + \exp\left\{\frac{Y_0}{2\bar{\varphi}_t}\right\} \right)$$

$$< \infty \quad \text{for } \alpha \in (-\frac{\delta}{2}, 0)$$

$$\bar{s}_t = \frac{s_t}{s_0} = \exp\left\{\int_0^t b_u du\right\}$$

$$\bar{\varphi}_t = s_0 (\varphi(t) - \varphi(0)) = \frac{1}{4} \int_0^t \frac{c_u^2}{\bar{s}_u} du$$

- first moment of SR process

$$\begin{aligned} E(Y_t | \mathcal{A}_0) &= E(Y_0 | \mathcal{A}_0) \exp \left\{ \int_0^t b_s ds \right\} \\ &\quad + \int_0^t \frac{\delta}{4} c_s^2 \exp \left\{ \int_s^t b_z dz \right\} ds \end{aligned}$$

- case $\delta = 4$

for $c_t^2 = c^2 > 0$ and $b_t = b < 0$

$$\lim_{t \rightarrow \infty} E(Y_t | \mathcal{A}_0) = -\frac{c^2}{b}$$

$$\lim_{t \rightarrow \infty} E(Y_t^{-1} | \mathcal{A}_0) = -2 \frac{b}{c^2}$$

- transition density

$$p(s, Y_s; t, Y_t) = \frac{p_\delta \left(\varphi(s), \frac{Y_s}{s_s}; \varphi(t), \frac{Y_t}{s_t} \right)}{s_t}$$

$$p(0, x; t, y) = \frac{1}{2\bar{s}_t \bar{\varphi}_t} \left(\frac{y}{x \bar{s}_t} \right)^{\frac{\nu}{2}} \exp \left\{ -\frac{x + \frac{y}{\bar{s}_t}}{2 \bar{\varphi}_t} \right\} I_\nu \left(\frac{\sqrt{x \frac{y}{\bar{s}_t}}}{\bar{\varphi}_t} \right)$$

for $0 < t < \infty$, $x, y \in (0, \infty)$,

where $\nu = \frac{\delta}{2} - 1$, $\bar{s}_t = \exp\{bt\}$ and $\bar{\varphi}_t = \frac{c^2}{4b}(1 - \frac{1}{\bar{s}_t})$

- stationary density

is gamma density

$$p_{Y_\infty}(y) = \frac{\left(\frac{-2b}{c^2}\right)^{\frac{\delta}{2}} y^{\frac{\delta}{2}-1}}{\Gamma\left(\frac{\delta}{2}\right)} \exp\left\{\frac{2b}{c^2}y\right\}$$

Minimal Market Model

Pl. (2001), special case of Pl.-Rendek model

Volatility Parametrization

- **Diversification Theorem** \implies

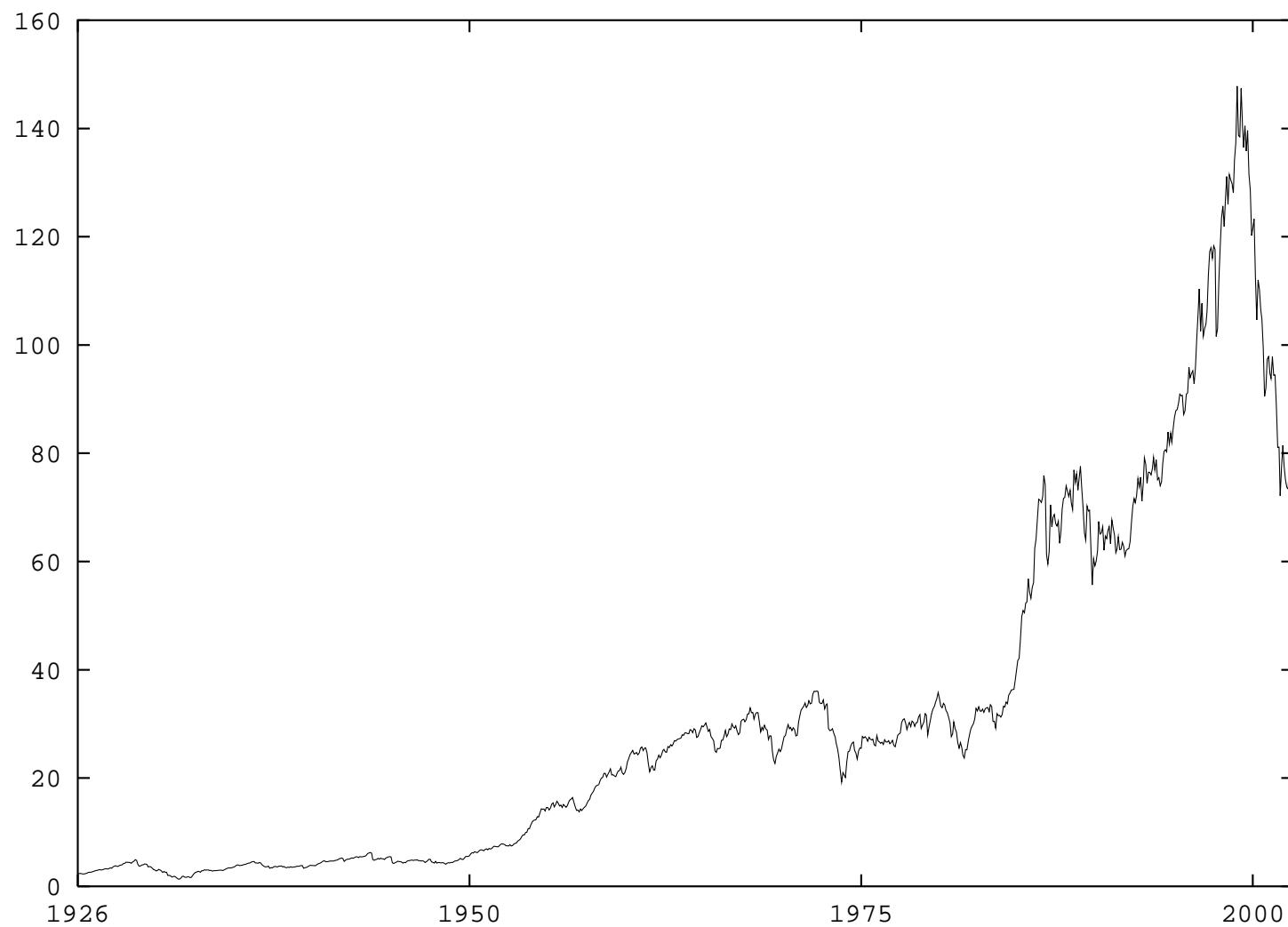
well diversified stock market index approximates NP

- **discounted NP**

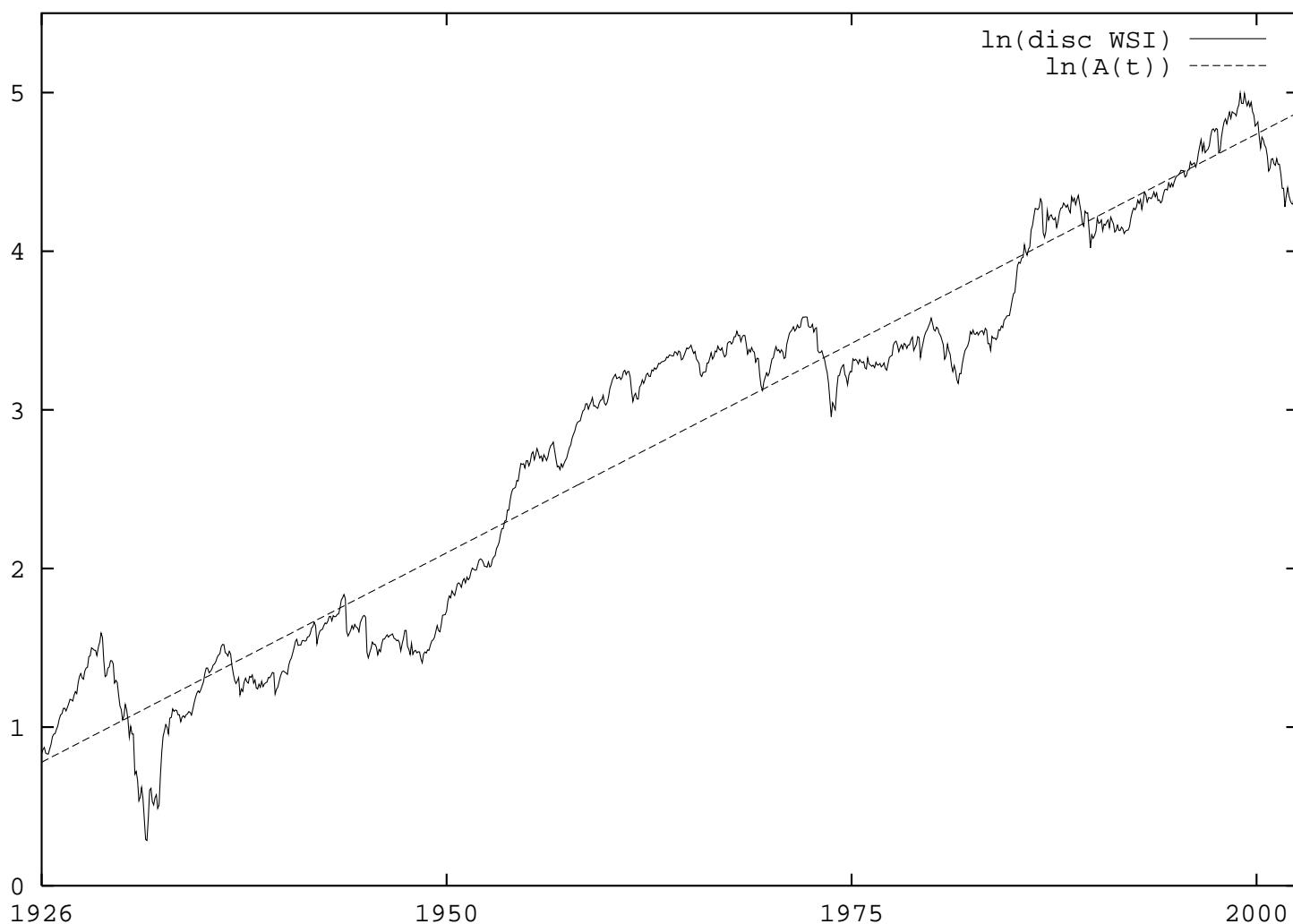
$$d\bar{S}_t^{\delta_*} = \bar{S}_t^{\delta_*} |\theta_t| (|\theta_t| dt + dW_t),$$

where

$$dW_t = \frac{1}{|\theta_t|} \sum_{k=1}^d \theta_t^k dW_t^k$$



Discounted NP



Logarithm of discounted NP

Drift Parametrization

- discounted NP drift

$$\alpha_t = \bar{S}_t^{\delta_*} |\theta_t|^2$$

strictly positive, predictable

\implies

- volatility

$$|\theta_t| = \sqrt{\frac{\alpha_t}{\bar{S}_t^{\delta_*}}}$$

leverage effect creates natural **feedback**
under reasonably independent drift

- discounted NP

$$d\bar{S}_t^{\delta_*} = \alpha_t dt + \sqrt{\bar{S}_t^{\delta_*} \alpha_t} dW_t$$

drift has economic meaning

- transformed time

$$\varphi_t = \frac{1}{4} \int_0^t \alpha_s ds$$

- squared Bessel process of dimension four

$$X_{\varphi_t} = \bar{S}_t^{\delta_*}$$

$$dW(\varphi_t) = \sqrt{\frac{\alpha_t}{4}} dW_t$$

$$dX_\varphi = 4 d\varphi + 2 \sqrt{X_\varphi} dW(\varphi)$$

Revuz & Yor (1999)

economically founded dynamics in φ -time

still no specific dynamics in t -time assumed

Time Transformed Bessel Process

$$d\sqrt{\bar{S}_t^{\delta_*}} = \frac{3\alpha_t}{8\sqrt{\bar{S}_t^{\delta_*}}} dt + \frac{1}{2} \sqrt{\alpha_t} dW_t$$

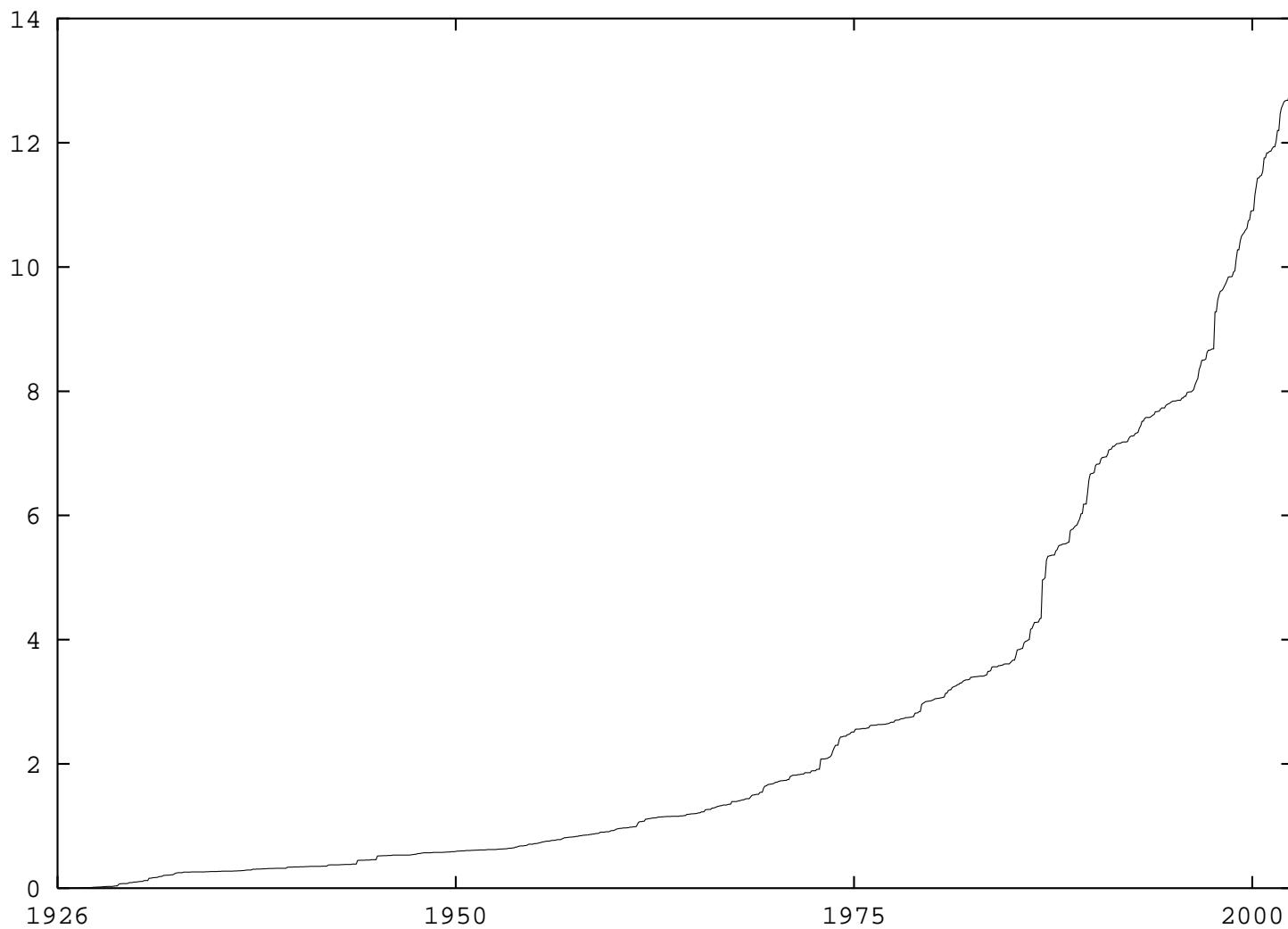
- quadratic variation

$$\left[\sqrt{\bar{S}_t^{\delta_*}} \right]_t = \frac{1}{4} \int_0^t \alpha_s ds$$

