



# Modern Approaches to Stochastic Volatility Calibration

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# Plan

- ▶ Generic method for volatility calibration
  - ▶ Markovian projection (MP)
  - ▶ Parameter averaging (PA)
- ▶ Examples
  - ▶ Options on baskets in local volatility models
  - ▶ Options on spreads in multi-stochastic volatility models
  - ▶ Short rate models
  - ▶ Forward Libor models
  - ▶ Long-dated FX

## Calibration

- ▶ Need fast methods for European options for calibration
- ▶ A number of SV models for interest rates and hybrids have been put forward recently, with various approaches to calibration
- ▶ Many of these approaches can be aggregated into what we call the Markovian Projection method:

a generic, powerful framework for deriving closed-form approximations to European option prices

**Step 1** Apply Markovian projection to  $S(\cdot)$ , a technique to replace a complicated process with a simple one, preserving European option prices

**Step 2** Approximate conditional expected values required

**Step 3** Apply parameter averaging techniques to obtain time-independent coefficients from time-dependent

**Step 4** Hopefully a simple model is obtained, use known results.

# The Markovian projection

Theorem (Dupire 97, Gyongy 86) Let  $X(t)$  be given by

$$dX(t) = \alpha(t) dt + \beta(t) dW(t), \quad (1)$$

where  $\alpha(\cdot)$ ,  $\beta(\cdot)$  are adapted bounded stochastic processes such that (1) admits a unique solution.

Define  $a(t, x)$ ,  $b(t, x)$  by

$$\begin{aligned} a(t, x) &= E(\alpha(t) | X(t) = x), \\ b^2(t, x) &= E(\beta^2(t) | X(t) = x), \end{aligned}$$

Then the SDE

$$\begin{aligned} dY(t) &= a(t, Y(t)) dt + b(t, Y(t)) dW(t), \\ Y(0) &= X(0), \end{aligned} \quad (2)$$

admits a weak solution  $Y(t)$  that has the same one-dimensional distributions as  $X(t)$ .

► See [Dup97], [Gyö86]

## The Markovian projection, cont

- Remark 1** Since  $X(\cdot)$  and  $Y(\cdot)$  have the same one-dimensional distributions, the prices of European options on  $X(\cdot)$  and  $Y(\cdot)$  for all strikes  $K$  and expiries  $T$  will be the same. Thus, for the purposes of European option valuation and/or calibration to European options, we can replace a potentially very complicated process  $X(\cdot)$  with a much simpler Markov process  $Y(\cdot)$ , which we call the Markovian projection of  $X(\cdot)$ .
- Remark 2** The process  $Y(\cdot)$  follows what is known as a “local volatility” process. The function  $b(t, x)$  is often called “Dupire’s local volatility”

## The Markovian projection, cont

- ▶ If  $X(\cdot)$  itself came from a local volatility model (perhaps complicated), then replacing it with a (simpler) local vol model is probably the right thing to do. But:
- ▶ Any process (including a stochastic volatility one) can be replaced by a local volatility process for the purposes of European option valuation. Is it a good idea?
- ▶ Requires calculations of conditional expected values. This is the hard bit. Approximations often necessary
- ▶ In approximations, better to replace “like for like”. Replace a (complicated) SV model with a (simpler) SV model.
  - ▶ Approximations to conditional expected values may be simpler
  - ▶ Errors of approximations will tend to “cancel out”
- ▶ Dupire-Gyongy theorem still works

**Corollary** If two processes have the same Dupire’s local volatility, the European option prices on both are the same for all strikes and expiries

# The Markovian projection for SV

- ▶ Let  $X_1(t)$  follow

$$dX_1(t) = b_1(t, X_1(t)) \sqrt{\zeta_1(t)} dW(t),$$

where  $\zeta_1(t)$  is some variance process.

- ▶ We would like to match the European option prices on  $X_1(\cdot)$  (for all expiries and strikes) in a model of the form

$$dX_2(t) = b_2(t, X_2(t)) \sqrt{\zeta_2(t)} dW(t),$$

where  $\zeta_2(t)$  is a different, and potentially simpler, variance process.

