C.Larédo

Introduction

Ground survey data

Inference for Incompletely Observed Branching Processes

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Introduction

Study of a mode for feral oilseed rape dynamics

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(1) Biological Problem

Context

- Many species can escape from fields and survive outside fields
- Raises numerous questions concerning
 - \Rightarrow their foundation and origine,
 - \Rightarrow their ability to persist,
 - \Rightarrow their dispersal vectors.
- problems linked to Theoretical Ecology: populations dynamics in a pertubated habitat.

- problems linked to Applied Ecology: environmental risks
 - \Rightarrow release of Genetically Modified Plants,
 - \Rightarrow escape of transgenes in the landscape.

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(2) Escape of a cultivated species

Echappement d'une espèce cultivée



Champ, jachères

Bords de champs, Bords de routes/chemins... Conséquences...

- modification des communautés des bordures (invasibilité, adaptation locale, compétitivité)

- flux de gènes (pollen et graines)





Tournesol





Colza





Sorgho

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(3) Reasons for choosing oilseed rape (*Brassica napus* .*L*)

- Environmental risks associated with cultivating transgenic oilseed rape (herbicide resistant).
- Abundant populations outside fields.
- Existence of wild species able to hybridize with feral plants.
- Persistence of seeds in the soil for several years: presence of a seed bank.
- Populations might be maintained by immigration: neighbouring fields or seeds released by trucks.

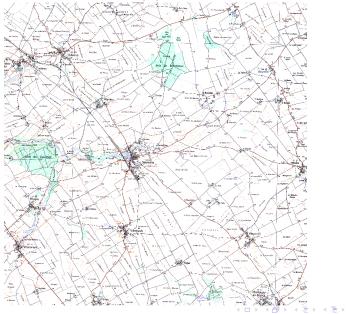
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(4) Region of Sélommes



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Experimental data Aims

(5) Production basin: Sélommes

- Ground survey of 500 feral populations on three roads and three paths
- Monthly observations from January 2001 to June 2003 et localization with G.P.S.
- Counts of the number of plants in each developmental stage within each population .

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• Observations of possible covariates: presence/absence of cultivated oilseedrape, same year, herbicide treatments, favourable Winter,..

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Experimental data Aims (6) Sélommes, Loir-et-Cher: Production basin for oilseed rape. January 2001- June 2003: suivi of cultivated fields and feral populations.

Map of the experiment with the three roads and three paths.



Fig. 2-12. : Map of the study area representing the three paths and three roads where crops and feral populations of oilseed rape were surveyed from January 2001 to June 2003.

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Experimenta data Aims

- Explore the processes involved in the dynamics of these populations.
- Link individual scale and population scale.
- Intrinsic randomness of survival and populations offsprings.
- Ecology \Rightarrow Stage structured models (Leslie matrices, Caswell 2001).
- Stochastic modelling for the dynamics of these populations
- Framework: Multitype branching processes with immigration in one of the types
- Parametric inference for the demographic parameters of the laws ruling the dynamics of these populations.
- Using all the data collected in Sélommes
- Problem: one type is never observed: the seeds
 - \Rightarrow New problem in Statistical Inference

(7) Aims

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Likelihood

Incomplete observations

Inference for Incompletely Observed Branching Processes Part 2

June 6, 2007

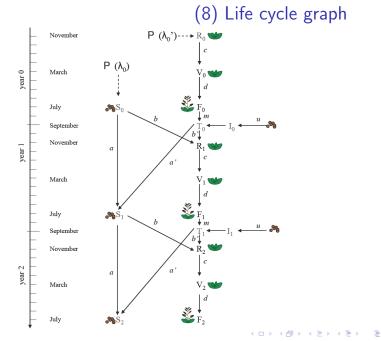
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(9) Annual plant model structured in 5 stages: $X_i = (S_i, T_i, R_i, V_i, F_i).$

- Seeds buried in the soil i.e. in the seed bank: S_i
- Seeds on the soil: T_i
- Rosettes before Winter or non vernalised rosettes: R_i
- Rosettes after Winter or vernalised rosettes: V_i
- Mature plants carrying pods: F_i.