- transformed time

$$\varphi_t = \left[\sqrt{\bar{S}_t^{\delta_*}} \right]_t$$

observable



Empirical quadratic variation of the square root of the discounted S&P 500

Stylized Minimal Market Model

Pl. (2001, 2002, 2006), Pl. & Rendek (2012)

- assume **discounted NP drift** as

$$\alpha_t = \alpha \exp \{ \eta t \}$$

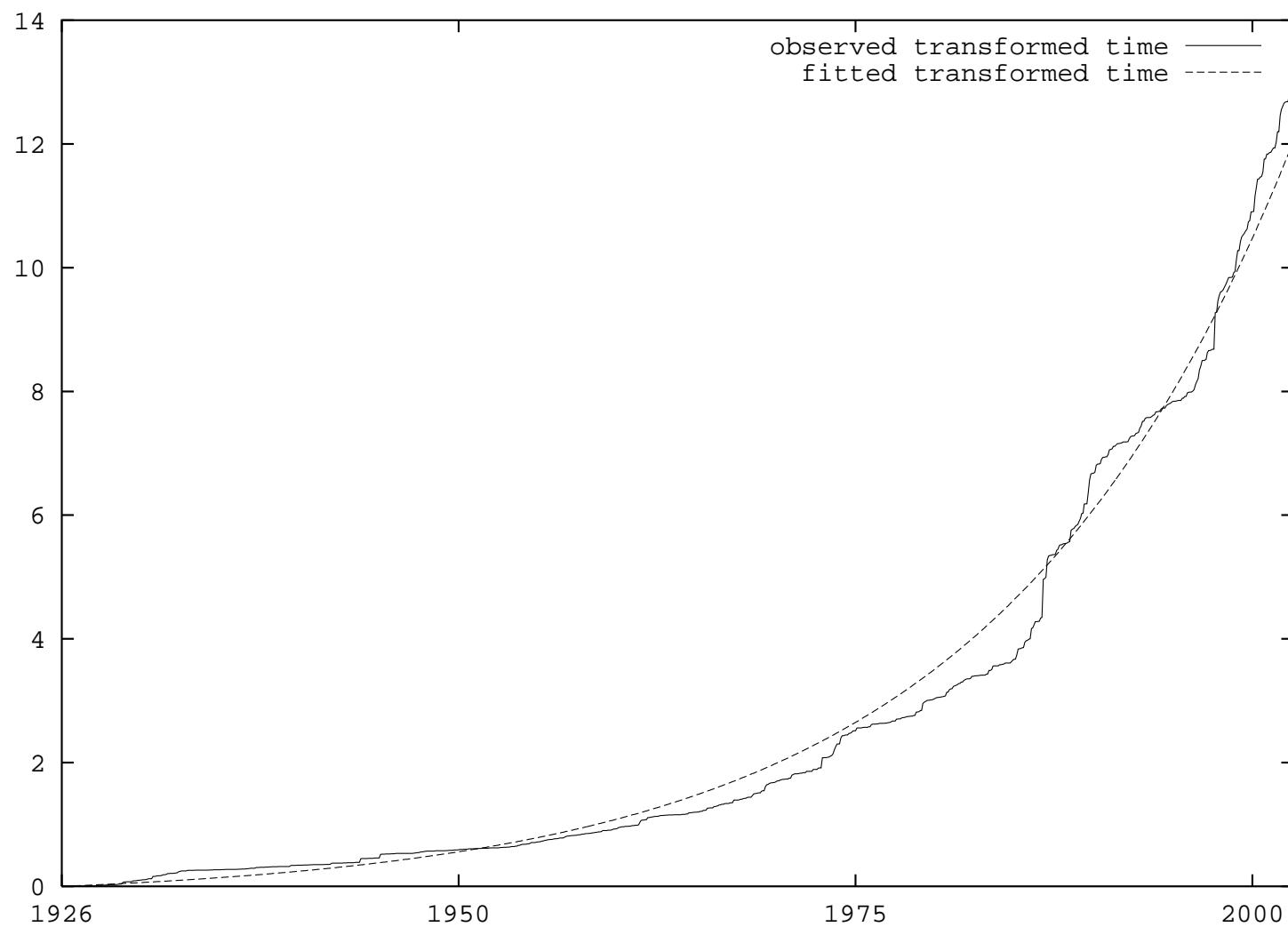
- initial parameter $\alpha > 0$

- net growth rate η

- transformed time

$$\varphi_t = \frac{\alpha}{4} \int_0^t \exp \{ \eta z \} dz$$

$$= \frac{\alpha}{4\eta} (\exp \{ \eta t \} - 1)$$



Fitted and observed transformed time

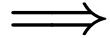
- normalized NP

$$Y_t = \frac{\bar{S}_t^{\delta_*}}{\alpha_t}$$

$$dY_t = (1 - \eta Y_t) dt + \sqrt{Y_t} dW_t$$

square root process of dimension four

\implies parsimonious model, long term viability
 confirmation of stylized Pl.-Rendek model
 Pl. & Rendek (2012)



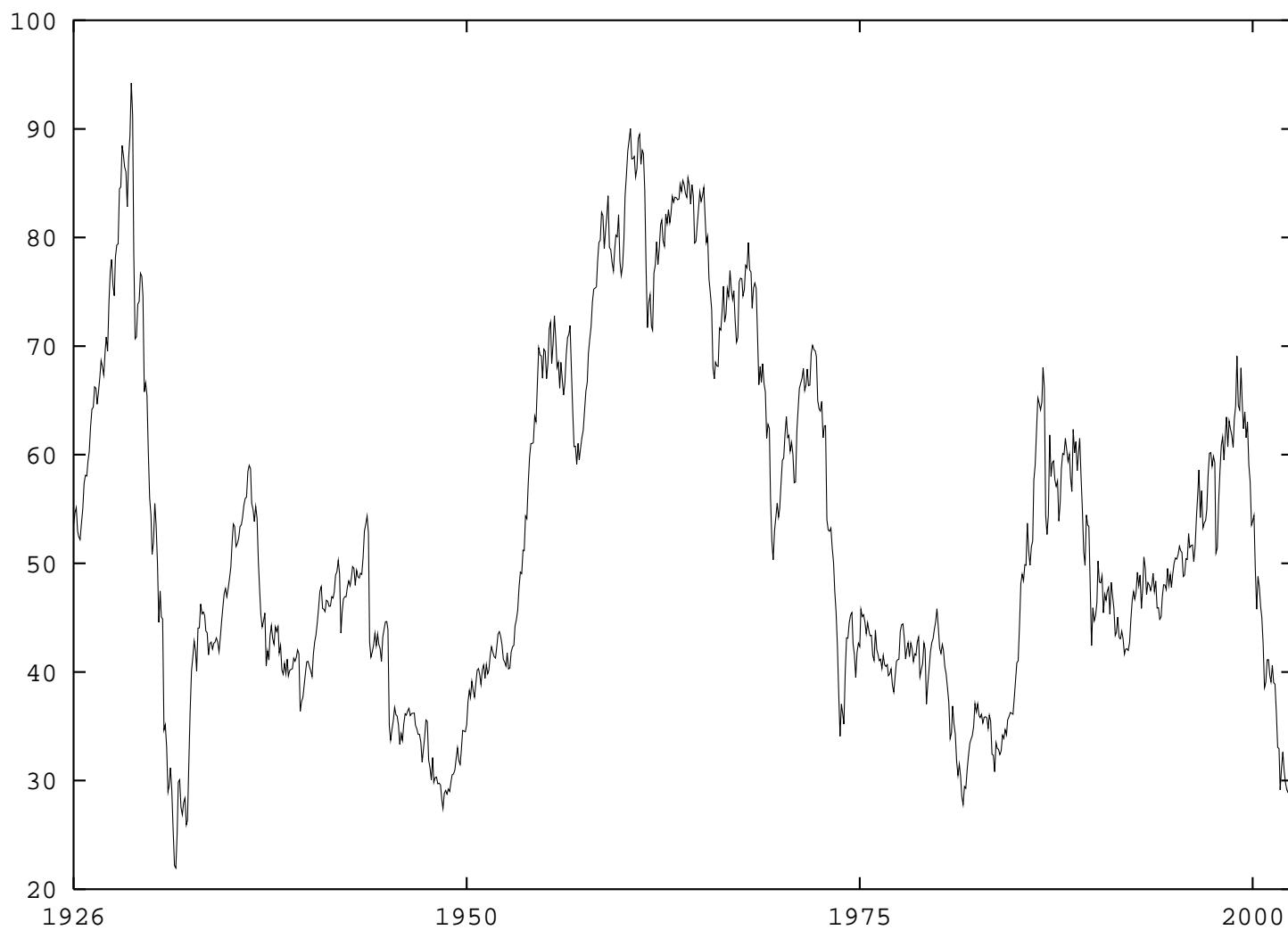
- discounted NP

$$\bar{S}_t^{\delta_*} = Y_t \alpha_t$$

- NP

$$S_t^{\delta_*} = S_t^0 \bar{S}_t^{\delta_*} = S_t^0 Y_t \alpha_t$$

- scaling parameter $\alpha = 0.043$
- net growth rate $\eta = 0.0528$
- reference level $\frac{1}{\eta} = 18.93939$
- speed of adjustment η
- half life time of major displacement $\frac{\ln(2)}{\eta} \approx 13$ years



Normalized NP

- **logarithm of discounted NP**

$$\ln(\bar{S}_t^{\delta_*}) = \ln(Y_t) + \ln(\alpha) + \eta t$$

- no need for extra volatility process in MMM
- realistic long term dynamics

Volatility of NP under the stylized MMM

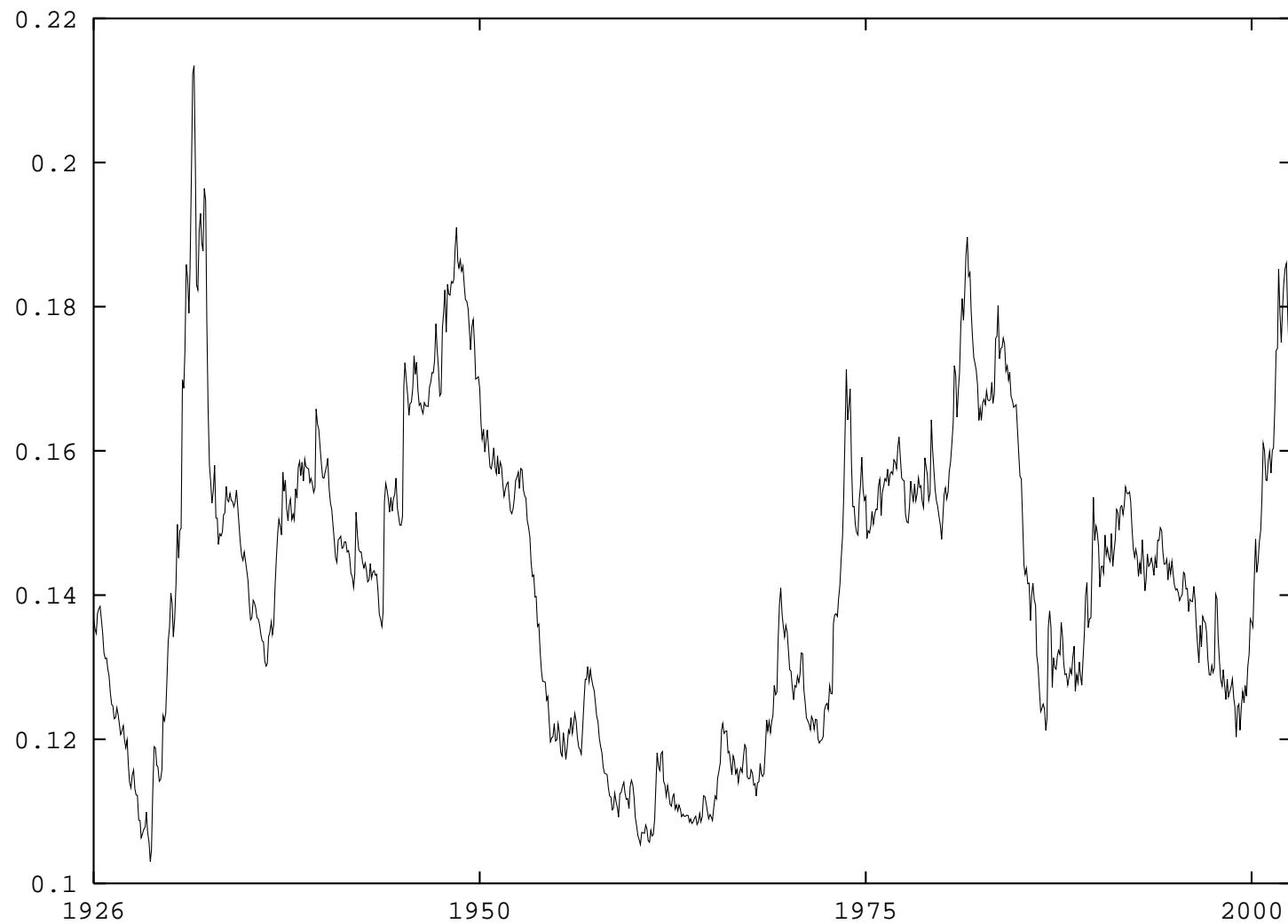
$$|\theta_t| = \frac{1}{\sqrt{Y_t}}$$

- squared volatility

$$d|\theta_t|^2 = d\left(\frac{1}{Y_t}\right) = |\theta_t|^2 \eta dt - (|\theta_t|^2)^{\frac{3}{2}} dW_t$$