- ▶ Then the Corollary and the Theorem imply that we need to set

$$b_2^2(t, x) = b_1^2(t, x) \frac{E(\zeta_1(t) | X_1(t) = x)}{E(\zeta_2(t) | X_2(t) = x)}. \quad (3)$$

- ▶ Error cancellation – whatever approximations are used for conditional expected values in (3), hopefully they will tend to cancel when we take the ratio

## Simple SV model

- ▶ After applying the MP method, often get the SDEs of the form

$$\begin{aligned} dz(t) &= \theta(1 - z(t)) dt + \gamma(t) \sqrt{z(t)} dV(t), \\ dS(t) &= (\beta(t)S(t) + (1 - \beta(t))S(0)) \sigma(t) \sqrt{z(t)} dW(t), \end{aligned} \quad (4)$$

- ▶ Or, rather, we apply the MP method with the goal of obtaining the SDEs in this form
  - ▶ Choose  $z$  to be the square root process
  - ▶ Linearize the volatility term of  $S$
- ▶ Why? When parameters are constant, this is the (shifted) Heston model, a model with very efficient numerical methods for European option valuation, see [AA02].
- ▶ How to replace time-dependent parameters with constant? Parameter averaging. Proofs and details in [Pit05b], [Pit05a]



## Example of averaging formula

- ▶ For motivation, consider a log-normal model with time-dependent volatility,

$$dS(t) = \sigma(t) S(t) dW(t).$$

- ▶ It is known that, an option value with expiry  $T_n$  in this model is equal to the Black-Scholes option value with “effective” volatility

$$\sigma_n = \left( \frac{1}{T_n} \int_0^{T_n} \sigma^2(t) dt \right)^{1/2}.$$

- ▶ Calibration by solving the following equations

$$\int_0^{T_n} \sigma^2(t) dt = \sigma_n^2 T_n, \quad n = 1, \dots, N.$$

Linear in  $\sigma^2(t)$ , trivial to solve.

- ▶ Direct link between “model” parameter  $\sigma(t)$  and “market” parameters ( $\sigma_n$ )

## Averaging volatility of variance

- ▶  $\int_0^T \sigma^2(t) z(t) dt$  is "realized variance"
- ▶ Curvature of the smile depends on the variance of realized variance (kurtosis, 4-th moment)
- ▶ Averaged vol of variance  $\eta$  (to  $T$ ) is obtained by solving

$$E \left( \int_0^T \sigma^2(t) z(t) dt \right)^2 = E \left( \int_0^T \sigma^2(t) \bar{z}(t) dt \right)^2,$$

where

$$\begin{aligned} dz(t) &= \theta(1 - z(t)) dt + \gamma(t) \sqrt{z(t)} dV(t), \\ d\bar{z}(t) &= \theta(1 - \bar{z}(t)) dt + \eta \sqrt{\bar{z}(t)} dV(t). \end{aligned}$$

# Averaging skew

- ▶ Fixed  $T$ , vol of variance already averaged (use constant  $\eta$ )

- ▶ Time-dependent skew

$$dS(t) = \sigma(t) (\beta(t) S(t) + (1 - \beta(t)) S(0)) \sqrt{z(t)} dW(t),$$

- ▶ Constant skew

$$d\bar{S}(t) = \sigma(t) (b\bar{S}(t) + (1 - b)\bar{S}(0)) \sqrt{z(t)} dW(t).$$

- ▶ Given  $\beta(\cdot)$ , find  $b$  such that option prices for different strikes (same expiry  $T$ ) are matched between two models

## Averaging skew, cont

- ▶ The main result. In the “small skew” limit,

$$b = \int_0^T \beta(t) w(t) dt,$$

where

$$w(t) = \frac{v^2(t) \sigma^2(t)}{\int_0^T v^2(t) \sigma^2(t) dt},$$

$$v^2(t) = E\left(z(t) (S_0(t) - x_0)^2\right).$$

- ▶ Comments:

- ▶ “Total skew”  $b$  is the average of “local skews”  $\beta(t)$  with weights  $w(t)$
- ▶ Weights proportional to total variance, i.e. local slope further away matters more
- ▶ Example: No SV ( $\eta = 0$ ), constant volatility  $\sigma(t) \equiv \sigma$ ,

$$b = (T^2/2)^{-1} \int_0^T t\beta(t) dt.$$

## Averaging volatility

- ▶ Approximate the dynamics of

$$dS(t) = \sigma(t) (bS(t) + (1-b)S(0)) \sqrt{z(t)} dW(t)$$

with

$$d\bar{S}(t) = \lambda (b\bar{S}(t) + (1-b)\bar{S}(0)) \sqrt{z(t)} dW(t).$$

- ▶ Can do numerically as in [Lew00], [AA02]: Do Fourier integral with integrand a solution to Riccati ODEs. Slow.
- ▶ Can use moment-matching

$$E(S(T) - S_0)^2 = E(\bar{S}(T) - S_0)^2, \quad \int_0^T \sigma^2(t) dt = \lambda^2 T.$$

Not always accurate

- ▶ Better: approximate a European option payoff locally with a function whose expectation can be computed in both models above; choose  $\lambda$  to match the two.

