# How to Model Operational Risk, if You Must

Paul Embrechts ETH Zürich (www.math.ethz.ch/~embrechts)

Based on joint work with V. Chavez-Demoulin, H. Furrer, R. Kaufmann, J. Nešlehová and G. Samorodnitsky

## Contents

- A. The New Accord (Basel II)
- B. Risk measurement methods for OP risks
- C. Loss distribution approach
- D. Some data
- E. The Capital Charge Problem
- F. Accuracy of VaR-estimates
- G. Conclusions

#### H. References

## A. The New Accord (Basel II)

- 1988: Basel Accord (Basel I): minimal capital requirements against credit risk, one standardised approach, Cooke ratio
- **1996**: Amendment to Basel I: market risk, internal models, netting
- **1999**: First Consultative Paper on the New Accord (Basel II)
- to date: CP3: Third Consultative Paper on the New Basel Capital Accord (www.bis.org/bcbs/bcbscp3.htmcp3)
- 2004: Revision: (final) version
- 2006–2007: full implementation of Basel II ([14])

### **Basel II: What is new?**

- Rationale for the New Accord: More flexibility and risk sensitivity
- **Structure** of the New Accord: Three-pillar framework:
  - Pillar 1: minimal capital requirements (risk measurement)
  - **2** Pillar 2: supervisory review of capital adequacy
  - Pillar 3: public disclosure

- Two options for the measurement of credit risk:
  - **s** Standard approach
  - % Internal rating based approach (IRB)
- Pillar 1 sets out the minimum capital requirements (Cooke Ratio):

$$\frac{\text{total amount of capital}}{\text{risk-weighted assets}} \ge 8\%$$

- MRC (minimum regulatory capital)  $\stackrel{\text{def}}{=} 8\%$  of risk-weighted assets
- Explicit treatment of operational risk

#### **Operational Risk**:

The risk of losses resulting from inadequate or failed internal processes, people and systems, or external events.

**Remark**: This definition includes legal risk, but excludes strategic and reputational risk.

**Note**: Solvency 2

- Notation:  $C_{OP}$ : capital charge for operational risk
- Target:  $C_{\text{OP}} \approx 12\%$  of MRC (down from initial 20%)
- Estimated total losses in the US (2001): \$50b

#### • Some examples

- **\*** 1977: Credit Suisse Chiasso-affair
- **#** 1995: Nick Leeson/Barings Bank, £1.3b
- **2001**: September 11
- # 2001: Enron (largest US bankruptcy so far)
- **2002:** Allied Irish, £450m

## **B. Risk measurement methods for OP risks**

Pillar 1 regulatory minimal capital requirements for operational risk: Three distinct approaches:

- Basic Indicator Approach
- **2** Standardised Approach
- Advanced Measurement Approach (AMA)

### **Basic Indicator Approach**

• Capital charge:

$$C_{\rm OP}^{\rm BIA} = \alpha \times GI$$

- $C_{\text{OP}}^{\text{BIA}}$ : capital charge under the Basic Indicator Approach
- GI: average annual gross income over the previous three years
- $\alpha = 15\%$  (set by the Committee based CISs)

### **Standardised Approach**

• Similar to the BIA, but on the level of each business line:

$$C_{\mathsf{OP}}^{\mathsf{SA}} = \sum_{i=1}^8 \beta_i \times GI_i$$

$$\beta_i \in [12\%, 18\%]$$
,  $i = 1, 2, \dots, 8$ 

• 8 business lines:

Corporate finance (18%) Trading & sales (18%) Agency Services (15%) Commercial banking (15%) Retail brokerage (12%)

Payment & Settlement (18%) Retail banking (12%) Asset management (12%)

- Some remarks from the proposal:
  - \* A bank must be able to demonstrate that its approach captures potentially severe "tail" loss events
  - \* A one year holding period and a 99.9% confidence interval (like internal ratings-based approach for credit risk)
  - \* Independent model validation
  - \* RC = EL + UL, however . . .
  - \* Sufficient granularity
  - \* Addition of risk measures
  - \* Correlation effects
  - \* Use of internal data, external data, scenario analysis and driving factors
  - \* Repetitive versus non-repetitive losses
  - \* Minimum 5-year (initially 3-year) internal loss observation period
  - \* De minimis gross loss threshold, e.g.  $\in$  10000
  - \* From partial to full AMA