3/2 volatility model

Pl. (1997), Lewis (2000)



Volatility of the NP under the stylized MMM

Transition Density of Stylized MMM

- transition density of **discounted NP** \bar{S}^{δ_*}

$$p(s, x; t, y) = \frac{1}{2(\varphi_t - \varphi_s)} \left(\frac{y}{x} \right)^{\frac{1}{2}} \exp \left\{ -\frac{x + y}{2(\varphi_t - \varphi_s)} \right\}$$

$$\times I_1 \left(\frac{\sqrt{xy}}{\varphi_t - \varphi_s} \right)$$

$$\varphi_t = \frac{\alpha}{4\eta} (\exp\{\eta t\} - 1)$$

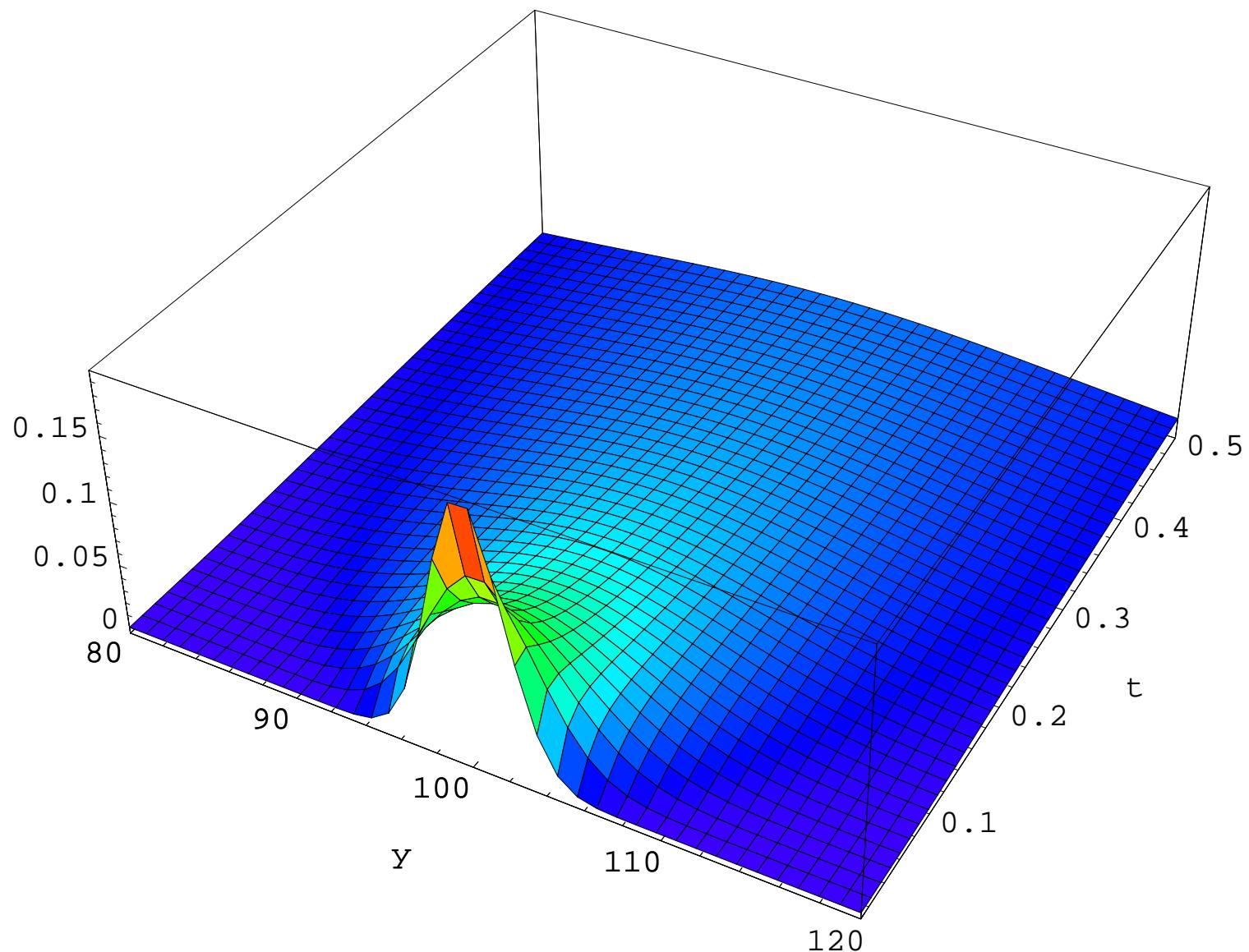
$I_1(\cdot)$ modified Bessel function of the first kind

- **non-central chi-square** distributed random variable:

$$\frac{y}{\varphi_t - \varphi_s} = \frac{\bar{S}_t^{\delta_*}}{\varphi_t - \varphi_s}$$

with $\delta = 4$ degrees of freedom and **non-centrality parameter**:

$$\frac{x}{\varphi_t - \varphi_s} = \frac{\bar{S}_s^{\delta_*}}{\varphi_t - \varphi_s}$$



Transition density of squared Bessel process for $\delta = 4$

Zero Coupon Bond under the stylized MMM

- zero coupon bond

$$P(t, T) = S_t^{\delta_*} E_t \left(\frac{1}{S_T^{\delta_*}} \right) = P_T^*(t) E_t \left(\frac{\bar{S}_t^{\delta_*}}{\bar{S}_T^{\delta_*}} \right)$$

with savings bond

$$P_T^*(t) = \exp \left\{ - \int_t^T r_s ds \right\}$$

$$\delta = 4 \implies$$

$$P(t, T) = P_T^*(t) \left(1 - \exp \left\{ -\frac{\bar{S}_t^{\delta_*}}{2(\varphi_T - \varphi_t)} \right\} \right) < P_T^*(t)$$

for $t \in [0, T]$, Pl. (2002)