## Averaging volatility, cont

- ▶ By conditioning on the realized variance

$$\mathbb{E} (S(T) - S_0)^+ = \mathbb{E} g \left( \int_0^T \sigma^2(t) z(t) dt \right),$$

where  $g$  is a known function.

- ▶ Approximate

$$g(x) \approx a + be^{-cx}$$

by matching the value and first two derivatives at

$$\zeta = \mathbb{E} \int_0^T \sigma^2(t) z(t) dt$$

- ▶ The problem reduced to finding  $\lambda$  such that

$$\mathbb{E} \exp \left( \frac{g''(\zeta)}{g'(\zeta)} \int_0^T \sigma^2(t) z(t) dt \right) = \mathbb{E} \exp \left( \lambda^2 \frac{g''(\zeta)}{g'(\zeta)} \int_0^T z(t) dt \right).$$

- ▶ Very fast and easy numerical search for  $\lambda$  (starting with a good initial guess  $\lambda^2 = T^{-1} \int_0^T \sigma^2(t) dt$ ).

## Direct calibration to market

- ▶ In equity/FX: Let  $\sigma_{\text{mkt}}(T, K)$  be market volatilities for all expiries  $T$  and strikes  $K$  (assumed known). Given an exogenous SV process  $z(t)$ , find  $b(t, x)$  such that the model

$$dS(t) = b(t, S(t)) \sqrt{z(t)} dW(t), \quad S(0) = S_0,$$

matches the market

- ▶ Define Dupire's market local volatility  $b_{\text{mkt}}(t, x)$  by the requirement that the local volatility model with  $b_{\text{mkt}}(t, x)$  matches the whole market. Easy to compute

$$b_{\text{mkt}}(t, x) = \frac{2\partial C/\partial t}{\partial^2 C/\partial x^2}.$$

- ▶ Then, from Theorem and Corollary,

$$b^2(t, x) = \frac{b_{\text{mkt}}^2(t, x)}{E(z(t) | S(t) = x)}. \quad (5)$$

- ▶ In practice  $E(z(t) | S(t) = x)$  is often computed numerically in a forward PDE in  $(S, z)$ . Slow and noisy.

## Direct calibration to market, cont

- ▶ Define a “proxy” process  $X(t)$  by

$$dX(t) = \tilde{b}(t, X(t)) \sqrt{z(t)} dW(t), \quad X(0) = S_0, \quad (6)$$

where  $\tilde{b}(t, x)$  is such that European options on  $X$  are easy to compute

- ▶ Define the “proxy” Dupire’s local volatility  $b_{\text{proxy}}(t, x)$  as before but for European options on  $X$  (not on market). Then

$$E(z(t) | X(t) = x) = \frac{b_{\text{proxy}}^2(t, x)}{\tilde{b}^2(t, x)}, \quad (7)$$

thus having a stochastic volatility model with cheaply-computable European option prices allows us to compute the conditional expected values easily.

- ▶ Combining the two results we get

$$b(t, x) = \tilde{b}(t, x) \times \frac{b_{\text{mkt}}(t, x)}{b_{\text{proxy}}(t, x)} \times \left( \frac{E(z(t) | X(t) = x)}{E(z(t) | S(t) = x)} \right)^{1/2}.$$



## Direct calibration to market, cont

- ▶ Choice 1: Approximate

$$E(z(t) | X(t) = x) = E(z(t) | S(t) = x)$$

- ▶ Choice 2: Link  $S(t)$  and  $X(t)$ .

- ▶ Define  $H(t, s)$  by the requirement that  $H(t, S(t))$  has the same  $dW$  term as  $dX$  ( $H$  a function of  $b, \tilde{b}$ )

- ▶ Then approximate

$$X(t) \approx H(t, S(t)),$$

$$E(z(t) | S(t) = x) \approx E(z(t) | X(t) = H(t, x)).$$

- ▶ Functional equation on  $b$ ,

$$b(t, x) = \tilde{b}(t, H(t, x)) \frac{b_{\text{mkt}}(t, x)}{b_{\text{proxy}}(t, H(t, x))}. \quad (8)$$

- ▶ Last derivation is an example of a clever way of computing conditional expectations
- ▶ Original result due to Forde ([For06]). More details in [Pit06a].