Model parameters

- $P(\text{seed in } S_i \rightarrow \text{non-vernalised rosette in } R_i) = b$
- $P((\text{seed in } T_i \rightarrow \text{non-vernalised rosette in } R_i) = b'$
- $P(\text{seed in } S_i \rightarrow \text{seed in } S_{i+1}) = a$
- $P(\text{seed in } T_i \rightarrow \text{seed in } S_{i+1}) = a'$
- $P(\text{non-vernalised rosette in } R_i \rightarrow \text{vernalised rosette in } V_i) = c$
- $P(\text{vernalised rosette in } V_i \rightarrow \text{mature plant in } F_i) = d$
- G(.): Offspring distribution of plants in F_i (\Rightarrow seeds in T_{i+1})
- I_{i+1} : Immigration r.v. distribution $\mu \iff \text{seeds in } T_{i+1}$).

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(10) Proposition:

- $X_i = (S_i, T_i, R_i, V_i, F_i)$ multitype branching process
- Initial distribution $\pi_0(x) = \pi_0(s, r, v, f, t)$

$$P(S_0 = s, T_0 = t) p_3(r/s, t) p_4(v/r) p_5(f/v)$$

with $p_4(v/r) = \mathcal{B}(r; c)(v), p_5(f/v) = \mathcal{B}(v; d)(f)$
 $p_3(r/s, t) = (\mathcal{B}(s; b) \star \mathcal{B}(t; b'))(r)$

• Transition kernel p(x; x')

$$p(x,x') = p_1(s'/s,t,r) p_2(t'/f) p_3(r'/s',t') p_4(v'/r') p_5(f'/v')$$
$$p_2(t'/f) = (G^{\star f} \star \mu)(t')$$
$$p_1(s'/s,t,r) = \frac{\mathcal{M}(s;a,b) \star \mathcal{M}(t;a',b'))(s',r')}{\mathcal{B}(s;b) \star \mathcal{B}(t;b'))(r)}$$

Notation: $\mathcal{M}(N; a, b, c)(i, j, k) = \mathcal{M}(N; a, b)(i, j)$ for $i + j \leq N$.

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(11) Notations

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Branching Processes Part 2

Incomplete observations

Parameters: $\theta = (\theta^1, \theta^2, \theta^3, c, d, a, b, a', b')$

- θ^1 distribution of (S_0, T_0) ,
- $\theta^2 \rightarrow$ offspring distribution $G(\theta^2,.)$,
- $\theta^3 \rightarrow \text{immigration distribution } \mu(\theta^3, .).$

Complete observations

- K = 300 independent populations during *n* years.
- Observations in population k at generation i: $x_i^k = (s_i^k, t_i^k, r_i^k, v_i^k, f_i^k).$
- Observations up to generation *n*: $O_{0:n}^k = (x_0^k, \dots, x_n^k)$.
- Whole observations up to time $n: O_{0:n} = (O_{0:n}^1, \dots, O_{0:n}^K).$

True value of the parameter: θ_0

(12) Asymptotics

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Three possible asymptotics:

- **1** K prescribed and $n \to \infty$
- **2** *n* prescribed and $K \to \infty$
- **3** $K \to \infty$ and $n \to \infty$

Here: K = 500 and n = 3.

- reasonable to choose (2)
- Other studies often belong to case (1)
- statistical inference also investigated in case (1) (here X_i subcritical branching with immigration \Rightarrow positive recurrent)

(13) Likelihood

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$$\log L(heta; O_{0:n}) = I(heta; O_{0:n}) = \sum_{i=0}^5 I_i(heta; O_{0:n})$$
 with

•
$$I_0(\theta; O_{0:n}) = I_0(\theta^1; O_{0:n}) = \sum_{k=1}^K \log p_{\theta^1}(s_0^k, t_0^k),$$

•
$$l_1(\theta; O_{0:n}) = \sum_{k=1}^{K} \sum_{i=0}^{n} \log(\mathcal{B}(s_i^k; b) \star \mathcal{B}(t_i^k; b'))(r_i^k)$$

•
$$l_2(\theta; O_{0:n}) = l_2(c; O_{0:n}) = \sum_{k=1}^{K} \sum_{i=0}^{n} \log \mathcal{B}(r_i^k; c)(v_i^k)$$

•
$$I_3(\theta; O_{0:n}) = I_3(\mathbf{d}; O_{0:n}) = \sum_{k=1}^{K} \sum_{i=0}^{n} \log \mathcal{B}(\mathbf{v}_i^k; \mathbf{d})(f_i^k)$$

•
$$I_4(\theta; O_{0:n}) = I_5(\theta^2, \theta^3; O_{0:n}) = \sum_{k=1}^K \sum_{i=0}^n \log((G^{*f_i^k} \star \mu)(t_{i+1}^k)).$$

