

GENERALIZED BETA REGRESSION MODELS FOR RANDOM LGD

Xinzheng Huang,
TU Delft

PORTFOLIO CREDIT LOSS

Credit loss of a portfolio consisting of n obligors

$$L = \sum_{i=1}^n EAD_i \times LGD_i \times D_i$$

- Default indicator D_i , Probability-of-Default (PD) $p_i = P(D_i = 1)$.
- Exposure-at-Default (EAD): the size of a loan.
- Loss-Given-Default (LGD): proportion of the EAD that will be lost if a default occurs. $LGD \in [0, 1]$.

PORTFOLIO CREDIT LOSS

$$L = \sum_{i=1}^n EAD_i \times LGD_i \times D_i$$

- ① the dependence between defaults: latent factor model that introduces systematic risk in default
- ② the dependence between LGDs
- ③ the dependence between PD and LGD

DEFAULT MODEL: VASICEK GAUSSIAN ONE-FACTOR MODEL

- The standardized asset log-return X_i is normally distributed.
- Default occurs if $X_i < \gamma_i$ where γ_i is a known default threshold.
- X is decomposed into a systematic part Y and an idiosyncratic part Z such that for obligor i ,

$$X_i = \sqrt{\rho} Y + \sqrt{1 - \rho} Z_i, \quad (1)$$

where Y and all Z_i are i.i.d standard normal random variables.

- ρ is called the asset correlation.

LGD MODELS

- Current Industry practice
 - Dependence between LGDs, dependence between PD and LGD ignored.
 - constant LGD: Vasicek model, CreditRisk⁺.
 - beta distributed LGD: KMV Portfolio Manager, CreditMetrics.
- Empirical evidence: Hu & Perraudin (2002); Altman et al. (2005)
 - LGD is positively correlated to the default rate; LGD is high when the default rate is high.
 - LGD is subject to systematic risk.