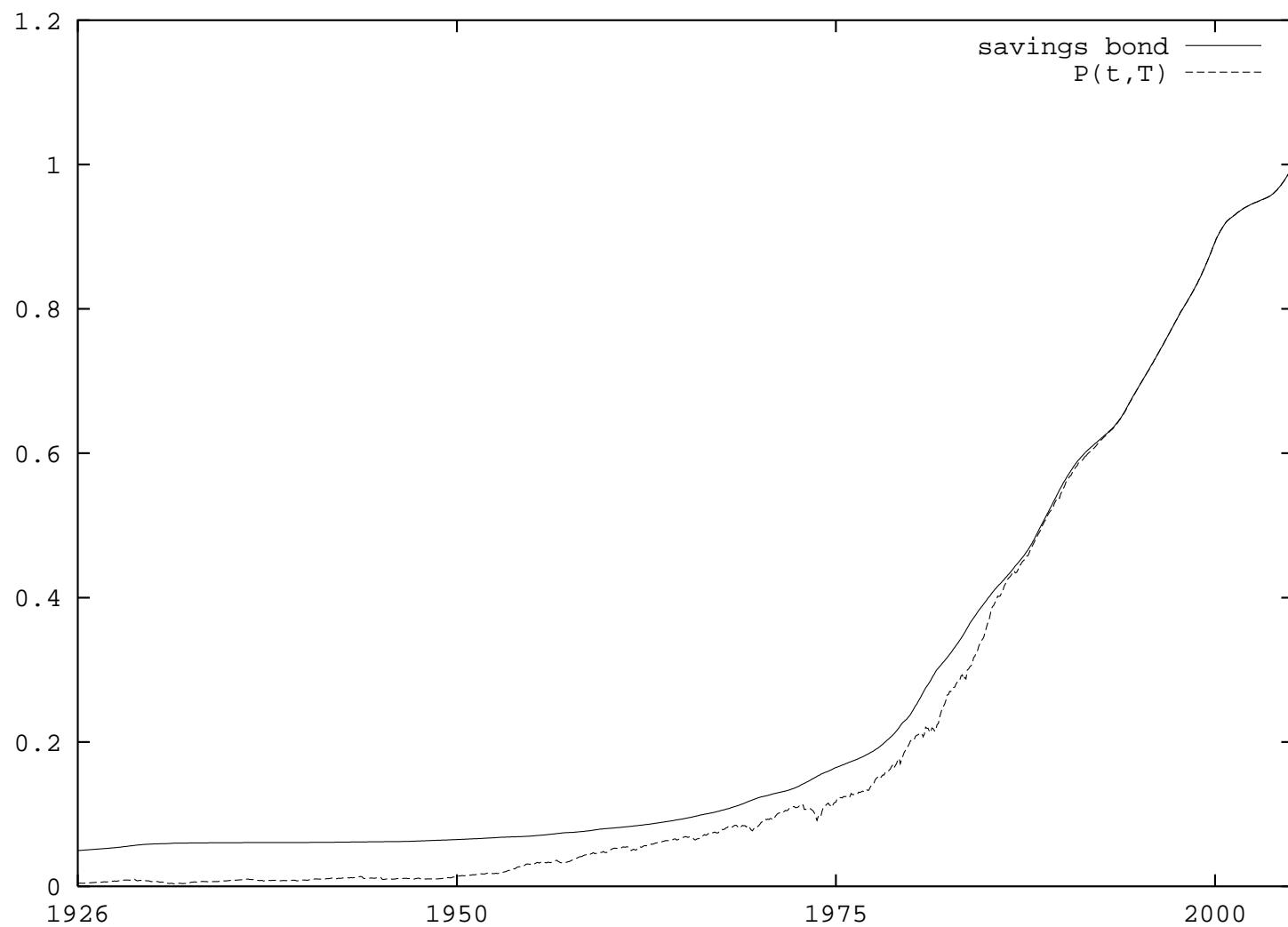
with time transform

$$\varphi_t = \frac{\alpha}{4\eta} (\exp\{\eta t\} - 1)$$

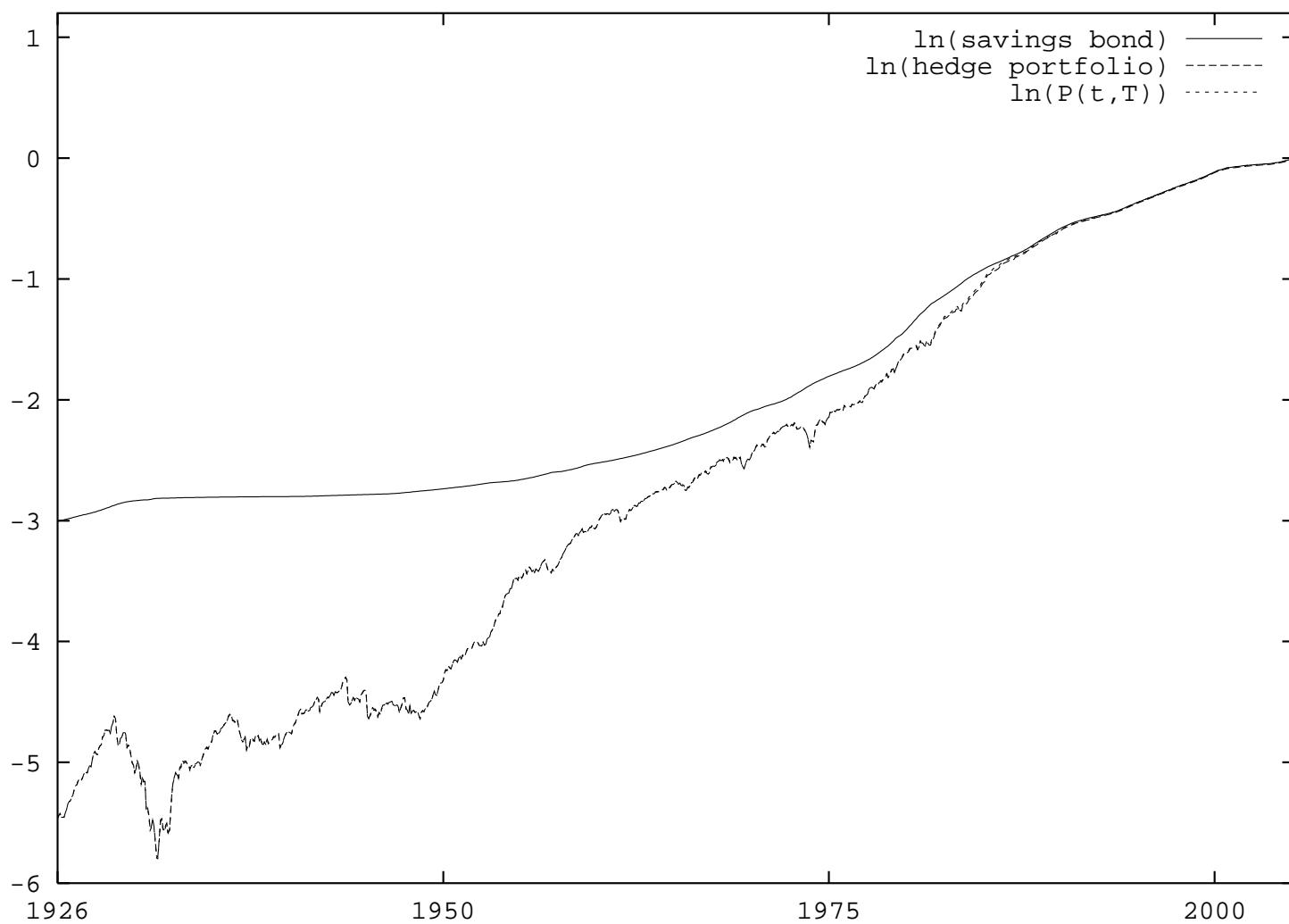
$$P(0, T) = 0.00423$$

$$P_T^*(0) = 0.04947$$

$$\frac{P(0, T)}{P_T^*(0)} \approx 0.0855$$



Zero coupon bond and savings bond



\ln from Zero coupon bond, hedge portfolio and savings bond

Forward Rates under the MMM

- forward rate

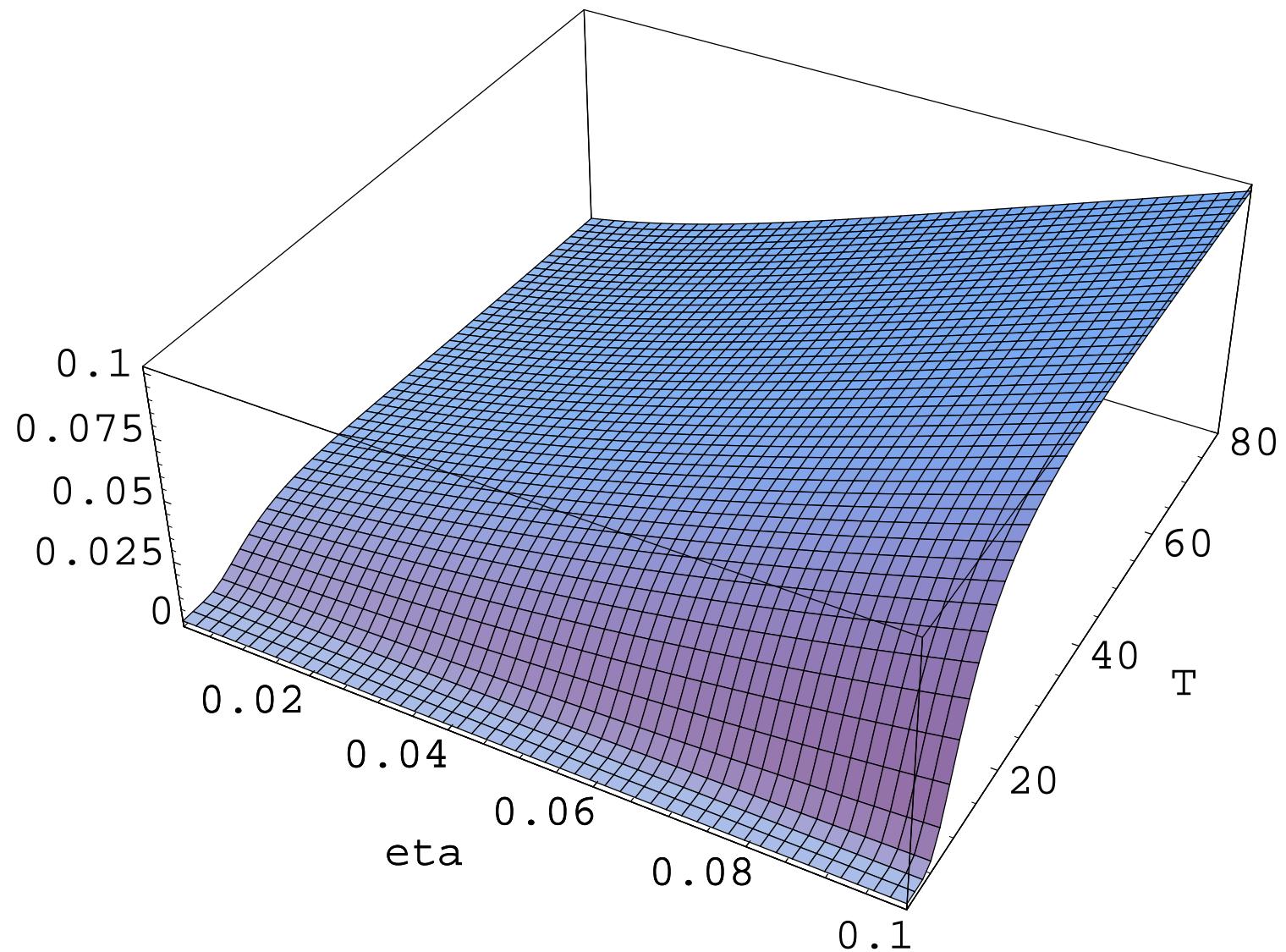
$$\begin{aligned} f(t, T) &= -\frac{\partial}{\partial T} \ln(P(t, T)) \\ &= r_T + n(t, T) \end{aligned}$$

- market price of risk contribution

$$\begin{aligned}
 n(t, T) &= -\frac{\partial}{\partial T} \ln \left(1 - \exp \left\{ -\frac{\bar{S}_t^{\delta_*}}{2(\varphi_T - \varphi_t)} \right\} \right) \\
 &= \frac{1}{\left(\exp \left\{ \frac{\bar{S}_t^{\delta_*}}{2(\varphi_T - \varphi_t)} \right\} - 1 \right)} \frac{\bar{S}_t^{\delta_*}}{(\varphi_T - \varphi_t)^2} \frac{\alpha_T}{8}
 \end{aligned}$$

$$\lim_{T \rightarrow \infty} n(t, T) = \eta$$

Pl. (2005), Miller & Pl. (2005)



Market price of risk contribution in dependence on η and T

Free Snack from Savings Bond

- potential existence of weak form of arbitrage?

borrow amount $P(0, T)$ from savings account

$$S_t^\delta = P(t, T) - P(0, T) \exp\{r t\}$$

such that $S_0^\delta = 0$

$$S_T^\delta = 1 - P(0, T) \exp\{r T\} > 0$$

lower bond

$$S_t^\delta \geq -P(0, T) \exp\{r t\}$$

- since S_t^δ may become negative
not strong arbitrage

Fundamental Theorem of Asset Pricing

Delbaen & Schachermayer (1998) \implies the MMM does not
admit an equivalent risk neutral probability measure

\implies

there is a **free lunch with vanishing risk**

Delbaen & Schachermayer (2006)

Absence of an Equivalent Risk Neutral Probability Measure

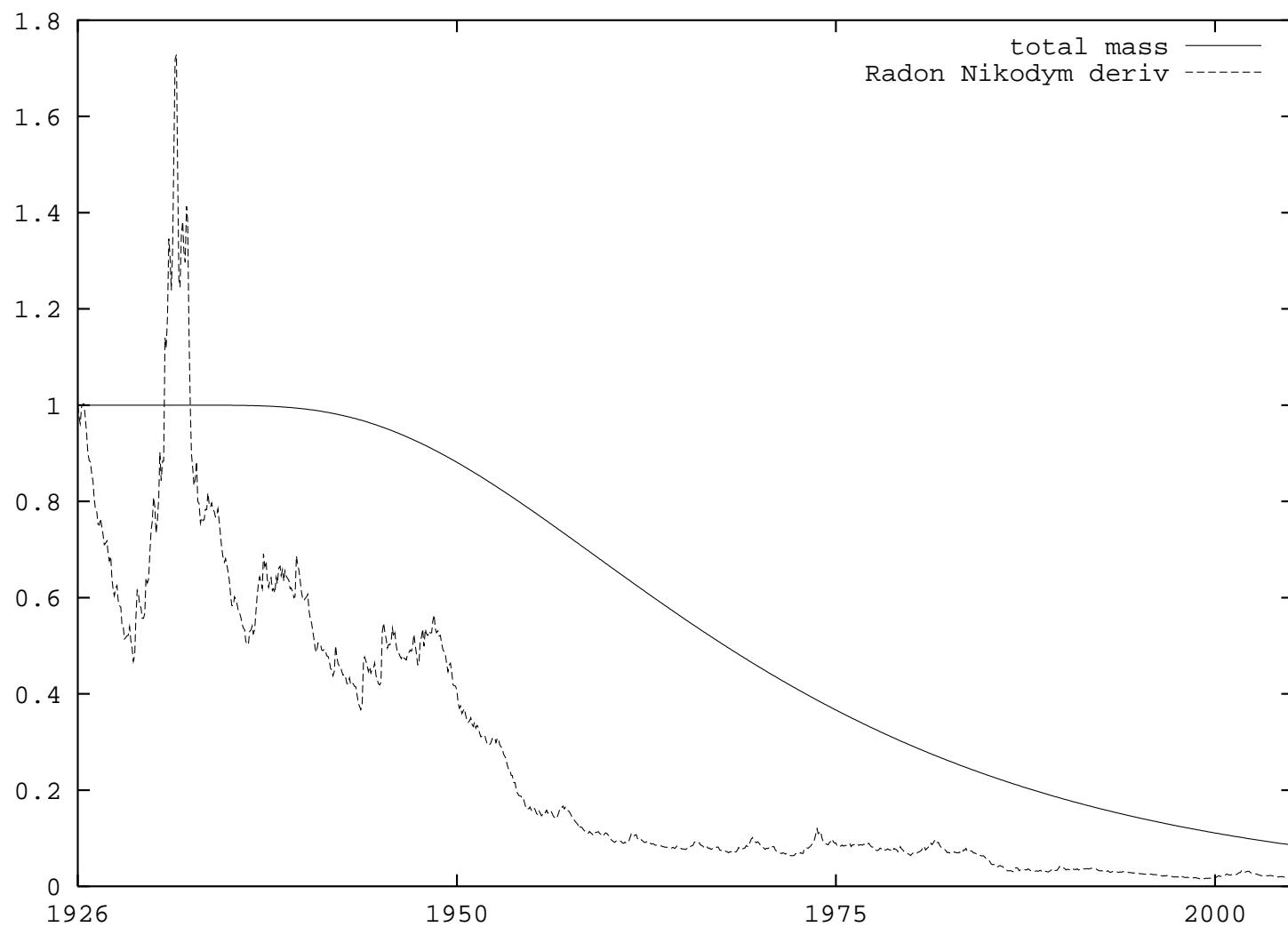
- fair zero coupon bond has a lower price than the savings bond

$$P(t, T) < P_T^*(t) = \exp \left\{ - \int_t^T r_s \, ds \right\} = \frac{S_t^0}{S_T^0}$$

- Radon-Nikodym derivative

$$\Lambda_t = \frac{\hat{S}_t^0}{\hat{S}_0^0} = \frac{\bar{S}_0^{\delta_*}}{\bar{S}_t^{\delta_*}}$$

strict supermartingale under \mathbb{M}^*



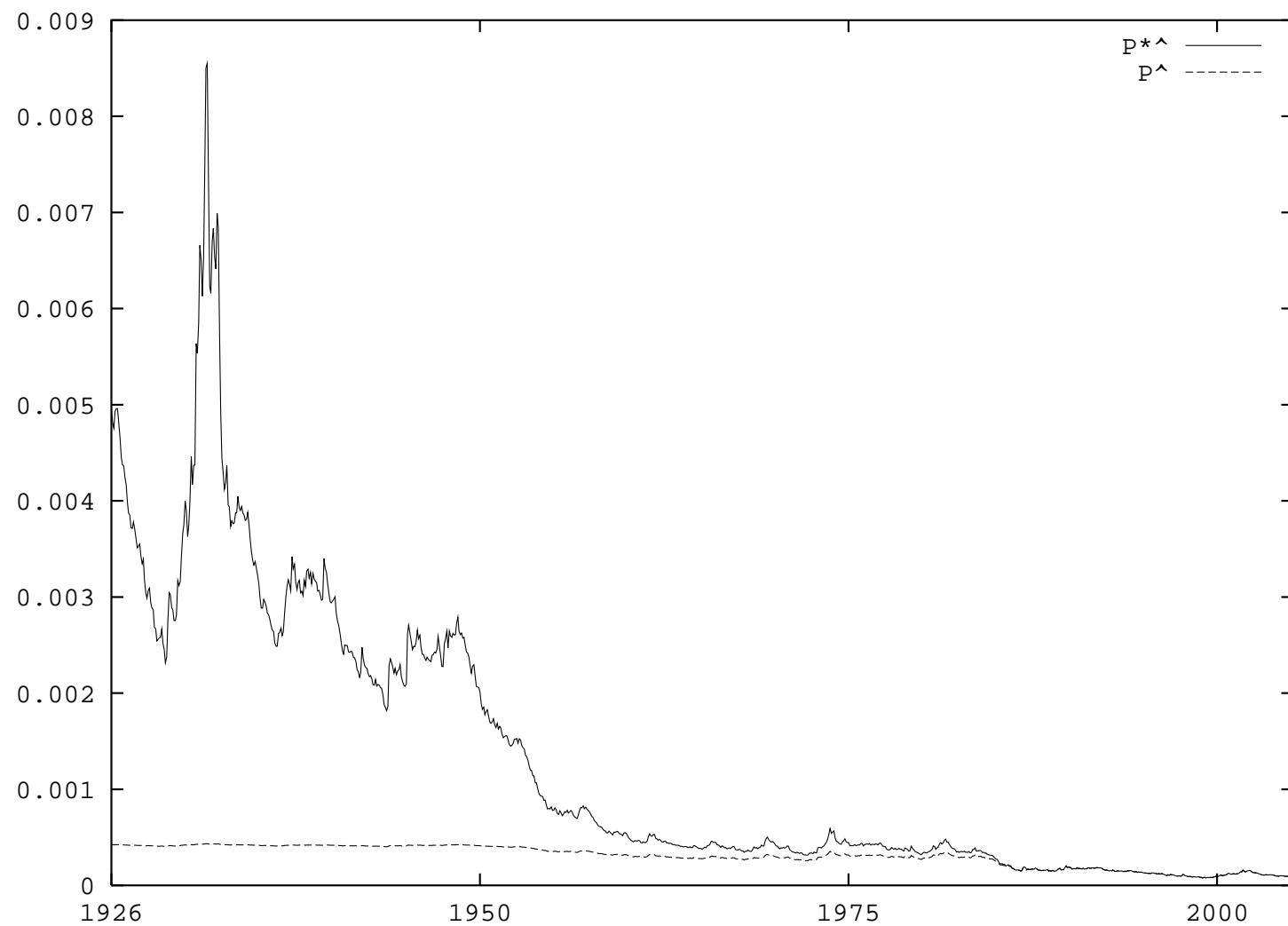
Radon-Nikodym derivative and total mass of candidate risk neutral measure

- hypothetical risk neutral measure

total mass:

$$P_{\theta,T}(\Omega) = E(\Lambda_T) = 1 - \exp \left\{ -\frac{\bar{S}_0^{\delta_*}}{2\varphi_T} \right\} < \Lambda_0 = 1$$

