

VB, EM and EP

Joris Mooij

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1 VB

1.1 Inequality

Suppose \mathbf{X} is an observed variable and \mathbf{Z} a latent variable. Then we can write

$$\ln p(\mathbf{X}) = \mathcal{L}(q) + KL(q||p_{\mathbf{Z}|\mathbf{X}})$$

where

$$\mathcal{L}(q) = \int q(\mathbf{Z}) \ln \frac{p(\mathbf{X}, \mathbf{Z})}{q(\mathbf{Z})} d\mathbf{Z}, \quad KL(q||p_{\mathbf{Z}|\mathbf{X}}) = - \int q(\mathbf{Z}) \ln \frac{p(\mathbf{Z}|\mathbf{X})}{q(\mathbf{Z})} d\mathbf{Z}$$

Then, because of the properties of the Kullback-Leibler divergence,

$$\ln p(\mathbf{X}) = \mathcal{L}(q) + KL(q||p_{\mathbf{Z}|\mathbf{X}}) \geq \mathcal{L}(q)$$

with equality if $q = p_{\mathbf{Z}|\mathbf{X}}$.

1.2 Variational Bayes

The VB approximation thus approximates the posterior $p_{\mathbf{Z}|\mathbf{X}}$ by

$$q^* = \arg \max_{q \in \mathcal{Q}} \mathcal{L}(q) = \arg \max_{q \in \mathcal{Q}} \int q(\mathbf{Z}) \ln \frac{p(\mathbf{X}, \mathbf{Z})}{q(\mathbf{Z})} d\mathbf{Z}$$

and the evidence $\ln p(\mathbf{X}) \approx \mathcal{L}(q^*)$.

As a special case, we take the family of distributions to factorize, i.e., we assume

$$q(\mathbf{Z}) = \prod_{i=1}^M q_i(\mathbf{Z}_i)$$

Then we get the VB update equation:

$$\begin{aligned} \ln q_i^*(\mathbf{Z}_i) &= \int \ln p(\mathbf{X}, \mathbf{Z}) q_{\setminus i}(\mathbf{Z}) d\mathbf{Z}_{\setminus i} + \text{const.} \\ &= \mathbb{E}_{q_{\setminus i}} \ln p(\mathbf{X}, \mathbf{Z}) + \text{const.} \end{aligned} \tag{1}$$

where

$$q_{\setminus i}(\mathbf{Z}) = \prod_{\substack{j=1 \\ j \neq i}}^M q_j(\mathbf{Z}_j).$$

2 EM

In EM, we distinguish latent variables $\boldsymbol{\theta}$ over which we *optimize* from latent variables \mathbf{Z} which we marginalize over. EM can be derived as a special case of VB, taking $q(\mathbf{Z}, \boldsymbol{\theta}) = q(\mathbf{Z})\delta(\boldsymbol{\theta} - \boldsymbol{\theta}_0)$. However, this will lead to a lower bound on the evidence of $-\infty$ because of the delta function. So it pays off to treat the $\boldsymbol{\theta}$ variables in a slightly different way. We write

$$\ln p(\mathbf{X}, \boldsymbol{\theta}) = \mathcal{L}(q, \boldsymbol{\theta}) + KL(q||p_{\mathbf{Z}|\mathbf{X}, \boldsymbol{\theta}})$$

where

$$\mathcal{L}(q, \boldsymbol{\theta}) = \int q(\mathbf{Z}) \ln \frac{p(\mathbf{X}, \mathbf{Z}, \boldsymbol{\theta})}{q(\mathbf{Z})} d\mathbf{Z}, \quad KL(q||p_{\mathbf{Z}|\mathbf{X}, \boldsymbol{\theta}}) = - \int q(\mathbf{Z}) \ln \frac{p(\mathbf{Z} | \mathbf{X}, \boldsymbol{\theta})}{q(\mathbf{Z})} d\mathbf{Z}$$

2.1 E-step

This is the maximization over $q(\mathbf{Z})$. The result depends on $\boldsymbol{\theta}$:

$$q_{\boldsymbol{\theta}}^*(\mathbf{Z}) = \arg \max_{q \in \mathcal{Q}} \mathcal{L}(q, \boldsymbol{\theta}) = \arg \max_{q \in \mathcal{Q}} \int q(\mathbf{Z}) \ln \frac{p(\mathbf{X}, \mathbf{Z}, \boldsymbol{\theta})}{q(\mathbf{Z})} d\mathbf{Z}$$