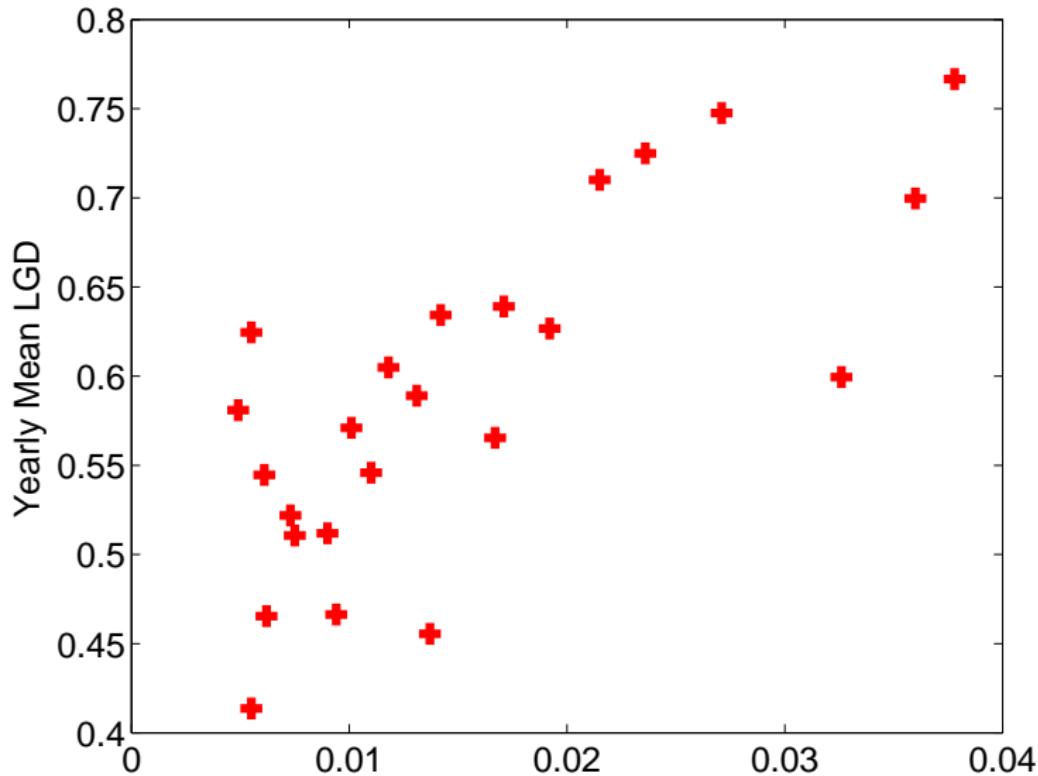


FIGURE: Yearly average default rate and mean LGD.

LGD MODELS

- Frye (2000)

$$LGD = \mu + \sigma \xi, \quad \xi = \sqrt{\bar{\rho}} Y + \sqrt{1 - \bar{\rho}} \varepsilon,$$

- Pykhtin (2003)

$$LGD = \left(1 - e^{\mu + \sigma \xi}\right)^+,$$

- Andersen & Sidenius (2004)

$$LGD = \Phi(\mu + \sigma \xi)$$

- Duellmann & Trapp (2004); Roesch & Scheule (2005)

$$LGD = \frac{1}{1 + \exp(\mu + \sigma \xi)}$$

- Giese (2006); Bruche & Gonzalez-Aguado (2008)

$$LGD \sim Beta(\alpha(Y), \beta(Y))$$

A new framework: Generalized Beta Regression Models

- Easy to interpret
- Accommodates better heteroscedastic errors
- Broad enough to fit the historical data and empirical testing of different models can be unified
- Allows efficient numerical procedures (normal approximation, saddlepoint approximation) to calculate the portfolio loss distribution.

PARAMETERIZATION OF A BETA DISTRIBUTION

- probability density function

$$f(x) = \frac{x^{\alpha-1}(1-x)^{\beta-1}}{B(\alpha, \beta)}.$$

- mean

$$\mu = E(X) = \frac{\alpha}{\alpha + \beta},$$

- variance

$$\sigma^2 = \text{Var}(X) = \frac{\alpha\beta}{(\alpha + \beta)^2(\alpha + \beta + 1)} = \frac{\mu(1 - \mu)}{\alpha + \beta + 1} = \frac{\mu(1 - \mu)}{\varphi + 1}.$$

- Dispersion parameter $\varphi = \alpha + \beta$.

BETA REGRESSION MODEL: FERRARI & CRIBARI-NETO (2004)

- Parameterization of a beta distribution: mean μ and dispersion φ .
- Mean Model (GLM)

- a linear predictor η

$$\eta = \alpha \zeta$$

where ζ is a vector of explanatory variables and α is a vector of the corresponding regression coefficients.

- a monotonic, differentiable link function g

$$g(\mu) = \eta.$$

- Dispersion parameter φ : a nuisance parameter

BETA REGRESSION MODEL

- A parsimonious one-factor model: $\zeta = [1, Y]'$, where Y is the factor also drives PD.

$$LGD \sim Beta\left(g^{-1}(a_1 + a_2 Y), \varphi\right)$$

- Choices of link function: $g^{-1}(\cdot) : \mathbb{R} \rightarrow [0, 1]$.
logit link

$$\mu = \frac{e^\eta}{1 + e^\eta}, \quad \eta = \log\left(\frac{\mu}{1 - \mu}\right),$$

the probit link

$$\mu = \Phi(\eta), \quad \eta = \Phi^{-1}(\mu).$$

cauchit, complementary log-log, scaled probit...