## **Advanced Measurement Approach (AMA)**

- Allows banks to use their internally generated risk estimates
- Preconditions: Bank must meet qualitative and quantitative standards before being allowed to use the AMA
- Risk mitigation via insurance possible ( $\leq 20\%$  of  $C_{OP}^{SA}$ )
- Incorporation of risk diversification benefits allowed
- "Given the continuing evolution of analytical approaches for operational risk, the Committee is not specifying the approach or distributional assumptions used to generate the operational risk measures for regulatory capital purposes."

• Example: Loss distribution approach © Paul Embrechts, 2005

### **Internal Measurement Approach**

• Capital charge (similar to Basel II model for Credit Risk):

$$C_{\mathsf{OP}}^{\mathsf{IMA}} = \sum_{i=1}^{8} \sum_{k=1}^{7} \gamma_{ik} e_{ik}$$

(first attempt)

- $e_{ik}$ : expected loss for business line i, risk type k $\gamma_{ik}$ : scaling factor
- 7 loss types: Internal fraud

External fraud Employment practices and workplace safety Clients, products & business practices Damage to physical assets Business disruption and system failures Execution, delivery & process management

### **C.** Loss Distribution Approach

• For each business line/loss type cell (i, k) one models

 $L_{i,k}^{T+1}$ : OP risk loss for business line *i*, loss type *k* over the future (one year, say) period [T, T+1]

$$L_{i,k}^{T+1} = \sum_{\ell=1}^{N_{i,k}^{T+1}} X_{i,k}^{\ell} \quad (\text{next period's loss for cell } (i,k))$$

Note that  $X_{i,k}^{\ell}$  is truncated from below

**Remark**: Look at the structure of the loss random variable  $L^{T+1}$ 

$$L^{T+1} = \sum_{i=1}^{8} \sum_{k=1}^{7} L_{i,k}^{T+1} \text{ (next period's total loss)}$$
$$= \sum_{i=1}^{8} \sum_{k=1}^{7} \sum_{\ell=1}^{N_{i,k}^{T+1}} X_{i,k}^{\ell}$$

### A methodological pause 1

$$L = \sum_{k=1}^{N} X_k$$
 (compound rv)

where  $(X_k)$  are the severities and N the frequency Models for  $X_k$ :

- gamma, lognormal, Pareto ( $\geq 0$ , skew)

Models for N:

- binomial (individual model)
- Poisson( $\lambda$ ) (limit model)
- negative binomial (randomize  $\lambda$  as a gamma rv)

• Choice of a risk measure g ( $\alpha \in (0,1)$  fixed)

$$C_{i,k}^{T+1,\mathsf{OR}} = g(L_{i,k}^{T+1}) = \begin{cases} F_{L_{i,k}^{T+1}}^{\leftarrow}(\alpha) = \operatorname{VaR}_{\alpha}(L_{i,k}^{T+1}) \\ \mathsf{ES}(L_{i,k}^{T+1}) = E(L_{i,k}^{T+1} | L_{i,k}^{T+1} > \operatorname{VaR}_{\alpha}(L_{i,k}^{T+1})) \end{cases}$$

- $VaR_{\alpha}$  is not coherent (example)
- $\mathsf{ES}_{\alpha}$  is coherent (modulo trivial change)

$$C^{T+1,\mathsf{OR}} = \sum_{i,k} g(L_{i,k}^{T+1})$$
 (perfect correlation)

- Why?
- Dependence effects (copulae)

 $VaR_{\alpha}$  is in general not coherent:

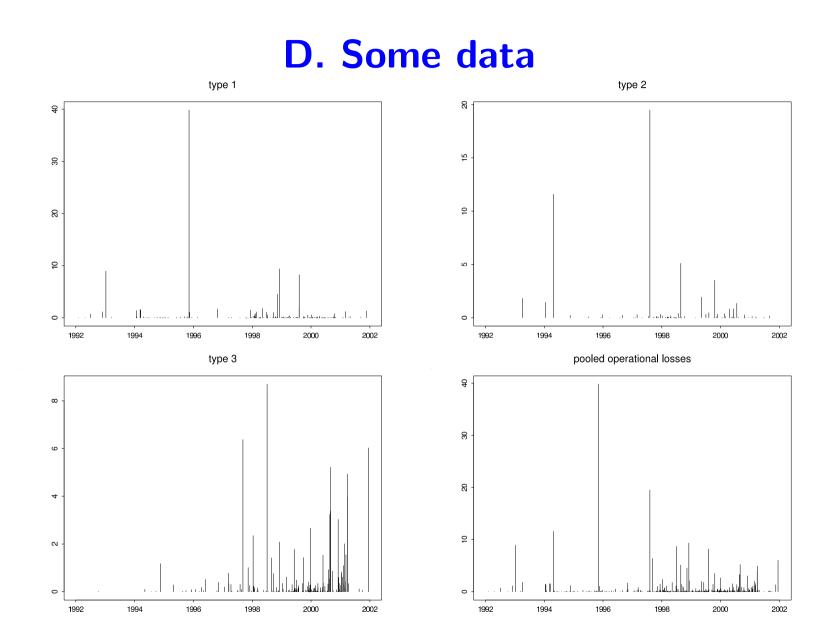
- 100 iid loans: 2%-coupon, 100 face value, 1% default probability (period: 1 year):

$$X_i = \begin{cases} -2 & \text{with probability 99\%} \\ 100 & \text{with probability 1\%} (\text{loss}) \end{cases}$$

- Two portfolios  $L_1 = \sum_{i=1}^{100} X_i$ ,  $L_2 = 100X_1$ 

$$-\underbrace{\operatorname{VaR}_{95\%}(L_1)}_{\operatorname{VaR}_{95\%}(\sum_{i=1}^{100} X_i)} > \underbrace{\operatorname{VaR}_{95\%}(100X_1)}_{\sum_{i=1}^{100} \operatorname{VaR}_{95\%}(X_i)}$$
(!)

- Hence the well-diversified portfolio  $L_1$  gets a higher (VaR-)risk charge than the very concentrated, "all eggs in one basket" portfolio  $L_2$
- Similar examples for heavy-tailed loss portfolios or special dependence
- This cannot happen when  $(X_1, \ldots, X_d)$  has a multivariate normal (or more generally, elliptical) distribution
- Link to Operational Risks: skewness



- Stylized facts about OP risk losses:
  - **\*** Loss amounts show extremes
  - **\*** Loss occurrence times are irregularly spaced in time (reporting bias, economic cycles, regulation, management interactions, structural changes, . . . )
  - % Non-stationarity (frequency(!), severity(?))
- Large losses are of main concern
- Repetitive versus non-repetitive losses
- Warning flag: observations are not in line with standard modelling assumptions

## A methodological pause 2

- severity models need to go beyond the classical models (binomial, homogeneous Poisson, negative binomial: → stochastic processes)
- as stochastic processes:
  - Poisson( $\lambda t$ ),  $\lambda > 0$  deterministic (1)
  - Poisson $(\lambda(t))$ ,  $\lambda(t)$  deterministic non-homogeneous Poisson, via time change  $\rightarrow$  (1)
  - Poisson( $\Lambda(t)$ ),  $\Lambda(t)$  stochastic process
    - double stochastic (or Cox-) process
    - basic model for credit risk
- a desert-island model: (NB,LN) (cover of [6])

## E. The Capital Charge Problem

• Estimate  $g_{\alpha}(L^{T+1})$  for  $\alpha$  large

Basel II:  $g_{\alpha} = \operatorname{VaR}_{\alpha}, \ \alpha \geq 99.9\%$  (reason)

- In-sample estimation of  $\operatorname{VaR}_{\alpha}(L^{T+1})$  for  $\alpha$  large is difficult, if not impossible (lack of data)
- Even for nice (repetitive) data one needs a structural model: Insurance Analytics ([12])

- Standard Actuarial Techniques
  - \* Analytic approximations (normal, translated gamma, Edgeworth, saddle-point, . . . )