•
$$I_5(\theta; O_{0:n}) = I_5(a, b, a', b'; O_{0:n}) =$$

,

$$\sum_{k=1}^{K} \sum_{i=0}^{n} \log(\frac{\mathcal{M}(s_i^k; a, b) \star \mathcal{M}(t_i^k; a', b')(s_{i+1}^k, r_i^k)}{\mathcal{B}(s_i^k; b) \star \mathcal{B}(t_i^k; b'))(r_i^k)})$$

Inference for Incompletely Observed Branching Processes Part 2

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(14) Maximum likelihood estimates

$$\hat{c} = \frac{\sum_{k=1}^{K} \sum_{i=0}^{n} v_{i}^{k}}{\sum_{k=1}^{K} \sum_{i=0}^{n} r_{i}^{k}}; \quad \hat{d} = \frac{\sum_{k=1}^{K} \sum_{i=0}^{n} f_{i}^{k}}{\sum_{k=1}^{K} \sum_{i=0}^{n} v_{i}^{k}}.$$
(1)

Under $P_{ heta_0}$ as $K o \infty$,

- (\hat{c}, \hat{d}) strongly consistent, asymptotically Gaussian at rate \sqrt{K} .
- $l_1 + l_5 = l_1' \rightarrow$ quasilikelihood $\tilde{l_1'}$: same results for (a, b, a', b').
- $I_4 \rightarrow$ branching part:
 - -loglikelihood: $\sum_{k=1}^{K} \sum_{i=0}^{n} (Log \left(G_{\theta^2}^{\star f_i^k} \star \mu_{\theta^3} \right) (t_{i+1}^k) \right)$
 - conditional least squares or variants: (Wei & Winnicki 1990) $\sum_{k=1}^{K} \sum_{i=0}^{n} (t_{i+1}^{k} m_{\theta^{2}} f_{i}^{k} u_{\theta^{3}})^{2},$ ($m_{\theta^{2}}$: mean of $G_{\theta^{2}}$ and $u_{\theta^{3}}$: mean of $\mu_{\theta^{3}}$).

• Consistent and asymptotically Gaussian estimators of $(m_{\theta^2}, u_{\theta^3})$ Conclusions: Standard study, estimation at rate \sqrt{K} . Remark: Asymptotics for Markov chains K = 1; $n \to \infty$: \Rightarrow would lead to similar results.

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Incomplete observations

Framework

Study for the Poisson case Likelihood

(15) Framework

Incomplete Observations

- Impossible in practice to observe S_i (nb of seeds in the seed bank) and T_i (nb of seeds on the soil).
- Requires to study the process $\{Y_i = (R_i, V_i, F_i); i = 1, ..., n\}$.
- (*Y_i*) is no longer Markov.
- (*Y_i*) is not linked to a Hidden Markov Model since (*S_i*, *T_i*) does not evolve independently

New statistical problem

- What parameters are identifiable when only (Y_i) is observed?
 (i.e. (y_i^k); i = 1...n; k = 1...K))
- How to estimate these parameters?
- Properties of these estimators?
- Non standard inference $pb \Rightarrow$ requires a specific study.

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Incomplete observations

Framework

Study for the Poisson case Likelihood

(16) Poisson case model

- * Informative example leading to explicit computations
- * Analogy with the Kalman filter

Assumptions

- Offspring distribution G: Poisson law $\mathcal{P}(m)$
- Immigration distribution μ in type T_i : Poisson $\mathcal{P}(u)$
- Initial distribution of S_0 (seeds in the seed bank): Poisson $\mathcal{P}(\sigma)$
- Initial distribution of T_0 (non-vernalised rosettes): Poisson $\mathcal{P}(\tau)$

• S_0 and T_0 are independent r. v.

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Framework

Study for the Poisson case

(17) Probabilistic properties

Notations

- Recap $\mathcal{F}_i = \sigma((S_k, T_k, R_k, V_k, F_k); k = 0, ..., i).$
- Define $\mathcal{G}_i = \sigma(R_k, V_k, F_k); k = 0, \dots, i).$
- Set $Y_k = (R_k, V_k, F_k)$.
- Set $\Lambda_0 = a\sigma + a'\tau$ and for $i \ge 1$,
- $\Lambda_i = a^i \Lambda_0 + a' u \frac{1-a^i}{1-a} + a' m (F_{i-1} + aF_{i-2} + a^2 F_{i-3} + \dots a^{i-1} F_0)$
- $\Lambda'_i = mF_i + u$ for $i \ge 0$,