## Basket modeling

- ▶ Consider a “simple” local volatility model for a basket

$$S(t) = \sum w_i S_i(t),$$

$$dS_i(t) = \varphi_i(S_i(t)) dW_i(t), \quad i = 1, \dots, I.$$

- ▶ Options on index  $S(\cdot)$ . Apply MP to write SDE for  $S$ . Start

$$dS(t) = \sigma(t) dW(t),$$

$$\sigma^2(t) = \sum_{n,m=1}^N w_n w_m \varphi_n(S_n(t)) \varphi_m(S_m(t)) \rho_{nm}.$$

- ▶ Then

$$\begin{aligned} dS(t) &= \varphi(t, S(t)) dW(t), \\ \varphi^2(t, x) &= E(\sigma^2(t) | S(t) = x). \end{aligned}$$

## Basket modeling, cont

- ▶ To compute  $E(\sigma^2(t) | S(t))$  use Gaussian approximation  $S_i \approx \bar{S}_i$ ,  $S \approx \bar{S}$ ,

$$d\bar{S}_i(t) = p_i dW_i(t), \quad d\bar{S}(t) = \sigma(0) dW(t),$$

$$p_i = \varphi_i(S_i(0)), \quad \sigma(0) = \sum_{n,m=1}^N w_n w_m p_n p_m \rho_{nm},$$

and linearization

$$\varphi_i(x) \approx p_i + q_i(x - S(0)), \quad q_i = \varphi_i'(S_i(0)).$$

- ▶ Then

$$E(\bar{S}_n(t) - S_n(0) | \bar{S}(t) = x) = \rho_n \frac{p_n}{p} (x - S(0)),$$

$$\rho_n \triangleq \langle d\bar{W}(t), dW_n(t) \rangle / dt = \frac{1}{p} \sum_{m=1}^N w_m p_m \rho_{nm}.$$

- ▶ See more in [Pit06a]. More accurate method in [ABOBF02].

## Spread options in SV model

- ▶ For spread options, important to use different SV process for each variable (see [Pit06c]),

$$\begin{aligned}dS_i(t) &= \varphi_i(S_i(t)) \sqrt{z_i(t)} dW_i(t), \quad i = 1, 2, \\dz_i(t) &= \theta(1 - z_i(t)) dt + \eta_i \sqrt{z_i(t)} dW_{2+i}(t), \quad z_i(0) = 1,\end{aligned}$$

with the correlations given by

$$\langle dW_i(t), dW_j(t) \rangle = \rho_{ij} \quad i, j = 1, \dots, 4.$$

- ▶ Denote

$$p_i = \varphi_i(S_i(0)), \quad q_i = \varphi'_i(S_i(0)).$$

- ▶ Write down  $dS(\cdot)$  for spread  $S = S_1 - S_2$
- ▶ Identify a suitable “spread variance” process  $z(\cdot)$
- ▶ Compute the skew function  $\varphi(\cdot)$  of the spread using the Markovian projection ideas above
- ▶ “Massage”  $z(\cdot)$  into the Heston form

## Process for the spread

- ▶ We have

$$dS_i(t) = \varphi_i(S_i(t)) \sqrt{z_i(t)} dW_i(t),$$

- ▶  $S = S_1 - S_2$ , then  $dS(t) = \sigma(t) dW(t)$ , where

$$\begin{aligned} \sigma^2(t) &= (\varphi_1(S_1(t)) u_1(t))^2 \\ &\quad - 2(\varphi_1(S_1(t)) u_1(t)) (\varphi_2(S_2(t)) u_2(t)) \rho_{12} \\ &\quad + (\varphi_2(S_2(t)) u_2(t))^2, \end{aligned}$$

$$\begin{aligned} dW(t) &= \frac{1}{\sigma(t)} (\varphi_1(S_1(t)) u_1(t) dW_1(t) \\ &\quad - \varphi_2(S_2(t)) u_2(t) dW_2(t)), \end{aligned}$$

$$u_i(t) = \sqrt{z_i(t)}, \quad i = 1, 2.$$

## Process for the variance of the spread

- ▶ Try to find a stochastic volatility process  $z(\cdot)$  such that the curvature of the smile of the spread  $S(\cdot)$  is explained by it, and the local volatility function is only used to induce the volatility skew
- ▶ To identify a suitable candidate for  $z(\cdot)$ , consider what the expression for  $\sigma^2(t)$  would be if  $\varphi_i(x)$ ,  $i = 1, 2$ , were constant functions.
- ▶ In this case, the expression for  $\sigma^2(t)$  above would not involve the processes  $S_i(\cdot)$ ,  $i = 1, 2$  and this is a good candidate for the stochastic variance process.
- ▶ We define (the division by  $\sigma^2(0)$  is to preserve the scaling  $z(0) = 1$ )