GENERALIZED BETA REGRESSION MODELS

- ① Jointly modeling mean and dispersion (JGLM)

$$h(\varphi) = b\zeta, \quad \varphi = e^{b\zeta}$$

$$LGD \sim Beta\left(g^{-1}(a_1 + a_2 Y), e^{b_1 + b_2 Y}\right)$$

- ② Mixed model: random effect in the linear predictor (GLMM)

$$g(\mu) = \eta = a\zeta + v, \quad v \sim N(0, \sigma_v^2)$$

$$LGD \sim Beta\left(g^{-1}(a_1 + a_2 Y + v), \varphi\right)$$

INFERENCE

- Method of moments and least squares
- Maximum likelihood

$$\ell(\mu, \varphi; x) = (\mu\varphi - 1)\log(x) + [(1 - \mu)\varphi - 1]\log(1 - x) + \\ + \log\Gamma(\varphi) - \log\Gamma(\mu\varphi) - \log\Gamma[(1 - \mu)\varphi]$$

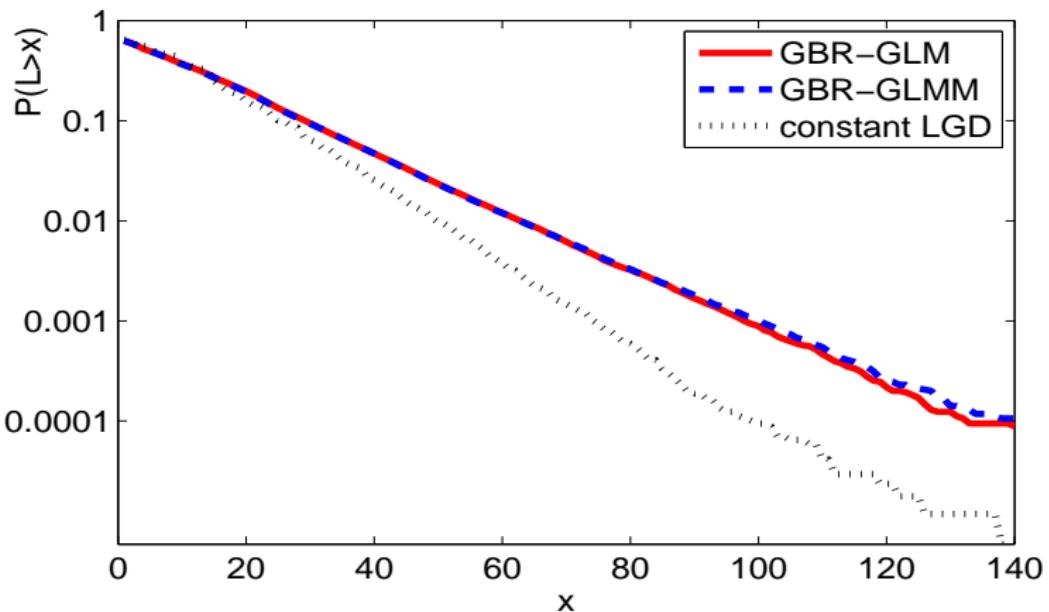
Marginal likelihood in GLMM

$$\ell_m = \sum_{t=1}^T \log \left(\int \prod_{k=1}^{K_t} L(a, \varphi, \zeta_t, v_t; x_{t,k}) p_{\sigma_v}(v_t) dv_t \right)$$

- Model selection: Likelihood ratio test, AIC, BIC

	GLM	JGLM	GLMM
a_1	0.3459 (0.0359)	0.3471	0.3319
a_2	-0.3213 (0.0298)	-0.3246	-0.3307
φ	3.0276 (0.1149)	-	3.3240
b_1	-	1.0879	-
b_2	-	-0.0306	-
σ_v	-	-	0.2943
-2ℓ	-402.74	-403.34	-468.78
AIC	-398.74	-395.34	-460.78
BIC	-381.67	-375.25	-440.69

TABLE: Estimates of the models with logit link for mean LGD.