Theorem

Under Assumptions (A1)-(A2), $Y_i = (R_i, V_i, F_i)$ satisfies

- initial distribution is $\tilde{\pi}_0(y) = \mathcal{P}(b\sigma + b'\tau) p_4(v/r) p_5(f/v)$
- conditional distribution $\mathcal{L}(Y_{i+1}/\mathcal{G}_i)$, $P(Y_{i+1} = (r', v', f')/\mathcal{G}_i) = \mathcal{P}(b\Lambda_i + b'\Lambda'_i)(r') p_4(v'/r') p_5(f'/v')$

Explicit dependence on the past up to time 0 through the r.v. F_i Rk: Conditionally on G_i , S_{i+1} and T_{i+1} independent $\mathcal{P}(\Lambda_i)$, $\mathcal{P}(\Lambda'_i)$.

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Framework Study for the Poisson case Likelihood

(18) Incomplete Model Likelihood

Notations Observations: $\tilde{O}_{0:n} = (y_i^k; i = 1 \dots n; k = 1 \dots K)$ Parameter: $\theta = (\sigma, \tau, m, u, c, d, a, b, a', b')$ Define: $\lambda_i^k(\theta)$ and $\lambda_i'^k(\theta)$ realizations of $\Lambda_i(\theta), \Lambda_i'(\theta)$ in population k. Define $\Phi_i = b\Lambda_i + b'\Lambda_i'$ and $\varphi_i^k(\theta) = b\lambda_i^k(\theta) + b'\lambda_i'^k(\theta)$

Likelihood for one population

• $L(\theta; y_0^k, \dots, y_n^k) = \tilde{\pi}_0(\theta; y_0^k) \prod_{i=0}^{n-1} P_{\theta}(Y_{i+1} = y_{i+1}^k/y_i^k, \dots, y_0^k).$

• $P_{\theta}(Y_{i+1} = y_{i+1}^k/y_i^k, \dots, y_0^k) = \mathcal{P}(\varphi_i^k(\theta))(r_{i+1}^k) p_4(\theta; v_{i+1}^k/r_{i+1}^k) p_5(\theta; f_{i+1}^k/r_{i+1}^k)$

Loglikelihood $\tilde{I}(\theta, \tilde{O}_{0:n})$ associated with $\tilde{O}_{0:n}$

- $\tilde{I}(\theta, \tilde{O}_{0:n}) = \sum_{i=0}^{4} \tilde{I}_i(\theta, \tilde{O}_{0:n})$ with
- $\tilde{l}_0(\theta, \tilde{O}_{0:n}) = \sum_{k=1}^K \log \mathcal{P}(b\sigma + b'\tau)(r_0^k).$
- $\tilde{l}_2(\theta, \tilde{O}_{0:n}) = l_2(c, O_{0:n}); \ \tilde{l}_3(\theta, \tilde{O}_{0:n}) = l_3(d, O_{0:n}).$
- It remains to study $\tilde{l}_0(\theta, \tilde{O}_{0:n})$ and $\tilde{l}_4(\theta, \tilde{O}_{0:n})$.

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(19)Study of $\tilde{I}_0(\theta, \tilde{O}_{0:n}), \tilde{I}_4(\theta, \tilde{O}_{0:n})$

Preliminaries

- Set $\mu = \Lambda_0 = a\sigma + a'\tau$ and $\nu = b\sigma + b'\tau$
- $\lambda_i^k = a^i \mu + a' u \frac{1-a^i}{1-a} + a' m (f_{i-1}^k + a f_{i-2}^k + a^2 f_{i-3}^k + \dots a^{i-1} f_0^k)$
- $\lambda_i^{\prime k} = (mf_i^k + u)$ and $\varphi_i^k = b\lambda_i^k + b'\lambda_i^{\prime k}$
- $\tilde{l_0}(\theta, \tilde{O}_{0:n}) = \tilde{l_0}(\nu; r_0^1, \dots r_0^k) \Rightarrow \nu$ identifiable
- MLE: $\hat{\nu} = \frac{\sum_{k=1}^{K} r_0^k}{K}$ consistent asympt. Gaussian at rate \sqrt{K} .

Estimation of $\theta = (\mu, m, u, a, b, a', b')$

(c,d) omitted now.

- All the difficulties are in the study of this last term
- $\tilde{l}_4(\theta, \tilde{O}_{0:n}) = \tilde{l}_4(\phi_i^k, \tilde{O}_{0:n}) = \sum_{i=1}^{K} (-\varphi_i^k + r_i^k \log \varphi_i^k)$
- What parameters are identifiable given that all the available information is contained in the φ_i^k?