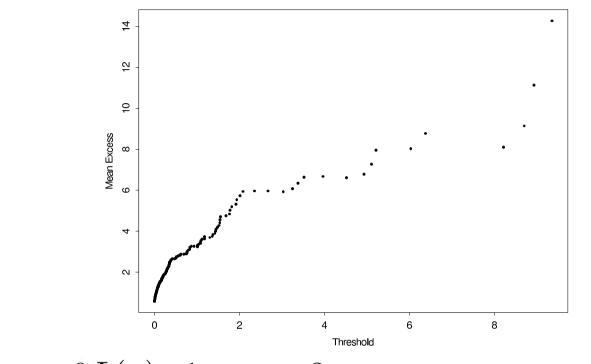
However: long-tailedness (Pareto, power tails)

 $P(X > x) \sim x^{-\alpha}L(x), x \text{ large}$ 

- \* Inversion methods (FFT)
- \* Recursive methods (Euler-Panjer)
- \* (Rare event) simulation
- \* Expert system Ansatz
- \* Extreme Value Theory (EVT) (confidence level  $\geq 99.9\%$ )
- \* Ruin theory

#### • Back to the data

pooled operational losses: mean excess plot



- $P(L > x) \sim x^{-\alpha}L(x), \ 1 < \alpha < 3$
- 20 80 rule
- one-claim-causes-ruin phenomenon ([1])

## **Summary**

- $\alpha\simeq 1$  and heavy-tailed loss-sizes, hence extremes matter
  - Extreme Value Theory (EVT) ([9])
- adding risk measures over different risk classes, hence dependence matters
  - Copulae  $(F_{\underline{X}}(\underline{x}) = C(F_1(x_1), \dots, F_d(x_d)))$  ([10])
- complicated loss-frequencies, hence point processes matter
  - double-stochastic (or Cox) processes ([4])
- full model analytically not tractable, hence
  - rare event simulation ([3])

### F. Accuracy of VaR-estimates

#### • Assumptions:

\* 
$$L_1, \ldots, L_n$$
 iid  $\sim F_L$   
\* For some  $\xi$ ,  $\beta$  and  $u$  large  $(G_{\xi,\beta}: \text{ GPD})$ :  
 $F_u(x) := \mathbb{P}[L - u \leq x | L > u] \sim G_{\xi,\beta(u)}(x), \ u \text{ large}$   
\* Use that:  $1 - F_L(x) = (1 - F_L(u)) (1 - F_u(x - u)), \quad x > u$ 

#### • Tail- and quantile estimate:

$$1 - \hat{F}_L(x) = \frac{N_u}{n} \left( 1 + \hat{\xi} \frac{x - u}{\hat{\beta}} \right)^{-1/\hat{\xi}}, \quad x > u$$

$$\widehat{\mathsf{VaR}}_\alpha = \hat{q}_\alpha = u - \frac{\hat{\beta}}{\hat{\xi}} \left( 1 - \left( \frac{N_u}{n(1 - \alpha)} \right)^{\hat{\xi}} \right)$$
(1)

• Idea: Comparison of estimated quantiles with the corresponding theoretical ones by means of a simulation study ([13], [7]).

#### • Simulation procedure:

- Choose  $F_L$  and fix  $\alpha_0 < \alpha < 1$ ,  $N_u$  (# of data points above u)
- **2** Calculate  $u = q_{\alpha_0}$  and the true value of the quantile  $q_{\alpha}$
- <sup>(3)</sup> Sample  $N_u$  independent points of  $F_L$  above u by the rejection method. Record the total number n of sampled points this requires
- ${\bf \Phi}$  Estimate  $\xi,\,\beta$  by fitting the GPD to the  $N_u$  exceedances over u by means of MLE
- **igodoldsymbol{\Theta}** Determine  $\hat{q}_{\alpha}$  according to (1)
- **6** Repeat N times the above to arrive at estimates of  $Bias(\hat{q}_{\alpha})$  and  $SE(\hat{q}_{\alpha})$
- ${\ensuremath{\textcircled{O}}}$  Require bias and standard error to be small  $\Rightarrow$  datasize