If we allow all possible distributions, then we simply obtain the posterior:

$$q_{\boldsymbol{\theta}}^*(\mathbf{Z}) = p(\mathbf{Z} | \mathbf{X}, \boldsymbol{\theta}).$$

2.2 M-step

This is the maximization over $\boldsymbol{\theta}$. The result depends on q :

$$\boldsymbol{\theta}_q^* = \arg \max_{\boldsymbol{\theta}} \int q(\mathbf{Z}) \ln \frac{p(\mathbf{X}, \mathbf{Z}, \boldsymbol{\theta})}{q(\mathbf{Z})} d\mathbf{Z}$$

If we take for $q = q_{\boldsymbol{\theta}}^*$ (the result of the standard E-step), then we get the standard formulation of the M-step:

$$\boldsymbol{\theta}^{\text{new}} = \arg \max_{\boldsymbol{\theta}} \int p(\mathbf{Z} | \mathbf{X}, \boldsymbol{\theta}^{\text{old}}) \ln \frac{p(\mathbf{X}, \mathbf{Z}, \boldsymbol{\theta})}{p(\mathbf{Z} | \mathbf{X}, \boldsymbol{\theta}^{\text{old}})} d\mathbf{Z}$$

2.3 Hybrid EM/VB

It should be obvious now how to derive hybrid versions of EM and VB. Simply take \mathcal{Q} to be a strict subset of all possible distributions on \mathbf{Z} . Then the standard E-step will be replaced by the VB updates for \mathbf{Z} , and the M-step does not use the exact posterior $p(\mathbf{Z} | \mathbf{X}, \boldsymbol{\theta}^{\text{old}})$ but the current VB approximation $q(\mathbf{Z})$.

3 EP

Given again observed variables \mathbf{X} and latents \mathbf{Z} , with joint probability

$$p(\mathbf{X}, \mathbf{Z}) = \prod_i f_i(\mathbf{Z}, \mathbf{X}).$$

We want to approximate the posterior

$$p(\mathbf{Z} | \mathbf{X}) = \frac{1}{p(\mathbf{X})} \prod_i f_i(\mathbf{Z}, \mathbf{X}) \approx \frac{1}{Z} \prod_i q_i(\mathbf{Z}) =: q(\mathbf{Z})$$

and model evidence

$$p(\mathbf{X}) = \int \prod_i f_i(\mathbf{Z}, \mathbf{X}) d\mathbf{Z} \approx \int \prod_i q_i(\mathbf{Z}) d\mathbf{Z}$$

The EP update for factor $q_i(\mathbf{Z})$ is obtained by considering:

$$q(\mathbf{Z})^{\text{new}} = \arg \min_{q \in \mathcal{Q}} KL \left(\frac{1}{Z_j} f_j(\mathbf{Z}, \mathbf{X}) q_{\setminus j}^{\text{old}}(\mathbf{Z}) \parallel q(\mathbf{Z}) \right)$$

where

$$q_{\setminus j}(\mathbf{Z}) := \prod_{i \neq j} q_i(\mathbf{Z}).$$

$$Z_j = \int f_j(\mathbf{Z}, \mathbf{X}) q_{\setminus j}(\mathbf{Z}) d\mathbf{Z}.$$

This optimization is easily done if all q_j are in an exponential family \mathcal{Q}_j , because then \mathcal{Q} is also an exponential family. The new approximate factor $q_j(\mathbf{Z})$ is then given by:

$$q_j^{\text{new}}(\mathbf{Z}) = Z_j \frac{q^{\text{new}}(\mathbf{Z})}{q_{\setminus j}(\mathbf{Z})}$$

3.1 Moment matching

The solution is obtained by matching the moments of the sufficient statistics. Suppose

$$\mathcal{Q} = \{h(\mathbf{Z})g(\boldsymbol{\eta}) \exp(\boldsymbol{\eta}^T \mathbf{u}(\mathbf{Z})) : \boldsymbol{\eta}\}$$

Then

$$q^* = \arg \min_{q \in \mathcal{Q}} KL(p \parallel q) \iff \mathbb{E}_{q^*}(\mathbf{u}(\mathbf{Z})) = \mathbb{E}_p(\mathbf{u}(\mathbf{Z})).$$