$$z(t) = \frac{1}{p^2} \left( (p_1 u_1(t))^2 - 2p_1 p_2 u_1(t) u_2(t) \rho_{12} + (p_2 u_2(t))^2 \right), \quad (9)$$

where

$$p = \sigma(0) = (p_1^2 - 2p_1 p_2 \rho_{12} + p_2^2)^{1/2}. \quad (10)$$

## Skew function of the spread

- ▶ By Corollary,

$$\varphi^2(t, \mathbf{x}) = \frac{E(\sigma^2(t) | S(t) = \mathbf{x})}{E(z(t) | S(t) = \mathbf{x})}. \quad (11)$$

- ▶ The expression for  $E(\sigma^2(t) | S(t) = \mathbf{x})$  is a linear combinations of the conditional expected values of the terms

$$\varphi_i(S_i(t)) \varphi_j(S_j(t)) u_i(t) u_j(t),$$

- ▶ Approximate to the first order by

$$p_i p_j \left( 1 + \frac{q_i}{p_i} (S_i(t) - S_i(0)) + \frac{q_j}{p_j} (S_j(t) - S_j(0)) + \dots \right).$$

- ▶ Use Gaussian approximation to compute conditional expected values

## Gaussian approximation

- ▶ Use  $\bar{X}$  to denote a Gaussian approximation to  $X$  for a generic  $X$ , then

$$\begin{aligned}E(S_i(t) - S_i(0) | S(t) = x) &\approx E(\bar{S}_i(t) - \bar{S}_i(0) | \bar{S}(t) = x) \\E(u_i(t) - 1 | S(t) = x) &\approx E(\bar{u}_i(t) - 1 | \bar{S}(t) = x),\end{aligned}$$

- ▶ Here (we ignore  $dt$  terms for  $du$ , although they may be included for more accurate approximations)

$$\begin{aligned}d\bar{S}_i(t) &= p_i dW_i(t), & d\bar{S}(t) &= p d\bar{W}(t), \\d\bar{u}_i(t) &= \frac{\eta_i}{2} dW_{2+i}(t), & d\bar{W}(t) &= \frac{1}{p} (p_1 dW_1(t) - p_2 dW_2(t)).\end{aligned}\tag{12}$$

- ▶ Then

$$\begin{aligned}E(\bar{S}_i(t) - \bar{S}_i(0) | \bar{S}(t) = x) &= \frac{p_i \rho_i}{p} (x - S(0)), \\E(\bar{u}_i(t) - 1 | \bar{S}(t) = x) &= \frac{\eta_i \rho_{2+i}}{2p} (x - S(0)),\end{aligned}$$



## Skew function of the spread

- ▶ Combining the results, we get the following approximation to the spread dynamics,

$$dS(t) = \varphi(S(t)) \sqrt{z(t)} dW(t),$$

- ▶ Here  $\varphi(x)$  is a function of the same type as  $\varphi_i(x)$  (linear or CEV) with

$$\varphi(S(0)) = p, \quad \varphi'(S(0)) = q,$$

where

$$p = (p_1^2 - 2p_1p_2\rho_{12} + p_2^2)^{1/2}$$
$$q \triangleq \frac{1}{p} (p_1\rho_1^2q_1 - p_2\rho_2^2q_2).$$

## Variance process for the spread

- ▶ The process for  $S$  is in a nice form. But  $z$  is not:

$$z(t) = \frac{1}{p^2} \left( p_1^2 z_1(t) - 2p_1 p_2 \sqrt{z_1(t) z_2(t)} \rho_{12} + p_2^2 z_2(t) \right).$$

- ▶ Compute  $dz$ ,

$$\begin{aligned} dz(t) &= \delta_1(t) dt + \delta_2(t) dt + \delta_3(t) dt \\ &\quad + \xi_1(t) dW_3(t) + \xi_2(t) dW_4(t), \end{aligned}$$

- ▶  $dW$  terms

$$\begin{aligned} \xi_1(t) &= \eta_1 \frac{p_1^2}{p^2} \left( \sqrt{z_1(t)} - \frac{p_2}{p_1} \rho_{12} \sqrt{z_2(t)} \right), \\ \xi_2(t) &= \eta_2 \frac{p_2^2}{p^2} \left( \sqrt{z_2(t)} - \frac{p_1}{p_2} \rho_{12} \sqrt{z_1(t)} \right). \end{aligned}$$