P_θ is **not** a probability measure

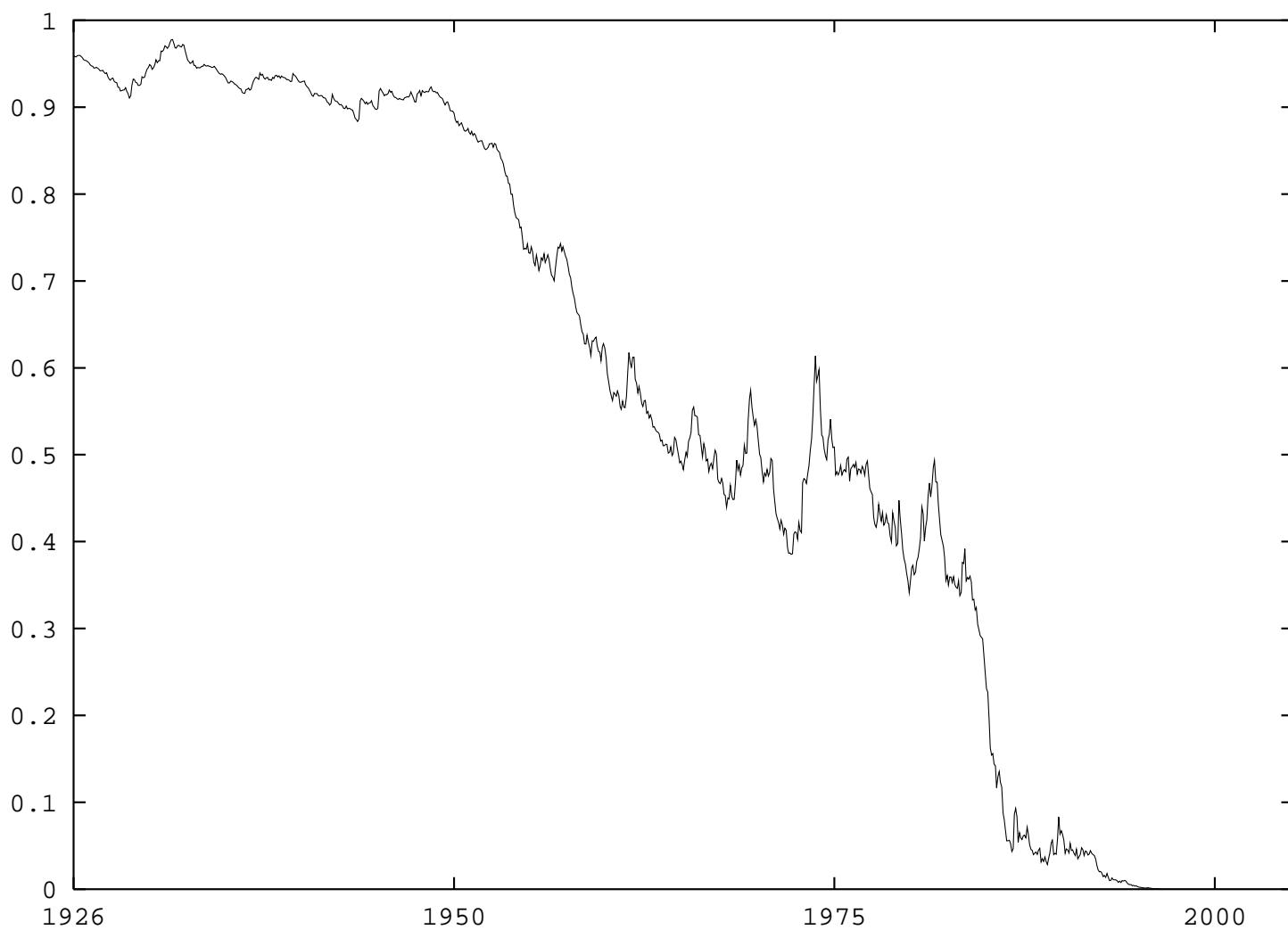


Benchmarked zero coupon bond

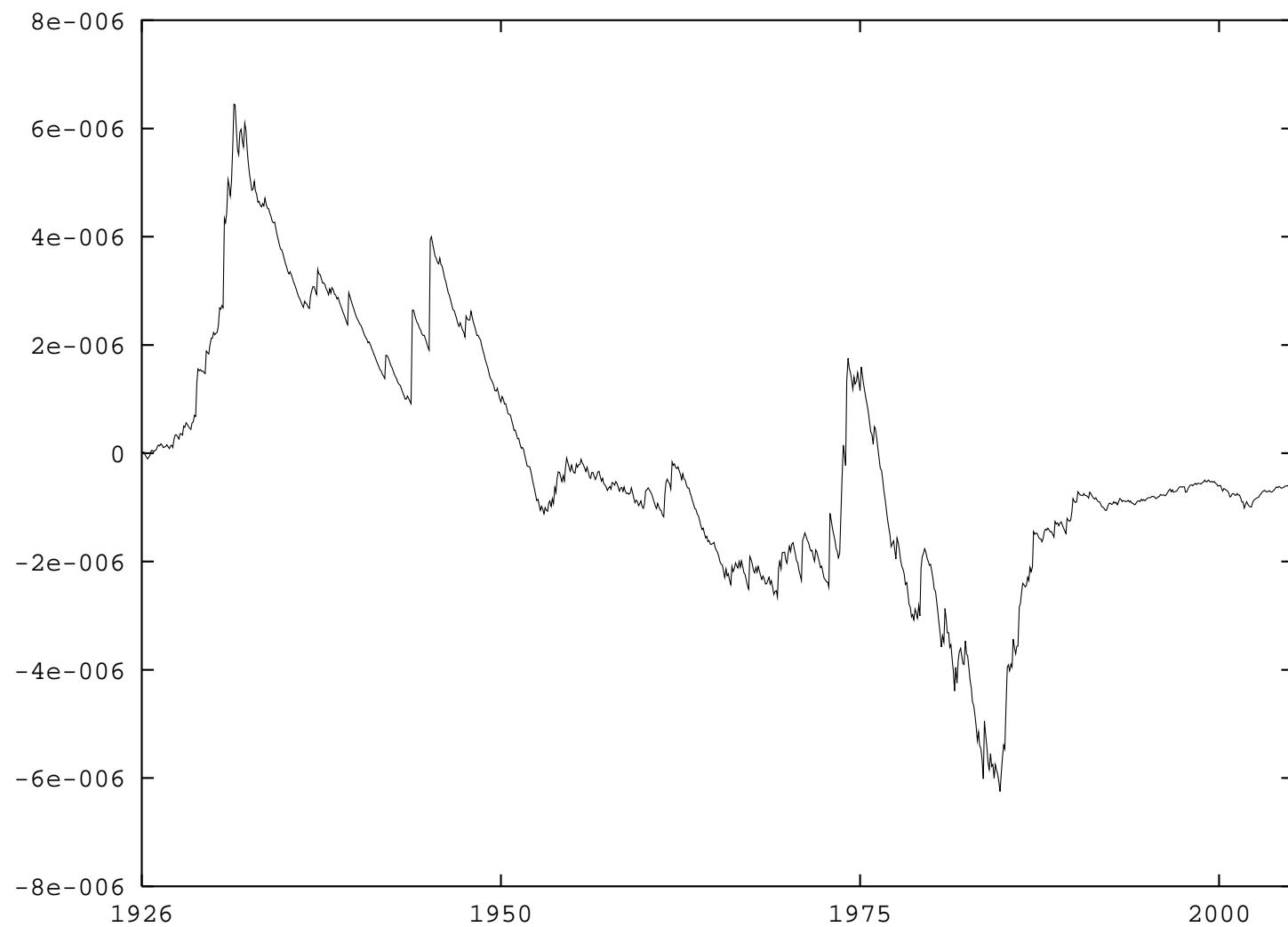
Hedge Simulation

- **delta** units in the NP

$$\begin{aligned}\delta_*(t) &= \frac{\partial P(t, T)}{\partial S_t^{\delta_*}} \\ &= \exp \left\{ - \int_0^T r_s \, ds \right\} \exp \left\{ - \frac{\bar{S}_t^{\delta_*}}{2(\varphi_T - \varphi_t)} \right\} \frac{1}{2(\varphi_T - \varphi_t)}\end{aligned}$$



Ratio in the NP



Benchmarked P&L for hedge portfolio

European Call Options under the MMM

real world pricing formula \implies

$$\begin{aligned} c_{T,K}(t, S_t^{\delta_*}) &= S_t^{\delta_*} E_t \left(\frac{(S_T^{\delta_*} - K)^+}{S_T^{\delta_*}} \right) \\ &= E_t \left(\left(S_t^{\delta_*} - \frac{K S_t^{\delta_*}}{S_T^{\delta_*}} \right)^+ \right) \\ &= S_t^{\delta_*} (1 - \chi^2(d_1; 4, \ell_2)) \\ &\quad - K \exp\{-r(T-t)\} (1 - \chi^2(d_1; 0, \ell_2)) \end{aligned}$$

$\chi^2(\cdot; \cdot, \cdot)$ non-central chi-square distribution

with

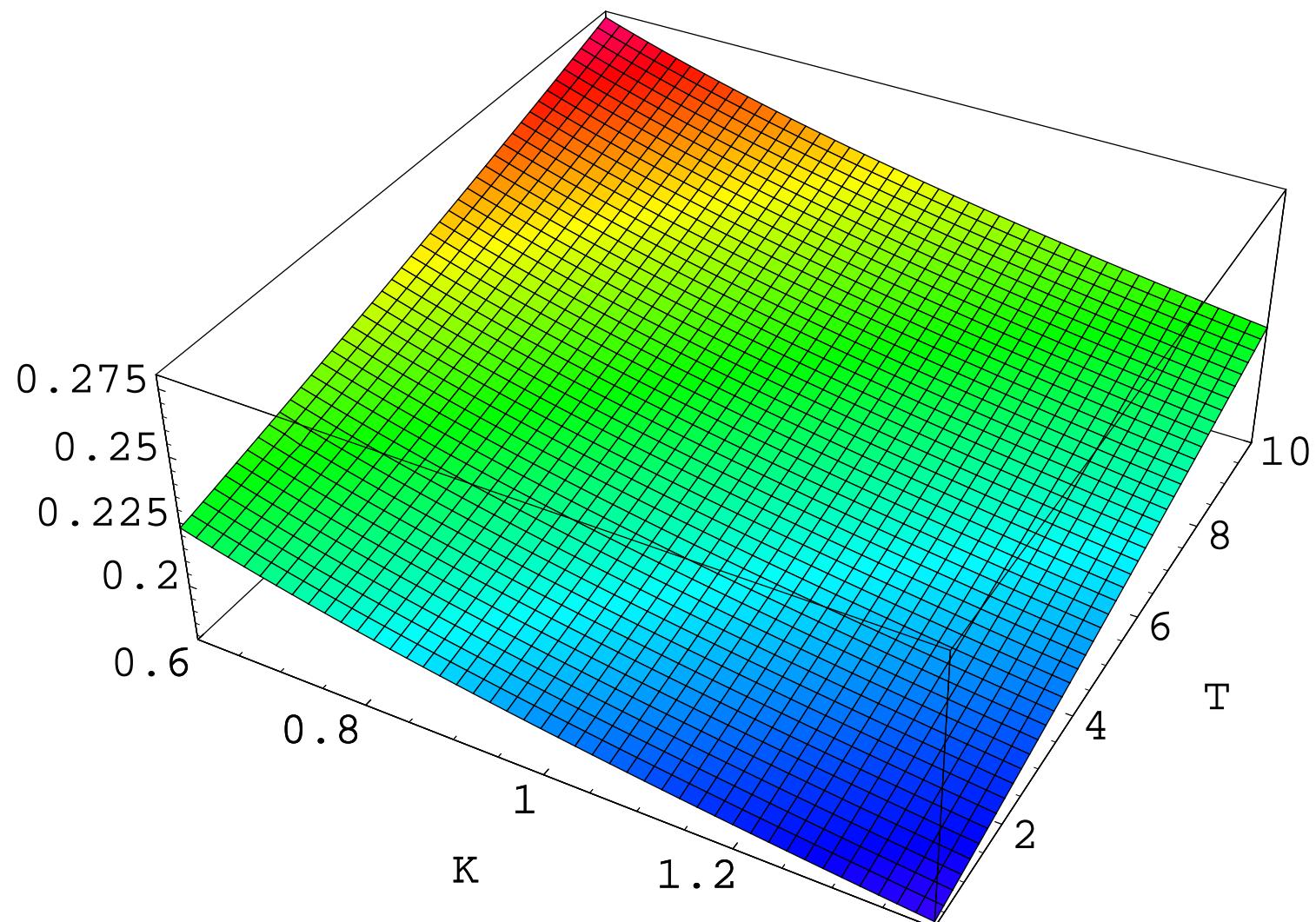
$$d_1 = \frac{4 \eta K \exp\{-r(T-t)\}}{S_t^0 \alpha_t (\exp\{\eta(T-t)\} - 1)}$$

and

$$\ell_2 = \frac{2 \eta S_t^{\delta_*}}{S_t^0 \alpha_t (\exp\{\eta(T-t)\} - 1)}$$

Hulley, Miller & Pl. (2005)

Miller & Pl. (2008)