SADDLEPOINT APPROXIMATION TO THE LOSS DISTRIBUTION

The conditional CGF and its derivatives up to second order are

$$\begin{aligned}\kappa(t, Y) &= \sum_{i=1}^n \log [1 - p_i + p_i {}_1F_1(\alpha_i, \alpha_i + \beta_i; w_i t)] . \\ \kappa'(t, Y) &= \sum_{i=1}^n \frac{w_i p_i {}_1F_1(\alpha_i + 1, \alpha_i + \beta_i + 1; w_i t)}{1 - p_i + p_i {}_1F_1(\alpha_i, \alpha_i + \beta_i; w_i t)} \frac{\alpha_i}{\alpha_i + \beta_i}, \\ \kappa''(t, Y) &= \sum_{i=1}^n \left\{ \frac{w_i^2 p_i \alpha_i (\alpha_i + 1) {}_1F_1(\alpha_i + 2, \alpha_i + \beta_i + 2; w_i t)}{(\alpha_i + \beta_i)(\alpha_i + \beta_i + 1)[1 - p_i + p_i {}_1F_1(\alpha_i, \alpha_i + \beta_i; w_i t)]} \right. \\ &\quad \left. - \frac{w_i^2 p_i^2 \alpha_i^2 {}_1F_1(\alpha_i + 1, \alpha_i + \beta_i + 1; w_i t)^2}{(\alpha_i + \beta_i)^2 [1 - p_i + p_i {}_1F_1(\alpha_i, \alpha_i + \beta_i; w_i t)]^2} \right\}.\end{aligned}$$

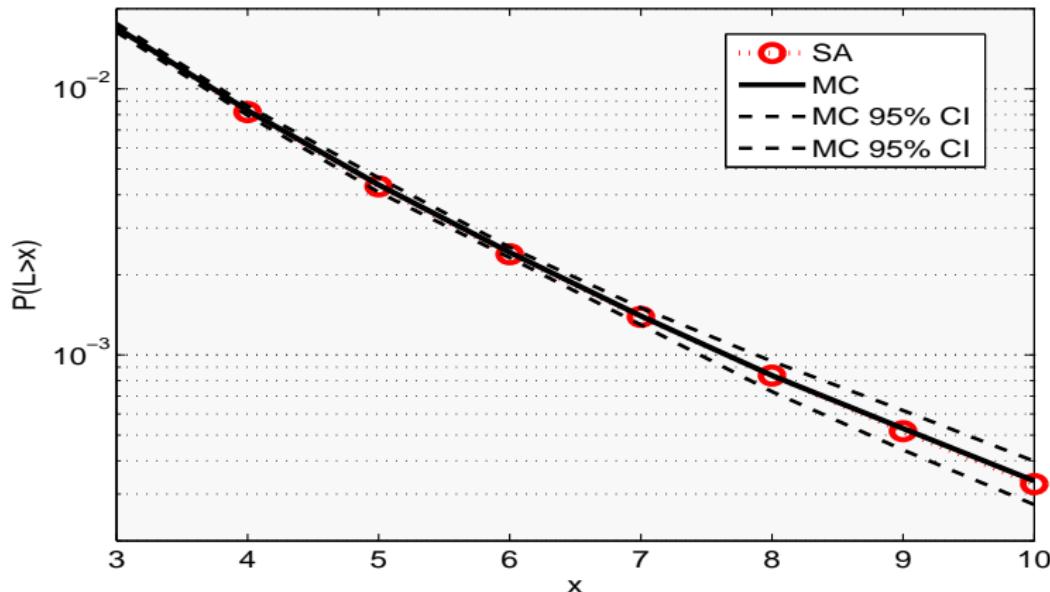


FIGURE: The loss distribution obtained from the saddlepoint approximation (SA) compared to results based on Monte Carlo (MC) simulation of two hundred thousand scenarios. The MC 95% confidence interval (CI) are based on the standard deviation calculated using 10 simulated sub-samples of 20 thousand scenarios each.