## Variance process for the spread, cont

► dt terms

$$\delta_1(t) = \theta \frac{p_1^2}{p^2} \left( 1 - \frac{p_2}{p_1} \rho_{12} \sqrt{\frac{z_2(t)}{z_1(t)}} \right) (1 - z_1(t)),$$

$$\delta_2(t) = \theta \frac{p_2^2}{p^2} \left( 1 - \frac{p_1}{p_2} \rho_{12} \sqrt{\frac{z_1(t)}{z_2(t)}} \right) (1 - z_2(t)),$$

$$\delta_3(t) = \frac{p_1 p_2 \rho_{12}}{4p^2} \left( \sqrt{\frac{z_2(t)}{z_1(t)}} \eta_1^2 - 2\eta_1 \eta_2 \rho_{34} + \sqrt{\frac{z_1(t)}{z_2(t)}} \eta_2^2 \right).$$

► Complicated expression, Not “closed” in  $z(\cdot)$

## Variance process for the spread, cont

- ▶ The curvature of the volatility smile (of options on  $S(\cdot)$ ) is driven by the variance of the stochastic variance
- ▶ It is preserved under the Markovian projection of  $z(\cdot)$  so can apply the Theorem again, now to the process for  $z(\cdot)$ !
- ▶ Formulas getting unwieldy: need to compute conditional expected values of the type  $E(\sqrt{z_i(t)z_j(t)} | z(t) = x)$  and  $E(\sqrt{z_i(t)/z_j(t)} | z(t) = x)$ , for which we would apply the Gaussian approximations
- ▶ Try something simpler:
  - ▶ replace  $\sqrt{z_1(t)}$ ,  $\sqrt{z_2(t)}$  in the dW terms with  $\sqrt{z(t)}$ ;
  - ▶ replace  $\sqrt{\frac{z_2(t)}{z_1(t)}}$ ,  $\sqrt{\frac{z_1(t)}{z_2(t)}}$  in dt terms with 1.
- ▶  $\delta_1(t) + \delta_2(t)$  becomes  $\theta(1 - z)$ ,

## Variance process for the spread, simple approximation

- ▶  $\delta_3(t)$  becomes

$$\gamma \triangleq \frac{p_1 p_2 \rho_{12}}{4p^2} (\eta_1^2 - 2\eta_1 \eta_2 \rho_{34} + \eta_2^2). \quad (13)$$

- ▶ The dW terms can be re-written as  $\eta \sqrt{z(t)} dB(t)$ , where

$$\eta^2 = \frac{1}{p^2} \left( (p_1 \eta_1 \rho_1)^2 - 2(p_1 \eta_1 \rho_1)(p_2 \eta_2 \rho_2) \rho_{34} + (p_2 \eta_2 \rho_2)^2 \right),$$
$$dB(t) = \frac{1}{\eta} (p_1 \eta_1 \rho_1 dW_3(t) - p_2 \eta_2 \rho_2 dW_4(t)).$$

- ▶ Altogether

$$dS(t) = \varphi(S(t)) \sqrt{z(t)} dW(t),$$
$$dz(t) = \theta \left( 1 + \frac{\gamma}{\theta} - z(t) \right) dt + \eta \sqrt{z(t)} dB(t).$$

- ▶ Linearize  $\varphi$  and apply Heston valuation formula to options on the spread S!

## Local volatility short rate model

- ▶ Simplest interest rate model: one-factor Gaussian (“Hull-White”)

$$r(t) = f(0, t) + x(t), \quad dx(t) = (\theta(t) - ax(t)) dt + \sigma(t) dW(t).$$

- ▶ Local-volatility extension: quasi-Gaussian (“Cheyette”)

$$\begin{aligned} dx(t) &= (y(t) - ax(t)) dt + \sigma(t, x(t), y(t)) dW(t), \\ dy(t) &= (\sigma^2(t, x(t), y(t)) - 2ay(t)) dt. \end{aligned}$$

- ▶ Swap rate (under swap measure),  $S(t) = S(t, x(t), y(t))$  for a known function  $S(t, x, y)$ ,

$$dS(t) = \left. \frac{\partial S(t, x, y)}{\partial x} \right|_{x=x(t), y=y(t)} \sigma(t, x(t), y(t)) dW^A(t).$$