Implied volatility surface for the stylized MMM

- **implied volatility**

in BS-formula adjust short rate to

$$\hat{r} = -\frac{1}{T-t} \ln(P(t, T))$$

otherwise put and call implied volatilities do not match

European Put Options under the MMM

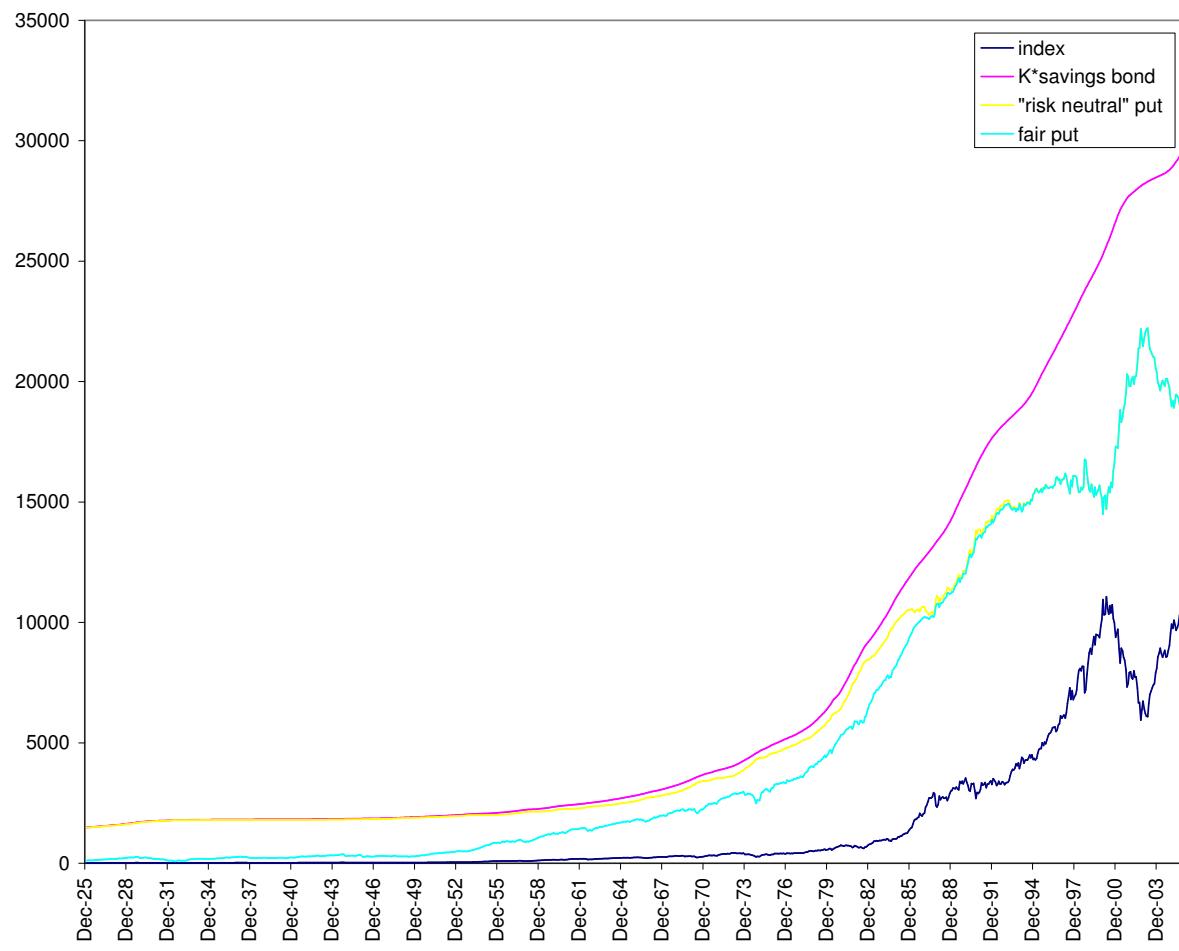
- fair put-call parity relation

$$p_{T,K}(t, \bar{S}_t^{\delta_*}) = c_{T,K}(t, \bar{S}_t^{\delta_*}) - S_t^{\delta_*} + K P(t, T)$$

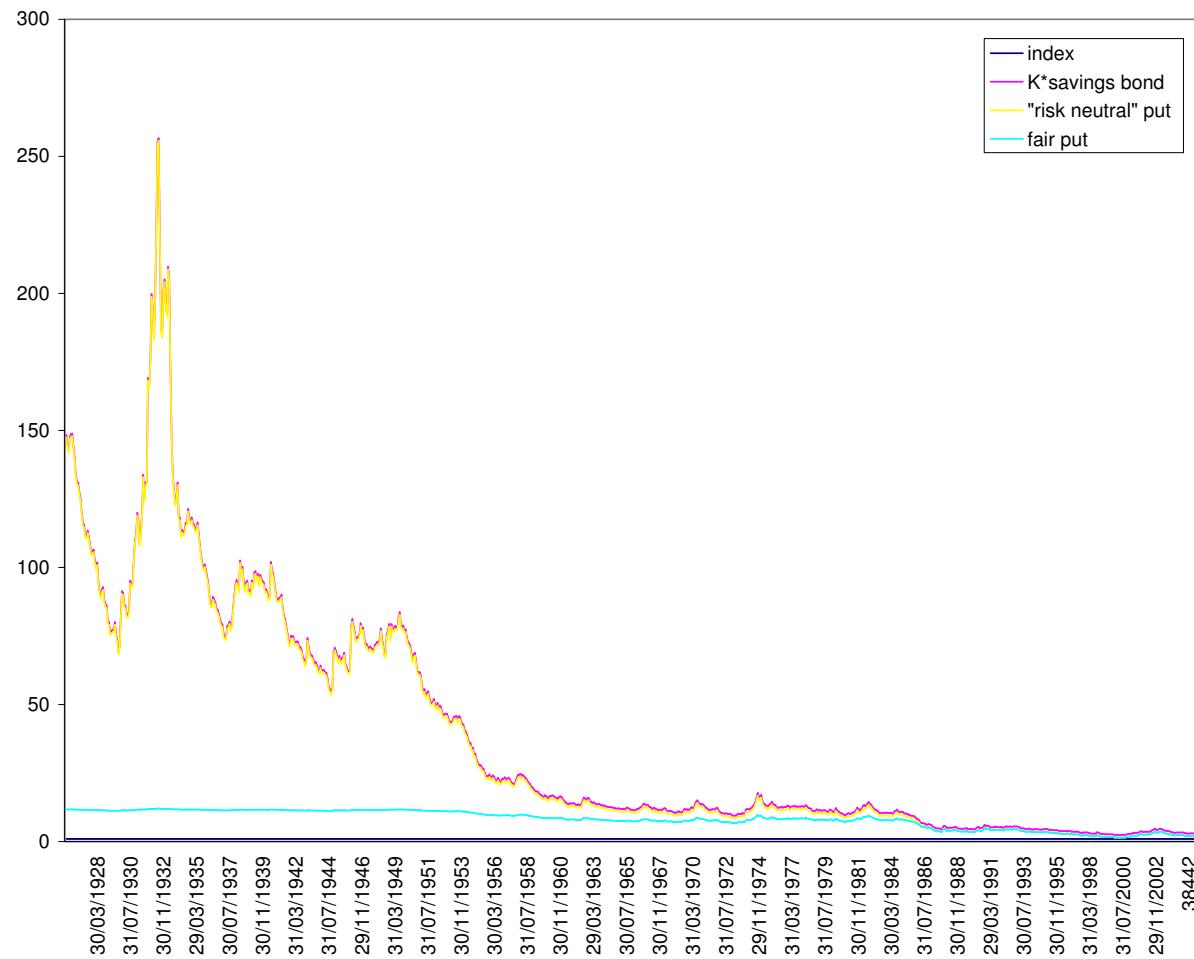
- European put option formula

$$\begin{aligned} p_{T,K}(t, S_t^{\delta_*}) &= -S_t^{\delta_*} (\chi^2(d_1; 4, \ell_2)) \\ &+ K \exp\{-r(T-t)\} (\chi^2(d_1; 0, \ell_2) - \exp\{-\ell_2\}) \end{aligned}$$

Hulley, Miller & Pl. (2005)



Risk neutral and fair put on index



Benchmarked “risk neutral” and fair put on index

- put-call parity **breaks down** if one uses the savings bond $P_T^*(t)$

$$p_{T,K}(t, \bar{S}_t^{\delta_*}) < c_{T,K}(t, \bar{S}_t^{\delta_*}) - S_t^{\delta_*} + K \exp\{-r(T-t)\}$$

for $t \in [0, T)$

Comparison to Hypothetical Risk Neutral Prices

- hypothetical risk neutral price $c_{T,K}^*(t, S_t^{\delta_*})$
of a European call option on the GOP
- benchmarked hypothetical risk neutral call price

$$\hat{c}_{T,K}^*(t, S_t^{\delta_*}) = \frac{c_{T,K}^*(t, S_t^{\delta_*})}{S_t^{\delta_*}}$$

local martingale

$\hat{c}_{T,K}^*(\cdot, \cdot)$ uniformly bounded

$\implies \hat{c}_{T,K}^*$ - martingale \implies

$$\hat{c}_{T,K}^*(t, S_t^{\delta_*}) = \hat{c}_{T,K}(t, S_t^{\delta_*})$$

\implies

$$c_{T,K}^*(t, S_t^{\delta_*}) = c_{T,K}(t, S_t^{\delta_*})$$

- hypothetical risk neutral put-call parity

$$p_{T,K}^*(t, S_t^{\delta_*}) = c_{T,K}^*(t, S_t^{\delta_*}) - S_t^{\delta_*} + K P_T^*(t)$$

since $P_T^*(t) > P(t, T) \implies$

$$p_{T,K}(t, S_t^{\delta_*}) < p_{T,K}^*(t, S_t^{\delta_*})$$

$$t \in [0, T)$$

Difference in Asymptotic Put Prices

- hypothetical risk neutral prices can become extreme if NP value tends towards zero
- **asymptotic fair zero coupon bond**

$$\lim_{\bar{S}_t^{\delta_*} \rightarrow 0} P_T(t, T) \stackrel{\text{a.s.}}{=} 0$$

for $t \in [0, T)$

- **asymptotic fair European call**

$$\lim_{\bar{S}_t^{\delta_*} \rightarrow 0} c_{T,K}(t, S_t^{\delta_*}) \stackrel{\text{a.s.}}{=} \lim_{\bar{S}_t^{\delta_*} \rightarrow 0} S_t^{\delta_*} \hat{c}_{T,K}(t, S_t^{\delta_*}) = 0$$

- asymptotic fair put

fair put-call parity \implies

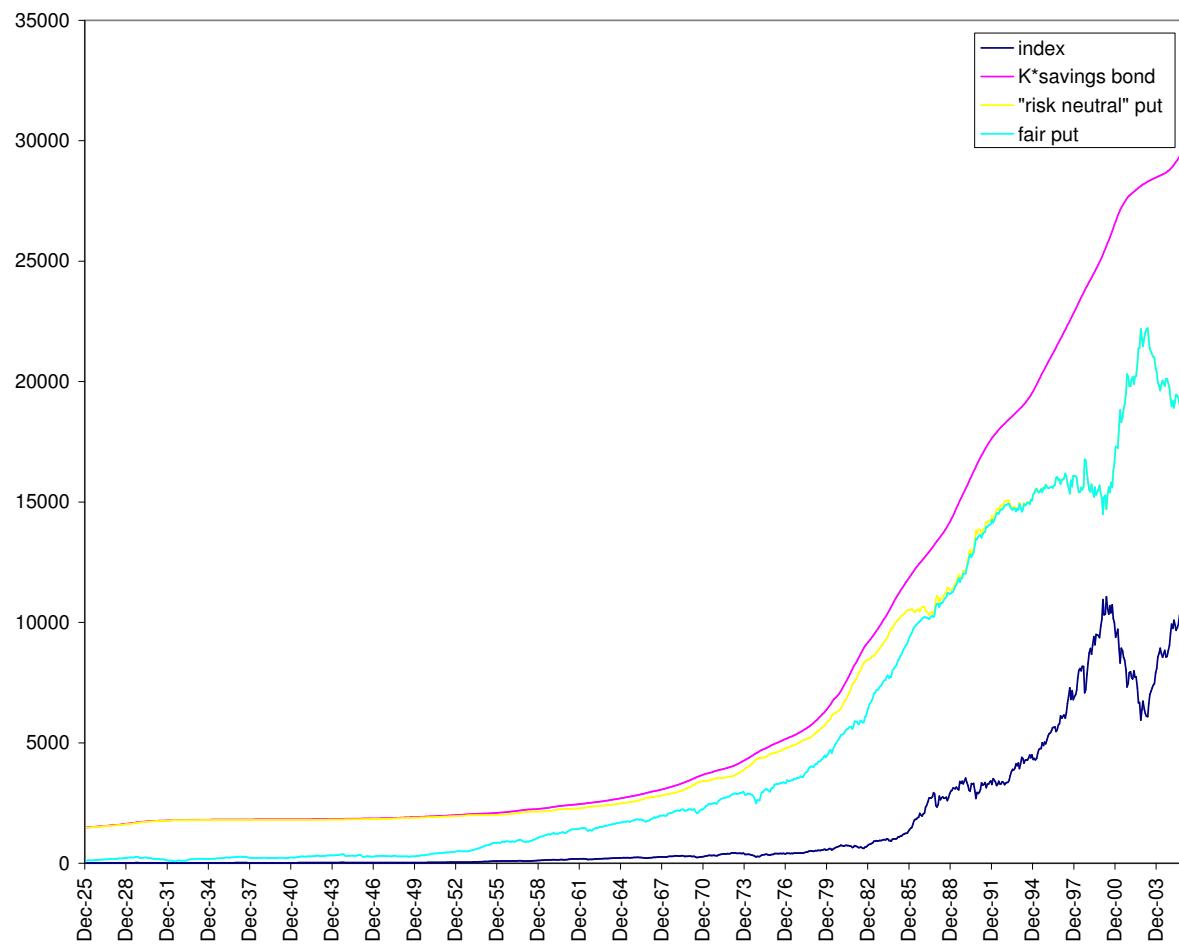
$$\lim_{\bar{S}_t^{\delta_*} \rightarrow 0} p_{T,K}(t, S_t^{\delta_*}) \stackrel{\text{a.s.}}{=} 0$$

- asymptotic hypothetical risk neutral put

hypothetical risk neutral put-call parity \implies

$$\begin{aligned}
 & \lim_{\bar{S}_t^{\delta_*} \rightarrow 0} p_{T,K}^*(t, S_t^{\delta_*}) \\
 & \stackrel{\text{a.s.}}{=} \lim_{\bar{S}_t^{\delta_*} \rightarrow 0} \left(p_{T,K}(t, S_t^{\delta_*}) + K \frac{S_t^0}{S_T^0} \exp \left\{ -\frac{\bar{S}_t^{\delta_*}}{2(\varphi_T - \varphi_t)} \right\} \right) \\
 & \stackrel{\text{a.s.}}{=} K \frac{S_t^0}{S_T^0} > 0
 \end{aligned}$$

dramatic differences can arise



“Risk neutral” and fair put on index

Pricing Annuities

Baldeaux and Pl. (2012)

Interest Indexed Payouts

$$\bar{Q}(t, T) = \bar{S}_t^* E_t \left(\frac{B_T}{S_T^*} \right)$$

$$\bar{S}_t^* = \frac{S_t^*}{B_t}$$

B_t - savings account

e.g.: Minimal Market Model (MMM), Pl. (2001)

\implies

$$\bar{Q}(t, T) = 1 - \exp \left\{ -\frac{2 \eta \bar{S}_t^*}{\alpha (\exp\{\eta T\} - \exp\{\eta t\})} \right\}$$

Interest Indexed Life Annuity

θ -entitlement level

\bar{T} -retirement date

$T_0 < T_1 < T_2 < \dots$

ξ_x - remaining life time of individual aged x at time 0

$$\begin{aligned}\bar{U}_{x,\bar{T}}^\theta(t) &= \bar{S}_t^* E_t \left(\sum_{T_i \geq \bar{T}} \mathbb{I}_{\{\xi_x > T_i\}} \frac{\theta B_{T_i}}{S_{T_i}^*} \right) \\ &= \theta \sum_{T_i \geq \bar{T}} P_t (\xi_x > T_i) \bar{Q}(t, T_i)\end{aligned}$$

Threshold Life Table

$$P(\xi_x > T) = \begin{cases} \exp \left\{ -\frac{b}{\ln(C)} (C^{x+T} - 1) \right\} & \text{for } T \leq N - x \\ q \left(1 + \lambda \left(\frac{x+T-N}{\omega} \right) \right)^{-\frac{1}{\lambda}} & \text{for } T > N - x \end{cases}$$

?)