### **Example:** Pareto distribution with $\alpha = 2$

$u = F^{\leftarrow}(x_q)$	lpha	Goodness of $\widehat{VaR}_{\alpha}$
q = 0.7	0.99	A minimum number of <b>100 exceedances</b> (corresponding to 333 observations) is required to ensure accuracy wrt bias and standard error.
	0.999	A minimum number of <b>200 exceedances</b> (corresponding to 667 observations) is required to ensure accuracy wrt bias and standard error.
q = 0.9	0.99	Full accuracy can be achieved with the minimum number <b>25</b> of <b>exceedances</b> (corresponding to 250 observations).
	0.999	A minimum number of <b>100 exceedances</b> (corresponding to 1000 observations) is required to ensure accuracy wrt bias and standard error.

## Summary

- Minimum number of observations increases as the tails become thicker ([13], [7]).
- Large number of observations necessary to achieve targeted accuracy.
- **Remember**: The simulation study was done under idealistic assumptions (iid, exact Pareto). Operational risk losses, however, typically do NOT fulfil these assumptions.

## **G.** Conclusions

- OP risk  $\neq$  market risk, credit risk
- OP risk losses resemble non-life insurance losses
- Actuarial methods (including EVT) aiming to derive capital charges are for the moment of limited use due to
  - % lack of data
  - % inconsistency of the data with the modelling assumptions
- OP risk loss databases must grow
- Sharing/pooling internal operational risk data? Near losses?

- Choice of risk measure: ES better than VaR
- Heavy-tailed ruin estimation for general risk processes ([11]): an interesting mathematical problem related to time change
- Alternatives?
  - Insurance. Example: FIORI, Swiss Re (Financial Institution Operating Risk Insurance)
  - % Securitization / Capital market products
- OP risk charges can not be based on statistical modelling alone
- Pillar 2 (overall OP risk management such as analysis of causes, prevention, . . . ) more important than Pillar 1

## **H.** References

[1] Asmussen, S. (2000). Ruin Probabilities. World Scientific.

- [2] Basel Committee on Banking Supervision. The New Basel Capital Accord. April 2003. BIS, Basel, Switzerland, www.bis.org/bcbs
- [3] Bucklew, J.A. (2004). Introduction to Rare Event Simulation. Springer.

[4] Chavez-Demoulin, V. and Embrechts, P. (2004). Smooth ex-

tremal models in finance and insurance. The Journal of Risk and Insurance 71(2), 183-199.

- [5] Chavez-Demoulin, V. Embrechts, P. and Nešlehová, J. (2005) Quantitative models for operational risk: extremes, dependence and aggregation. Preprint, ETH Zürich.
- [6] Cruz, M.G. (2002). Modeling, Measuring and Hedging Operational Risk. A Quantitative Approach. Wiley, New York.
- [7] Embrechts P., Furrer H.J., and Kaufmann R. (2003). Quantifying regulatory capital for operational risk. Derivatives Use, Trading and Regulation 9(3), 217-233.

[8] Embrechts, P., Kaufmann, R., and Samorodnitsky, G. (2004).

Ruin theory revisited: stochastic models for operational risk. Risk Management for Central Bank Foreign Reserves, 243-261.

- [9] Embrechts, P., Klüppelberg, C., and Mikosch, T. (1997). *Modelling Extremal Events for Insurance and Finance*. Springer.
- [10] Embrechts, P., McNeil, A., and Straumann, D. (2002). Correlation and dependence in risk management: properties and pitfalls.
   In: *Risk Management: Value at Risk and Beyond*, ed. M.A.H. Dempster, Cambridge University Press, Cambridge, 176-223.
- [11] Embrechts, P., and Samorodnitsky, G. (2003). Ruin problem and how fast stochastic processes mix. Annals of Applied Probability, Vol. 13, 1-36.

- [12] McNeil, A.J., Frey, R., and Embrechts, P. (2005) *Quantitative Risk Management: Concepts, Techniques and Tools.* Princeton University Press.
- [13] McNeil, A.J., and Saladin, T. (1997) The peaks over thresholds method for estimating high quantiles of loss distributions. *Proceedings of XXVIIth International ASTIN Colloquium*, Cairns, Australia, 23-43.
- [14] The Economist (2003). *Blockage in Basel.* Vol. 369, No 8344, October 2003.