## Local volatility short rate model, cont

- ▶ Markovian projection (preserves European swaptions)

$$dS(t) = \eta(t, S(t)) dW^A(t),$$

$$\eta^2(t, S) = E^A \left( \left( \frac{\partial S(t, x(t), y(t))}{\partial x} \right)^2 \sigma^2(t, x(t), y(t)) \middle| S(t) = S \right)$$

- ▶ Let  $y^*(t) = E^A(y(t))$ ,  $\xi(t, s)$  is the inverse of  $S(t, x, y^*(t))$  in  $x$ . Then

$$\eta^2(t, S) \approx \left( \frac{\partial S(t, x, y^*(t))}{\partial x} \bigg|_{x=\xi(t, S)} \right) \sigma(t, \xi(t, S), y^*(t)).$$

- ▶ Local-volatility model for  $S$  with a known  $\eta$ . Apply parameter averaging (on skew and vol), then shifted-lognormal formula to get option prices.

## Stochastic volatility short rate model

- ▶ Stochastic-volatility extension: quasi-Gaussian SV

$$\begin{aligned}dx(t) &= (y(t) - ax(t)) dt + \sqrt{z(t)}\sigma(t, x(t), y(t)) dW(t), \\dy(t) &= (z(t)\sigma^2(t, x(t), y(t)) - 2ay(t)) dt, \\dz(t) &= \theta(1 - z(t)) dt + \gamma(t)\sqrt{z(t)}dV(t).\end{aligned}$$

- ▶ Same results (use the same  $z(\cdot)$  in (3)), after MP:

$$\begin{aligned}dS(t) &= \eta(t, S(t))\sqrt{z(t)}dW^A(t), \\dz(t) &= \theta(1 - z(t)) dt + \gamma(t)\sqrt{z(t)}dV(t).\end{aligned}$$

- ▶ Linearize  $\eta(t, S)$ , apply PA on skew, vol, vol of vol.
- ▶ See [And05]



# Forward Libor model with time-dependent skews

- ▶  $L_n(t)$  are spanning forward Libor rates

$$dL_n(t) = \psi_n(t, L_n(t)) dW_n^{T_{n+1}}(t), \quad n = 1, \dots, N-1.$$

- ▶ Swap rate ( $S = S_{n,m}$ ) dynamics

$$\begin{aligned} dS(t) &= \sum_{k=n}^{n+m-1} \frac{\partial S(t)}{\partial L_k(t)} \psi_k(t, L_k(t)) dW_k^A(t) \\ &= \Sigma(t, \bar{L}(t)) dW_n^A(t), \end{aligned}$$

$$\Sigma^2(t, \bar{L}(t)) = \sum_{k,k'} \frac{\partial S(t)}{\partial L_k(t)} \frac{\partial S(t)}{\partial L_{k'}(t)} \psi_k(t, L_k(t)) \psi_{k'}(t, L_{k'}(t)) \rho_{kk'}.$$

- ▶ By MP

$$\begin{aligned} \eta(t, S) &= (E^A(\Sigma^2(t, \bar{L}(t)) | S(t) = S))^{1/2} \\ &\approx E^A(\Sigma(t, \bar{L}(t)) | S(t) = S) \end{aligned}$$

## Forward Libor model with time-dependent skews, cont

- ▶ Linearize

$$\Sigma(t, \bar{L}(t)) = \Sigma(t, E^A \bar{L}(t)) + [\nabla \Sigma(t, E^A \bar{L}(t))]^\top (\bar{L}(t) - E^A \bar{L}(t))$$

- ▶ Approximate  $\bar{L}(t), S(t)$  with Gaussian processes (use “hats”)

$$E^A(\bar{L}(t) - E^A \bar{L}(t) | S(t)) \approx \langle \hat{L}(t), \hat{L}(t) \rangle^{-1} \langle \hat{L}(t), \hat{S}(t) \rangle (S(t) - S(0))$$

- ▶ Then

$$dS(t) = (a(t) + b(t)(S - S(0))) dW_n^A(t)$$

$$a(t) = \Sigma(t, E^A \bar{L}(t)),$$

$$b = [\nabla \Sigma(t, E^A \bar{L}(t))]^\top \langle \hat{L}(t), \hat{L}(t) \rangle^{-1} \langle \hat{L}(t), \hat{S}(t) \rangle.$$

- ▶ Shifted lognormal process for  $S$  with time-dependent coeffs (skew is the weighted average of Libor skews), apply PA, and we are done.