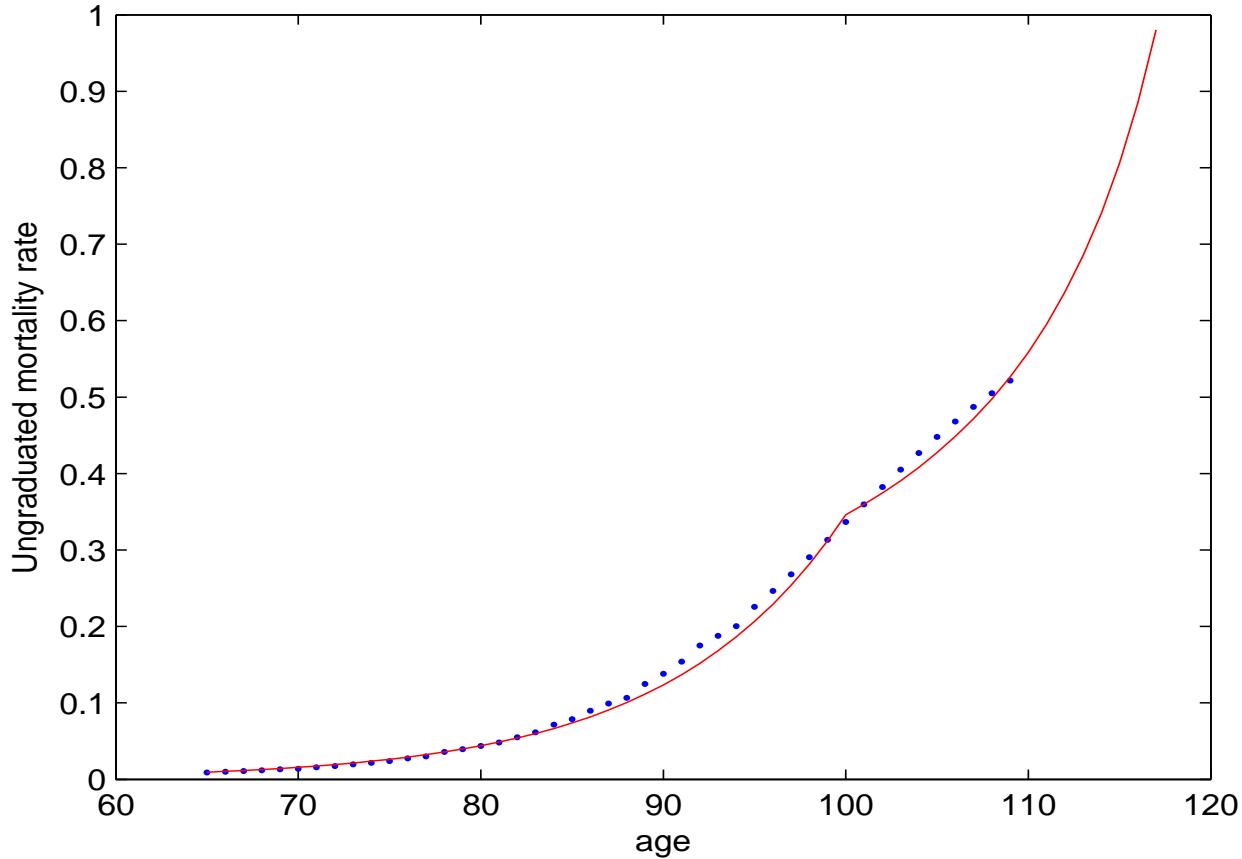


Figure 1: Mortality rates fitted to mortality data of the US population in 2007 with threshold at $N = 93$.

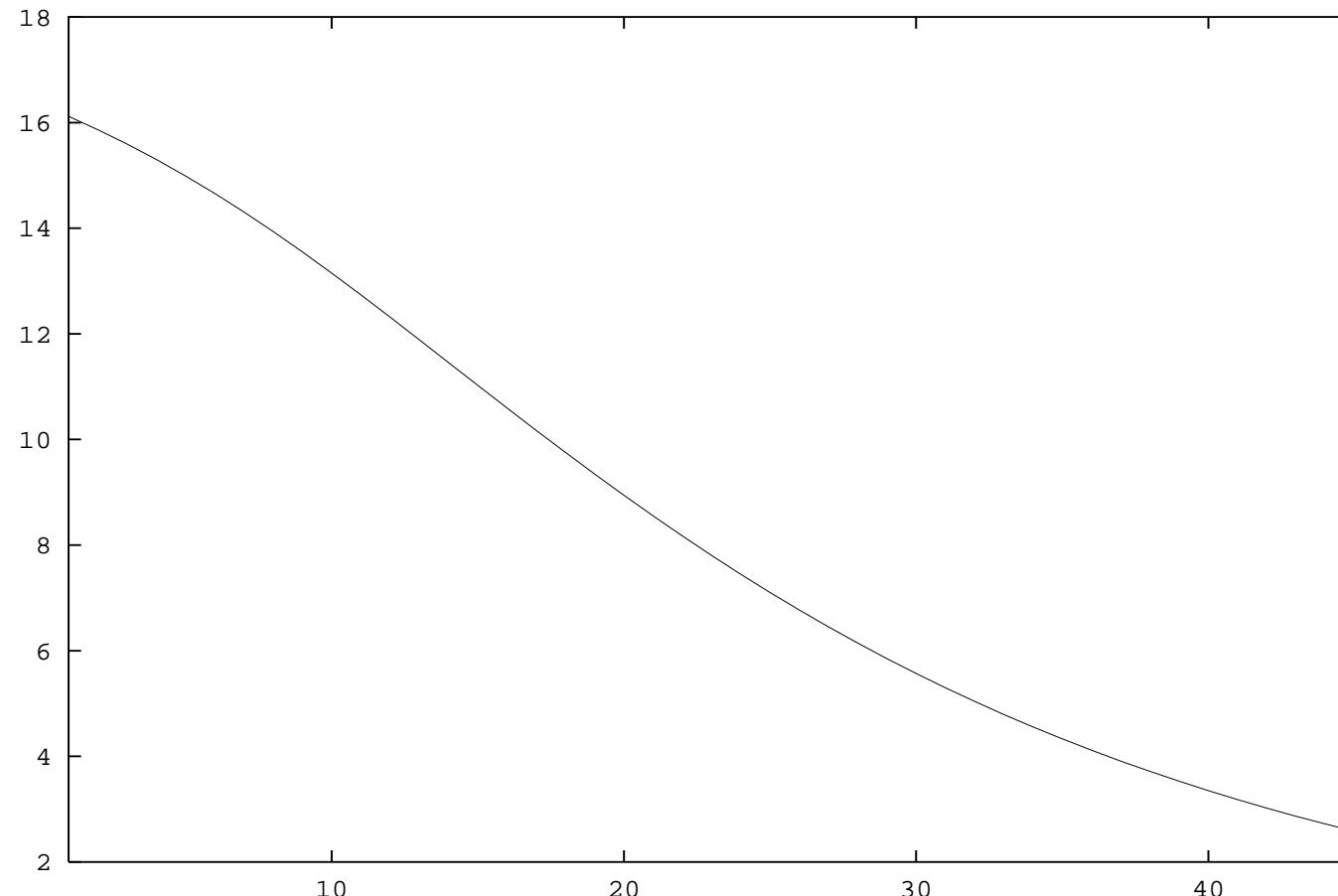


Figure 2: Discounted fair interest indexed life annuity as function of time to retirement.

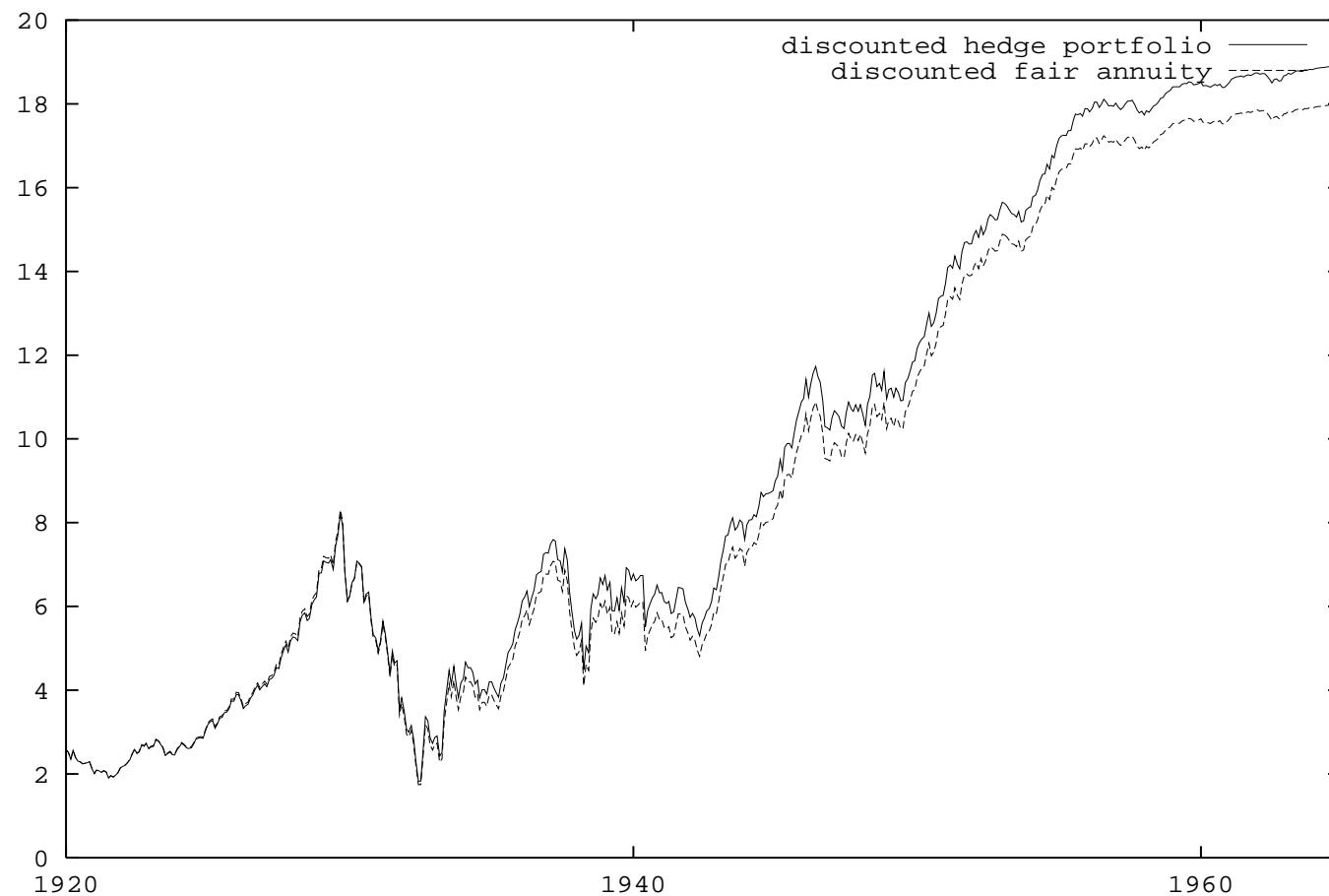


Figure 3: Discounted hedging portfolio and discounted fair interest indexed life annuity evolving over the years.

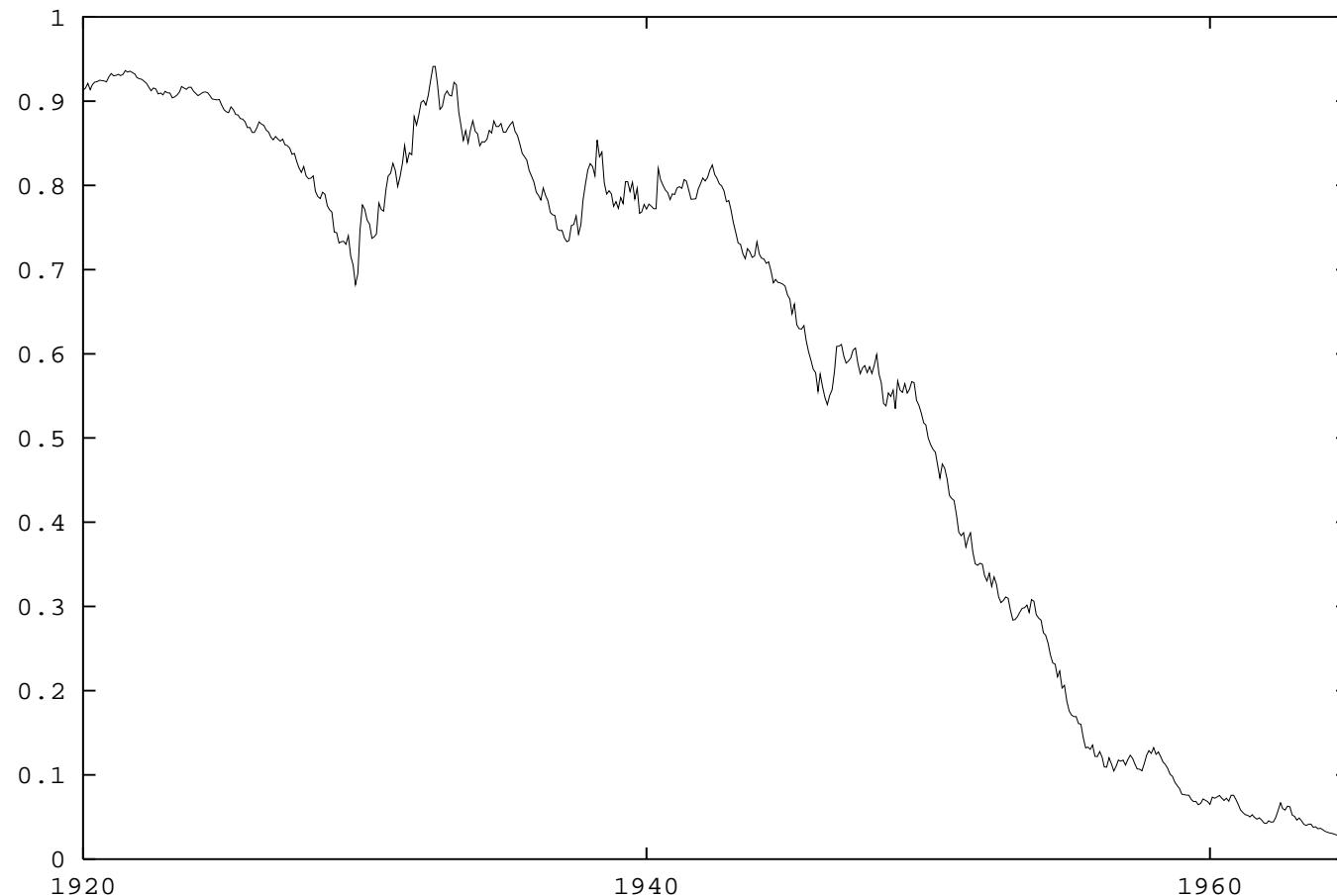


Figure 4: Fraction of the hedging portfolio invested in the benchmark portfolio.