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(20)Estimating θ from $\tilde{l}_4(\theta, \tilde{O}_{0:n})$

Define $\mathcal{K}(P, Q)$ as the Kullback-Leibler information of Q w.r.t. PRecap If $P \sim \mathcal{P}(\lambda_0), \ Q \sim \mathcal{P}(\lambda), \ \mathcal{K}(P, Q) = \lambda - \lambda_0 - \lambda_0 (\log \lambda - \log \lambda_0).$

Theorem

Let $heta_0$ be the true parameter value. Then, almost surely under $P_{ heta_0}$, as $K \to +\infty$

$$\frac{1}{\mathcal{K}}\tilde{l}_4(\theta,\tilde{O}_{0:n})\to -E_{\theta_0}\sum_{i=0}^{n-1}\mathcal{K}(\mathcal{P}(\Phi_i(\theta_0),\Phi_i(\theta)).$$

 Φ_i : random variables depending on θ and on the r.v. F_0, \ldots, F_i .

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(21) Identifiability

Corollary

- 1 If n = 0, only $\nu = b\sigma + b'\tau$ is identifiable
- 2 If n = 1, the identifiable parameters are: $\nu, b\mu + b'u, b'm$
- **3** If n = 2, the identifiable parameters are: $\nu, b\mu + b'u, b'm, ab\mu + b'u + a'bu, \frac{a'b}{b'}$
- **4** If $n \ge 3$, the identifiable parameters are: $\nu, b\mu, b'u, b'm, a, \frac{a'b}{b'}$.

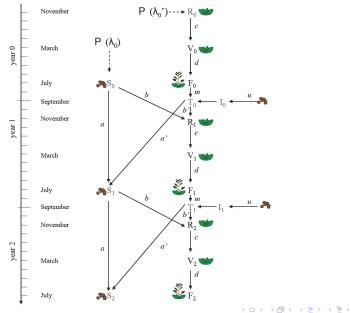
Consequences

- Only the combinations appearing in (3) can be estimated.
- Natural here to use E.M. or Bayesian approaches: ongoing work.
- Very important to ecologists: how parameters are linked using the available information ⇒ impossible with E.M
- Observations collected in the ground survey: n = 2.
- Many observed populations $K = 500 \Rightarrow$ rate \sqrt{K} .
- K large \Rightarrow Ability to introduce covariates in the estimation.

Inference for incompletely observed branching processes

Dynstoch 2007

Life cycle graph



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Estimation of the parameters Taking the affiles of the

		Estimates of the x_i 's	5	
	"All"	"With crop"	"Without crop"	
	n = 595	n = 55	n = 540	
\hat{x}_1	2.8 [0.9, 4.8]	7.4 [1.6, 13.3]	1.8 [0.3, 3.4]	<i>x</i> ₁
\hat{x}_2	1.5 [-0.1, 3.1]	-2.4 [-22.6, 17.8]	1.5 [0.17, 2.9]	<i>x</i> ₂
\hat{x}_3	16.4 [5.7, 27.1]	7.8 [-8.3, 23.8]	18.2 [7.7, 28.6]	<i>x</i> ₃
\hat{x}_4	12.5 [3.1, 21.8]	119.0 [12.4, 225.7]	6.4 [0.3, 12.4]	<i>x</i> ₄
$\hat{\tau}$	22.7	17.9	19.8	

Link with the model parameters

- $x_1 = b'm$: "Efficient fecundity"
- $x_2 = a'bm$:" Efficient delayed fecundity "

• $x_3 = ub' + b\lambda_0$: seeds in the seed bank + immigrating seeds.

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•
$$x_4 = ub' + a'um + a\lambda_0$$

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Estimated values for the model parameters

Known values from the bibliography (Claessen, data from U.K.)

- Incorporation in the seed bank: $\hat{a'} = 0.006$
- Annual survival in the seed bank: $\hat{a} = 0.15$
- Emergence rate from the seed bank: $\hat{b'} = 0.0043$

Derived estimated values for the other model parameters

- R \rightarrow V: $\hat{c} = 0.31$ (favourable Winter); $\hat{c} = 0.14$ (hiver non favorable);
- V \rightarrow F : $\hat{d} = 0.05$
- Offspring distribution G : mean $\hat{m} = 700$
- Immigration: $\hat{u} = 110$ seeds/m (with crop); $\hat{u} = 25$ seeds/m (without crop)

- S \rightarrow R: $\hat{b} = 0.36$
- Seeds in the seed bank at time 0: $\hat{\lambda_0} = 25$.