## Forward Libor model with SV

- ▶ Use the same SV process  $z(\cdot)$  for all Libor rates

$$dL_n(t) = \psi_n(t, L_n(t)) \sqrt{z(t)} dW_n^{T_{n+1}}(t), \quad n = 1, \dots, N-1.$$

- ▶ Same results, get

$$\begin{aligned} dS(t) &= (a(t) + b(t)(S - S(0))) \sqrt{z(t)} dW_n^A(t) \\ dz(t) &= \theta(1 - z(t)) dt + \gamma(t) \sqrt{z(t)} dV(t). \end{aligned}$$

same  $a(\cdot)$ ,  $b(\cdot)$ .

- ▶ See [Pit05a]

## Interest-rate/FX hybrids

- ▶ Interest rates in two currencies + a process for FX

$$dP_d(t, T) / P_d(t, T) = r_d(t) dt + \sigma_d(t, T) dW_d(t), \quad (14)$$

$$dP_f(t, T) / P_f(t, T) = r_f(t) dt + \sigma_f(t, T) dW_f(t),$$

$$dS(t) / S(t) = (r_d(t) - r_f(t)) dt + \gamma(t, S(t)) dW_S(t),$$

- ▶ The “standard” Gaussian framework is recovered by choosing the function  $\gamma(t, x)$  that is independent of  $x$ ,  $\gamma(t, x) = \gamma(t)$ .
- ▶ FX skew via the local volatility function  $\gamma(t, x)$ .
- ▶ Skew very important for FX hybrids, eg PRDC
- ▶ Use a parametric form of the local volatility function

$$\gamma(t, x) = \nu(t) \left( \frac{x}{L(t)} \right)^{\beta(t)-1}. \quad (15)$$

- ▶  $\nu(t)$  is the relative volatility function,  $\beta(t)$  is a time-dependent constant elasticity of variance parameter and  $L(t)$  is a time-dependent scaling constant (“level”).

## Interest-rate/FX hybrids, cont

- ▶ Market – options on forward FX,  $S(T) = F(T, T)$ ,

$$F(t, T) = S(t) / D(t, T), \quad D(t, T) = P_d(t, T) / P_f(t, T).$$

- ▶ Under domestic T-forward measure,

$$\begin{aligned} dF(t, T) / F(t, T) &= \sigma_f(t, T) dW_d^T(t) - \sigma_d(t, T) dW_d^T(t) \\ &\quad + \gamma(t, F(t, T) D(t, T)) dW_S^T(t). \end{aligned} \tag{16}$$

- ▶ Single stochastic driver

$$dF(t, T) / F(t, T) = \Lambda(t, F(t, T) D(t, T)) dW_F(t), \tag{17}$$

where

$$\begin{aligned} \Lambda(t, x) &= (a(t) + b(t) \gamma(t, x) + \gamma^2(t, x))^{1/2}, \\ a(t) &= \dots, b(t) = \dots \end{aligned}$$

- ▶ If  $\gamma(t, x)$  is a function of time  $t$  only, then the  $\Lambda(t, F(t, T) D(t, T)) = \Lambda(t)$  is also a deterministic function of time, and  $F(T, T)$  is lognormal

## Interest-rate/FX hybrids, cont

- ▶ In general case – use MP:

$$\tilde{\Lambda}^2(t, x) = E_0^T (\Lambda^2(t, F(t, T) D(t, T)) | F(t, T) = x).$$

- ▶ Approximate :

$$\begin{aligned}\hat{\Lambda}(t, x) &\approx (a(t) + b(t) \hat{\gamma}(t, x) + \hat{\gamma}^2(t, x))^{1/2}, \\ \hat{\gamma}(t, x) &= \nu(t) \left( \frac{x D_0(t, T)}{L(t)} \right)^{\beta(t)-1} \\ &\quad \times \left( 1 + (\beta(t) - 1) r(t) \left( \frac{x}{F(0, T)} - 1 \right) \right),\end{aligned}$$





here  $r(t)$  is a “regression” coefficient of discount bond ratio to the forward FX.

- ▶ Local volatility model with time-dependent skew, use PA. FX forward approximately shifted-lognormal. See details in [Pit06b].

# Conclusions

- ▶ We have presented a generic method for calibrating models with smile, consisting of
  - ▶ Markovian projection, and
  - ▶ Parameter averaging
- ▶ The method can be applied to a wide variety of models: baskets, spreads, interest rate models, interest rate/FX models, interest rate/equity models, etc
- ▶ While the application of the method can be more, or less, successful depending on the technical difficulties encountered on each step, at least we have a plan of attack applicable to any model

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