Targeted Pension

$\Gamma_l(t)$ -account of l th member in units of interest indexed life annuity

- **total discounted value**

$$\sum_{l=1}^{M_{t_i}} \Gamma_l(t_i) \bar{U}_{x_l, \bar{T}_l}^{\theta_{t_i}}(t_i) = \delta_{\text{total}}^0(t_i) + \delta_{\text{total}}^*(t_i) \bar{S}_{t_i}^*$$

\implies entitlement level

$$\theta_{t_i} = \frac{\delta_{\text{total}}^0(t_i) + \delta_{\text{total}}^* \bar{S}_{t_i}^*}{\sum_{l=1}^{M_i} \Gamma_l(t_i) \sum_{T_k \geq \bar{T}_l} P_{t_i}(\xi_{x_l} > T_k) \bar{Q}(t_i, T_k)}$$

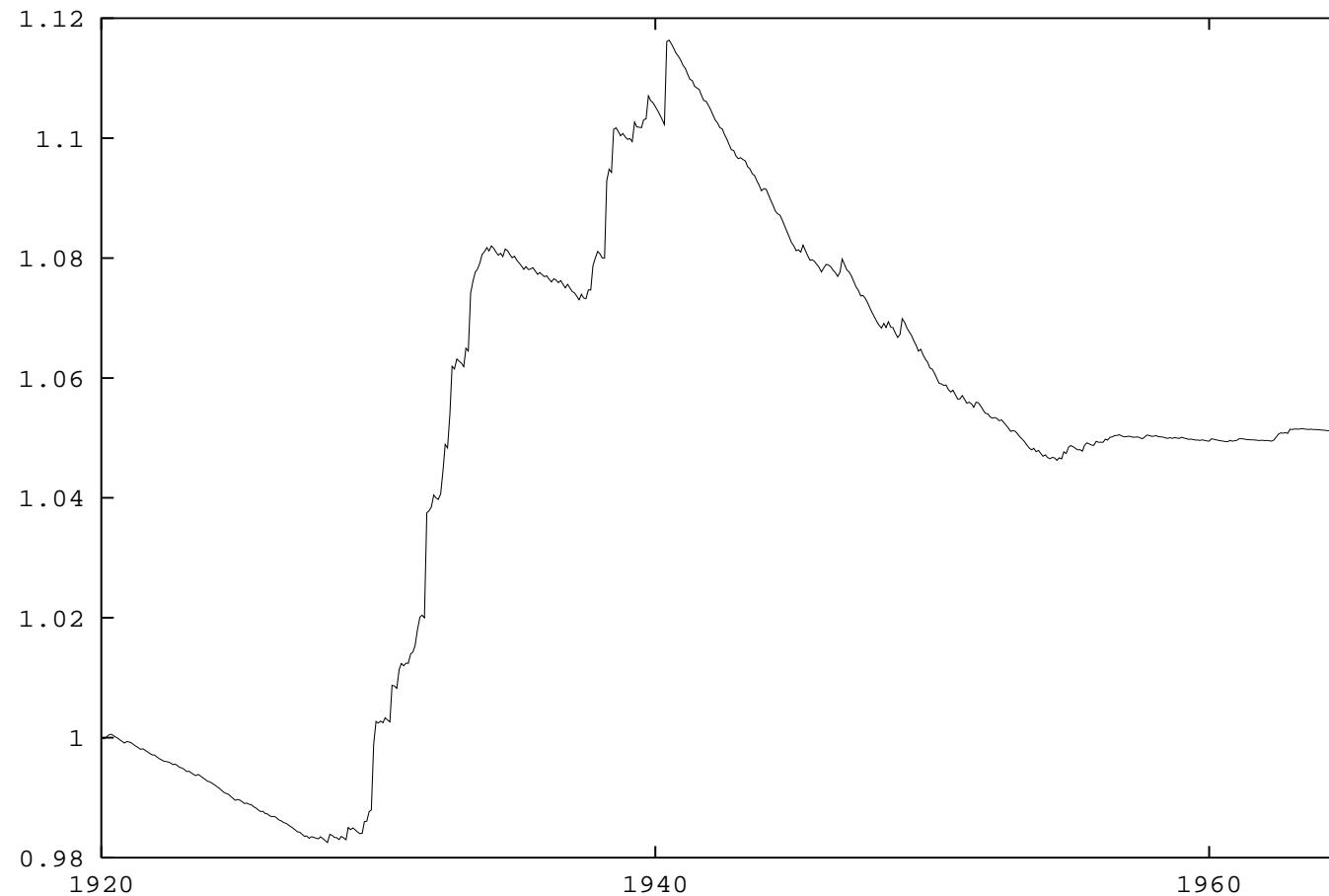


Figure 5: Entitlement level over the years.

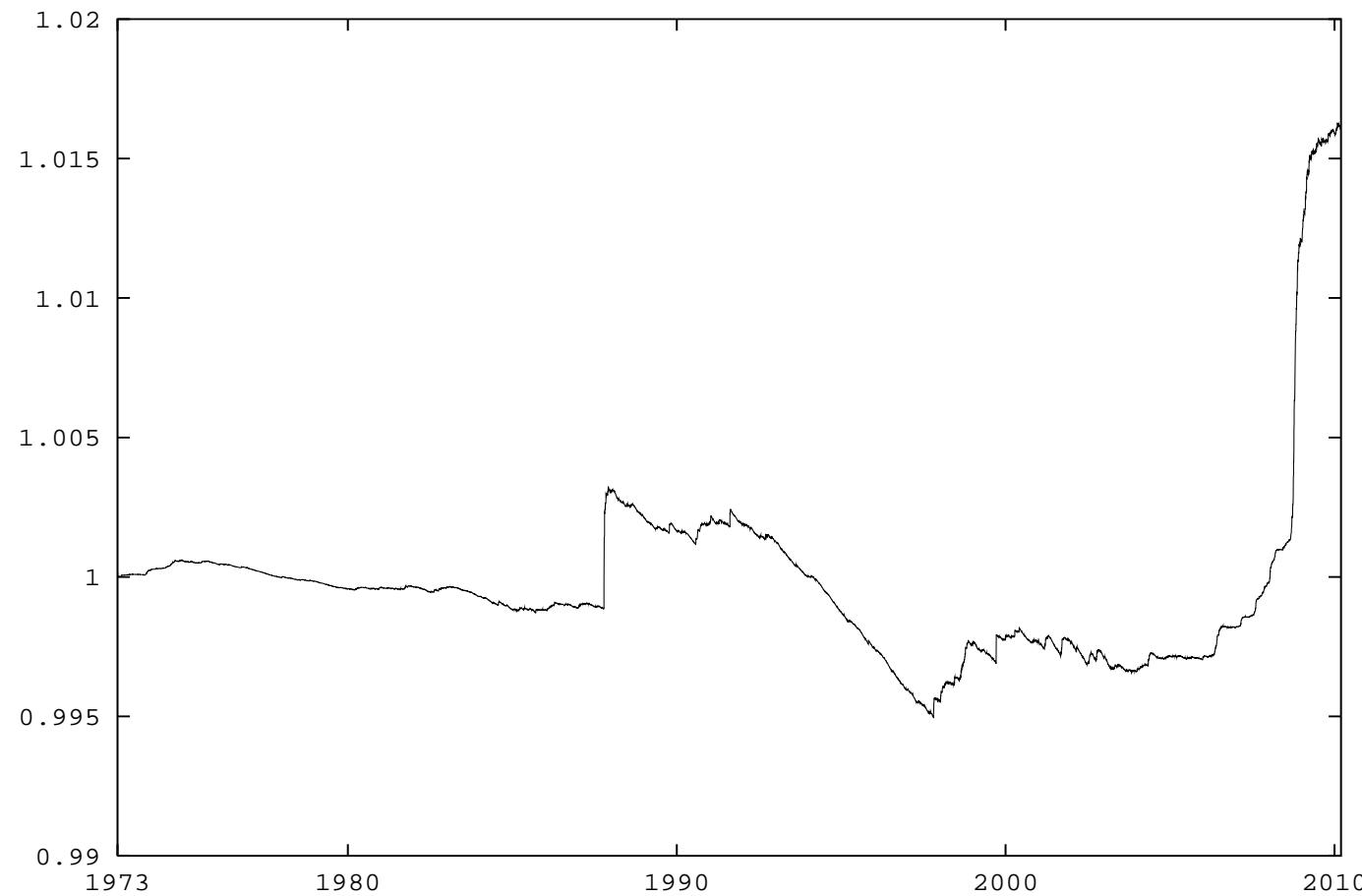


Figure 6: Entitlement level using fast growing portfolio and daily hedging.

Variable Annuity

Guaranteed Minimum Death Benefits

Marquardt, Pl. & Jaschke (2008)

- payout to the policyholder

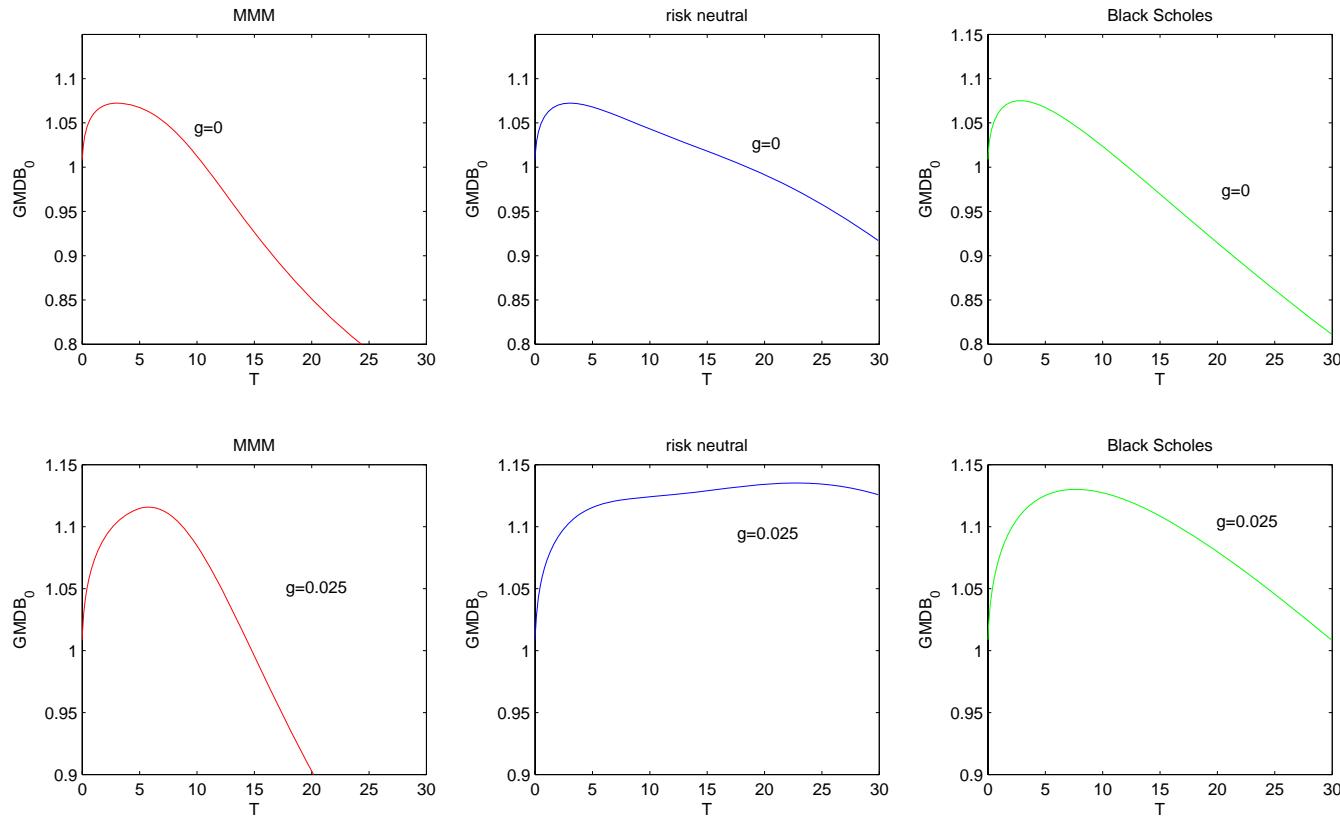
$$\max(e^{g\tau} V_0, V_\tau)$$

τ time of death

$g \geq 0$ is the guaranteed instantaneous growth rate

V_0 is the initial account value

V_τ is the unit value of the policyholder's account at time of death τ
embedded put option



Present value of the GMDB under the real world pricing formula (left), the risk neutral pricing formula (middle) and the Black Scholes formula (right) for $\eta = 0.05$, $\alpha_0 = 0.05$, $r = 0.05$, $\xi = 0.01$ and $Y_0 = 20$.

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